1 Binary Child given one binary parent

Let X and Y be two binary variables, where the conditional probability of Y given X is expressed as follows:

$$p(Y|X) = p_{y|x}^{I(X=0)I(Y=0)} p_{\bar{y}|x}^{I(X=0)I(Y=1)} p_{y|\bar{x}}^{I(X=1)I(Y=0)} p_{\bar{y}|\bar{x}}^{I(X=1)I(Y=1)}$$

This conditional probability can be expressed in different exponential forms as follows:

$$\ln p(Y|X) = \theta(X)^T s(Y) - A(X)$$

$$= \begin{pmatrix} I(X=0) \ln p_{y|x} + I(X=1) \ln p_{y|\bar{x}} \\ I(X=0) \ln p_{\bar{y}|x} + I(X=1) \ln p_{\bar{y}|\bar{x}} \end{pmatrix}^T \begin{pmatrix} I(Y=0) \\ I(Y=1) \end{pmatrix} - 0$$

$$\ln p(Y|X) = \theta(Y)^T s(X) - A(Y)
= \begin{pmatrix} I(Y=0) \ln p_{y|x} + I(Y=1) \ln p_{\bar{y}|x} \\ I(Y=0) \ln p_{y|\bar{x}} + I(Y=1) \ln p_{\bar{y}|\bar{x}} \end{pmatrix}^T \begin{pmatrix} I(X=0) \\ I(X=1) \end{pmatrix} - 0$$

$$\ln p(Y|X) = \theta^T s(Y,X) - A(\theta)$$

$$= \begin{pmatrix} \ln p_{y|x} \\ \ln p_{\bar{y}|x} \\ \ln p_{y|\bar{x}} \\ \ln p_{\bar{y}|\bar{x}} \end{pmatrix}^T \begin{pmatrix} I(Y=0)I(X=0) \\ I(Y=1)I(X=0) \\ I(Y=0)I(X=1) \\ I(Y=1)I(X=1) \end{pmatrix} - 0$$

2 Multinomial Child given a set of multinomial parents

Let us Y be a multinomial variable whose state space is $\{y_1, ..., y_k\}$. Let us also be $X_1, ..., X_p$ the set of parents of variable Y, all them also multinomial. The log-conditional probability of Y given X can be expressed as follows:

$$\ln p(Y|X_1,\ldots,X_n) =$$

Similarly the above conditional probability can be expressed in the following exponential forms:

$$\ln p(Y|X_1,...,X_p) = \theta(X_1,...,X_p)^T s(Y) - A(X_1,...,X_p)$$

$$\ln p(Y|X_1,\ldots,X_p) = \theta(Y,X_1,\ldots,X_{L-1},X_L,\ldots,X_p)^T s(X_L) - A(Y,X_1,\ldots,X_{L-1},X_L,\ldots,X_p)$$

$$\ln p(Y|X_1,\ldots,X_p) = \theta^T s(Y,X_1,\ldots,X_p) - A(\theta)$$

3 Normal Child given a set of normal parents

Let us Y be a normal variable and X_1, \ldots, X_p the set of parents of variable Y, all them also normal. The log-conditional probability of Y given X can be expressed as follows:

$$\ln p(Y|X_1,\dots,X_p) = \ln \left(\frac{1}{\sigma \sqrt{2(\beta_0 + \sum_i \beta_i X_i)}} e^{-\frac{(x - (\beta_0 + \sum_i \beta_i X_i))^2}{2\sigma^2}} \right)$$

Similarly the above conditional probability can be expressed in the following exponential forms:

$$\ln p(Y|X) = \theta(X)^T s(Y) - A(\theta(X)) + h(X)$$

$$= \left(\frac{\frac{\mu_{Y|X}}{\sigma^2}}{\frac{1}{2\sigma^2}}\right)^T \left(\frac{Y}{Y^2}\right) - \left(\frac{\mu_{Y|X}^2}{2\sigma^2} + \ln \sigma\right) + \ln \frac{1}{\sqrt{2\mu_{Y|X}}}$$

where $\mu_{Y|X} = \beta_0 + \sum_i \beta_i X_i$

$$\ln p(Y|X) = \theta(Y)^T s(X) - A(\theta(Y)) + h(X)$$

$$= \begin{pmatrix} -\frac{\beta_1^2}{2\sigma^2} \\ \cdots \\ -\frac{\beta_p^2}{2\sigma^2} \\ \frac{\beta_1(Y - \beta_0)}{\sigma^2} \\ \cdots \\ \frac{\beta_p(Y - \beta_0)}{\sigma^2} \\ -\frac{\beta_1\beta_2}{\sigma^2} \cdots \\ -\frac{\beta_1\beta_p}{\sigma^2} \cdots \\ -\frac{\beta_p-1\beta_p}{\sigma^2} \end{pmatrix} \begin{pmatrix} X_1^2 \\ \cdots \\ X_p \\ X_1X_2 \\ \cdots \\ X_1X_2 \\ \cdots \\ X_1X_p \\ \cdots \\ X_{p-1}X_p \end{pmatrix}$$

$$\ln p(Y|X) = \theta^T s(Y,X) - A(\theta) + h(X)$$

$$\begin{pmatrix} \frac{-1}{2\sigma^2} \\ -\beta_1^2 \\ 2\sigma^2 \\ \frac{-\beta_2^2}{2\sigma^2} \\ \frac{\beta_1}{\sigma^2} \\ \cdots \\ \frac{\beta_p}{\sigma^2} \\ \frac{-\beta_0\beta_1}{\sigma^2} \\ \cdots \\ \frac{-\beta_0\beta_1}{\sigma^2} \\ \cdots \\ \frac{-\beta_0\beta_p}{\sigma^2} \\ \frac{-\beta_1\beta_2}{\sigma^2} \\ \cdots \\ \frac{-\beta_1\beta_p}{\sigma^2} \\ \cdots \\ \frac{-\beta_1\beta_p}{\sigma^2} \\ \cdots \\ \frac{-\beta_1\beta_p}{\sigma^2} \\ \cdots \\ X_1X_2 \\ \cdots \\ X_1X_2 \\ \cdots \\ X_1X_2 \\ \cdots \\ X_1X_2 \\ \cdots \\ X_1X_p \\ \cdots \\ X_1X_p \\ \cdots \\ X_p \\ X_1X_p \\ \cdots \\ X_{p-1}X_p \end{pmatrix}$$