

# Mathematical Modeling and Computation in Finance

## With Exercises and Python and MATLAB Computer Codes

C.W. Oosterlee & L.A. Grzelak

Solutions to exercises from Chapter 1

<https://QuantFinanceBook.com>

Ex 1.1. The cumulative distribution function for the standard normal distribution is given by,

$$F_{\mathcal{N}_{(0,1)}}(x) = \int_{-\infty}^x f(z)dz,$$

where  $f(z) = \frac{e^{-z^2/2}}{\sqrt{2\pi}}$  is an even function. Changing the integration variable,  $y = -z$ , gives us,

$$F_{\mathcal{N}_{(0,1)}}(x) = \int_{-x}^{\infty} f(y)dy.$$

Hence

$$\begin{aligned} F_{\mathcal{N}_{(0,1)}}(x) + F_{\mathcal{N}_{(0,1)}}(-x) &= \left( \int_{-\infty}^x f(z)dz + \int_x^{\infty} f(z)dz \right) \\ &= \int_{-\infty}^{\infty} f(z)dz = 1. \end{aligned}$$

Ex. 1.3. We have,

$$\begin{aligned} \mathbb{E}[Y] &= \mathbb{E}[a + bX] = a + b\mu, \\ \text{Var}[Y] &= \mathbb{E}[(bX - b\mu)^2] = \text{Var}[bX] = (b\sigma)^2, \\ \mathbb{E}[e^X] &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^x e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{(x-(\mu+\sigma^2))^2}{2\sigma^2}} e^{\mu+\sigma^2/2} dx \\ &= e^{\mu+\sigma^2/2}. \end{aligned}$$

Ex. 1.5. a. By linearity, we find,

$$\mathbb{E}[\bar{X}] = \frac{1}{n} \sum_{k=1}^n \mathbb{E}[X_k] = \mu.$$

b.

$$\begin{aligned} \text{Var}[\bar{X}] &= \frac{1}{n^2} \mathbb{E} \left[ \sum_i \sum_j X_i X_j \right] - \mu^2 \\ &= \frac{1}{n^2} \sum_i \left( \mathbb{E}[X_i^2] + \sum_i \sum_{j \neq i} \mathbb{E}[X_i X_j] \right) - \mu^2. \end{aligned}$$

For each variable,  $\mathbb{E}[X_i^2] = \mu^2 + \sigma^2$ , and since the random variables are independent, it follows that  $\mathbb{E}[X_i X_j] = \mathbb{E}[X_i]\mathbb{E}[X_j]$ . Hence,

$$\begin{aligned}\mathbb{V}\text{ar}[\bar{X}] &= \frac{1}{n^2} \mathbb{E} \left[ \sum_i \sum_j X_i X_j \right] - \mu^2 \\ &= \frac{1}{n^2} (n(\mu^2 + \sigma^2) + n(n-1)\mu^2) - \mu^2 = \sigma^2/n.\end{aligned}$$

c. We find,

$$\begin{aligned}\sum_{k=1}^N (X_k - \bar{X})^2 &= \sum_{k=1}^N (X_k^2 - 2X_k \bar{X} + \bar{X}^2) \\ &= \sum_{k=1}^N X_k^2 - 2N\bar{X}^2 + N\bar{X}^2 \\ &= \sum_{k=1}^N X_k^2 - N\bar{X}^2.\end{aligned}$$

d. Using the result of part c),

$$\mathbb{E}[v_N^2] = \frac{1}{N-1} \left( \sum_{k=1}^N \mathbb{E}[X_k^2] - N\mathbb{E}[\bar{X}^2] \right) = \frac{1}{N-1} (N(\sigma^2 + \mu^2) - N(\sigma^2/N + \mu^2)) = \sigma^2.$$

Ex. 1.7. By using Leibniz' rule for  $zW(z)$ , we obtain,

$$d(zW(z)) = dzW(z) + z dW(z).$$

Integrating both sides gives,

$$\int_0^t W(z) dz = \int_0^t d(zW(z)) - \int_0^t z dW(z).$$

The first integral on the right-hand side can be expressed as,

$$\int_0^t d(zW(z)) = tW(t) = t \int_0^t dW(z).$$

Hence

$$\int_0^t W(z) dz = \int_0^t (t-z) dW(z).$$

Ex. 1.9. a. The SDE must not contain a drift term. The differential process of the integral equation is found to be,

$$\begin{aligned}dX(t) &= d(g(t)W(t)) - g'(t)W(t)dt \\ &= g'(t)W(t)dt - g(t)dW(t) - g'(t)W(t)dt \\ &= -g(t)dW(t),\end{aligned}$$

where the differentiation, denoted by  $d$ , is in the Itô sense. Hence, the process is a martingale.

b. Express the term  $e^{2t}W(t)$  in its integral form. as follows,

$$\begin{aligned}d(e^{2t}W(t)) &= 2e^{2t}W(t)dt + e^{2t}dW(t) \\ e^{2T}W(T) &= \int_{t=T}^{t=0} (2e^{2t}W(t)dt + e^{2t}dW(t)),\end{aligned}$$

where  $W(0) = 0$  is used in the second equation. Taking expectations at both sides and using  $\mathbb{E}[\int_{t=T}^{t=0} e^{2t} dW(t)] = 0$ , because infinitesimal increments of Brownian motion are governed by a normal distribution with zero mean, gives,

$$\mathbb{E}[e^{2T} W(T)] = \mathbb{E}\left[\int_{t=T}^{t=0} 2e^{2t} W(t) dt\right].$$

Ex. 1.11. Integrating the differential process  $dX(t)$ , we get

$$X(T) = \mu T + \sigma W(T) + x_0.$$

Substituting this into the integral, gives us,

$$\int_0^T X(t) dt = \mu T^2/2 + x_0 T + \int_0^T \sigma W(t) dt.$$

We now use the identity derived in Exercise 1.7,

$$\int_0^T X(t) dt = \mu T^2/2 + x_0 T + \sigma \int_0^T (T-t) dW(t).$$

The integral  $\int_0^T (T-t) dW(t)$  can be discretized, giving,

$$\int_0^T (T-t) dW(t) = \lim_{N \rightarrow \infty} \sum_{i=0}^N (T-t) \Delta W_i,$$

where  $\Delta W_i = W_{i+1} - W_i$ , and each of these increments follows a normal distribution. The sum of normal random variables is also a normal random variable. Its mean and variance are given by,

$$\mathbb{E}\left[\int_0^T X(t) dt\right] = \mu T^2/2 + x_0 T.$$

$$\begin{aligned} \mathbb{V}\text{ar}\left[\int_0^T X(t) dt\right] &= \mathbb{E}\left[\left(\mu T^2/2 + x_0 T + \sigma \int_0^T (T-t) dW(t)\right)^2\right] - (\mu T^2/2 + x_0 T)^2 \\ &= (\mu T^2/2 + x_0 T)^2 + \sigma(\mu T^2/2 + x_0 T) \mathbb{E}\left[\int_0^T (T-t) dW(t)\right] \\ &\quad + \mathbb{E}\left[\left(\sigma \int_0^T (T-t) dW(t)\right)^2\right]. \end{aligned}$$

The second term in the second equation vanishes. Finally, with the Itô isometry on the final integral, we find,

$$\mathbb{E}\left[\left(\sigma \int_0^T (T-t) dW(t)\right)^2\right] = \sigma^2 \int_0^T \mathbb{E}[(T-t)^2] dt = \frac{1}{3} \sigma^2 T^3.$$

Therefore,

$$\int_0^T X(t) dt \sim \mathcal{N}\left(x_0 T + \frac{1}{2} \mu T^2, \frac{1}{3} \sigma^2 T^3\right).$$