

## Continuous Uniform Distribution

- $X$  can only take on values from  $[a, b]$
- Same probability for every interval of a fixed length
- $X \sim \text{Uniform}(a, b)$

Probability density function: since all points are equally likely,  $f(x) = k$  for all  $a \leq x \leq b$

$$\int_a^b f(x) dx = 1$$

$$\Rightarrow \int_a^b k dx = 1$$

$$\Rightarrow k = \frac{1}{b-a}$$

$$\text{CDF: } F(x) = P(X \leq x)$$

$$= \int_a^x f(x) dx$$

$$= \int_a^x \frac{1}{b-a} dx$$

$$= \frac{x-a}{b-a}$$

$$\text{Mean: } \frac{a+b}{2}$$

$$\text{Variance: } \frac{(b-a)^2}{12}$$

Example: How to transform a r.v. with a general continuous distribution to obtain a uniformly distributed one.

Let  $X$  be a continuous r.v. with p.d.f. of the form

$$f(x) = 0.1e^{-0.1x} \text{ for } x > 0.$$

Let the new r.v. be given by

$$Y = e^{-0.1X}.$$

Show that  $Y$  has a uniform distribution on  $[0,1]$ .

$$f(x) = 0.1e^{-0.1x}$$

$$\begin{aligned} F_x(x) &= P(X \leq x) = \int_0^x 0.1e^{-0.1u} du \\ &= 1 - e^{-0.1x} \end{aligned}$$

$$\text{We know that } Y = e^{-0.1x}$$

$$\begin{aligned} F_Y(y) &= P(Y \leq y) \\ &= P(e^{-0.1x} \leq y) \\ &= P(X > \frac{\ln(y)}{0.1}) \\ &= 1 - P(X \leq \frac{\ln(y)}{0.1}) \\ &\vdots \end{aligned}$$

$$f_Y(y) = 1$$

### Exponential Distribution

In a Poisson process in which events occur at a rate  $\lambda$  over time, let  $X$  represent the length of time we wait for the first event occurrence.

Since  $X = \text{TIME}$  (not number of occurrences, which is discrete), exponential is *continuous*

Ex. Calls to a fire station follow a Poisson process, while the *length of time between consecutive calls* follows an exponential distribution.

Let  $X \sim \text{Exponential}(\lambda)$

$$\text{PDF: } f(x) = \lambda e^{-\lambda x}$$

$$\begin{aligned} \text{CDF: } F(x) &= P(X \leq x) = \int_0^x \lambda e^{-\lambda u} du \\ &= 1 - e^{-\lambda x} \end{aligned}$$

### Link between a Poisson process and an exponential distribution

Suppose that lightning strikes occur according to a Poisson process with rate  $\lambda$ .

Now, let  $X$  represent the time until the first lightning strike.

$$F(X) = P(X \leq x) = P(\text{time until the first strike} \leq x)$$

Since we are calculating the probability of one strike in a time interval  $\leq x$ , then there must be at least one strike in  $(0, x)$ . The number of strikes by  $x$  has a Poisson distribution with mean  $\mu = \lambda x$ .

$$\text{So } F(x) = P(\text{at least one strike by time } x)$$

$$F(x) = 1 - P(\text{no strikes by time } x)$$

$$= 1 - \frac{e^{-\lambda x} (\lambda x)^0}{0!}$$

$$= 1 - e^{-\lambda x}$$

$$\text{PDF: } f(x) = F'(x) = \lambda e^{-\lambda x}$$

Alternate form: for  $\theta > 0$ :

$$f(x) = \frac{1}{\theta} e^{-x/\theta} \quad \therefore \lambda = \frac{1}{\theta}$$

0 otherwise

$$F(x) = 1 - e^{-x/\theta}$$

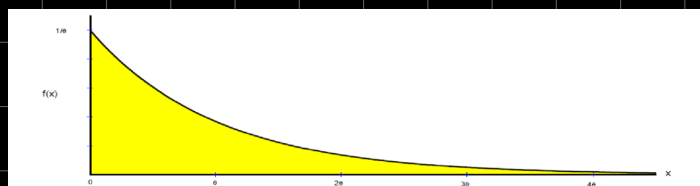


Figure 8.9: Graph of the probability density function of a *Exponential* ( $\theta$ ) random variable

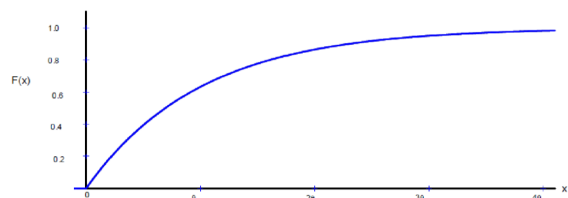


Figure 8.10: Cumulative distribution function for a *Exponential* ( $\theta$ ) random variable

### Deriving the mean and variance

We could do this by using integration by parts

Or we could use a gamma function:

$$\Gamma(\alpha) = \int_0^{\infty} x^{\alpha-1} e^{-x} dx \quad (\alpha > 0)$$

Properties:

$$\Gamma(1) = \int_0^{\infty} e^{-x} dx = 1$$

$$\begin{aligned} \Gamma(\alpha) &= (\alpha-1)\Gamma(\alpha-1) \\ &= (\alpha-1)! \quad \star \end{aligned}$$

$$\Gamma(1/2) = \sqrt{\pi}$$

Let  $X \sim \text{Exponential}(\theta)$

$$E(X) = \int_0^{\infty} x \cdot f(x) dx$$

$$= \int_0^{\infty} x \cdot \frac{1}{\theta} e^{-x/\theta} dx$$

Using  $y = \frac{x}{\theta} \Rightarrow x = \theta y, dx = \theta dy$

$$E(X^2) = \int_0^{\infty} \theta^2 y^2 \cdot \frac{1}{\theta} \cdot e^{-y} \cdot \theta dy$$

$$= \theta^2 \int_0^{\infty} y^2 e^{-y}$$

$$= \theta^2 \cdot \Gamma(3)$$

$$= \theta^2 \cdot 2!$$

$$= 2\theta^2$$

The same substitution gives us  $E(X) = \theta$

$$\Rightarrow \text{Var}(X) = E(X^2) - E(X)^2$$

$$= 2\theta^2 - \theta^2$$

$$= \theta^2$$

$$\Rightarrow \text{sd}(X) = \theta$$

### Memoryless Property of Exponential Distributions

$$P(X > b+c \mid X > b) = \frac{P(X > b+c \cap X > b)}{P(X > b)}$$

$$= \frac{e^{-(b+c)/\theta}}{e^{-b/\theta}}$$

$$= e^{-c/\theta}$$

$$= P(X > c)$$

Ex. The amount of time in hours that a computer survives before breaking down is exponentially distributed with a mean of 100 hours.

Let  $C$  = event that if a computer survives more than 100 hours, it survives an additional 50 hours

Let  $A$  = event that computer survives more than 100 hours

$B$  = event that a computer survives more than 150 hours

$$P(C) = P(B|A) = P(\text{survives } 100-50 \text{ hours}) = P(X > 50)$$

“Forgets” that it survived the last 100 hours