

MATH20812: PRACTICAL STATISTICS I

SEMESTER 2

NOTES ON CHI-SQUARE GOODNESS OF FIT TESTS

Chi-Square Test of Goodness of Fit

Consider a population with a characteristic taking values $1, 2, \dots, k$. Let P_i denote the probability that a randomly chosen observation has characteristic i . For a random sample of size n , let N_i denote the number that has characteristic i .

Characteristic	Probability	Count
1	p_1	N_1
2	p_2	N_2
3	p_3	N_3
\vdots	\vdots	\vdots
k	p_k	N_k
Totals	1	n

We wish to test the null hypothesis

$$H_0 : P_1 = p_1, P_2 = p_2, \dots, P_k = p_k \quad (1)$$

versus

$$H_1 : P_i \neq p_i \text{ for some } i. \quad (2)$$

The rule to reject H_0 is:

$$t = \sum_{i=1}^k \frac{(N_i - np_i)^2}{np_i} > \chi_{k-1, \alpha}^2. \quad (3)$$

This rule is based on an approximation and will be good if $np_i \geq 1$ for each i and at least 80% of the values of np_i exceed 5. The corresponding p-value is $\Pr(\chi_{k-1}^2 \geq t)$.

Test of Independence

Consider a population with each observation classified according to two distinct characteristics X and Y – characteristic X taking r levels and characteristic Y taking s levels. Let $P_{ij} = \Pr(X = i, Y = j)$ denote the probability that a random chosen observation takes the i th level of characteristic X and j th level of characteristic Y . For a random sample of size n , let N_{ij} denote the number in the sample that take the i th level of characteristic X and j th level of characteristic Y .

Level of X	Level of Y				Totals
	1	2	\dots	s	
1	N_{11}	N_{12}	\dots	N_{1s}	$N_{1\cdot}$
2	N_{21}	N_{22}	\dots	N_{2s}	$N_{2\cdot}$
\vdots	\dots	\dots	\dots	\dots	\dots
\vdots	\dots	\dots	\dots	\dots	\dots
r	N_{r1}	N_{r2}	\dots	N_{rs}	$N_{r\cdot}$
Totals	$N_{\cdot 1}$	$N_{\cdot 2}$	\dots	$N_{\cdot s}$	$N_{\cdot\cdot} = n$

We wish to test the hypothesis that X and Y are independent, i.e.

$$H_0 : p_{ij} = p_{i\cdot} p_{\cdot j} \quad (4)$$

versus

$$H_1 : p_{ij} \neq p_{i\cdot} p_{\cdot j} \text{ for some } i \text{ and } j, \quad (5)$$

where

$$p_{i\cdot} = \sum_{j=1}^s p_{ij} = \Pr(X = i) \quad (6)$$

and

$$p_{\cdot j} = \sum_{i=1}^r p_{ij} = \Pr(Y = j). \quad (7)$$

The rule to reject H_0 is:

$$t = \sum_{i=1}^r \sum_{j=1}^s \frac{(N_{ij} - N_{i\cdot} N_{\cdot j} / n)^2}{N_{i\cdot} N_{\cdot j} / n} > \chi_{(r-1)(s-1), \alpha}^2. \quad (8)$$

The corresponding p-value is $\Pr(\chi_{(r-1)(s-1)}^2 \geq t)$.

Measures of Dependence

If X and Y turn out to be dependent then two measures of dependence are:

$$C = \sqrt{\frac{t}{n + t}} \quad (\text{Contingency Coefficient}) \quad (9)$$

and

$$V = \sqrt{\frac{t}{n \min(r-1, s-1)}} \quad (\text{Cramer's } V). \quad (10)$$

Test of Homogeneity

Consider independent random samples of size $N_{.1}, N_{.2}, \dots, N_{.s}$ from s different populations. The observations from each sample are classified according to r levels of a characteristic X . Let $P_{ij} = \Pr(X = i \mid Y = j)$ denote the probability that a randomly chosen observation from the j th population takes the i th level of characteristic X . Let N_{ij} denote the number in the j th sample taking the i th level of characteristic X .

Level of X	Population				Total
	1	2	\dots	s	
1	N_{11}	N_{12}	\dots	N_{1s}	$N_{1.}$
2	N_{21}	N_{22}	\dots	N_{2s}	$N_{2.}$
\dots	\dots	\dots	\dots	\dots	\dots
\dots	\dots	\dots	\dots	\dots	\dots
r	N_{r1}	N_{r2}	\dots	N_{rs}	$N_{r.}$
Sample Size	$N_{.1}$	$N_{.2}$	\dots	$N_{.s}$	$N_{..} = n$

We wish to test the hypothesis of homogeneity, i.e.

$$H_0 : p_{ij} \text{ depends only on } i \quad (11)$$

versus

$$H_1 : p_{ij} \text{ depends on } i \text{ and } j \quad (12)$$

The rejection rule and the p-value are the same as for the test of independence.

Appendix: Likelihood Ratio Statistic

Consider a sample x_1, x_2, \dots, x_n with joint probability density (mass) function $f(x_1, x_2, \dots, x_n)$, parameterized by Θ (denoting a vector of parameters). Suppose we wish to test the hypothesis $H_0 : \Theta \in \Theta_0$ versus $H_0 : \Theta \notin \Theta_0$. A useful result in statistics is the likelihood ratio test. It goes like this. Let L_1 denote the maximum of $f(x_1, x_2, \dots, x_n)$ over all possible values of Θ , whether H_0 is true or not. Let L_2 denote the maximum of $f(x_1, x_2, \dots, x_n)$ over all $\Theta \in \Theta_0$. Take the ratio $\lambda = 2 \log(L_1/L_2)$, which is known as the likelihood ratio statistic. For n large enough and when H_0 is true this ratio has the chi-square distribution with degrees of freedom equal to the number of *free* parameters in Θ minus the number of free parameters in Θ_0 . Hence we can reject H_0 if the value of λ exceeds the table chi-square value.