

last lecture: $P(Y \geq X) = \sum_{x=1}^{\infty} P(X=x, Y \geq x)$
 $= \dots = \sum_{x=1}^{\infty} P(X=x) \underbrace{P(Y \geq x)}_{?}$

So what is $P(Y \geq x)$?

$$\begin{aligned}
 P(Y \geq x) &= 1 - P(Y < x) \\
 &= 1 - \sum_{y=1}^{x-1} P(Y=y) \\
 &= 1 - \sum_{y=1}^{x-1} p (1-p)^{y-1} \quad b/c \quad Y \sim \text{geom}_1(p) \\
 &\stackrel{m=y-1}{=} 1 - p \sum_{m=0}^{x-2} (1-p)^m \\
 &= 1 - p \frac{1 - (1-p)^{x-1}}{1 - (1-p)} \\
 &= (1-p)^{x-1}
 \end{aligned}$$

$$\sum_{k=0}^{n-1} r^k = \frac{1-r^n}{1-r}; \text{ here } r=1-p < 1$$

So now back to our problem

Pg 2

$$\begin{aligned}
 P(Y \geq X) &= \sum_{x=1}^{\infty} (1-p)^x p (1-p)^{x-1} \\
 &= p \sum_{x=1}^{\infty} (1-p)^{2(x-1)} \\
 &= \frac{p}{1-(1-p)^2} \\
 &= \frac{p}{1-(1+p^2-2p)} \\
 &= \frac{1}{(p-2)} = \frac{1}{(2-p)}
 \end{aligned}$$

b/c $\sum_{k=0}^{\infty} r^k = \frac{1}{1-r}$
here $r = (1-p)^2 < 1$

$$(a) P(\underbrace{\min(X, Y)}_{\text{let's call this } Z} > z) \stackrel{\text{why?}}{=} P(X > z, Y > z)$$

$$\stackrel{\text{independence}}{=} P(X > z)P(Y > z)$$

$$P(Z > z) = (1-p)^{z-1}(1-p)^{z-1} = \left\{ (1-p)^2 \right\}^{z-1}$$

$$\therefore Z \sim \text{geom}_1\left(1 - (1-p)^2\right). \quad \leftarrow \text{think about this}$$

$$(c) P(X+Y=z) = \sum_{x=1}^{z-1} P(X=x, X+Y=z) \quad \text{for any } z \geq 2$$

↑ Why ? pg 3

$$= \sum_{x=1}^{z-1} P(X=x, Y=z-x)$$

$$\stackrel{\text{indep.}}{=} \sum_{x=1}^{z-1} P(X=x) P(Y=z-x)$$

$$= \sum_{x=1}^{z-1} p(1-p)^{x-1} p(1-p)^{z-x-1}$$

$$= \dots (z-1)p^2(1-p)^{z-2}$$

Simply use the sum of geometric series trick.

$$(d) P(Y=y | X+Y=z) \stackrel{\text{defn}}{=} \frac{P(Y=y, X+Y=z)}{P(X+Y=z)}$$

$$P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{P(A, B)}{P(B)}$$

Why?

think of $A \equiv Y=y$
 $B \equiv X+Y=z$!

$$= \frac{P(X=z-y, Y=y)}{P(X+Y=z)} \stackrel{\text{indep.}}{=} \frac{P(X=z-y) P(Y=y)}{p(z-1)p^2(1-p)^{z-2}}$$

$$= \dots = \frac{1}{z-1}$$

Now let us attempt to work out problems using similar concepts but for continuous random variables!

(Q) X and Y are independent RVs each w/ an exponential (λ) distribution. Find the pdf of (i) $Z = X + Y$; and (ii) $W = Y - X^2$

Soln :- Clearly we expect $f_{XY}(x,y) = f_X(x)f_Y(y)$
 $\Rightarrow f_{XY}(x,y) = \begin{cases} \lambda e^{-\lambda x} \lambda e^{-\lambda y} & ; \text{ if } x, y \geq 0 \\ 0 & ; \text{ o.w.} \end{cases}$

$$(i) f_Z(z) = \frac{d}{dz} F_Z(z) \text{ where } F_Z(z) = P(Z \leq z)$$

$$F_Z(z) = P(X + Y \leq z) = P(Y \leq -X + z)$$

$$\text{Let } z \geq 0; \quad F_Z(z) = P(Y \leq -X + z) = \int_0^z \int_0^{-x+z} f_{XY}(x,y) dy dx$$

You should think of the outer integral as summing over all partitioning subsets (eg. $\sum_{x=1}^{\infty} P(\cdot, \cdot)$) of pg 1.

But for all $x > z \quad P(Y \leq -x+z)$

$$= P(Y \leq (\text{re } \text{no})) = 0 \cdot (\text{Why?})$$

(think of the def'n of $\exp(\lambda) \text{ D}^n$.)

$$\therefore F_Z(z) = \int_0^z \int_0^{(-x+z)} \lambda e^{-\lambda x} \lambda e^{-\lambda y} dy dx$$

$$= \begin{cases} 1 - e^{-\lambda z} - \lambda z e^{-\lambda z}; & \forall z \geq 0 \\ 0 & ; z < 0 \end{cases}$$

$$P(X \leq x)$$

$$= \int_0^x f_X(x) dx$$

$$= \int_0^x \lambda e^{-\lambda x} dx$$

for all $x > 0$.

but 0 $\forall x < 0$

$$\text{Consequently, } f_Z(z) = \begin{cases} 0 & ; z < 0 \\ \lambda z e^{-\lambda z}; & z \geq 0 \end{cases}$$

part (ii) :- Next Lecture!

Review Concepts :- for continuous RVs

<u>Marginal Cdfs</u>	$F_X(x) = P(X \leq x) = \int_{-\infty}^x \int_{-\infty}^{\infty} f_{XY}(x,y) dy dx$
	$f_X(x) = \int_{-\infty}^{\infty} f_{XY}(x,y) dy \Rightarrow F_X(x) = P(X \leq x) = \int_{-\infty}^x f_X(x) dx$