

Residual Sum of Squares

- Its the sum of rest of $e_i = y_i - \hat{y}_i$

$$RSS = e_1^2 + e_2^2 + \dots + e_n^2 = \sum_{i=1}^n e_i^2$$

- The RSS expression is derived from the [Maximum Likelihood Estimation](#) by assuming that the error terms takes on a [normal distribution](#) bell curve

$$\hat{\beta}^{\text{MLE}} = \arg \max_{\beta} \sum_{i=1}^n \log[p(X_i|\beta)]$$

RSS Derivation

1. Assumption of Linear Regression

$$Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i$$

- $\varepsilon_i \sim \mathcal{N}(0, \sigma^2)$
- i.i.d.(independent,identically distributed)

1. Likelihood Function

$$L(\beta, \sigma^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x_i - \mu)^2 / 2\sigma^2}$$

Where :

- $x_i = Y_i$ The observed Data
- $\mu = \hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$ The expected value
- The [normal distribution](#) is on Y

$$L(\beta, \sigma^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(Y_i - \beta_0 - \beta_1 X_i)^2 / 2\sigma^2}$$

$$L(\beta, \sigma^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\sum (y_i - \hat{y}_i)^2 / 2\sigma^2}$$

- $(y_i - \hat{y}_i)^2 = \sum_{i=1}^n (y_i - \hat{y}_i)^2$

Log Likelihood

$$\begin{aligned} \log(L(\beta, \sigma^2)) &= \sum_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} - \sum (y_i - \hat{y}_i)^2 / 2\sigma^2 \\ &= -\frac{n}{2} \log(2\pi\sigma^2) - \sum_{i=1}^n (y_i - \hat{y}_i)^2 / 2\sigma^2 \end{aligned}$$

Maximazing with respect to β_0

$$\begin{aligned} \frac{dL}{d\hat{\beta}_0} &= \frac{1}{\sigma^2} \sum (y_i - \hat{\beta}_0 - \hat{\beta}_1 X_i) \\ &= \frac{1}{\sigma^2} \sum_{i=1}^n (y_i - \hat{\beta}_1 X_i) - \hat{\beta}_0 \end{aligned}$$

Solving For 0

$$\frac{1}{\sigma^2} \sum_{i=1}^n (y_i - \hat{\beta}_1 X_i) - n\hat{\beta}_0 = 0$$

$$\begin{aligned} n\hat{\beta}_0 &= \sum y_i - \hat{\beta}_1 X_i = n\bar{y} - \hat{\beta}_1 n\bar{x} \\ \hat{\beta}_0 &= \bar{y} - \hat{\beta}_1 \bar{x} \end{aligned}$$

Maximizing with respect to β_1

$$\frac{dL}{d\hat{\beta}_1} = -\frac{1}{2\sigma^2} \sum (-2X_i(y_i - \hat{\beta}_0 - \hat{\beta}_1 X_i))$$

$$\frac{dL}{d\hat{\beta}_1} = \frac{1}{\sigma^2} \sum (X_i(y_i - \bar{y} + \hat{\beta}_1 \bar{x} - \hat{\beta}_1 X_i))$$

Solving For 0

$$\frac{1}{\sigma^2} \sum (X_i(y_i - \bar{y} + \hat{\beta}_1 \bar{x} - \hat{\beta}_1 X_i)) = 0$$

$$\sum X_i(y_i - \hat{y}) - \hat{\beta}_1 \sum X_i(x_i - \bar{x}) = 0$$

$$\hat{\beta}_1 = \frac{\sum y_i - \hat{y}}{\sum x_i - \bar{x}}$$

$$\hat{\beta}_1 = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sum (x_i - \bar{x})^2}$$

$$\hat{\beta}_1 = \frac{\text{Cov}(X, Y)}{\text{Var}(X)}$$

- $f^n \frac{d}{df^n} = n \frac{d}{df} f^{n-1}$
- When maximizing a **Likelihood function** and solving for 0 we ignore the constant terms $\frac{1}{\sigma^2}$

Maximizing with respect to σ^2

$$\frac{dL}{d\hat{y}_i} = -\frac{n}{2\sigma^2} + \frac{1}{2(\sigma^2)^2} \sum (y_i - \hat{y}_i)^2$$

Solving For 0

$$-\frac{n}{2\sigma^2} + \frac{1}{2(\sigma^2)^2} \sum (y_i - \hat{y}_i)^2 = 0$$

$$\frac{n}{2\sigma^2} = \frac{1}{2(\sigma^2)^2} \sum (y_i - \hat{y}_i)^2$$

$$\frac{2(\sigma^2)^2}{2\sigma^2} = \frac{\sum_{i=1}^n (y_i - \hat{y}_i)^2}{n}$$

$$\sigma^2 = \frac{\sum_{i=1}^n (y_i - \hat{y}_i)^2}{n} = \frac{\text{RSS}}{n} = \text{MSE}$$