Chapter 3. Panel Threshold Regression Models School of Economics and Management - University of Geneva

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In econometrics, threshold regression models are a category of regime-switching models in which

- The slope parameters vary according to a "regime" switching mecanism that depends on a threshold variable,
- The regime is observable ex-post, contrary to the Markovian regime switching models.

Definition (threshold regression model)

A typical threshold regression model is given by

$$y_t = \alpha + \beta_0' x_t + \beta_1' x_t h(q_t; \theta) + \varepsilon_t$$

where β_0 and β_1 are $K \times 1$ vectors, q_t is a **threshold variable**, θ a vector of parameters and $h\left(q_t;\theta\right)$ a **transition function**.

Example (threshold regression model)

If the transition function is a binary function such that

$$h\left(q_{t};c
ight) = \left\{egin{array}{l} 1 ext{ if } q_{t} \geq c \ 0 ext{ if } q_{t} < c \end{array}
ight.,$$

then the model is simply defined by

$$y_t = \alpha + \beta_0' x_t + \beta_1' x_t \mathbb{I}_{(q_t \ge c)} + \varepsilon_t$$

where $\mathbb{I}_{(.)}$ is the indicator function and c is a location parameter.

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Example (threshold regression model)

If the transition function is a binary function, we have a **regime-switching mechanism** for the slope parameters that depends on the threshold variable and a location parameter \boldsymbol{c}

$$y_t = \left\{ egin{array}{l} lpha + \left(eta_0' + eta_1'
ight) x_t + arepsilon_t ext{ if } q_t \geq c \ lpha + eta_0' x_t + arepsilon_t ext{ if } q_t < c \end{array}
ight.,$$

In this simple case, we have two "regimes" for the slope parameters, i.e. β_0 and $\beta_0+\beta_1.$

Remarks

- The transition function $h(q_t; \theta)$ may be smoother than a binary function.
- ullet In general, the transition function $h\left(q_t; heta
 ight)$ is assumed to verified

$$0 \leq h\left(q_t;\theta\right) \leq 1$$

Sometimes, we have

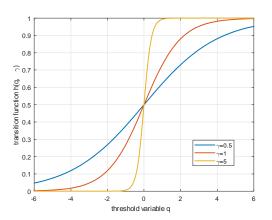
$$\underset{q_{t}\rightarrow+\infty}{\lim}h\left(q_{t};\theta\right)=1\quad\underset{q_{t}\rightarrow-\infty}{\lim}h\left(q_{t};\theta\right)=0$$

Example (Logistic transition function)

A logistic transition function is defined as:

$$h\left(q_{t};\gamma,c
ight)=rac{1}{1+\exp\left(-\gamma\left(q_{t}-c
ight)
ight)},\ \ \gamma>0$$

Figure: Logistic transition function with c = 0



Panel data

- The threshold regressions in panel data models allow to model the heterogeneity of the slope parameters.
- These models give a parametric approach of the heterogeneity which is associated to an economic "story" (interpretation).

The outline of this chapter is the following:

Section 1: Introduction

Section 2: Panel Threshold Regression (PTR) Model

Section 3: Panel Smooth Threshold Regression (PSTR) Model

Section 2

The Panel Threshold Regression (PTR) Model

Objectives

- Introduce the panel threshold regression (PTR) model.
- Understand the link with heterogeneous panel models.
- Understand the link with time varying parameters panel models.
- Understand the differences with random coefficient models.

- The Panel Threshold Regression (PTR) model has been introduced by Hansen (1999).
- In this paper, threshold regression methods are developed for non-dynamic panels with individual fixed effects.
- Hansen, B. E. 1999. Threshold effects in non-dynamic panels: estimation, testing, and inference, *Journal of Econometrics*, 93, 334-368

Definition (panel threshold regression model)

The panel threshold regression (PTR) model is defined as

$$y_{it} = \alpha_i + \beta_1' x_{it} \mathbb{I}_{(q_{it} \le c)} + \beta_2' x_{it} \mathbb{I}_{(q_{it} > c)} + \varepsilon_{it}$$

where the dependent variable y_{it} is scalar, α_i is a fixed effect, the threshold variable q_{it} is scalar, the regressor x_{it} is a k vector, and $\mathbb{I}_{(.)}$ is the indicator function and c is a threshold parameter.

Assumptions

- The threshold variable is exogeneous or at least predetermined $(q_{it} = y_{i,t-d} \text{ with } d \ge 1)$.
- For the identification of β_1 and β_3 , it is required that the elements of x_{it} are not time invariant.
- The threshold variable q_{it} is not time invariant.
- The error ε_{it} is assumed to be i.i.d. with $\mathbb{E}\left(\varepsilon_{it}\right)=0$ and $\mathbb{V}\left(\varepsilon_{it}\right)=\sigma_{\varepsilon}^{2}$.

Panel Threshold Regression model

An alternative specification of the PTR is

$$y_{it} = \alpha_i + \beta'_0 x_{it} + \beta'_3 x_{it} \mathbb{I}_{(q_{it} \le c)} + \varepsilon_{it}$$

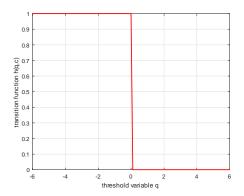
or equivalently

$$y_{it} = \alpha_i + \beta'_0 x_{it} + \beta'_3 x_{it} h(q_{it}; c) + \varepsilon_{it}$$

whith $h\left(q_{it};c
ight)=\mathbb{I}_{\left(q_{it}\leq c
ight)}$ the binary transition function

$$y_{it} = \begin{cases} \alpha_i + (\beta'_0 + \beta'_3) x_{it} + \varepsilon_{it} & \text{if } q_{it} \leq c \\ \alpha_i + \beta'_3 x_{it} + \varepsilon_{it} & \text{if } q_{it} > c \end{cases},$$

Figure: Transition function $h(q_{it}; c)$ with c = 0



Definition (heterogeneous panel data model)

The PTR model can be viewed as an **heterogeneous** and **time-varying parameters** panel data model

$$y_{it} = \alpha_i + \beta'_{it} x_{it} + \varepsilon_{it} = \begin{cases} \alpha_i + \beta'_1 x_{it} + \varepsilon_{it} & \text{if } q_{it} \leq c \\ \alpha_i + \beta'_2 x_{it} + \varepsilon_{it} & \text{if } q_{it} > c \end{cases}$$

where the marginal effect (slope parameters) satisfy

$$rac{\partial y_{it}}{\partial x_{it}} = eta_{it} = \left\{ egin{array}{l} eta_1 & ext{if } q_{it} \leq c \ eta_2 & ext{if } q_{it} > c \end{array}
ight.,$$

Heterogeneous panel data model

$$y_{it} = \begin{cases} \alpha_i + \beta'_1 x_{it} + \varepsilon_{it} & \text{if } q_{it} \leq c \\ \alpha_i + \beta'_2 x_{it} + \varepsilon_{it} & \text{if } q_{it} > c \end{cases},$$

At a given time t, two cross-section units i and j may have two different slope parameters

$$rac{\partial y_{it}}{\partial x_{it}} = eta_1
eq rac{\partial y_{jt}}{\partial x_{jt}} = eta_2 \quad ext{if } q_{it} \leq c ext{ and } q_{jt} > c.$$

Time-varying parameter panel data model

$$y_{it} = \begin{cases} \alpha_i + \beta'_1 x_{it} + \varepsilon_{it} & \text{if } q_{it} \leq c \\ \alpha_i + \beta'_2 x_{it} + \varepsilon_{it} & \text{if } q_{it} > c \end{cases},$$

A given cross-section unit i may have different slope parameters at different dates t and s

$$rac{\partial y_{it}}{\partial x_{it}} = eta_1
eq rac{\partial y_{is}}{\partial x_{is}} = eta_2 \quad ext{if } q_{it} \leq c ext{ and } q_{is} > c.$$

Example (Marginal effect and PTR)

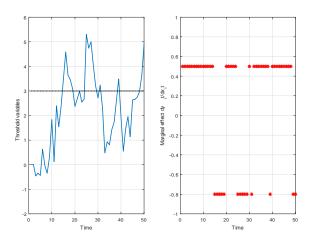
Consider a PTR model with K=1 regressor such that

$$y_{it} = \alpha_i + 0.5 x_{it} \mathbb{I}_{(q_{it} \le 3)} - 0.8 x_{it} \mathbb{I}_{(q_{it} > 3)} + \varepsilon_{it}$$

Then, the marginal effect of the regressor x_{it} on y_{it} is equal to

$$rac{\partial y_{it}}{\partial x_{it}} = 0.5 \mathbb{I}_{(q_{it} \le 3)} - 0.8 \mathbb{I}_{(q_{it} > 3)} = \left\{ egin{array}{l} 0.5 & \text{if } q_{it} \le 3 \\ -0.8 & \text{if } q_{it} > 3 \end{array}
ight.,$$

Figure: Marginal effect with K=1, $\beta_1=0.5$, $\beta_2=-0.8$ and c=3



Panel Threshold Regression model

The PTR model is an alternative to the random coefficient model

Panel Threshold Regression $eta_{it} = f\left(q_{it};c
ight)$ eta_{it} is a constant term

Economic interpretation (through q_{it})

Random Coefficient Model β_i i.i.d. $(\overline{\beta}, \Delta)$ β_i is a random variable

No economic interpretation

Estimation by NLS

The PTR model can be rewritten as:

$$y_{it} = \alpha_i + \beta' x_{it} (c) + \varepsilon_{it}$$

$$x_{it}\left(c\right) = \left(\begin{array}{c} x_{it} \mathbb{I}_{\left(q_{it} \leq c\right)} \\ x_{it} \mathbb{I}_{\left(q_{it} > c\right)} \end{array}\right) \qquad \beta = \left(\begin{array}{c} \beta_1 \\ \beta_2 \end{array}\right)$$

Estimation by NLS

In order to eliminate the individual effects, we apply the Within transformation

$$y_{it}^{*} = \beta' x_{it}^{*} (c) + \varepsilon_{it}^{*}$$

$$y_{it}^{*} = y_{it} - \overline{y}_{i} \quad \overline{y}_{i} = T^{-1} \sum_{t=1}^{T} y_{it}$$

$$x_{it}^{*} (c) = x_{it} (c) - \overline{x}_{i} (c) \quad \overline{x}_{i} (c) = T^{-1} \sum_{t=1}^{T} x_{it} (c)$$

$$\varepsilon_{it}^{*} = \varepsilon_{it} - \overline{\varepsilon}_{i} \quad \overline{\varepsilon}_{i} = T^{-1} \sum_{t=1}^{T} \varepsilon_{it}$$

Let us define

$$y_{i}^{*} = \begin{pmatrix} y_{i,1}^{*} \\ y_{i,2}^{*} \\ \cdots \\ y_{it}^{*} \end{pmatrix} \quad X_{i}^{*}\left(c\right) = \begin{pmatrix} x_{1,i,1}\left(c\right)' \\ x_{1,i,2}\left(c\right)' \\ \cdots \\ x_{1,it}\left(c\right)' \end{pmatrix} \quad \varepsilon_{i}^{*} = \begin{pmatrix} \varepsilon_{i,1}^{*} \\ \varepsilon_{i,2}^{*} \\ \cdots \\ \varepsilon_{it}^{*} \end{pmatrix}$$

$$Y^* = \begin{pmatrix} y_1^* \\ y_2^* \\ \dots \\ y_n^* \end{pmatrix} \qquad X^* (c) = \begin{pmatrix} X_1^* (c) \\ X_2^* (c) \\ \dots \\ X_n^* (c) \end{pmatrix} \qquad \varepsilon^* \\ (T_{n,1}) = \begin{pmatrix} \varepsilon_1^* \\ \varepsilon_2^* \\ \dots \\ \varepsilon_n^* \end{pmatrix}$$

Definition (given threshold)

For any **given threshold** c, the slope coeffcient β can be estimated by ordinary least squares (OLS).

$$\widehat{\beta}(c) = \left(X^*(c)'X^*(c)\right)^{-1}X^*(c)'Y$$

The vector or residuals is given by

$$\widehat{\varepsilon}^{*}(c) = Y^{*} - X^{*}(c)\widehat{\beta}(c)$$

and the sum of squared errors is

$$SSR(c) = \widehat{\varepsilon}^*(c)'\widehat{\varepsilon}^*(c)$$

Definition (Threshold estimation)

The estimation if the threshold parameter c is obtained by minimization of the concentrated sum of squared

$$\widehat{c} = \operatorname*{arg\,min}_{c \in \Theta} \mathit{SSR}\left(c\right)$$

Remarks

- It is undesirable for a threshold c to be selected which sorts too few observations into one or the other regime.
- This possibility can be excluded by restricting the search in to values of c such that a minimal percentage of the observations (say, 1% or 5%) lie in each regime.

$$c \in \Theta = \left[\mathsf{quantile}\left(\{q_{it}\}_{i=1,t=1}^{\textit{N},\textit{T}}, 0.05\right), \mathsf{quantile}\left(\{q_{it}\}_{i=1,t=1}^{\textit{N},\textit{T}}, 0.95\right)\right]$$

ullet Given \widehat{c} , we can compute the estimates for eta as

$$\widehat{eta} = eta\left(\widehat{c}
ight) = \left(egin{array}{c} eta_1\left(\widehat{c}
ight) \ eta_2\left(\widehat{c}
ight) \end{array}
ight)$$



Boosting idea dissemination and research reproducibility



```
PURPOSE: Hansen (1999) "Threshold effect in non dynamic panels: Estimation, Testing, and Inference",
Journal of Econometrics 93, 345-368
Usage: Estimate a two regimes Threshold Panel, and test the effect of threshold
Function: Hansen 2reg(Y,Q,X,T,min reg,CL,simulation,affi)
Where:
          Y: the dependant variable, size: vector (N*T,1)
          O: the transition variable, size: vector (N*T.1)
          X : the regressor, size : matrix (N*T,K)
     : Time Dimension
min req : Trimming parameter, if is empty : minreq = 0.05
     : confidence level for gamma, if is empty : CL= 0.95
simulation: number of simulation for test the threshold effect, if is empty: simulation = 0
affi = 1 if the user want to display the results
RETURNS:
```

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RETURNS:
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res.gam : estimated threshold parameter gamma;
res.IC gam : confidence intervals of the threshold
```

res.effets fixes : fixed effects ;

res.coef cent : estimated slope parameters (in each regimes);

res.coer_cent . escimated slope parameters (in each regimes);

res.coef_stat : give the coefficients, the statistic and a second statistic robust to heteroskedastic res.residus : estimated residuals:

res.rss : sum of squared resid;

res.F1 : F1 Statistic of the test H0: no threshold against H1 one threshold (only if simulate

res.pvalue : pvalue of F1 statistic (only if simulaton>0)

```
==== Model with one threshold ===
_____
Cross-unit dimension N = 76
Time dimension T = 35
Estimated threshold parameter (gamma) = -2.7731
Confidence Interval (95%) on gamma = [-3.1245, -2.7278]
Estimated slope parameters (first column) and t-stats (second column) for the FIRST regime : q(t)<=-2.773
  0.5067
          30.8901
  0.1641 10.1197
  0.0139 0.7037
Estimated slope parameters (first column) and t-stats (second column) for the SECOND regime : q(t)>-2.773
  0.5572 40.3585
  0.1046 5.8240
  0.0955 9.2412
```

```
SSR = 44.4338

Estimated residual variance = 0.0172

Number of simulation (for F1 statistic pvalue) = 100

F1 test statistic = 112.1812

pvalue = 0.0000
```

Definition (extension to r + 1 regimes)

A general specification of the PTR can be proposed with r+1>2 regimes or r threshold parameters $c_1, ... c_r$.

$$y_{it} = \alpha_i + \sum_{j=1}^r \beta'_j x_{it} \mathbb{I}_{(c_{j-1} < q_{it} \le c_j)} + \varepsilon_{it}$$

with $c_0 = -\infty$ and $c_{r+1} = +\infty$.

Example (PTR with three regimes)

For instance, a three-regimes model (r=3) with 2 threshold parameters c_1 and c_2 is given by

$$y_{it} = \alpha_{i} + \beta'_{1} x_{it} \mathbb{I}_{(q_{it} \leq c_{1})} + \beta'_{2} x_{it} \mathbb{I}_{(c_{1} < q_{it} \leq c_{2})}$$
$$+ \beta'_{3} x_{it} \mathbb{I}_{(q_{it} > c_{2})} + \varepsilon_{it}$$

Example (PTR with four regimes)

For instance, a three-regimes model (r=4) with 3 threshold parameters c_1, c_2 and c_3 is given by

$$y_{it} = \alpha_i + \beta_1' x_{it} \mathbb{I}_{(q_{it} \le c_1)} + \beta_2' x_{it} \mathbb{I}_{(c_1 < q_{it} \le c_2)}$$
$$+ \beta_3' x_{it} \mathbb{I}_{(c_2 < q_{it} \le c_3)} + \beta_4' x_{it} \mathbb{I}_{(q_{it} > c_3)} + \varepsilon_{it}$$

Inference

 If one comes to test whether the threshold effect is statically significant in the model with two regimes, the null hypothesis is

$$H_0: \beta_1 = \beta_2$$

- This null hypothesis corresponds to the hypothesis of no threshold effect.
- Under H_0 the model is then equivalent to a linear model.

$$y_{it} = \alpha_i + \beta' x_{it} + \varepsilon_{it}$$

with
$$\beta = \beta_1 = \beta_2$$
.



Inference

- ullet The null hypothesis $H_0:eta_1=eta_2$ can be tested by a standard test.
- If we note S_0 the sum of squared of the linear model, the approximate likelihood ratio test of H_0 is based on:

$$F_1 = \frac{SSR_0 - SRR_1(\widehat{c})}{\widehat{\sigma}^2}$$

where $\hat{\sigma}^2$ denotes a convergent estimate of σ^2 .

Inference

- The main problem is that under the null, the threshold parameter c is not identified (nuisance parameter).
- Consequently, the asymptotic distribution of F_1 is not standard and, in particular, does not correspond to a chi-squared distribution.
- This issue has been largely studied in literature devoted to threshold models, notably since the seminal paper by Davies (1977, 1987).
- One solution is to use a bootstrap procedure to determine the asymptotic distribution of the statistic F_1 .

Number of regimes (thresholds)

- The same kind of inference procedure can be applied in order to determining the number of thresholds.
- A likelihood ratio test of one threshold versus two thresholds is based on the statistic

$$F_{2} = \frac{SSR_{1}\left(\widehat{c}\right) - SSR_{2}\left(\widehat{c}_{1}, \widehat{c}_{2}\right)}{\widehat{\sigma}^{2}}$$

where \hat{c}_1 and \hat{c}_2 denote the threshold estimates of the model with three regimes , and SSR_2 (\hat{c}_1, \hat{c}_2) denotes the corresponding residual sum of squares.

Number of regimes (thresholds)

- The hypothesis of one threshold is rejected in favor of two thresholds if F_2 is larger than its critical value.
- The corresponding asymptotic p-value can be approximated by bootstrap simulations (Hansen, 1999).
- If the model with two thresholds (three regimes) is not rejected accepted, we to test the hypothesis of two thresholds (three regimes) against the alternative of three thresholds (four regimes).

Number of regimes (thresholds)

ullet The corresponding likelihood ratio statistic, denoted $F_{3.}$, is defined as:

$$F_{3} = \frac{SSR_{2}\left(\widehat{c}_{1},\widehat{c}_{2}\right) - SSR_{3}\left(\widehat{c}_{1},\widehat{c}_{2},\widehat{c}_{3}\right)}{\widehat{\sigma}^{2}}$$

where $SSR_3(\hat{c}_1, \hat{c}_2, \hat{c}_3)$ denotes the residual sum of squares of the model with four regimes and three threshold parameters.

• Thus, a sequential procedure based on F_1 , F_2 , F_3 , etc. allows to determining the number of regimes

Example

Candelon, Colletaz, Hurlin (2012) investigate the threshold effects in the productivity of infrastructure investment in developing countries.

$$y_{it} = \begin{cases} a_i + \alpha_1 k_{it} + \beta_1 h_{it} + \gamma_1 x_{it} + \varepsilon_{it} & \text{if } q_{it} \leq \lambda \\ a_i + \alpha_2 k_{it} + \beta_2 h_{it} + \gamma_2 x_{it} + \varepsilon_{it} & \text{if } q_{it} > \lambda. \end{cases}$$

where y_{it} is the aggregate added value, k_{it} is physical capital, h_{it} is human capital, x_{it} is infrastructure stock.



Candelon B., Colletaz G., Hurlin C. (2013), Network Effects and Infrastructure Productivity in Developing Countries, *Oxford Bulletin of Economics and Statistics*, 75(6), 887-913.

Table 3. Tests for Threshold Effects: Model A, $q_{it} = x_{it}^{12}$

	Roads	Electricity	Telephones	Railways
T				
Test for single threshold				
F_1	90.2	112.1	59.5	128.4
P-value	0.00	0.00	0.00	0.00
1% Critical Values	13.7	14.3	14.3	14.9
5% Critical Values	15.3	15.5	16.8	16.5
10% Critical Values	19.3	20.6	22.0	21.3
Test for double threshold				
F_2	84.7	74.3	82.6	122.6
P-value	0.00	0.00	0.00	0.00
1% Critical Values	50.4	19.8	17.2	30.7
5% Critical Values	55.2	22.7	19.2	36.9
10% Critical Values	68.2	26.2	25.1	42.2
Test for triple threshold				
F_3	41.7	43.8	55.7	85.9
P-value	0.00	0.00	0.00	0.00
1% Critical Values	13.6	13.6	13.6	13.2
5% Critical Values	15.2	15.6	15.7	15.5
10% Critical Values	18.0	21.2	19.2	18.4

Notes: P-values and critical values are computed from 300 simulations. F1 denotes the Fisher type statistic associated to the test of the null of no threshold against one threshold. F2 corresponds to the test one threshold against two thresholds and F3 corresponds to the test of two thresholds against three thresholds.

Table 4. Four Regimes Panel Models¹³. Model A: $q_{it} = x_{it}$

	Roads	Electricity	Telephones	Railways
Regime 1: $q_{it} \leq \lambda_1$				
Physical Capital per Worker	0.575 (32.30)	0.519 (26.78)	0.431 (21.43)	0.613 (43.70)
Human Capital per Worker	0.053 (2.67)	0.145 (8.57)	0.102	0.565 (15.62)
Infrastructure per Worker	$0.122 \ (3.967)$	0.022 (1.08)	-0.099 (-2.36)	$0.184 \ (6.54)$
Regime 2: $\lambda_1 < q_{it} \le \lambda_2$				
Physical Capital per Worker	0.403 (12.34)	0.754 (22.06)	0.389 (18.48)	0.597 (42.53)
Human Capital per Worker	0.077 (1.89)	0.185 (5.75)	0.108 (6.18)	0.400 (14.25)
Infrastructure per Worker	-1.682 (-5.67)	$0.655 \ (7.68)$	0.220 (14.44)	-0.096 (-1.68)
Regime 3: $\lambda_2 < q_{it} \leq \lambda_3$				
Physical Capital per Worker	0.567 (34.93)	0.584 (38.39)	0.465 (14.51)	0.625 (42.21)
Human Capital per Worker	0.208 (10.62)	0.067 (3.39)	-0.168 (-4.17)	0.228 (9.48)
Infrastructure per Worker	$0.166 \ (8.79)$	0.135 (12.09)	0.132	-0.157 (-6.06)

Regime 4: $q_{it} > \lambda_3$				
Physical Capital per Worker	0.574 (32.28)	0.548 (34.94)	0.380 (14.03)	0.526 (28.97)
Human Capital per Worker	0.104 (2.53)	0.252 (7.36)	0.126 (3.27)	0.716 (12.35)
Infrastructure per Worker	$0.129 \atop (5.03)$	$0.041 \ (2.50)$	0.210 (11.13)	-0.014 (-0.54)
Threshold Estimates				
First Threshold $\widehat{\lambda}_1$	-0.917	-3.317	1.337	-1.427
Second Threshold $\widehat{\lambda}_2$	-0.642	-2.954	4.314	-0.882
Third Threshold $\widehat{\lambda}_3$	1.248	-0.519	4.808	0.975
Residual Sum of Squares	12.11	42.01	23.27	26.61

Notes: The t-statistics in parenthesis are computed with an estimator of the covariance matrix robust to heteroskedasticity. The confidence intervals for the threshold parameters are not reported. See Appendix A1, for the confidence intervals in a model with three regimes.



Key Concepts Section 2

- Panel Threshold Regression Model
- 4 Heterogeneous and time-varying parameters
- Non Linear Least Squares
- Inference for the number of regimes
- Davies problem

Section 2

The Panel Smooth Threshold Regression (PSTR) Model

Objectives

- Introduce the panel smooth threshold regression (PSTR) model.
- Understand the link with heterogeneous panel models.
- Understand the link with time varying parameters panel models.

- The Panel Smooth Threshold Regression (PSTR) model has been introduced by Gonzalez, Teräsvirta and van Dijk (2005).
- This model can be viewed as a generalization of the PTR model.



Definition (PSTR model)

The PSTR with two extreme regimes can be defined as

$$y_{it} = \alpha_i + \beta'_0 x_{it} + \beta'_1 x_{it} g(q_{it}; \gamma, c) + \varepsilon_{it}$$

where $g\left(q_{it};\gamma,c\right)$ is a transition function, q_{it} a threshold variable, c a location parameter and γ a slope parameter.

Transition function

- The transition function $g\left(q_{it};\gamma,c\right)$ is a continuous function of the observed variable q_{it}
- The transition function is normalized to be bounded between 0 and 1.

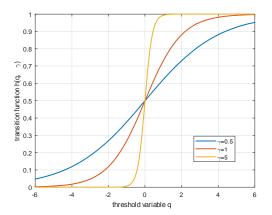
$$0 \le g(q_{it}; \gamma, c) \le 1$$

Definition

Gonzalez, Teräsvirta and van Dijk (2005) consider a logistic transition function

$$g\left(q_{it};\gamma,c
ight)=rac{1}{1+\exp\left(-\gamma\left(q_{it}-c
ight)
ight)},\ \ \gamma>0$$

Figure: Logistic transition function with c = 0



Remark 1

If $\gamma \to \infty$, the logistic function tends to an indicator function and the PSTR corresponds to a **PTR model**

$$\lim_{\gamma \to \infty} g\left(q_{it}; \gamma, c\right) = \mathbb{I}_{(q_{it} \ge c)} = \begin{cases} 1 \text{ if } q_{it} \ge c \\ 0 \text{ if } q_{it} < c \end{cases},$$

$$y_{it} = \alpha_i + \beta_0' x_{it} + \beta_1' x_{it} \mathbb{I}_{(q_{it} \ge c)} + \varepsilon_{it}$$

Remark 2

If $\gamma \to 0$, the logistic function tends to an indicator function and the PSTR corresponds to a **linear panel model**

$$\lim_{\gamma
ightarrow \infty} g\left(q_{it}; \gamma, c
ight) = rac{1}{2} \hspace{0.5cm} orall q_{it}$$

$$y_{it} = \alpha_i + \beta' x_{it} + \varepsilon_{it}$$
 $\beta = \frac{\beta_0 + \beta_1}{2}$

Definition (heterogeneous panel data model)

The PSTR model can be viewed as an **heterogeneous** and **time-varying parameters** panel data mode since the marginal effect (slope parameters) satisfy

$$\frac{\partial y_{it}}{\partial x_{it}} = \beta_{it} = \beta_0 + \beta_1 g(q_{it}; \gamma, c)$$

as soon as q_{it} does not belong to x_{it} , with by convention

$$\beta_0 \le \beta_{it} \le \beta_0 + \beta_1$$

Remark 3

Even there are two "extreme" regimes/values for the slope parameters, in fact there is an **infinity of regimes** (possible values) for the slope parameters β_{it} given the observed values for q_{it} :

$$eta_{it} = eta_0 + eta_1 g\left(q_{it}; \gamma, c
ight)$$
 $eta_0 \leq eta_{it} \leq eta_0 + eta_1$

Definition (generalization to r+1 regimes)

The PSTR model can be generalised to r+1 extreme regimes as follows:

$$y_{it} = \alpha_i + \beta'_0 x_{it} + \sum_{j=1}^r \beta'_j x_{it} \ g_j(q_{it}; \gamma_j, c_j) + \varepsilon_{it}$$

where the r transition functions $g_j(q_{it}; \gamma_j, c_j)$ depend on the slope parameters γ_j and on the location parameters c_j .

Marginal effects

If the transition variable does not belong to the set of regressors, the slope parameters become:

$$eta_{it} = rac{\delta y_{it}}{\delta x_{it}} = eta_0 + \sum_{i=1}^r eta_j \ g_j(q_{it}; \gamma_j, c_j)$$

Marginal effects

If the transition variable does not belong to the set of regressors, the slope parameters become:

$$\beta_{it} = \frac{\delta y_{it}}{\delta x_{it}} = \beta_0 + \sum_{j=1}^r \beta_j \ g_j(q_{it}; \gamma_j, c_j) + \sum_{j=1}^r \beta_j \ \frac{\delta g_j(q_{it}; \gamma_j, c_j)}{\delta q_{it}} q_{it}$$

Estimation

- The estimation of the parameters of the PSTR model consists of eliminating the individual effects α_i and then in applying NLS to the transformed model
- See Gonzàlez et al. (2005) or Colletaz and Hurlin (2006), for more details.
- González, A., Teräsvirta, T., van Dijk, D., 2005. Panel smooth transition regression model. Working Paper Series in Economics and Finance, vol. 604.
- Colletaz, G., Hurlin, C., 2006. Threshold effects in the public capital productivity: an international panel smooth transition approach. *Document de Recherche du Laboratoire d'Economie d'Orléans*. 2006-1.

Inference

- Gonzàlez et al. (2005) propose a testing procedure in order (i) to test the linearity against the PSTR model and (ii) to determine the number, r, of transition functions, i.e. the number of extreme regimes which is equal to r+1.
- Testing the linearity in a PSTR model can be done by testing

$$H_0: \gamma = 0$$
 or $H_0: \beta_0 = \beta_1$

• But in both cases, the test will be non standard since under H_0 the PSTR model contains unidentified nuisance parameters.

Inference

• A possible solution is to replace the transition function $g_j(q_{it}, \gamma_j, c_j)$ by its **first-order Taylor expansion** around $\gamma = 0$ and to test an equivalent hypothesis in the auxiliary regression

$$y_{it} = \alpha_i + \theta_0 x_{it} + \theta_1 x_{it} q_{it} + \varepsilon_{it}$$

- The parameters θ_i are proportional to the slope parameter γ of the transition function.
- Thus, testing the linearity of the model against the PSTR simply consists of testing

$$H_0: \theta_1 = 0$$

Definition (linearity/homogeneity test)

The F-statistic for the homogeneity assumption H_0 : $\theta_1=0$ is then defined by:

$$LM_F = (SSR_0 - SSR_1)/[SSR_0/(TN - N - 1)]$$

with SSR_0 the panel sum of squared residuals under H_0 (linear panel model with individual effects) and SSR_1 the panel sum of squared residuals under H_1 (PSTR model with two regimes). Under the null hypothesis, the F-statistic has an approximate F(1, TN-N-1) distribution.

Choice of number of transitions

- The logic is similar when it comes to testing the number of transition functions in the model or equivalently the number of extreme regimes.
- We use a sequential approach by testing the null hypothesis of no remaining nonlinearity in the transition function.

Choice of number of transitions

- We want to test whether there is one transition function $(H_0: r=1)$ or whether there are at least two transition functions $(H_0: r=1)$.
- Let us assume that the model with r=2 is defined as:

$$y_{it} = \alpha_i + \beta_0 x_{it} + \beta_1 x_{it} g_1(q_{it}; \gamma_1, c_1) + \beta_2 x_{it} g_2(q_{it}; \gamma_2, c_2) + \varepsilon_{it}$$

• The logic of the test consists in replacing the second transition function by its first-order Taylor expansion around $\gamma_2=0$ and then in testing linear constraints on the parameters in

$$y_{it} = \alpha_i + \beta_0 x_{it} + \beta_1 x_{it} g_1(q_{it}; \gamma_1, c_1) + \theta_1 x_{it} q_{it} + \epsilon_{it}^*$$

and the test of no remaining nonlinearity is simply defined by H_0 : $\theta_1 = 0$.

Choice of number of transitions

• Let us denote SSR_0 the panel sum of squared residuals under H_0 , i.e. in a PSTR model with one transition function.

$$y_{it} = \alpha_i + \beta_0 x_{it} + \beta_1 x_{it} g_1(q_{it}; \gamma_1, c_1) + \varepsilon_{it}$$

 Let us denote SSR₁ the sum of squared residuals of the transformed model.

$$y_{it} = \alpha_i + \beta_0 x_{it} + \beta_1 x_{it} g_1(q_{it}; \gamma_1, c_1) + \theta_1 x_{it} q_{it} + \epsilon_{it}^*$$

• The F-statistic LM_f can be calculated in the same way by adjusting the number of degrees of freedom.

Testing procedure

- Given a PSTR model with $r = r^*$, we test the null $H_0: r = r^*$ against $H_1: r = r^* + 1$.
- ② If H_0 is not rejected the procedure ends.
- **3** Otherwise, the null hypothesis H_0 : $r = r^* + 1$ is tested against H_1 : $r = r^* + 2$.
- **①** The testing procedure continues until the first acceptance of H_0 .
- Given the sequential aspect of this testing procedure, at each step of the procedure the significance level must be reduced by a constant factor $0<\rho<1$ in order to avoid excessively large models. Gonzàlez et al. (2005) suggest $\rho=0.5$.



PURPOSE: Gonzalez, Terascirta and Van Dijk (2004), "Panel Smooth Transition Regression Model and An Application to Investment Under Credit Constraint", Stockholm School of Economics

```
Usage: Estimate a Panel Smooth Threshold Panel

Function: res=STAR_Panel(Y,Q,X,N0,m0,rmax,condini_user)

Where:

Y : the dependant variable, size, vector (N*T,1)
Q : the transition variable, size, vector (N*T,1)
X : the regressor, size, matrix (N*T,K)
N : number of individuals
m : number of location parameters
rmax : maximum number of transition functions authorised (r<=rmax)
condini_user : initial conditions given by the user
```

RETURNS:

```
res.balanced
                   : the message 'Balanced Sample' indicates a balanced sample
                   : cross-section dimension
res.N
                   : time dimension. For unbalanced panel, give the T dimension for each c
res.T
res.m
                   : number of location parameters for each transition (fixed by the user)
                   : Optimal number of transition function
res.r
                   : Estimated slope parameter (for each transition function)
res.gam
res.c
                   : Estimated location parameter (for each transition function)
res.beta
                   : Estimated slope parameters for the explicative variables (by column f
res.beta std
                  : Standard errors (corrected for heteroskedasticity, by column for each
res.beta tstat
                   : t-statistics (by column for each transition)
res.beta std nc:
                   : Standard errors (not corrected for heteroskedasticity, by column for
res.rss
                   : RSS
res.fixed
                   : Estimated fixed effects
res.resid
                  : Residuals
                  : Values of the transition function (N*T,r) matrix, where r is the numbe
res.q
res.exitflag
                   : Equal to 1 if there is convergence
res.output
                   : Informations on the convergence
                   : Individual coefficients for each explicative variable (NT,K) matrix
res.coef indi
res.nbparam
                   : Number of parameters
res.ATC
                   : Akaike criteria
                  : BIC criteria
res.BTC
```

```
*******************

*** LINEARITY Tests ***

*******************

HO: Linear Model H1: PSTR model with at least one Threshold Variable (r=1)

Wald Tests (LM): W = 84.104 pvalue = 0.000

Fisher Tests (LMF): F = 45.876 pvalue = 0.000

LRT Tests (LRT): LRT = 89.161 pvalue = 0.000
```

TESTING THE NUMBER OF REGIMES: TESTS OF NO REMAINING NON-LINEARITY ***

```
Initial Conditions : Assumed Number of Thresholds r = 1    Number of Regressions = 270
Initial Conditions on (c,gamma)
    5.0000   -1.4357

Estimation of the Model with r = 1 and m = 1 : Convergence = 1    RSS = 3.982
RSS under H1 = 3.956

H0: PSTR with r = 1 against H1: PSTR with at least r = 2
    Wald Tests (LM):    W = 4.872    pvalue = 0.088
    Fisher Tests (LMF):    F = 2.364    pvalue = 0.095
    LRT Tests (LRT):    LRT = 4.888    pvalue = 0.087
Given the choices of rmax = 2 and m = 1, the OPTIMAL (LMF criterion) NUMBER OF THRESHOLD FUNCTIONS is r = 1
```

```
*** FINAL ESTIMATION OF PSTR MODEL ***
***************
 Final Estimation of the Model with r = 1 and m = 1 by NLS ***
 Initial Conditions on (gamma,c) :
  4.8632 -1.4664
RSS = 3.982 Convergence = 1
ATC = -5.221 BTC = -5.184
Estimated slope parameter of the transition function (one for for each transition function)
  4.8632
Estimated location parameters (per column for each transition function)
 -1.4664
```

```
Estimated location parameters (per column for each transition function)
-1.4664

Estimated slope parameters (per column for each transition function)
0.3460 -0.0974
0.4430 -0.3796

Standard Errors of estimated slope parameters corrected fo heteroskedasticity (per column f
0.0234 0.0199
0.0352 0.0342

t-statistics based on corrected standard errors (per column for each transition function)
14.8021 -4.9051
12.5958 -11.1053

OUPUT: Individual Elasticities for each explicative variable (first column is the cross sec
```

Example

Hurlin, Rabaud and Fouquau (2008) considers a PSTR model for determining the relative influence of five factors on the Feldstein and Horioka result for OECD countries.

$$I_{it} = \alpha_i + \beta_0 S_{it} + \beta_1 S_{it} g(q_{it}; c) + \epsilon_{it}$$

They consider five main factors (as potential threshold variable) generally considered in this literature: (i) economic growth, (ii) demography, (iii) degree of openness, (iv) country size and (v) current account balance.



Hurlin C., Rabaud I., and Fouquau J. (2008) The Feldstein-Horioka Puzzle: a PSTR Approach. *Economic Modelling*, 25, 284-299.

Table 2 LM_E tests for remaining nonlinearity

Model Threshold variable	Model A	Model B	Model C
	Growth	Openness	Size
H_0 : $r=0$ vs. H_1 : $r=1$	73.08(0.00)	293.1(0.00)	6.54(0.01)
H_0 : $r=1$ vs. H_1 : $r=2$	0.637(0.42)	0.071(0.79)	0.48(0.49)
H_0 : $r=2$ vs. H_1 : $r=3$	-	-	_
Model	Model D	Model E	Model F
Threshold variable	pop<15 years	pop>64 years	Cur. account
H_0 : $r=0$ vs. H_1 : $r=1$	186.5(0.00)	211.1(0.00)	141.3(0.00)
H_0 : r=1 vs. H_1 : r=2	3.21(0.07)	0.075(0.78)	0.006(0.94)
H_0 : $r=2$ vs. H_1 : $r=3$	_ ` `	_ ` `	_ ` `

Notes: For each model (i.e. for each threshold variable), the testing procedure works as follows. First, test a linear model (r=0) against a model with one threshold (r=1). If the null hypothesis is rejected, test the single threshold model against a double threshold model (r=2). The procedure goes on until the hypothesis no additional threshold is not rejected. The corresponding LM_F statistic has an asymptotic F[1,TN-N-(r+1)] distribution under H_0 . The corresponding P_0 -values are reported in parentheses.

Table 5 Parameter estimates for the final PSTR models

Specification Threshold variable	Model A Growth	Model B Openness	Model C Size	Model D pop<15 years	Model E pop>64 years	Model F Cur. account
Parameter β_0	0.487(0.04)	0.885(0.02)	0.292(0.05)	0.378(0.03)	0.600(0.04)	0.893(0.08)
Parameter β_1	0.126(0.01)	-0.678(0.03)	0.492(0.05)	0.215(0.01)	-0.371(0.02)	-0.676(0.07)
Location parameters c	2.74(1.81)	87.2(16.1)	0.52(0.04)	20.7(0.48)	14.2(0.848)	-2.63(3.34)
Slopes parameters y	0.774(1.70)	0.037(0.01)	73.6(116)	0.579(1.38)	0.547(0.04)	0.109(0.10)
AIC criterion	2.102	1.907	2.068	1.957	1.939	1.700
Schwarz criterion	2.123	1.927	2.068	1.978	1.959	1.730
Number of obs.	936	960	960	960	960	639

Notes: The standard errors for coefficients in parentheses are corrected for heteroskedasticity. The standard errors for the smooth parameter γ and the threshold parameter c are computed by evaluating the Hessian matrix except for two transition variables: the share of under 15s in total population, and the growth. In this two cases, we use the outer product of gradients. For each model, the number of transition functions r is determined by a sequential testing procedure (see Table 1). For each transition function, the estimated location parameters c and the corresponding estimated slope parameter γ are reported. The P\$TR parameters can not be directly interpreted as elasticities.

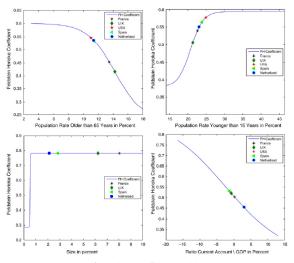


Fig. 2. Estimated FH coefficients PSTR models.

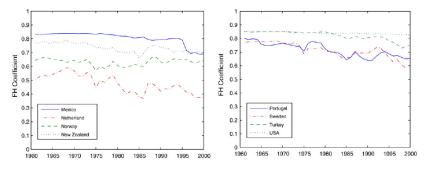


Fig. 3. Estimated FH individual coefficients: PSTR Model B.

Key Concepts Section 3

- Panel Smooth Threshold Regression Model
- Marginal effects (slope parameters)
- Non Linear Least Squares
- Inference for the number of regimes

End of Chapter 3

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