

# Computational Statistics-Report

David Niederkofler, Erlend Lokna

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```
mydata<-read.table("Report2_Dataset.txt", header=FALSE)
```

## Statistical Analysis of Covariates

It is important to mention the use of notation before we proceed. We will in this section use the notation  $\mathbf{x} = \{x_1, x_2, \dots, x_n\}$  as our covariate vector.

### Ascicles

#### 1.1 Model selection

Since the Ascicles - covariate has a 0-1 outcome we can assume that it is Bernoulli distributed with parameter  $\theta$ . A natural conjugate prior for the Bernoulli distribution is the Beta distribution. The posterior beta distribution for the parameter is given by

$$Beta(\theta|a + \sum_{i=1}^n x_i, b + n - \sum_{i=1}^n x_i)$$

#### 1.2 Results

The following results were found using the posterior beta distribution with  $a=1$  and  $b=1$  (Uniform distributed) for the ascicles data:

```
## Posterior mean: 0.08227848
```

```
## Posterior mode: 0.07961783
```

```
## Centered 95% Confidence Interval: [ 0.05235453 , 0.1119428 ]
```

With the following HPD interval:

```
##      lower      upper
## 0.05074464 0.10984480
## attr("credMass")
## [1] 0.95
```

### 1. Sex

The sex of the patients is encoded in a binary variable, where 0 means *male* and 1 means *female*.

#### 1.1 Model selection

We assume a Bernoulli model  $Ber(\theta)$  for the sex of the patient conditional on one parameter  $\theta$ , the probability of the patient to be female. The density function is given by

$$f(x|\theta) = \theta^x (1 - \theta)^{1-x}, \tag{1}$$

where  $x \in \{0, 1\}$ . As a prior distribution for  $\theta$  we use the natural conjugate family of the Bernoulli distribution, namely the Beta distribution,  $Beta(a, b)$ , with two shape parameters  $a = b = 2$  to give more weight to the middle of the interval  $[0, 1]$ , knowing how females and males are represented in the general population. The density is given by

$$h(\theta) = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \theta^{a-1} (1-\theta)^{b-1}, \quad (2)$$

for  $\theta \in [0, 1]$ .

## 1.2 Results

From the given dataset we get the sample size  $n$  and the sum of the observations  $s$ :

```
n<-length(mydata$V6[!is.na(mydata$V6)])
s<-sum(mydata$V6)
n
## [1] 312
s
```

```
## [1] 276
```

Therefore the posterior distribution is  $Beta(2 + s, 2 + n - s)$ , which turns out to be  $Beta(278, 38)$ . From that we get

```
## Posterior mean: 0.8797468
## Posterior mode: 0.8821656
## Centered 95% Confidence Interval: [ 0.8417454 , 0.9132003 ]
```

And the HPD confidence Interval calculates to:

```
tst<-rbeta(1e5,278,38)
hdi(tst)

##      lower      upper
## 0.8433744 0.9145678
## attr(,"credMass")
## [1] 0.95
```

## 2. Spiders

The presence of spiders is encoded in a Binary variable, where 1 means spiders are present.

### 1.1 Model selection

We assume a Bernoulli model  $Ber(\theta)$  for the presence of spiders in patients conditional on one parameter  $\theta$ , the probability of the presence of spiders in the patient. The density function is given as stated earlier. As a prior distribution for  $\theta$  we use the natural conjugate family of the Bernoulli distribution, namely the Beta distribution,  $Beta(a, b)$ , with two shape parameters  $a = b = 1$ , because we have no prior information. The density is given as above.

### 1.2 Results

From the given dataset we get the sample size  $n$  and the sum of the observations  $s$ :

```
n<-length(mydata$V9[!is.na(mydata$V9)])
s<-sum(mydata$V9)
n
```

```
## [1] 312
```

```
s
```

```
## [1] 90
```

Therefore the posterior distribution is  $Beta(1 + s, 1 + n - s)$ , which turns out to be  $Beta(91, 223)$ . From that we get

```
## Posterior mean: 0.2911392
```

```
## Posterior mode: 0.2898089
```

```
## Centered 95% Confidence Interval: [ 0.2410228 , 0.341131 ]
```

And the HPD confidence interval calculates to:

```
tst<-rbeta(1e5,91,223)
hdi(tst)
```

```
##      lower      upper
## 0.2392598 0.3393446
## attr(,"credMass")
## [1] 0.95
```

### 3. Hepatomegaly

The presence of hepatomegaly is encoded in a Binary variable, where 1 means hepatomegaly is present.

#### 1.1 Model selection

We assume a Bernoulli model  $Ber(\theta)$  for the presence of hepatomegaly in the patient, conditional on one parameter  $\theta$ , the probability of the presence of hepatomegaly in the patient. The density function is given as stated earlier. As a prior distribution for  $\theta$  we use the natural conjugate family of the Bernoulli distribution, namely the Beta distribution,  $Beta(a, b)$ , with two shape parameters  $a = b = 1$ , because we have no prior information. The density is given as above.

#### 1.2 Results

From the given dataset we get the sample size  $n$  and the sum of the observations  $s$ :

```
n<-length(mydata$V8[!is.na(mydata$V8)])
s<-sum(mydata$V8)
n
```

```
## [1] 312
```

```
s
```

```
## [1] 160
```

Therefore the posterior distribution is  $Beta(1 + s, 1 + n - s)$ , which turns out to be  $Beta(161, 153)$ . From that we get

```
## Posterior mean: 0.5126582
```

```
## Posterior mode: 0.5127389
```

```
## Centered 95% Confidence Interval: [ 0.4575015 , 0.5678225 ]
```

And the HPD confidence interval calculates to:

```
tst<-rbeta(1e5,161,153)
hdi(tst)
```

```
##      lower      upper
## 0.4572249 0.5666820
## attr(,"credMass")
## [1] 0.95
```

## 4. Histologic stage

The Histologic stage of the disease is a number in  $\{1, 2, 3, 4\}$ , where the stage increases with severeness. We will give here the frequencies of the stages in the dataset.

```
##      1      2      3      4
## 16    67   120   109
```

We see that, most patients have been diagnosed in the last to stages of the disease.

## 5. Age

The age of the patient in days.

### 5.1 Model selection

The data seems to follow a poisson distribution  $Poi(\lambda)$ . Using the non informative Jeffreys prior, we can derive that the posterior for the parameter  $\lambda$  is Gamma distributed.

$$\theta|x \sim \text{Gamma}(\alpha = \frac{1}{2} + \sum_{i=1}^n x_i, \beta = n)$$

### 5.2 Results

$$\theta|x \sim \text{Gamma}(\frac{1}{2} + s, n)$$

```
## posterior distribution: Gamma( 5700066 , 312 )
## mean: 18269.44
## variance: 58.55591
## HPD intervall:
##      lower      upper
## 18254.75 18284.78
## attr(,"credMass")
## [1] 0.95
```

## 6. Cholesterol

### 6.1 Model selection

We assume that the data is sampled from a poisson,  $Poi(\lambda)$ , distribution, and we use the non informative Jeffreys prior for the rate parameter in the bayesian analysis.

$$\mathbf{x} \sim Poi(\lambda)$$

$$h(\lambda) \propto \lambda^{-\frac{1}{2}}$$

## 6.2 Results

```
## posterior distribution: Gamma( 104941.5 , 312 )
## mean: 336.351
## variance: 1.078048
## HPD intervall:

##      lower      upper
## 334.3068 338.3873
## attr(,"credMass")
## [1] 0.95
```

## 7. Urine

### 7.1 Model selection

$$\mathbf{x} \sim Poi(\lambda)$$

$$h(\lambda) \propto \lambda^{-\frac{1}{2}}$$

### 7.2 Results

```
## posterior distribution: Gamma( 30271.5 , 312 )
## mean: 97.02404
## variance: 0.3109745
## HPD intervall:

##      lower      upper
## 95.93102 98.11615
## attr(,"credMass")
## [1] 0.95
```

## 8 SGOT

### 8.1 Model selection

$$\mathbf{x} \sim Poi(\lambda)$$

$$h(\lambda) \propto \lambda^{-\frac{1}{2}}$$

### 8.2 Results

```
poisson_jeffrey(mydata$V16)
```

```
## posterior distribution: Gamma( 38238.08 , 312 )
## mean: 122.5579
## variance: 0.3928139
## HPD intervall:

##      lower      upper
## 121.3463 123.8056
## attr(,"credMass")
## [1] 0.95
```

## 9. Plateles

### 9.1 Model selection

$$\mathbf{x} \sim Poi(\lambda)$$

$$h(\lambda) \propto \lambda^{-\frac{1}{2}}$$

### 9.2 Results

```
poisson_jeffrey(mydata$V18)
```

```
## posterior distribution: Gamma( 80676.5 , 312 )
## mean: 258.5785
## variance: 0.8287773
## HPD intervall:
##      lower      upper
## 256.8205 260.3943
## attr(,"credMass")
## [1] 0.95
```

## 10. Prothrombin

### 10.1 Model selection

$$\mathbf{x} \sim Poi(\lambda)$$

$$h(\lambda) \propto \lambda^{-\frac{1}{2}}$$

### 10.2 Results

```
poisson_jeffrey(mydata$V19)
```

```
## posterior distribution: Gamma( 3346.9 , 312 )
## mean: 10.72724
## variance: 0.03438219
## HPD intervall:
##      lower      upper
## 10.36086 11.08776
## attr(,"credMass")
## [1] 0.95
```

## Appendix

### Bernoulli/Beta

A natural conjugate prior for the Bernoulli distribution is the Beta distribution.

$$f(x_i|\theta) = \theta^{x_i}(1-\theta)^{1-x_i}$$

$$L(\mathbf{x}|\theta) = \theta^{\sum_{i=1}^n x_i} (1-\theta)^{n-\sum_{i=1}^n x_i}$$

$$h(\theta) = Beta(a, b)$$

We proceed by calculating the posterior distribution for  $\theta$

$$h(\theta|\mathbf{x}) \propto L(\mathbf{x}|\theta)h(\theta) = \theta^{\sum_{i=1}^n x_i} (1-\theta)^{n-\sum_{i=1}^n x_i} \frac{1}{B(a, b)} \theta^{a-1} (1-\theta)^{b-1} I(0 < \theta < 1)$$

$$\propto \text{Beta}(\theta|a + \sum_{i=1}^n x_i, b + n - \sum_{i=1}^n x_i)$$

## Poisson/Gamma

If our data  $X_1, \dots, X_n$  are iid  $\text{Poisson}(\lambda)$  distributed then a  $\text{gamma}(\alpha, \beta)$  prior on  $\lambda$  is a conjugate prior. The Likelihood function is:

$$L(\lambda|\mathbf{x}) = \prod_{i=1}^n \frac{e^{-\lambda} \lambda^{\sum_{i=1}^n x_i}}{x_i!} = \frac{e^{-\lambda} \lambda^{\sum_{i=1}^n x_i}}{\prod_{i=1}^n x_i!}$$

Our gamma prior has the expression:

$$h(\lambda) = \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda}$$

Using bayes rule we find the following posterior:

$$h(\lambda|\mathbf{x}) \propto h(\lambda)L(\mathbf{x}|\lambda) \propto \lambda^{\sum_{i=1}^n x_i + \alpha - 1} e^{-(n+\beta)\lambda}$$

$$\propto \text{gamma}(\sum_{i=1}^n x_i + \alpha, n + \beta)$$

## Poisson/Jeffreys prior

The density distribution for poisson is equal to

$$f(n|\lambda) = e^{-\lambda} \frac{\lambda^n}{n!}$$

The jeffreys prior  $h(\theta)$  is a non informative prior distribution for a parameter space and its proportionality is expressed as

$$h(\theta) \propto \sqrt{\det I(\theta)}$$

$$I(\theta) = -E\left[\frac{\partial^2}{\partial \theta^2} \log f(x|\theta)\right] = \frac{1}{\theta}$$

And the following jeffreys prior is thus

$$h(\theta) \propto \theta^{-\frac{1}{2}} I_{\theta>0}$$

The posterior is calculated as follows

$$h(\theta|x) \propto f(\mathbf{x}|\theta)h(\theta) \propto e^{-n\theta} \theta^{-\frac{1}{2} + \sum_{i=1}^n x_i}$$

which is in fact a gamma distribution

$$\theta|x \sim \text{Gamma}(\alpha = \frac{1}{2} + \sum_{i=1}^n x_i, \beta = n)$$