



# Two group intervals

Statistical Inference

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# Independent group $t$ confidence intervals

- Suppose that we want to compare the mean blood pressure between two groups in a randomized trial; those who received the treatment to those who received a placebo
- We cannot use the paired  $t$  test because the groups are independent and may have different sample sizes
- We now present methods for comparing independent groups

# Notation

- Let  $X_1, \dots, X_{n_x}$  be iid  $N(\mu_x, \sigma^2)$
- Let  $Y_1, \dots, Y_{n_y}$  be iid  $N(\mu_y, \sigma^2)$
- Let  $\bar{X}$ ,  $\bar{Y}$ ,  $S_x$ ,  $S_y$  be the means and standard deviations
- Using the fact that linear combinations of normals are again normal, we know that  $\bar{Y} - \bar{X}$  is also normal with mean  $\mu_y - \mu_x$  and variance  $\sigma^2 (\frac{1}{n_x} + \frac{1}{n_y})$
- The pooled variance estimator  $S_p^2 = \frac{(n_x - 1) S_x^2 + (n_y - 1) S_y^2}{(n_x + n_y - 2)}$  is a good estimator of  $\sigma^2$

# Note

- The pooled estimator is a mixture of the group variances, placing greater weight on whichever has a larger sample size
- If the sample sizes are the same the pooled variance estimate is the average of the group variances
- The pooled estimator is unbiased 
$$E[S_p^2] = \frac{(n_x - 1) E[S_x^2] + (n_y - 1) E[S_y^2]}{n_x + n_y - 2} = \frac{(n_x - 1)\sigma^2 + (n_y - 1)\sigma^2}{n_x + n_y - 2}$$
- The pooled variance estimate is independent of  $\bar{Y} - \bar{X}$  since  $S_x$  is independent of  $\bar{X}$  and  $S_y$  is independent of  $\bar{Y}$  and the groups are independent

# Result

- The sum of two independent Chi-squared random variables is Chi-squared with degrees of freedom equal to the sum of the degrees of freedom of the summands
- Therefore 
$$(n_x + n_y - 2) S_p^2 / \sigma^2 = (n_x - 1) S_x^2 / \sigma^2 + (n_y - 1) S_y^2 / \sigma^2 \\ = \chi^2_{n_x - 1} + \chi^2_{n_y - 1} = \chi^2_{n_x + n_y - 2}$$

# Putting this all together

- The statistic  $\frac{\bar{Y} - \bar{X} - (\mu_y - \mu_x)}{\sigma \sqrt{\frac{1}{n_x} + \frac{1}{n_y}}} \sim \frac{\bar{Y} - \bar{X} - (\mu_y - \mu_x)}{S_p \sqrt{\frac{1}{n_x} + \frac{1}{n_y}}}$  is a standard normal divided by the square root of an independent Chi-squared divided by its degrees of freedom
- Therefore this statistic follows Gosset's t distribution with  $n_x + n_y - 2$  degrees of freedom
- Notice the form is (estimator - true value) / SE

# Confidence interval

- Therefore a  $(1 - \alpha) \times 100\%$  confidence interval for  $\mu_y - \mu_x$  is  $\bar{Y} - \bar{X} \pm t_{\{n_x + n_y - 2, 1 - \alpha/2\}} S_p \sqrt{\left(\frac{1}{n_x} + \frac{1}{n_y}\right)^{1/2}}$
- Remember this interval is assuming a constant variance across the two groups
- If there is some doubt, assume a different variance per group, which we will discuss later

# Example

Based on Rosner, Fundamentals of Biostatistics

- Comparing SBP for 8 oral contraceptive users versus 21 controls
- $\bar{X}_{\{OC\}} = 132.86$  mmHg with  $s_{\{OC\}} = 15.34$  mmHg
- $\bar{X}_{\{C\}} = 127.44$  mmHg with  $s_{\{C\}} = 18.23$  mmHg
- Pooled variance estimate

```
sp <- sqrt((7 * 15.34^2 + 20 * 18.23^2)/(8 + 21 - 2))  
132.86 - 127.44 + c(-1, 1) * qt(0.975, 27) * sp * (1/8 + 1/21)^0.5
```

```
## [1] -9.521 20.361
```



```
data(sleep)
x1 <- sleep$extra[sleep$group == 1]
x2 <- sleep$extra[sleep$group == 2]
n1 <- length(x1)
n2 <- length(x2)
sp <- sqrt(((n1 - 1) * sd(x1)^2 + (n2 - 1) * sd(x2)^2)/(n1 + n2 - 2))
md <- mean(x1) - mean(x2)
semd <- sp * sqrt(1/n1 + 1/n2)
md + c(-1, 1) * qt(0.975, n1 + n2 - 2) * semd
```

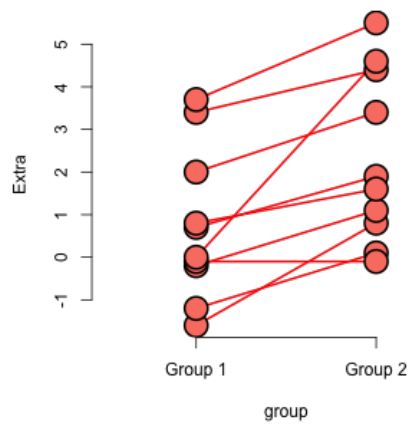
```
## [1] -3.3639 0.2039
```

```
t.test(x1, x2, paired = FALSE, var.equal = TRUE)$conf
```

```
## [1] -3.3639 0.2039
## attr("conf.level")
## [1] 0.95
```

```
t.test(x1, x2, paired = TRUE)$conf
```

# Ignoring pairing



# Unequal variances

- Under unequal variances  $\bar{Y} - \bar{X} \sim N\left(\mu_y - \mu_x, \frac{s_x^2}{n_x} + \frac{\sigma_y^2}{n_y}\right)$
- The statistic  $\frac{\bar{Y} - \bar{X} - (\mu_y - \mu_x)}{\sqrt{\frac{s_x^2}{n_x} + \frac{\sigma_y^2}{n_y}}}$  approximately follows Gosset's t distribution with degrees of freedom equal to  $\frac{(S_x^2 / n_x + S_y^2 / n_y)^2}{\left\{ \left( \frac{S_x^2}{n_x} \right)^2 / (n_x - 1) + \left( \frac{S_y^2}{n_y} \right)^2 / (n_y - 1) \right\}}$

# Example

- Comparing SBP for 8 oral contraceptive users versus 21 controls
- $\bar{X}_{\{OC\}} = 132.86$  mmHg with  $s_{\{OC\}} = 15.34$  mmHg
- $\bar{X}_{\{C\}} = 127.44$  mmHg with  $s_{\{C\}} = 18.23$  mmHg
- $df=15.04$ ,  $t_{\{15.04, .975\}} = 2.13$
- Interval  $132.86 - 127.44 \pm 2.13 \sqrt{\frac{15.34^2}{8} + \frac{18.23^2}{21}} = [-8.91, 19.75]$
- In R, `t.test(..., var.equal = FALSE)`

# Comparing other kinds of data

- For binomial data, there's lots of ways to compare two groups
  - Relative risk, risk difference, odds ratio.
  - Chi-squared tests, normal approximations, exact tests.
- For count data, there's also Chi-squared tests and exact tests.
- We'll leave the discussions for comparing groups of data for binary and count data until covering glms in the regression class.
- In addition, Mathematical Biostatistics Boot Camp 2 covers many special cases relevant to biostatistics.