

Exercise Sheet 1

Exercise 1: Estimating the Bayes Error (10 + 10 + 10 P)

The Bayes decision rule for the two classes classification problem results in the Bayes error

$$P(\text{error}) = \int P(\text{error} \mid \mathbf{x}) p(\mathbf{x}) d\mathbf{x},$$

where $P(\text{error} \mid \mathbf{x}) = \min [P(\omega_1 \mid \mathbf{x}), P(\omega_2 \mid \mathbf{x})]$ is the probability of error for a particular input \mathbf{x} . Interestingly, while class posteriors $P(\omega_1 \mid \mathbf{x})$ and $P(\omega_2 \mid \mathbf{x})$ can often be expressed analytically and are integrable, the error function has discontinuities that prevent its analytical integration, and therefore, direct computation of the Bayes error.

- (a) Show that the full error can be upper-bounded as follows:

$$P(\text{error}) \leq \int \frac{2}{\frac{1}{P(\omega_1 \mid \mathbf{x})} + \frac{1}{P(\omega_2 \mid \mathbf{x})}} p(\mathbf{x}) d\mathbf{x}.$$

Note that the integrand is now continuous and corresponds to the harmonic mean of class posteriors weighted by $p(\mathbf{x})$.

- (b) Show using this result that for the univariate probability distributions

$$p(x \mid \omega_1) = \frac{\pi^{-1}}{1 + (x - \mu)^2} \quad \text{and} \quad p(x \mid \omega_2) = \frac{\pi^{-1}}{1 + (x + \mu)^2},$$

the Bayes error can be upper-bounded by:

$$P(\text{error}) \leq \frac{2 P(\omega_1) P(\omega_2)}{\sqrt{1 + 4\mu^2 P(\omega_1) P(\omega_2)}}$$

(Hint: you can use the identity $\int \frac{1}{ax^2 + bx + c} dx = \frac{2\pi}{\sqrt{4ac - b^2}}$ for $b^2 < 4ac$.)

- (c) Explain how you would estimate the error if there was no upper-bounds that are both tight and analytically integrable. Discuss following two cases: (1) the data is low-dimensional and (2) the data is high-dimensional.

Exercise 2: Bayes Decision Boundaries (15 + 15 P)

One might speculate that, in some cases, the generated data $p(x \mid \omega_1)$ and $p(x \mid \omega_2)$ is of no use to improve the accuracy of a classifier, in which case one should only rely on prior class probabilities $P(\omega_1)$ and $P(\omega_2)$ assumed here to be strictly positive.

For the first part of this exercise, we assume that the data for each class is generated by the univariate Laplacian probability distributions:

$$p(x \mid \omega_1) = \frac{1}{2\sigma} \exp\left(-\frac{|x - \mu|}{\sigma}\right) \quad \text{and} \quad p(x \mid \omega_2) = \frac{1}{2\sigma} \exp\left(-\frac{|x + \mu|}{\sigma}\right).$$

where $\mu, \sigma > 0$.

- (a) Determine for which values of $P(\omega_1), P(\omega_2), \mu, \sigma$ the optimal decision is to always predict the first class (i.e. under which conditions $P(\text{error} \mid x) = P(\omega_2 \mid x) \quad \forall x \in \mathbb{R}$).
- (b) Repeat the exercise for the case where the data for each class is generated by the univariate Gaussian probability distributions:

$$p(x \mid \omega_1) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right) \quad \text{and} \quad p(x \mid \omega_2) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x + \mu)^2}{2\sigma^2}\right).$$

where $\mu, \sigma > 0$.

Exercise 3: Programming (40 P)

Download the programming files on ISIS and follow the instructions.

Exercise 1: Estimating the Bayes Error (10 + 10 + 10 P)

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where $P(\text{error} | \mathbf{x}) = \min [P(\omega_1 | \mathbf{x}), P(\omega_2 | \mathbf{x})]$ is the probability of error for a particular input \mathbf{x} . Interestingly, while class posteriors $P(\omega_1 | \mathbf{x})$ and $P(\omega_2 | \mathbf{x})$ can often be expressed analytically and are integrable, the error function has discontinuities that prevent its analytical integration, and therefore, direct computation of the Bayes error.

(a) Show that the full error can be upper-bounded as follows:

$$P(\text{error}) \leq \int \frac{2}{\frac{1}{P(\omega_1 | \mathbf{x})} + \frac{1}{P(\omega_2 | \mathbf{x})}} p(\mathbf{x}) d\mathbf{x}.$$

Note that the integrand is now continuous and corresponds to the harmonic mean of class posteriors weighted by $p(\mathbf{x})$.

Solutions:

$\therefore P(\text{error} | \mathbf{x}), P(\omega_1 | \mathbf{x})$ and $P(\omega_2 | \mathbf{x})$ are all ≥ 0

\therefore we only need to show :

$$P(\text{error} | \mathbf{x}) \leq \frac{2}{\frac{1}{P(\omega_1 | \mathbf{x})} + \frac{1}{P(\omega_2 | \mathbf{x})}}$$

$$\therefore P(\text{error} | \mathbf{x}) = \min [P(\omega_1 | \mathbf{x}), P(\omega_2 | \mathbf{x})]$$

\therefore

① we first suppose that $P(\omega_1 | \mathbf{x}) \leq P(\omega_2 | \mathbf{x})$ (some part of the density function)

$$\therefore P(\text{error} | \mathbf{x}) = P(\omega_1 | \mathbf{x})$$

$$\therefore P(\omega_1 | \mathbf{x}) = \frac{2}{\frac{2}{P(\omega_1 | \mathbf{x})}} \leq \frac{2}{\frac{1}{P(\omega_1 | \mathbf{x})} + \frac{1}{P(\omega_2 | \mathbf{x})}}$$

since $\frac{1}{P(\omega_2 | \mathbf{x})} \leq \frac{1}{P(\omega_1 | \mathbf{x})}$

② We can also suppose that $P(\omega_2 | \mathbf{x}) \geq P(\omega_1 | \mathbf{x})$

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then following the same proof procedure:

we can also proof that:

$$p(\omega_2|x) \leq \frac{2}{\frac{1}{p(\omega_1|x)} + \frac{1}{p(\omega_2|x)}}$$

$$\therefore \text{Since in both 2 cases } P(\text{error}|x) \leq \frac{2}{\frac{1}{p(\omega_1|x)} + \frac{1}{p(\omega_2|x)}}$$

\therefore proofed.

(b) Show using this result that for the univariate probability distributions

$$p(x | \omega_1) = \frac{\pi^{-1}}{1 + (x - \mu)^2} \quad \text{and} \quad p(x | \omega_2) = \frac{\pi^{-1}}{1 + (x + \mu)^2},$$

the Bayes error can be upper-bounded by:

$$P(\text{error}) \leq \frac{2 P(\omega_1) P(\omega_2)}{\sqrt{1 + 4\mu^2 P(\omega_1) P(\omega_2)}}$$

(Hint: you can use the identity $\int \frac{1}{ax^2+bx+c} dx = \frac{2\pi}{\sqrt{4ac-b^2}}$ for $b^2 < 4ac$.)

Solution:

$$p(\text{error}) \leq \int \frac{2}{\frac{1}{p(\omega_1|x)} + \frac{1}{p(\omega_2|x)}} p(x) dx$$

$$= \int \frac{2 \times 1}{\left(\frac{1}{p(\omega_1|x)} + \frac{1}{p(\omega_2|x)}\right) \frac{1}{p(x)}} dx$$

$$= \int \frac{2}{\frac{1}{p(x|\omega_1)p(\omega_1)} + \frac{1}{p(x|\omega_2)p(\omega_2)}} dx$$

$$= \int \frac{1 + (x - \mu)^2}{\pi^{-1} p(\omega_1)} + \frac{1 + (x + \mu)^2}{\pi^{-1} p(\omega_2)} dx$$

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$$= \int \frac{2}{\frac{(1+(x-\mu)^2)P(w_2) + (1+(x+\mu)^2)P(w_1)}{2\pi^{-1}P(w_1)P(w_2)}} dx$$

$$= \int \frac{(P(w_1)+P(w_2)) + x^2 + \mu^2 + 2x\mu(P(w_1)-P(w_2))}{2\pi^{-1}P(w_1)P(w_2)} dx$$

$$= 2\pi^{-1}P(w_1)P(w_2) \int \frac{1}{x^2 + 2\mu(P(w_1)-P(w_2))x + \mu^2 + 1} dx$$

obviously $4\mu^2(P(w_1)-P(w_2))^2 < 4(1+\mu^2)$
 $b^2 < 4ac$

$$= 2\pi^{-1}P(w_1)P(w_2) \frac{2\pi}{\sqrt{4(1+\mu^2) - 4\mu^2(P(w_1)-P(w_2))^2}}$$

$$= \frac{2P(w_1)P(w_2)}{\sqrt{1+\mu^2 - \mu^2(P(w_1)-P(w_2))^2}}$$

since $(P(w_1)+P(w_2)) = 1$

$$= \frac{2P(w_1)P(w_2)}{\sqrt{1+\mu^2(P(w_1)+P(w_2))^2 - \mu^2(P(w_1)-P(w_2))^2}}$$

$$= \frac{2P(w_1)P(w_2)}{\sqrt{1+4\mu^2P(w_1)P(w_2)}}$$

proved.

(c) Explain how you would estimate the error if there was no upper-bounds that are both tight and analytically integrable. Discuss following two cases: (1) the data is low-dimensional and (2) the data is high-dimensional.

Solution:

(1) Low dimensional

we can accumulate the error using density diagram.

日期: $\int f(x)p(x)dx \xrightarrow{\text{Discretize}} \sum_i f(x_i)p(x_i) = P(\text{error})$

1-dim / 2-dim / 3-dim

(2) High Dimensional Data:

We can't integrate the density function in every dimension so we can take multiple samples, then count the number of all wrong classifications and calculate the percentage.

Exercise 2: Bayes Decision Boundaries (15 + 15 P)

One might speculate that, in some cases, the generated data $p(x | \omega_1)$ and $p(x | \omega_2)$ is of no use to improve the accuracy of a classifier, in which case one should only rely on prior class probabilities $P(\omega_1)$ and $P(\omega_2)$ assumed here to be strictly positive.

For the first part of this exercise, we assume that the data for each class is generated by the univariate Laplacian probability distributions:

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where $\mu, \sigma > 0$.

- (a) Determine for which values of $P(\omega_1), P(\omega_2), \mu, \sigma$ the optimal decision is to always predict the first class (i.e. under which conditions $P(\text{error} | x) = P(\omega_2 | x) \quad \forall x \in \mathbb{R}$).

Solution:

$$P(\omega_1 | x) = \frac{P(x | \omega_1) P(\omega_1)}{P(x)}$$

↓ take log, slides 10/25

$$P(\omega_2 | x) = \frac{P(x | \omega_2) P(\omega_2)}{P(x)}$$

↓ take log, slides 10/25

$$\log P(\omega_1 | x) = \log \frac{P(x | \omega_1) P(\omega_1)}{P(x)} \quad \log P(\omega_2 | x) = \log \frac{P(x | \omega_2) P(\omega_2)}{P(x)}$$

$$\downarrow \text{remove all constant value}$$
$$\log P(x | \omega_1) P(\omega_1)$$

$$\parallel$$
$$\log P(x | \omega_1) + \log P(\omega_1)$$
$$\parallel$$

$$-\frac{|x - \mu|}{\sigma} + \log P(\omega_1)$$

$$\downarrow$$
$$\log P(x | \omega_2) P(\omega_2)$$

$$\parallel$$
$$\log P(x | \omega_2) + \log P(\omega_2)$$
$$\parallel$$

$$-\frac{|x + \mu|}{\sigma} + \log P(\omega_2)$$

∴ if class 1 dominate the classifier, then we should have.

$$-\frac{|x - \mu|}{\sigma} + \log P(\omega_1) > -\frac{|x + \mu|}{\sigma} + \log P(\omega_2)$$

∴ according to the value of x we can remove the absolute sign by expanding the following cases.

① $x \geq \mu$

$$\therefore -\frac{x - \mu}{\sigma} + \log P(\omega_1) > -\frac{x + \mu}{\sigma} + \log P(\omega_2)$$

日期: $\frac{2\mu}{\sigma} > \log \frac{P(w_2)}{P(w_1)} \Rightarrow -\frac{2\mu}{\sigma} < \log \frac{P(w_1)}{P(w_2)}$

∴ the relationship of $\mu, \sigma, P(w_2), P(w_1)$ when $x \geq \mu$ is shown above.

② $-\mu < x < \mu$

$$-\frac{\mu-x}{\sigma} + \log P(w_1) > -\frac{x+\mu}{\sigma} + \log P(w_2)$$

$$\therefore \frac{2x}{\sigma} > \log \frac{P(w_2)}{P(w_1)} \rightarrow -\frac{2x}{\sigma} < \log \frac{P(w_1)}{P(w_2)}$$

↓

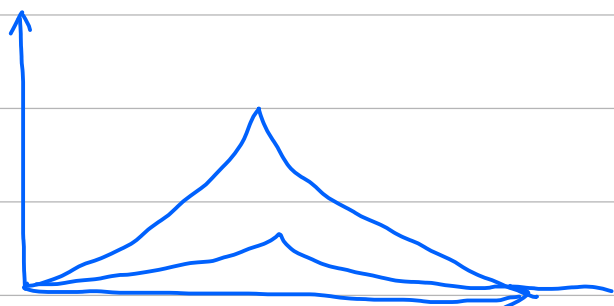
$$\frac{2\mu}{\sigma} < \log \frac{P(w_1)}{P(w_2)}$$

③ $x \leq -\mu$

$$-\frac{\mu-x}{\sigma} + \log P(w_1) > \frac{x+\mu}{\sigma} + \log P(w_2)$$

$$\therefore -\frac{2\mu}{\sigma} > \log \frac{P(w_2)}{P(w_1)} \Rightarrow \frac{2\mu}{\sigma} < \log \frac{P(w_1)}{P(w_2)}$$

$$\left\{ \begin{array}{l} -\frac{2\mu}{\sigma} < \log \left(\frac{P(w_1)}{P(w_2)} \right) \\ \frac{2\mu}{\sigma} < \log \left(\frac{P(w_1)}{P(w_2)} \right) \\ \frac{2\mu}{\sigma} < \log \left(\frac{P(w_1)}{P(w_2)} \right) \end{array} \right\} \text{ combine } \rightarrow \frac{2\mu}{\sigma} < \log \frac{P(w_1)}{P(w_2)}$$



(b) Repeat the exercise for the case where the data for each class is generated by the univariate Gaussian probability distributions:

$$p(x | \omega_1) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right) \quad \text{and} \quad p(x | \omega_2) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x+\mu)^2}{2\sigma^2}\right).$$

where $\mu, \sigma > 0$.

Solution:

$$-\frac{(x-\mu)^2}{2\sigma^2} + \log P(\omega_1) > -\frac{(x+\mu)^2}{2\sigma^2} + \log P(\omega_2)$$

\Downarrow

$$\frac{2\mu x}{2\sigma^2} + \log P(\omega_1) > -\frac{2\mu x}{2\sigma^2} + \log P(\omega_2)$$

$$\frac{2\mu x}{\sigma^2} > \log\left(\frac{P(\omega_2)}{P(\omega_1)}\right)$$

$$-x < \frac{\sigma^2}{2\mu} \log\left(\frac{P(\omega_1)}{P(\omega_2)}\right) \quad (1)$$

\therefore if we want to guarantee to take class one then we should try to make (1) hold.

