

概率统计第四次作业

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3.1

证明 $E(X) = np$:

$$\because X \sim B(n, p)$$

$$\therefore P(X = k) = \binom{n}{k} p^k (1-p)^{n-k}$$

$$\therefore E(X) = \sum_{k=0}^{\infty} k \cdot P(X = k) = \sum_{k=0}^{\infty} k \binom{n}{k} p^k (1-p)^{n-k} = (1-p)^n \sum_{k=1}^{\infty} k \binom{n}{k} \left(\frac{p}{1-p}\right)^k$$

$$\text{令 } x = \frac{p}{1-p}, \text{ 则}$$

$$\because \sum_{k=0}^{\infty} \binom{n}{k} x^k = (1+x)^n$$

两边求导得

$$\therefore \sum_{k=1}^{\infty} k \binom{n}{k} x^{k-1} = n(1+x)^{n-1}$$

$$\therefore \sum_{k=1}^{\infty} k \binom{n}{k} x^k = nx(1+x)^{n-1}$$

$$\therefore E(X) = (1-p)^n \cdot n \cdot \frac{p}{1-p} \cdot \left(1 + \frac{p}{1-p}\right)^{n-1} = np$$

证明 $\text{Var}(X) = np(1-p)$:

$$\begin{aligned} \therefore E(X^2) &= \sum_{k=0}^{\infty} k^2 \cdot P(X = k) \\ &= \sum_{k=0}^{\infty} k^2 \binom{n}{k} p^k (1-p)^{n-k} \\ &= \sum_{k=0}^{\infty} k(k-1) \binom{n}{k} p^k (1-p)^{n-k} + \sum_{k=0}^{\infty} k \binom{n}{k} p^k (1-p)^{n-k} \\ &= (1-p)^n \sum_{k=1}^{\infty} k(k-1) \binom{n}{k} \left(\frac{p}{1-p}\right)^k + (1-p)^n \sum_{k=1}^{\infty} k \binom{n}{k} \left(\frac{p}{1-p}\right)^k \end{aligned}$$

$$\because \sum_{k=1}^{\infty} k \binom{n}{k} x^{k-1} = n(1+x)^{n-1}$$

$$\therefore \sum_{k=2}^{\infty} k(k-1) \binom{n}{k} x^{k-2} = n(n-1)(1+x)^{n-2}$$

$$\therefore \sum_{k=2}^{\infty} k(k-1) \binom{n}{k} x^k = n(n-1)x^2(1+x)^{n-2}$$

$$\therefore E(X^2) = (1-p)^n \cdot n(n-1) \left(\frac{p}{1-p} \right)^2 \left(1 + \frac{p}{1-p} \right)^{n-2} + np = n(n-1)p^2 + np = np(1-p) + n^2p^2$$

$$\therefore \text{Var}(X) = E(X^2) - E(X)^2 = np(1-p)$$

3.2

证明 $E(X) = \frac{1}{p}$:

$$\therefore P(X = k) = (1-p)^{k-1}p$$

$$\therefore E(X) = \sum_{k=1}^{\infty} k \cdot P(X = k) = \sum_{k=1}^{\infty} k(1-p)^{k-1}p = p \sum_{k=1}^{\infty} k(1-p)^{k-1}$$

$$\therefore \sum_{k=0}^{\infty} x^k = \frac{1}{1-x}$$

$$\therefore \sum_{k=1}^{\infty} kx^{k-1} = \frac{1}{(1-x)^2}$$

$$\therefore \sum_{k=1}^{\infty} k(1-p)^{k-1} = \frac{1}{p^2}$$

$$\therefore E(X) = p \cdot \frac{1}{p^2} = \frac{1}{p}$$

证明 $\text{Var}(X) = \frac{1-p}{p^2}$:

$$\therefore E(X^2) = \sum_{k=1}^{\infty} k^2(1-p)^{k-1}p = p \sum_{k=2}^{\infty} k(k-1)(1-p)^{k-1} + p \sum_{k=1}^{\infty} k(1-p)^{k-1}$$

$$\therefore \sum_{k=1}^{\infty} kx^{k-1} = \frac{1}{(1-x)^2}$$

$$\therefore \sum_{k=2}^{\infty} k(k-1)x^{k-2} = \frac{2}{(1-x)^3}$$

$$\therefore \sum_{k=2}^{\infty} k(k-1)x^{k-1} = \frac{2x}{(1-x)^3}$$

$$\therefore E(X^2) = p \cdot \frac{2(1-p)}{p^3} + \frac{1}{p} = \frac{1}{p} + \frac{2(1-p)}{p^2} = \frac{2-p}{p^2}$$

$$\therefore \text{Var}(X) = E(X^2) - E(X)^2 = \frac{2-p}{p^2} - \frac{1}{p^2} = \frac{1-p}{p^2}$$

3.3

证明 $E(X) = \frac{r}{p}$:

$$\therefore P(X = k) = \binom{k-1}{r-1} p^r (1-p)^{k-r}$$

$$\therefore E(X) = \sum_{k=r}^{\infty} k \cdot \binom{k-1}{r-1} p^r (1-p)^{k-r} = r p^r \sum_{k=r}^{\infty} \binom{k}{r} (1-p)^{k-r}$$

$$\therefore \sum_{k=r}^{\infty} \binom{k}{r} (1-p)^{k-r} = \sum_{k=r}^{\infty} \binom{k+1-1}{r+1-1} (1-p)^{k-r} = p^{-(r+1)}$$

$$\therefore E(X) = r p^r \cdot p^{-(r+1)} = \frac{r}{p}$$

证明 $\text{Var}(X) = \frac{r(1-p)}{p^2}$:

$$\therefore E(X^2) = \sum_{k=r}^{\infty} k^2 \cdot \binom{k-1}{r-1} p^r (1-p)^{k-r} = r(r+1) p^r \sum_{k=r}^{\infty} \binom{k+1}{r+1} (1-p)^{k-r} - r p^r \sum_{k=r}^{\infty} \binom{k}{r} (1-p)^{k-r}$$

$$\therefore \sum_{k=r}^{\infty} \binom{k+1}{r+1} (1-p)^{k-r} = \sum_{k=r}^{\infty} \binom{k+2-1}{r+2-1} (1-p)^{k-r} = p^{-(r+2)}$$

$$\therefore E(X^2) = r(r+1) p^r \cdot p^{-(r+2)} - \frac{r}{p}$$

$$\therefore \text{Var}(X) = E(X^2) - E(X)^2 = r(r+1) p^r \cdot p^{-(r+2)} - \frac{r}{p} - \frac{r^2}{p^2} = \frac{r(1-p)}{p^2}$$

3.4

证明 $E(X) = \lambda$:

$$\therefore P(X = k) = \frac{\lambda^k}{k!} e^{-\lambda}$$

$$\therefore E(X) = \sum_{k=0}^{\infty} k \cdot \frac{\lambda^k}{k!} e^{-\lambda} = \lambda e^{-\lambda} \sum_{k=1}^{\infty} \frac{\lambda^{k-1}}{(k-1)!} = \lambda e^{-\lambda} \cdot e^{\lambda} = \lambda$$

证明 $\text{Var}(X) = \lambda$:

$$\therefore E(X^2) = \sum_{k=0}^{\infty} k^2 \cdot \frac{\lambda^k}{k!} e^{-\lambda} = \sum_{k=2}^{\infty} k(k-1) \cdot \frac{\lambda^k}{k!} e^{-\lambda} + \sum_{k=1}^{\infty} k \cdot \frac{\lambda^k}{k!} e^{-\lambda} = \lambda^2 e^{-\lambda} \cdot e^{\lambda} + \lambda = \lambda^2 + \lambda$$

$$\therefore \text{Var}(X) = E(X^2) - E(X)^2 = \lambda$$

3.5

设 Y 表示一个的叶节点的高度.

设 X_i 表示第 i 轮该叶节点是否被选中, 选中时 $X_i = 1$, 未选中时 $X_i = 0$.

因为在第 i 轮一共有 i 个节点, 因此该节点被选中的概率 $P(X_i = 1) = \frac{1}{i}$, 则 $E(X_i) = 1 \cdot P(X_i = 1) = \frac{1}{i}$

$$\therefore E(Y) = E\left(\sum_{i=1}^k X_i\right) = \sum_{i=1}^k E(X_i) = \sum_{i=1}^k \frac{1}{i} \approx \ln k$$

3.6

有放回:

$$\therefore P(X = k) = \frac{\sum_{i=1}^5 \binom{5}{i} (k-1)^{5-i}}{10^5}$$

$$\therefore P(X = 1) = 0.00001, P(X = 2) = 0.00031, P(X = 3) = 0.00211, P(X = 4) = 0.00781, P(X = 5) = 0.02101, P(X = 6) = 0.04651, P(X = 7) = 0.09031, P(X = 8) = 0.15961, P(X = 9) = 0.26281, P(X = 10) = 0.40951$$

X	1	2	3	4	5	6	7	8	9	10
P	0.00001	0.00031	0.00211	0.00781	0.02101	0.04651	0.09031	0.15961	0.26281	0.40951

无放回:

$$\therefore P(X = k) = \frac{\binom{k-1}{4}}{\binom{10}{5}}$$

X	5	6	7	8	9	10
P	$\frac{1}{252}$	$\frac{5}{252}$	$\frac{5}{84}$	$\frac{5}{36}$	$\frac{5}{18}$	$\frac{1}{2}$

3.7

令 $X \sim B(n, 0.99)$, 则 X 服从参数为 n 和 0.99 的二项分布.

$$\therefore P(X \geq k) = \sum_{k=k}^n \binom{n}{k} \times 0.99^k \times 0.01^{n-k}$$

令 $n = 102$, 则

$$\therefore P(X \geq 100) = \sum_{k=100}^{102} \binom{102}{k} \times 0.99^k \times 0.01^{102-k} = 0.916911014889440$$

令 $n = 103$, 则

$$\therefore P(X = 100) = \sum_{k=100}^{103} \binom{103}{k} \times 0.99^k \times 0.01^{103-k} = 0.979758767886053$$

所以 $x = 103 - 100 = 3$. 即 x 最小值是 3.

3.8

2. (1)

$$\therefore P(X = k) = \frac{\binom{k-1}{2}}{\binom{5}{3}}$$

$$\therefore P(X = 3) = 0.1, P(X = 4) = 0.3, P(X = 5) = 0.6$$

X	3	4	5
P	0.1	0.3	0.6

2. (2)

$$\therefore P(X = k) = \frac{\binom{6-k}{1}}{\binom{6}{2}}$$

$$\therefore P(X = 1) = \frac{1}{3}, P(X = 2) = \frac{4}{15}, P(X = 3) = \frac{1}{5}, P(X = 4) = \frac{2}{15}, P(X = 5) = \frac{1}{15}$$

X	1	2	3	4	5
P	$\frac{1}{3}$	$\frac{4}{15}$	$\frac{1}{5}$	$\frac{2}{15}$	$\frac{1}{15}$

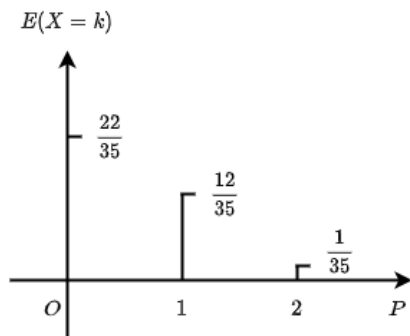
3. (1)

$$\therefore P(X = k) = \frac{\binom{2}{k} \binom{13}{3-k}}{\binom{15}{3}}$$

$$\therefore P(X = 0) = \frac{22}{35}, P(X = 1) = \frac{12}{35}, P(X = 2) = \frac{1}{35}$$

X	0	1	2
P	$\frac{22}{35}$	$\frac{12}{35}$	$\frac{1}{35}$

3. (2)



3.9

2.

设 Y 为每次抽取 10 件产品进行检验时发现的次品数. 则 $Y \sim B(10, 0.1)$

$$\therefore P(Y = 0) = \binom{10}{0} \times 0.1^0 \times 0.9^{10} = 0.9^{10}$$

$$\therefore P(Y = 1) = \binom{10}{1} \times 0.1^1 \times 0.9^9 = 0.9^9$$

$$\therefore P(Y > 1) = 1 - 0.9^{10} - 0.9^9 = 1 - 1.9 \times 0.9^9$$

即发现超过 1 件次品的概率为 $p = 1 - 1.9 \times 0.9^9$

则我们有 $X \sim (4, p)$, 即 $X \sim (4, 1 - 1.9 \times 0.9^9)$

$$\therefore E(X) = 4p = 4 - 7.6 \times 0.9^9$$

3.

$$\therefore P(X = k) = \frac{\sum_{i=1}^3 \binom{3}{i} (4-k)^{3-i}}{4^3}$$

$$\therefore P(X = 1) = \frac{37}{64}, P(X = 2) = \frac{19}{64}, P(X = 3) = \frac{7}{64}, P(X = 4) = \frac{1}{64}$$

X	1	2	3	4
P	$\frac{37}{64}$	$\frac{19}{64}$	$\frac{7}{64}$	$\frac{1}{64}$

$$\therefore E(X) = 1 \times \frac{37}{64} + 2 \times \frac{19}{64} + 3 \times \frac{7}{64} + 4 \times \frac{1}{64} = \frac{25}{16}$$

3.10

4. (1)

因为 $\sum_{j=1}^{\infty} \left| (-1)^{j+1} \frac{3^j}{j} \cdot \frac{2}{3^j} \right| = \sum_{j=1}^{\infty} \frac{2}{j}$ 调和级数发散, 即 $E(X)$ 并不绝对收敛.

所以 X 的数学期望并不存在.

4. (2)

$$\therefore P(X = k) = \left(\prod_{i=1}^{k-1} \frac{i}{i+1} \right) \cdot \frac{1}{k+1}$$

$$\therefore E(X) = \sum_{k=1}^{\infty} k \cdot \left(\prod_{i=1}^{k-1} \frac{i}{i+1} \right) \cdot \frac{1}{k+1} = \sum_{k=1}^{\infty} \frac{1}{k+1} \text{ 调和级数发散.}$$

6. (1)

$$E(X) = -2 \times 0.4 + 0 \times 0.3 + 2 \times 0.3 = -0.2$$

$$E(X^2) = 4 \times 0.4 + 0 \times 0.3 + 4 \times 0.3 = 2.8$$

$$E(3X^2 + 5) = 3E(X^2) + 5 = 3 \times 2.8 + 5 = 13.4$$

6. (2)

$$\therefore X \sim \pi(\lambda)$$

$$\therefore P(X = k) = \frac{\lambda^k}{k!} e^{-\lambda}$$

$$\therefore E\left(\frac{1}{X+1}\right) = \sum_{k=0}^{\infty} \frac{1}{k+1} \cdot \frac{\lambda^k}{k!} e^{-\lambda} = \frac{e^{-\lambda}}{\lambda} \sum_{k=0}^{\infty} \frac{\lambda^{k+1}}{(k+1)!} = \frac{(e^{\lambda} - 1) e^{-\lambda}}{\lambda}$$