Linear Models for Statistical Natural Language Processing

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Introduction

This is a collection of notes that I use for teaching Georgia Tech Computer Science 4650 and 7650, "Natural Language." The notes focus on what I view as a core subset of the field of natural language processing, unified by the concept of linear models. This includes approaches to document classification, word sense disambiguation, sequence labeling (part-of-speech tagging and named entity recognition), parsing, coreference resolution, relation extraction, discourse analysis, and, to a limited degree, language modeling and machine translation. The theme was inspired by Fernando Pereira's EMNLP 2008 keynote, "Are linear models right for language." The notes are heavily influenced by several other good resources (e.g., Manning and Schütze, 1999; Jurafsky and Martin, 2009; Figueiredo et al., 2013; Collins, 2013), but for various reasons I wanted to create something of my own.

¹You can see a version of this talk — not the one I saw — online at vimeo.com/30676245

Notation

w_n	word token at position n
\boldsymbol{x}_i	a vector of feature counts for instance <i>i</i> , often word counts
N	number of training instances
V	number of words in vocabulary
$oldsymbol{ heta}$	a vector of weights
y_i	the label for instance i
$oldsymbol{y}$	vector of labels across all instances
\mathcal{Y}	set of all possible labels
K	number of possible labels $K = \# \mathcal{Y} $
$oldsymbol{f}(oldsymbol{x}_i, y_i)$	feature vector for instance i with label y_i
P(A)	probability function of event A
$p_B(b)$	the marginal probability of random variable B taking value b
M	length of a sequence (of words or tags)
$\mathcal{T}(oldsymbol{w})$	the set of possible tag sequences for the word sequence w
\Diamond	the start symbol
	the stop symbol
λ	the amount of regularization

Linear classification and features

Suppose you want to build a spam detector. Spam vs. Ham. How would you do it, using only the text in the email?

One solution is to represent document i as a column vector of word counts: $x_i = [0 \ 1 \ 1 \ 0 \ 0 \ 2 \ 0 \ 1 \ 13 \ 0 \dots]^\mathsf{T}$, where $x_{i,j}$ is the count of word j in document i. Suppose the size of the vocabulary is V, so that the length of x_i is also V.

We've thrown out grammar, sentence boundaries, paragraphs — everything but the words! But this could still work. If you see the word *free*, is it spam or ham? How about *c1al1s*? How about *Bayesian*? One approach would be to define a "spamminess" score for every word in the dictionary, and then just add them up. This is also a commonly-used approach to sentiment analysis, where each word is scored as one of $\{1,0,-1\}$, with 1 indicating positive sentiment and -1 indicating negative sentiment.

These scores are called **weights**, written θ , and we'll spend a lot of time later talking about where they come from. But for now, let's generalize: suppose we want to build a multi-way classifier to distinguish stories about sports, celebrities, music, and business. Each label is an element y_i in a set of K possible labels \mathcal{Y} . Then for any pair $\langle \mathbf{x}_i, y_i \rangle$, we can define a *feature vector* $\mathbf{f}(\mathbf{x}_i, y_i)$, such that:

$$\boldsymbol{f}(\boldsymbol{x}, y = 0) = [\boldsymbol{x}_i^\mathsf{T} \ \boldsymbol{0}_{V(K-1)}^\mathsf{T}]^\mathsf{T}$$
(3.1)

$$\boldsymbol{f}(\boldsymbol{x}, y = 1) = [\boldsymbol{0}_{V}^{\mathsf{T}} \, \boldsymbol{x}_{i}^{\mathsf{T}} \, \boldsymbol{0}_{V(K-2)}^{\mathsf{T}}]^{\mathsf{T}}$$
(3.2)

$$\boldsymbol{f}(\boldsymbol{x}, y = 2) = [\mathbf{0}_{2V}^{\mathsf{T}} \, \boldsymbol{x}_i^{\mathsf{T}} \, \mathbf{0}_{V(K-3)}^{\mathsf{T}}]^{\mathsf{T}}$$
(3.3)

$$\dots$$
 (3.4)

$$\boldsymbol{f}(\boldsymbol{x},K) = [\mathbf{0}_{V(K-1)}^{\mathsf{T}} \ \boldsymbol{x}_i^{\mathsf{T}}]^{\mathsf{T}},\tag{3.5}$$

where $\mathbf{0}_{VK}$ is a column vector of VK zeros. Often we'll add an **offset** feature at

the end of x, which is always 1; we then have to also add an extra zero to each of the zero vectors. This gives the entire feature vector f(x, y) a length of (V + 1)K.

Now, given a vector of weights, $\theta \in \mathcal{R}^{(V+1)K}$, we can compute the inner product $\theta^T f(x, y)$. Then for any document x_i , we can predict a label \hat{y} as

$$\hat{y} = \arg\max_{y} \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}_{i}, y)$$
 (3.6)

We could just set the weights by hand. If we wanted to distinguish, say, English from Spanish, we could just use English and Spanish dictionaries, and set each weight to 1. For example,

$$\begin{array}{ll} \theta_{\rm english, bicycle} = 1 & \theta_{\rm spanish, bicycle} = 0 \\ \theta_{\rm english, bicicleta} = 0 & \theta_{\rm spanish, bicicleta} = 1 \\ \theta_{\rm english, con} = 1 & \theta_{\rm spanish, con} = 1 \\ \theta_{\rm english, ordinateur} = 0 & \theta_{\rm spanish, ordinateur} = 0 \end{array}$$

Similarly, if we want to distinguish positive and negative sentiment, we could use positive and negative *sentiment lexicons*, which are defined by expert psychologists (Tausczik and Pennebaker, 2010). You'll try this in Project 1.

But it's usually not easy to set the weights by hand. Instead, we will learn them from data. For example, suppose that an email user has manually labeled thousands of messages as "spam" or "not spam"; or a newspaper may label its own articles as "business" or "fashion." Such **instance labels** are a typical form of labeled data that we will encounter in NLP. In **supervised machine learning**, we use instance labels to automatically set the weights for a classifier. An important tool for this is probability.

3.1 Review of basic probability

This section is inspired/borrowed from Manning and Schütze (1999).

- **Formally**: When we write $P(\dot)$, this denotes a function $P: \mathcal{F} \to [0,1]$ from an **event space** \mathcal{F} to a **probability**. A probability is a real number between zero and one, with zero representing impossibility and one representing certainty.
- The probabilities of disjoint event sets are additive: $A_i \cap A_j = \emptyset \Rightarrow P(A_i \cup A_j) = P(A_i) + P(A_j)$. This is a restatement of the Third Axiom of probability.

- For example, you might ask what is the probability of two heads on three coin flips. There are eight possible series of three flips HHH, HHT, ..., and each is an equally likely event. Of these events, three meet the criterion, HHT, HTH, THH. So the probability is $\frac{3}{8}$.
- More generally, $P(A_i \cup A_j) = P(A_i) + P(A_j) P(A_i \cap A_j)$. This can be derived from the third axiom.

$$P(A_i \cup A_j) = P(A_i) + P(A_j - (A_i \cap A_j))$$
(3.7)

$$P(A_{i}) = P(A_{i} - (A_{i} \cap A_{i})) + P(A_{i} \cap A_{i})$$
(3.8)

$$P(A_{i} - (A_{i} \cap A_{i})) = P(A_{i}) - P(A_{i} \cap A_{i})$$
(3.9)

$$P(A_i \cup A_j) = P(A_i) + P(A_j) - P(A_i \cap A_j)$$
(3.10)

• If the probability $P(A \cap B) = P(A)P(B)$, then the events A and B are *independent*, written $A \perp B$.

Conditional probability and Bayes' Rule

A conditional probability is an expression like $P(A \mid B)$, where we are interested in the probability of A conditioned on B happening.

- Conditional probability: $P(A \mid B) = P(A \cap B)/P(B)$
- If $P(A \cap B \mid C) = P(A \mid C)P(B \mid C)$, then the events A and B are **conditionally independent**, written $A \perp B \mid C$.
- Chain rule: $P(A \cap B) = P(A \mid B)P(B)$, which is just a rearrangement of terms.
- We can apply the chain rule multiple times:

$$P(A \cap B \cap C) = P(A \mid B \cap C)P(B \cap C)$$
$$= P(A \mid B \cap C)P(B \mid C)P(C)$$

We'll do this a lot later in the course.

• Bayes' rule follows from the Chain rule: $P(A \mid B) = P(A \cap B)/P(B) = P(B \mid A)P(A)/P(B)$

Often we want the maximum a posteriori (MAP) estimate

$$\begin{split} \hat{B} &= \arg\max_{B} P(B \mid A) \\ &= \arg\max_{B} P(A \mid B)P(B)/P(A) \\ &\propto \arg\max_{B} P(A \mid B)P(B) \end{split}$$

- We don't need to normalize the probability because P(A) is the same for all values of B.
- If we do need to compute the conditional $P(A \mid B)$, we can compute P(A) by summing over $P(A \cap B) + P(A \cap \overline{B})$, where $B \cap \overline{B} = \emptyset$ and $B \cup \overline{B} = \Omega$, the entire sample space (such that $P(\Omega) = 1$).
- More generally, if $\bigcup_i B_i = \Omega$ and $\forall_{i,j}, B_i \cap B_j = \emptyset$, then $P(A) = \sum_i P(A \mid B_i) P(B_i)$.

Example Manning and Schütze (1999) have a nice example of Bayes Rule (Bayes Law) in a linguistic setting.

- Suppose one is interested in a rare syntactic construction, perhaps parasitic gaps, which occurs on average once in 100,000 sentences.
 - (An example of a sentence with a parasitic gap is *Which class did you attend* __ without registering for __? -JE)
- Lana Linguist has developed a complicated pattern matcher that attempts to identify sentences with parasitic gaps. Its pretty good, but it's not perfect:
 - If a sentence has a parasitic gap, it will say so with probability 0.95 (this is the **recall** -JE).
 - If it doesn't, it will wrongly say it does with probability 0.005 (this is the **false positive rate**, the additive inverse of **precision** -JE).
- Suppose the test says that a sentence contains a parasitic gap. What is the probability that this is true?
- (This example is usually framed in terms of tests for rare diseases. -JE)
 - (c) Jacob Eisenstein 2014-2015. Work in progress.

Solution: Let *G* be the event of a sentence having a parasitic gap, and *T* be the event of the test being positive.

$$P(G \mid T) = \frac{P(G \mid T)P(T)}{P(G \mid T)P(T) + P(G \mid \overline{T})P(\overline{T})}$$

$$= \frac{0.95 \times 0.00001}{0.95 \times 0.00001 + 0.005 \times 0.99999} \approx 0.002$$
(3.11)

$$= \frac{0.95 \times 0.00001}{0.95 \times 0.00001 + 0.005 \times 0.99999} \approx 0.002$$
 (3.12)

Random variables

A random variable takes on a specific value in \mathbb{R}^n , typically with n=1, but not always. Discrete random variables can take values only in some countable subset of \mathbb{R} .

- Recall the coin flip example. The number of heads, H, can be viewed as a discrete random variable, $H \in [0, 1, 2, 3]$.
- The probability mass associated with each number is $\{\frac{1}{8}, \frac{3}{8}, \frac{3}{8}, \frac{1}{8}\}$.
- This set of numbers represents the **probability distribution** over H, written P(H = h) = p(h).
- To indicate that the RV H is distributed as p(h), we write $H \sim p(h)$.
- The function p(h) is called a probability mass function (pmf) if h is discrete, and a probability **density** function (pdf) if *h* is continuous.
- If we have more than one variable, we can write a joint probability p(a, b) =P(A=a,B=b).
- We can write a **marginal** probability $p_A(a) = \sum_b p(a, b)$.
- Random variables are independent iff $p_A(a,b) = p_A(a)p_B(b)$.
- We can write a conditional probability as $p(a \mid b) = \frac{p(a,b)}{p_R(b)}$.
 - (c) Jacob Eisenstein 2014-2015. Work in progress.

Expectations

Sometimes we want the **expectation** of a function, such as $E[g(x)] = \sum_{x \in \mathcal{X}} g(x)p(x)$. Expectations are easiest to think about in terms of probability distributions over discrete events:

- If it is sunny, Marcia will eat three ice creams.
- If it is rainy, she will eat only one ice cream.
- There's a 80% chance it will be sunny.
- The expected number of ice creams she will eat is $0.8 \times 3 + 0.2 \times 1 = 2.6$.

If the random variable *X* is continuous, the sum becomes an integral:

$$E[g(x)] = \int_{\mathcal{X}} g(x)p(x)dx \tag{3.13}$$

For example, a fast food restaurant in Quebec gives a 1% discount on french fries for every degree below zero. Assuming they used a thermometer with infinite precision, the expected price would be an integral over all possible temperatures.

3.2 Naïve Bayes

Back to classification! A Naïve Bayes classifier chooses the weights θ to maximize the *joint* probability of a labeled dataset, $p(x_{1:N}, y_{1:N})$, where $\langle x_i, y_i \rangle$ is a labeled instance.

We first need to define the probability p(x, y). We'll do that through a "generative model," which describes a hypothesized stochastic process that has generated the observed data.¹

- For each document i,
 - draw the label $y_i \sim \text{Categorical}(\mu)$
 - draw the vector of counts $x_i \sim \text{Multinomial}(\phi_{y_i})$,

¹We'll see a lot of different generative models in this course. They are a helpful tool because they clearly and explicitly define the assumptions that underly the form of the probability distribution.

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The first thing this generative model tells us is that we can treat each document independently: the probability of the whole dataset is equal to the product of the probabilities of each individual document. The observed word counts and document labels are independent and identically distributed (IID).

$$p(\boldsymbol{x}, \boldsymbol{y}; \mu, \phi) = \prod_{i} p(\boldsymbol{x}_i, y_i; \mu, \phi)$$
(3.14)

This means that the words in each document are **conditionally independent** given the parameters μ and ϕ .

When we write $y_i \sim \text{Categorical}(\mu)$, that means y_i is a stochastic draw from a categorical distribution with **parameter** μ . A categorical distribution is just like a weighted die: $p_{\text{cat}}(y;\mu) = \mu_y$, where μ_y is the probability of the outcome Y = y. We require $\sum_y \mu_y = 1$ and $\forall_y, \mu_y \geq 0$.

A multinomial distribution is only slightly more complex:

$$p_{\text{mult}}(\boldsymbol{x};\phi) = \frac{\left(\sum_{j} x_{j}\right)!}{\prod_{j} x_{j}!} \prod_{j} \phi_{j}^{x_{j}}$$
(3.15)

We again require that $\sum_j \phi_j = 1$ and $\forall j, \phi_j \geq 0$. The first part of the equation doesn't depend on ϕ , and can usually be ignored. Can you see why we need the first part at all?²

We can write $p(x_i | y_i; \phi)$ to indicate the conditional probability of word counts x_i given label y_i , with parameter ϕ , which is equal to $p_{\text{mult}}(x_i; \phi_{y_i})$.

By specifying the multinomial distribution, we are working with *multinomial naïve Bayes* (MNB). Why "naïve"? Because the multinomial distribution treats each word token independently: the probability mass function factorizes across the counts.³ We'll see this more clearly later, when we show how MNB is an example of linear classification.

Another version of Nave Bayes

Consider a slight modification to the generative story of NB:

²Technically, a multinomial distribution requires a second parameter, the total number of counts (the number of words in the document). Even more technically, that number should be treated as a random variable, and drawn from some other distribution. But none of that matters for classification.

³You can plug in any probability distribution to the generative story and it will still be naïve Bayes, as long as you are making the "naïve" assumption that your features are generated independently.

- For each document *i*
 - Draw the label $y_i \sim \text{Categorical}(\mu)$
 - For each word $n \leq D_i$
 - * Draw the word $w_{i,n} \sim \text{Categorical}(\phi_{y_i})$

This is not quite the same model as multinomial Naive Bayes (MNB): it's a product of categorical distributions over words, instead of a multinomial distribution over word counts. This means we would generate the words in order, like $p_W(multinomial)p_W(Naive)p_W(Bayes)$. Formally, this is a model for the joint probability p(w, y), not p(x, y).

However, as a classifier, it is identical to MNB. The final probabilities are reduced by a factor corresponding to the normalization term in the multinomial, $\frac{(\sum_j x_j)!}{\prod_j x_j!}$. This means that the resulting probabilities for a given x are different. However, none of this has anything to do with the label y or the parameters ϕ . The ratio of probabilities between any two labels y_1 and y_2 will be identical, as will the maximum likelihood estimates for the parameters μ and ϕ (defined later).

Prediction

The Naive Bayes prediction rule is to choose the label y which maximizes $p(x, y; \phi, \mu)$:

$$\begin{split} \hat{y} = & \arg\max_{y} \mathbf{p}(\boldsymbol{x}, y; \mu, \phi) \\ = & \arg\max_{y} \mathbf{p}(\boldsymbol{x} \mid y; \phi) \mathbf{p}(y; \mu) \\ = & \arg\max_{y} \log \mathbf{p}(\boldsymbol{x} \mid y; \phi) + \log \mathbf{p}(y; \mu) \end{split}$$

Converting to logarithms makes the notation easier. It doesn't change the prediction rule because the log function is monotonically increasing.

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Now we can plug in the probability distributions from the generative story.

$$\begin{split} \log \mathbf{p}(\boldsymbol{x}, y; \mu, \phi) = & \arg \max_{y} \log \mathbf{p}(\boldsymbol{x} \mid y; \phi) + \log \mathbf{p}(y; \mu) \\ = & \log \left[\frac{\left(\sum_{j} x_{j}\right)!}{\prod_{j} x_{j}!} \prod_{j} \phi_{y, j}^{x_{j}} \right] + \log \mu_{y} \\ = & \log \frac{\left(\sum_{j} x_{j}\right)!}{\prod_{j} x_{j}!} + \sum_{j} x_{j} \log \phi_{y, j} + \log \mu_{y} \\ \propto & \sum_{j} x_{j} \log \phi_{y, j} + \log \mu_{y} \\ = & \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}, y), \end{split}$$

where

$$oldsymbol{ heta} = {[oldsymbol{ heta^{(1)}}^\mathsf{T}, oldsymbol{ heta^{(2)}}^\mathsf{T}, \dots, oldsymbol{ heta^{(K)}}^\mathsf{T}]^\mathsf{T}} \ oldsymbol{ heta^{(y)}} = [\log \phi_{y,1} \, \log \phi_{y,2} \, \dots \, \log \phi_{y,M} \log \mu_y]^\mathsf{T}$$

and f(x, y) is a vector of word counts and an offset, padded by zeros for the labels not equal to y (see equations 3.1-3.5). This ensures that the inner product $\theta^T f(x, y)$ only activates the features in $\theta^{(y)}$, which are what we need to compute the joint log-probability $\log p(x, y)$ for each y.

Estimation

The parameters of a multinomial distribution have a simple interpretation: they're the expected frequency for each word. Based on this interpretation, it's tempting to set the parameters empirically, as

$$\phi_{y,j} = \frac{\sum_{i:Y_i = y} x_{i,j}}{\sum_{j'} \sum_{i:Y_i = y} x_{i,j'}} = \frac{\text{count}(y,j)}{\sum_{j'} \text{count}(y,j')}$$
(3.16)

In NLP this is called a *relative frequency estimator*. It can be justified more rigorously as a *maximum likelihood estimate*.

As in prediction, we want to maximize the joint likelihood of the data,

$$L = \sum_{i} \log p_{\text{mult}}(\boldsymbol{x}_i; \boldsymbol{\phi}_{y_i}) + \log p_{\text{cat}}(y_i; \mu)$$
(3.17)

Since p(y) is unrelated to ϕ , we can forget about it for now. But before we can just optimize L, we have to deal with a constraint:

$$\sum_{j} \phi_{y,j} = 1 \tag{3.18}$$

We'll do this by adding a Lagrange multiplier. Here's the resulting Lagrangian:

$$\ell[\phi_y] = \sum_{i:Y_i = y} \sum_j x_{ij} \log \phi_{y,j} + \lambda(\sum_j \phi_{y,j} - 1)$$
(3.19)

We solve by setting $\frac{\partial \ell}{\phi_i} = 0$.

$$0 = \sum_{i:Y_i=y} x_{i,j}/\phi_{y,j} - \lambda$$

$$\lambda \phi_{y,j} = \sum_{i:Y_i=y} x_{i,j}$$

$$\phi_{y,j} \propto \sum_{i:Y_i=y} x_{i,j} = \sum_i \delta(Y_i = y)x_{i,j}$$

$$= \frac{\sum_{i:Y_i=y} x_{i,j}}{\sum_{j'} \sum_{i:Y_i=y} x_{i,j'}}$$

Similarly, $\mu_y \propto \sum_i \delta(Y_i = y)$, where $\delta(Y_i = y) = 1$ if $Y_i = y$ and 0 otherwise.

Smoothing and MAP estimation

If data is sparse, you can end up with values of $\phi = 0$, allowing a single feature to completely veto a label. This is undesirable, because it imposes high **variance**: depending on what data happens to be in the training set, we could get vastly different classification rules.

One solution is Laplace smoothing: adding "pseudo-counts" of α to each estimate, and then normalize.

$$\phi_{y,j} = \frac{\alpha + \sum_{i:Y_i = y} x_{i,j}}{\sum_{j'} \alpha + \sum_{i:Y_i = y} x_{i,j'}} = \frac{\alpha + \operatorname{count}(i,j)}{V\alpha + \sum_{j'} \operatorname{count}(i,j')}$$
(3.20)

Laplace smoothing has a nice Bayesian justification, in which we extend the generative story to include ϕ as a random variable (rather than as a parameter). The resulting estimate is called *maximum a posteriori*, or MAP.

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Smoothing reduces **variance**, but it takes us away from the maximum-likelihood estimate: it imposes a **bias** (towards uniform probabilities). Machine learning theory shows that errors on held out data result from the sum of bias and variance. Techniques for reducing variance typically increase the bias, so there is a **biasvariance tradeoff**.

- Unbiased classifiers **overfit** the training data, yielding poor performance on unseen data.
- But if we set a very large smoothing value, we can **underfit** instead. In the limit of $\alpha \to \infty$, we have zero variance: it is the same classifier no matter what data we see! But the bias of such a classifier will be high.
- Navigating this tradeoff is hard. But in general, as you have more data, variance is less of a problem, so you just go for low bias.

Training, testing, and tuning (development) sets

We'll soon talk about more learning algorithms, but whichever one we apply, we will want to report its accuracy. Really, this is an educated guess about how well the algorithm will do on new data in the future.

To do this, we need to hold out a separate "test set" from the data that we use for estimation (i.e., training, learning). Otherwise, if we measure accuracy on the same data that is used for estimation, we will badly overestimate the accuracy we're likely to get on new data. See http://xkcd.com/1122/ for a cartoon related to this idea.

Many learning algorithms also have "tuning" parameters:

- the smoothing pseudo-counts α in Naive Bayes
- the regularization λ in logistic regression
- the slack weight C in the support-vector machine

All of these tuning parameters really do the same thing: they navigate the biasvariance tradeoff. Where is the best position on this tradeoff curve? It's hard to tell in advance. Sometimes it is tempting to see which tuning parameter gives the best performance on the test set, and then report that performance. Resist this temptation! It will also lead to overestimating accuracy on truly unseen future

data. For that reason, this is a sure way to get your research paper rejected. Instead, you should split off a piece of your training data, called a "development set" (or "tuning set").

Sometimes, people average across multiple test sets and/or multiple development sets. One way to do this is to divide your data into "folds," and allow each fold to be the development set one time. This is called **K-fold cross-validation**. In the extreme, each fold is a single data point. This is called **leave-one-out**.

The Naivety of Naive Bayes

Naive Bayes is very simple to work with. Estimation and prediction can done in closed form, and the nice probabilistic interpretation makes it relatively easy to extend the model in various ways.

But Naive Bayes makes assumptions which seriously limit its accuracy, especially in NLP.

• The multinomial distribution assumes that each word is generated independently of all the others (conditioned on the parameter ϕ_y). Formally, we assume conditional independence:

$$p(na\"{i}ve, Bayes; \phi) = p(na\"{i}ve; \phi)p(Bayes; \phi). \tag{3.21}$$

 But this is clearly wrong, because words "travel together." Question for you, is it:

$$p(na\"{i}ve\ Bayes) > p(na\"{i}ve)p(Bayes)$$
 (3.22)

or...

$$p(na\"{i}ve\ Bayes) < p(na\"{i}ve)p(Bayes)$$
 (3.23)

Apply the chain rule!

Traffic lights Dan Klein makes this point with an example about traffic lights. In his hometown of Pittsburgh, there is a 1/7 chance that the lights will be broken, and both lights will be red. There is a 3/7 chance that the lights will work, and the north-south lights will be green; there is a 3/7 chance that the lights work and the east-west lights are green.

The *prior* probability that the lights are broken is 1/7. If they are broken, the conditional likelihood of each light being red is 1. The prior for them not being broken is 6/7. If they are not broken, the conditional likelihood of each being light being red is 1/2.

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Now, suppose you see that both lights are red. According to Naive Bayes, the probability that the lights are broken is $1/7 \times 1 \times 1 = 1/7 = 4/28$. The probability that the lights are not broken is $6/7 \times 1/2 \times 1/2 = 6/28$. So according to naive Bayes, there is a 60% chance that the lights are not broken!

What went wrong? We have made an independence assumption to factor the probability $P(R,R \mid \text{not-broken}) = P_{\text{north-south}}(R \mid \text{not-broken})P_{\text{east-west}}(R \mid \text{not-broken})$. But this independence assumption is clearly incorrect, because $P(R,R \mid \text{not-broken}) = 0$.

Less Naive Bayes? Of course we could decide not to make the naive Bayes assumption, and model P(R,R) explicitly. But this idea does not scale when the feature space is large (as it often is in NLP). The number of possible feature configurations grows exponentially, so our ability to estimate accurate parameters will suffer from high variance. With an infinite amount of data, we'd be fine (in theory, maybe not in practice); but we never have that. Naive Bayes accepts some bias (because of the incorrect modeling assumption) in exchange for lower variance.

3.3 Recap

- Bag-of-words representation f(x, y)
- Classification as a dot-product $\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}, y)$
- Naive Bayes
 - Define $p(\boldsymbol{x}, \boldsymbol{y})$ via a generative model
 - Prediction: $\hat{y} = \arg \max_{y} \mathbf{p}(\mathbf{x}_i, y)$
 - Learning:

$$\theta = \arg \max_{\theta} p(\boldsymbol{x}, \boldsymbol{y}; \boldsymbol{\theta})$$

$$p(\boldsymbol{x}, \boldsymbol{y}; \boldsymbol{\theta}) = \prod_{i} p(\boldsymbol{x}_{i}, y_{i}; \boldsymbol{\theta}) = \prod_{i} p(\boldsymbol{x}_{i} | y_{i}) p(y_{i})$$

$$\phi_{y,j} = \frac{\sum_{i:Y_{i}=y} x_{ij}}{\sum_{i:Y_{i}=y} \sum_{j} x_{ij}}$$

$$\mu_{y} = \frac{\text{count}(Y = y)}{N}$$

This gives the maximum-likelihood estimator (MLE; same as relative frequency estimator)

• Bias-variance tradeoff: MLE is high-variance, so add smoothing pseudo counts α . This reduces variance but adds bias.

Chapter 4 Sentiment analysis

Todo: add notes about sentiment analysis here

Discriminative learning

5.1 Features

Naive Bayes is a simple classifier, where the weights are learned based on the joint probability of labels and words. It includes an independence assumption: all features are mutually independent, conditioned on the label.

- We have defined a **feature function** f(x, y), which corresponds to "bag-of-words" features. While these features do violate the independence assumption, the violation is relatively mild.
- We may be interested in other features, which violate independence more severely. Can you think of any?
 - Prefixes, e.g. anti-, im-, un-
 - Punctuation and capitalization
 - Bigrams, e.g. not good, not bad, least terrible, ...

Rich feature sets generally cannot be combined with Naive Bayes because the distortions resulting from violations of the independence assumption overwhelm the additional power of better features.

$$p(not \ bad \ food | y) \approx p(not | y)p(bad | y)p(food | y)$$
(5.1)

$$p(not \ bad \ food | y) \not\approx p(not | y) p(bad | y) p(not \ bad | y) p(food | y)$$
(5.2)

To use these features, we will need learning algorithms that do not rely on an independence assumption.

5.2 Perceptron

In NB, the weights can be interpreted as parameters of a probabilistic model. But this model requires an independence assumption that usually does not hold, and limits our choice of features.

Why not forget about probability and learn the weights in an error-driven way?

- Until converged, at each iteration t
 - Select an instance i
 - Let $\hat{y} = \arg\max_{y} \boldsymbol{\theta}_{t}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}_{i}, y)$
 - If $\hat{y} = y_i$, do nothing
 - If $\hat{y} \neq y_i$, set $\boldsymbol{\theta}_{t+1} \leftarrow \boldsymbol{\theta}_t + \boldsymbol{f}(\boldsymbol{x}_i, y_i) \boldsymbol{f}(\boldsymbol{x}_i, \hat{y})$

Basically we are saying: if you make a mistake, increase the weights for features which are active with the correct label y_i , and decrease the weights for features which are active with the guessed label \hat{y} .

This seems like a cheap heuristic, right? Will it really work? In fact, there is some nice theory for the perceptron.

- If there is a set of weights that correctly separates your data, then your data is **separable**.
- ullet Formally, your data is (linearly) separable if there exists a set of weights ullet such that

$$\forall \boldsymbol{x}_i, y_i, \, \boldsymbol{\theta}^\mathsf{T} \boldsymbol{f}(\boldsymbol{x}_i, y_i) > \max_{y' \neq y_i} \boldsymbol{\theta}^\mathsf{T} \boldsymbol{f}(\boldsymbol{x}_i, y')$$
 (5.3)

- If your data is linearly separable, it can be proven that the perceptron algorithm will eventually find a separator.
- What if your data is not separable?
 - the number of errors is bounded...
 - but the algorithm will thrash. That is, the weights will cycle between different values, and will never converge.

The perceptron is an **online** learning algorithm.

• This means that it adjusts the weights after every example.

- This is different from Naïve Bayes, which computes corpus statistics and then sets the weights in a single operation. This is a batch learning algorithm.
- Other algorithms are **iterative**, in that they perform multiple updates to the weights, but are also **batch**, in that they have to use all the training data to compute the update. We'll mention two of those algorithms later.

Voted (averaged) perceptron

One solution to thrashing is to average the weights across all iterations:

$$\overline{\boldsymbol{\theta}} = \frac{1}{T} \sum_{t=1}^{T} \boldsymbol{\theta}_t$$

$$y = \arg \max_{y} \overline{\boldsymbol{\theta}}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}, y)$$

There is some analysis showing that voting can improve generalization (Freund and Schapire, 1999; Collins, 2002). However, this rule as described here is not practical. Can you see why not, and how to fix it?

5.3 Loss functions and large-margin classification

Naive Bayes chooses the weights θ by maximizing the likelihood p(x, y). This can be seen, equivalently, as maximizing the log-likelihood (due to the monotonicity of the log function), and as **minimizing** the negative log-likelihood. This negative log-likelihood can therefore be viewed as a **loss function**, which is minimized:

$$\log p(\boldsymbol{x}, \boldsymbol{y}; \boldsymbol{\theta}) = \sum_{i=1}^{N} \log p(\boldsymbol{x}_i, y_i; \boldsymbol{\theta})$$
 (5.4)

$$\ell_{NB}(\boldsymbol{\theta}; \boldsymbol{x}_i, y_i) = -\log p(\boldsymbol{x}_i, y_i; \boldsymbol{\theta})$$
(5.5)

$$\hat{\boldsymbol{\theta}} = \arg\min_{\boldsymbol{\theta}} \sum_{i=1}^{N} \ell_{\text{NB}}(\boldsymbol{\theta}, \boldsymbol{x}_i, y_i)$$
 (5.6)

This may seem confusing and backwards, but loss functions provide a very general framework in which to compare many approaches to machine learning.

For example, even though the perceptron is not a probabilistic model, it is also trying to minimize a **loss function**:

$$\ell_{\text{perceptron}}(\boldsymbol{\theta}; \boldsymbol{x}_i, y_i) = \begin{cases} 0, & y_i = \arg \max_y \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(x_i, y) \\ 1, & \text{otherwise} \end{cases}$$
(5.7)

This loss function has some pros and cons in comparison with Naive Bayes.

- ℓ_{NB} can suffer **infinite** loss on a single example, which suggests it will overemphasize some examples, and underemphasize others.
- $\ell_{\text{perceptron}}$ treats all errors equally. It only cares if the example is correct, and not about how confident the classifier was. Since we usually evaluate on accuracy, this is a better match.
- $\ell_{\text{perceptron}}$ is non-convex¹ and discontinuous. Finding the global optimum is intractable when the data is not separable.

We can fix this last problem by defining a loss function that behaves more nicely. To do this, let's define the *margin* as

$$\gamma(\boldsymbol{\theta}; \boldsymbol{x}_i, y_i) = \boldsymbol{\theta}^\mathsf{T} \boldsymbol{f}(\boldsymbol{x}_i, y_i) - \max_{y \neq y_i} \boldsymbol{\theta}^\mathsf{T} \boldsymbol{f}(\boldsymbol{x}_i, y)$$
 (5.8)

Then we can write a convex and continuous "hinge loss" as

$$\ell_{\text{hinge}}(\boldsymbol{\theta}; \boldsymbol{x}_i, y_i) = \begin{cases} 0, & \gamma(\boldsymbol{\theta}; \boldsymbol{x}_i, y_i) \ge 1, \\ 1 - \gamma(\boldsymbol{\theta}; \boldsymbol{x}_i, y_i), & \text{otherwise} \end{cases}$$
(5.9)

Equivalently, we can write $\ell_{\text{hinge}}(\boldsymbol{\theta}; \boldsymbol{x}_i, y_i) = (1 - \gamma(\boldsymbol{\theta}; \boldsymbol{x}_i, y_i))_+$, where $(x)_+$ indicates the positive part of x.

Essentially, we want a *margin* of at least 1 between the score for the true label and the best-scoring alternative, which we have written \hat{y} .

The hinge and perceptron loss functions are shown in Figure 5.1.

¹As a reminder, a function f is convex iff $\alpha f(x_i) + (1 - \alpha)f(x_j) \ge f(\alpha x_i + (1 - \alpha)x_j)$, for all $\alpha \in [0,1]$ and for all x_i and x_j on the domain of the function. Convexity implies that any local minimum is also a global minimum, and there are a wide array of techniques for optimizing convex functions (Boyd and Vandenberghe, 2004)



Figure 5.1: Hinge and perceptron loss functions

Large-margin online classification

Note that we can write $\theta = su$, where $||u||_2 = 1$. Think of s as the magnitude and u as the direction of the vector θ . If the data is separable, there are many values of s which attain zero hinge loss. For generality, we will try to make the smallest magnitude change to θ possible.²

At step t, we optimize:

$$\boldsymbol{\theta}_{t+1} = \arg\min_{\boldsymbol{\theta}} \frac{1}{2} ||\boldsymbol{\theta} - \boldsymbol{\theta}_t||^2 \ s.t. \ \ell_{\text{hinge}}(\boldsymbol{\theta}; \boldsymbol{x}_i, y_i) = 0$$
 (5.10)

Assuming that the constraint can be satisfied (i.e., the problem is linearly separabe), the optimal solution is found at,

$$\boldsymbol{\theta}_{t+1} = \boldsymbol{\theta}_t + \tau_t(\boldsymbol{f}(y_i, \boldsymbol{x}_i) - \boldsymbol{f}(\hat{y}, \boldsymbol{x}_i))$$
 (5.11)

$$\tau_t = \frac{\ell(\boldsymbol{\theta}; \boldsymbol{x}_i, y_i)}{||\boldsymbol{f}(x_i, y_i) - \boldsymbol{f}(x_i, \hat{y})||^2},$$
(5.12)

²In the support vector machine (without slack variables), we choose the smallest magnitude weights that satisfy the constraint of zero hinge loss. Pegasos is an online algorithm for training SVMs (Shwartz et al., 2007); it is similar to Passive-Aggressive.

where again \hat{y} is the best scoring y according to θ_t . This solution can be obtained by introducing τ_t as a Lagrange multiplier for the constraint in (5.10).

If the data is not lienarly separable, there will be instances for which we can't meet this constraint. To deal with this, we introduce a "slack" variable ξ_i . We use the slack variable to trade off between the constraint (having a large margin) and the objective (having a small change in θ). The tradeoff is controlled by a parameter C.

$$\min w \frac{1}{2} ||\boldsymbol{\theta} - \boldsymbol{\theta}_t||^2 + C\xi_t$$

$$s.t.\ell_{\text{hinge}}(\boldsymbol{\theta}; \boldsymbol{x}_i, y_i) \le \xi_t, \xi_t \ge 0$$
(5.13)

The solution to 5.13 is,

$$\boldsymbol{\theta}_{t+1} = \boldsymbol{\theta}_t + \tau_t(\boldsymbol{f}(y_i, \boldsymbol{x}_i) - \boldsymbol{f}(\hat{y}, \boldsymbol{x}_i))$$
 (5.14)

$$\tau_t = \min \left(C, \frac{\ell(\boldsymbol{\theta}; \boldsymbol{x}_i, y_i)}{||\boldsymbol{f}(\boldsymbol{x}_i, y_i) - \boldsymbol{f}(\boldsymbol{x}_i, \hat{y})||^2} \right), \tag{5.15}$$

- If *C* is 0, then infinite slack is permitted, and the weights will never change.
- As $C \to \infty$, no slack is permitted, and the optimization is identical to equation 5.10 and 5.12.

This algorithm is called "Passive-Aggressive" (PA; Crammer et al., 2006), because it is passive when the margin constraint is satisfied, but it aggressively changes the weights to satisfy the constraints if necessary.³

- PA is error-driven like the perceptron, but is more stable to violations of separability, like the averaged perceptron.
- PA allows more explicit control than the Averaged Perceptron, due to the C parameter. When C is small, we make very conservative adjustments to θ from each instance, because the slack variables aren't very expensive. When C is large, we make large adjustments to avoid using the slack variables.
- You can also apply weight averaging to PA.

³A related algorithm without slack variables is called MIRA, for Margin-Infused Relaxed Algorithm (Crammer and Singer, 2003).

• Support vector machines (SVMs) are another learning algorithm based on the hinge loss (Burges, 1998), but they try to minimize the norm of the weights, rather than the norm of the change in the weights. They are typically trained in batch style, meaning that they have to read all the training instances in to compute each update. However, SVMs can also be trained in an online fashion (Shwartz et al., 2007). The LXMLS lab guide provides a simpler online learning algorithm, based on stochastic subgradient descent (Figueiredo et al., 2013).

Pros and cons of Perceptron and PA

- Perceptron and PA are error-driven, which means they usually do better in practice than naive Bayes.
- They are also online, which means we can learn without having our whole
 dataset in memory at once. NB can also be estimated online, in the sense
 that you can stream the data and store the counts.
- The original perceptron doesn't behave well if the data is not separable, and doesn't make it easy to control model complexity.
- All these models lack a probabilistic interpretation. Probabilities are useful because they quantify the classification certainty, allowing us to compute expected utility, and to incorporate the classifier in more complex probabilistic models.

5.4 Logistic regression

Logistic regression is error-driven like the perceptron, but probabilistic like Naive Bayes. This is useful in case we want to quantify the uncertainty about a classification decision.

Recall that NB selects weights to optimize the joint probability p(y, x).

- In NB, we factor this as p(y, x) = p(x|y)p(y).
- But we could equivalently write p(y, x) = p(y|x)p(x).

Since we always know x, we really care only about p(y|x). Logistic regression optimizes this directly. To do this, we have to define the probability function

differently. We define the conditional probability directly, as,

$$p(y|\mathbf{x}) = \frac{\exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\mathbf{x}, y)\right)}{\sum_{y' \in \mathcal{Y}} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\mathbf{x}, y')\right)}$$
(5.16)

$$\log p(y|\boldsymbol{x}) = \sum_{i=1}^{N} \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}_{i}, y_{i}) - \log \sum_{y' \in \mathcal{Y}} \exp \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}_{i}, y')$$
 (5.17)

$$\hat{\boldsymbol{\theta}} = \arg \max_{\boldsymbol{\theta}} \sum_{i=1}^{N} \log p(y_i | \boldsymbol{x}_i; \boldsymbol{\theta})$$
 (5.18)

Inside the sum, we have the (additive inverse of) the **logistic loss**.

• In binary classification, we can write this as

$$\ell_{\text{logistic}}(\boldsymbol{\theta}; \boldsymbol{x}_i, y_i) = -(y_i \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{x}_i - \log(1 + \exp \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{x}_i))$$
 (5.19)

• In multi-class classification, we have,⁴

$$\ell_{\text{logistic}}(\boldsymbol{\theta}; \boldsymbol{x}_i, y_i) = -(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}_i, y_i) - \log \sum_{y' \in \mathcal{V}} \exp \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}_i, y'))$$
(5.20)

The logistic loss is shown in Figure 5.2. Because it is smooth and convex, we can optimize it through gradient steps:

⁴The log-sum-exp term is very common in machine learning. It is numerically instable because you can underflow if the inner product is small, and overflow if the inner product is large. Libraries like scipy contain special functions for computing logsumexp, but with some thought, you should be able to see how to create an implementation that is numerically stable.

⁽c) Jacob Eisenstein 2014-2015. Work in progress.



Figure 5.2: Hinge, perceptron, and logistic loss functions

$$\ell = \sum_{i=1}^{N} \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}_i, y_i) - \log \sum_{y' \in \mathcal{Y}} \exp \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}_i, y')$$
 (5.21)

$$\frac{\partial \ell}{\partial \boldsymbol{\theta}} = \sum_{i=1}^{N} \boldsymbol{f}(\boldsymbol{x}_i, y_i) - \frac{\sum_{y' \in \mathcal{Y}} \exp \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}_i, y') \boldsymbol{f}(\boldsymbol{x}_i, y')}{\sum_{y'' \in \mathcal{Y}} \exp \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}_i, y'')}$$
(5.22)

$$= \sum_{i=1}^{N} \boldsymbol{f}(\boldsymbol{x}_{i}, y_{i}) - \sum_{y' \in \mathcal{Y}} \frac{\exp \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}_{i}, y')}{\sum_{y'' \in \mathcal{Y}} \exp \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}_{i}, y'')} \boldsymbol{f}(\boldsymbol{x}_{i}, y')$$
(5.23)

$$= \sum_{i=1}^{N} \boldsymbol{f}(\boldsymbol{x}_i, y_i) - \sum_{y' \in \mathcal{Y}} p(y'|\boldsymbol{x}_i; \boldsymbol{\theta}) \boldsymbol{f}(\boldsymbol{x}_i, y')$$
 (5.24)

$$=\sum_{i=1}^{N} \boldsymbol{f}(\boldsymbol{x}_i, y_i) - E[\boldsymbol{f}(\boldsymbol{x}_i, y)]$$
(5.25)

This gradient has a pleasing interpretation as the difference between the observed counts and the expected counts.⁵ Compare this gradient with the percep-

⁵Recall that the definition of an expected value $E[f(x)] = \sum_{x} f(x)p(x)$

tron and PA update rules.

The bias-variance tradeoff is handled by penalizing large θ in the objective, adding a term of $\frac{\lambda}{2}||\theta||_2^2$. This is called L2 regularization, because of the L2 norm. It can be viewed as placing a 0-mean Gaussian prior on θ .

This penalty contributes a term of $\lambda \theta$ to the gradient, so we have,

$$\ell = \sum_{i=1}^{N} \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}_{i}, y_{i}) - \log \sum_{y' \in \mathcal{Y}} \exp \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}_{i}, y') + \frac{\lambda}{2} ||\boldsymbol{\theta}||_{2}^{2}$$
 (5.26)

$$\frac{\partial \ell}{\partial \boldsymbol{\theta}} = \sum_{i=1}^{N} \boldsymbol{f}(\boldsymbol{x}_i, y_i) - E[\boldsymbol{f}(\boldsymbol{x}_i, y)] - \lambda \boldsymbol{\theta}.$$
 (5.27)

Optimization

Batch optimization In batch optimization, you keep all the data in memory and iterate over it many times.

- The logistic loss is smooth and convex, so we can find the global optimum using gradient descent. But in practice, this can be very slow.
- Second-order (Newton) optimization would incorporate the inverse Hessian. The Hessian is

$$H_{i,j} = \frac{\partial^2}{\partial w_i \partial w_j} \ell, \tag{5.28}$$

but this matrix is usually too big to deal with.

• In practice, people usually apply **quasi-Newton optimization**, which approximates the Hessian matrix. The specific method that is particularly popular is L-BFGS⁶ NLP people usually treat L-BFGS as a black box; you will typically pass it a pointer to a function that computes the likelihood and gradient. L-BFGS is provided in scipy.optimize.

Online optimization In online optimization, you consider one example (or a "mini-batch" of a few examples) at a time. *Stochastic gradient descent* makes a

⁶A friend of mine told me you can remember the order of the letters as "Large Big Friendly Giants." Does this help you?

stochastic online approximation to the overall gradient:

$$\boldsymbol{\theta}^{(t+1)} \leftarrow \boldsymbol{\theta}^{(t)} - \eta_t \nabla_{\boldsymbol{\theta}} \ell(\boldsymbol{\theta}^{(t)}, \boldsymbol{x}, \boldsymbol{y})$$
 (5.29)

$$=\boldsymbol{\theta}^{(t)} - \eta_t(\lambda \boldsymbol{\theta}^{(t)} - \sum_{i}^{N} \boldsymbol{f}(\boldsymbol{x}_i, y_i) - E[\boldsymbol{f}(\boldsymbol{x}_i, y)])$$
 (5.30)

$$= (1 - \lambda \eta_t) \boldsymbol{\theta}^{(t)} + \eta_t \sum_{i}^{N} \boldsymbol{f}(\boldsymbol{x}_i, y_i) - E[\boldsymbol{f}(\boldsymbol{x}_i, y)]$$
 (5.31)

$$\approx (1 - \lambda \eta_t) \boldsymbol{\theta}^{(t)} + N \eta_t \left(\boldsymbol{f}(\boldsymbol{x}_{i(t)}, y_{i(t)}) - E[\boldsymbol{f}(\boldsymbol{x}_{i(t)}, y)] \right)$$
 (5.32)

where η_t is the **stepsize** at iteration t, and $\langle x_{i(t)}, y_{i(t)} \rangle$ is the instance selected at iteration t. (So here we are setting the mini-batch size equal to one.) As always, N is the total number of instances. As above, the expectation is equal to a weighted sum over the labels,

$$E[\boldsymbol{f}(\boldsymbol{x}_{i(t)}, y)] = \sum_{y' \in \mathcal{Y}} p(y'|\boldsymbol{x}_{i(t)}; \boldsymbol{\theta}) \boldsymbol{f}(\boldsymbol{x}_{i(t)}, y').$$
 (5.33)

- Note how similar this update is to the perceptron!
- If we set $\eta_t = \eta_0 t^{-\alpha}$ for $\alpha \in [1, 2]$, we have guaranteed convergence.
- We can also just fix η_t to a small value, like 10^-3 . (This is what we will do in the problem set.)
- In either case, we could tune this parameter on a development set. However, it would be acceptable to just find a value that gives a good regularized log-likelihood on the training set, since this parameter relates to the quality of the optimization, and not the generalization capability of the classifier.
- In theory, we select $\langle x_{i(t)}, y_{i(t)} \rangle$ at random, but in practice we usually just iterate through the dataset.
- We can fold N into η and λ , so that $\eta^* = N\eta$ and $\lambda^* = \lambda \frac{\eta^*}{N}$. This gives the more compact form,

$$(1 - \lambda^* \eta_t^*) \boldsymbol{\theta}^{(t)} + \eta_t^* \left(\boldsymbol{f}(\boldsymbol{x}_{i(t)}, y_{i(t)}) - E[\boldsymbol{f}(\boldsymbol{x}_{i(t)}, y)] \right)$$
(5.34)

For more on stochastic gradient descent, as applied to a number of different learning algorithms, see (Zhang, 2004) and (Bottou, 1998). Murphy (2012) traces SGD to a 1978 paper by GT's own Arkadi Nemirovski (Nemirovski and Yudin, 1978). You can find several recent chapters about online optimization in the edited volume by Sra et al. (2012).

Adagrad Recent work has shown that you can often learn more quickly by using an adaptive step-size, which is different for every feature (Duchi et al., 2011). Specifically, in the **Adagrad** algorithm (adaptive gradient), you keep track of the sum of the squares of the gradients for each feature, and rescale the learning rate by its inverse:

$$\boldsymbol{g}_{t} = -\boldsymbol{f}(\boldsymbol{x}_{i}, y_{i}) + \sum_{y' \in \mathcal{V}} p(y' \mid \boldsymbol{x}_{i}) \boldsymbol{f}(\boldsymbol{x}_{i}, y_{i}) + \lambda \boldsymbol{\theta}$$
 (5.35)

$$\mathbf{g}_{t} = -\mathbf{f}(\mathbf{x}_{i}, y_{i}) + \sum_{y' \in \mathcal{Y}} \mathbf{p}(y' \mid \mathbf{x}_{i}) \mathbf{f}(\mathbf{x}_{i}, y_{i}) + \lambda \boldsymbol{\theta}$$

$$\theta_{j}^{(t+1)} \leftarrow \theta_{j}^{(t)} - \frac{\eta}{\sqrt{\sum_{t'=1}^{t} g_{t,j}^{2}}} g_{t,j},$$

$$(5.36)$$

where j iterates over features in f(x, y). The effect of this is that features with consistently large gradients are updated more slowly. Another way to view this update is that rare features are taken more seriously, since their sum of squared gradients will be smaller. Adagrad seems to require less careful tuning of η_i and Dyer (2014) reports that $\eta = 1$ works for a wide range of problems.

Note that the Adagrad update can apply to any smooth loss function, including the hinge loss defined in Equation 5.9.

Names

Logistic regression is so named because in the binary case where $y \in \{0,1\}$, we are performing a regression of x against y, after passing the inner product $\theta^{\top}x$ through a logistic transformation. You could always do a linear regression, but this would ignore the fact that the *y* is limited to a few values.

- Logistic regression is also called **maximum conditional likelihood** (MCL), because it maximizes... the conditional likelihood $p(y \mid x)$.
- Logistic regression can be viewed as part of a larger family, called generalized linear models. If you use R, you are probably familiar with glmnet.
- Logistic regression is also called maximum entropy, especially in the earlier NLP literature (Berger et al., 1996). This is due to an alternative formulation, which tries to find the maximum entropy probability function that satisfies moment-matching constraints.

The moment matching constraints specify that the empirical counts of each label-feature pair should match the expected counts:

$$\forall j, \sum_{i=1}^{N} f_j(\boldsymbol{x}_i, y_i) = \sum_{i=1}^{N} \sum_{y \in \mathcal{Y}} p(y \mid \boldsymbol{x}_i; \boldsymbol{\theta}) f_j(\boldsymbol{x}_i, y)$$
(5.37)

Note that this constraint will be met exactly when the derivative of the likelihood function (equation 5.25) is equal to zero. However, this will be true for many values of θ . Which should we choose?

The entropy of a conditional likelihood function P(Y|X) is

$$H(P) = -\sum_{x \in \mathcal{X}} \tilde{p}(x) \sum_{y \in \mathcal{Y}} p(y|x) \log p(y|x),$$
 (5.38)

where $\tilde{p}(x)$ is the *empirical probability* of x. We compute an empirical probability by summing over all the instances in training set.

If the entropy is large, this function is smooth across possible values of y; if it is small, the function is sharp. The entropy is zero if p(y|x)=1 for some particular Y=y and zero for everything else. By saying we want maximum-entropy classifier, we are saying we want to make the least commitments possible, while satisfying the moment-matching constraints:

$$\max_{\boldsymbol{\theta}} \qquad -\sum_{\boldsymbol{x}} \tilde{p}(\boldsymbol{x}) \sum_{y} p(y|\boldsymbol{x}; \boldsymbol{\theta}) \log p(y|\boldsymbol{x}; \boldsymbol{\theta})$$
 (5.39)

s.t.
$$\forall j, \sum_{i=1}^{N} f_j(\boldsymbol{x}_i, y_i) = \sum_{i=1}^{N} \sum_{y} p(y|\boldsymbol{x}_i; \boldsymbol{\theta}) f_j(\boldsymbol{x}_i, y)$$
 (5.40)

Now, the solution to this constrained optimization problem is identical to the maximum conditional likelihood (logistic-loss) formulation we've considered in the previous section.

This view of logistic regression is arguably a little dated, but it's useful to understand what's going on. The information-theoretic concept of entropy will pop up again a few times in the course. For a tutorial on maximum entropy, see http://www.cs.cmu.edu/afs/cs/user/aberger/www/html/tutorial/tutorial.html.

5.5 Summary of learning algorithms

- Naive Bayes. pros: easy and probabilistic. cons: arguably optimizes wrong objective; usually has poor accuracy, especially with overlapping features.
- **Perceptron and PA**. pros: easy, online, and error-driven. cons: not probabilistic. this can be bad in pipeline architectures, where the output of one system becomes the input for another.
- Logistic regression. pros: error-driven and probabilistic. cons: batch learning requires black-box software; hinge loss sometimes yields better accuracy than logistic loss.

What about non-linear classification?

The feature spaces that we consider in NLP are usually huge, so non-linear classification can be quite difficult. When the feature dimension V is larger than the number of instances N — often the case in NLP — you can always learn a linear classifier that will perfectly classify your training instances. This makes selecting an appropriate **non-linear** classifier especially difficult. Nonetheless, there are some approaches to non-linear learning in NLP:

- You can add **features**, such as bigrams, which are non-linear combinations of other features. For example, the base feature *⟨coffee house⟩* will not fire unless both features *⟨coffee⟩* and *⟨house⟩* also fire.
- Another option is to apply non-linear transformations to the feature vector. Recall that the feature function f(x, y) may be composed of a vector of word counts, padded by zeros. We can think of these word counts as basic features, and apply non-linear transformations, such as $x \circ x$ or |x|.
- There is some work in NLP on using kernels for strings, bags-of-words, sequences, trees, etc. Kernelized learning algorithms are outside the scope of this class (Collins and Duffy, 2001; Zelenko et al., 2003). Kernel-based learning can be seen as a generalization of algorithms such *k*-nearest-neighbors, which classifies instances by considering the labels of the *k* most similar instances in the training set (Hastie et al., 2009).

⁷Assuming your feature matrix is full-rank.

⁽c) Jacob Eisenstein 2014-2015. Work in progress.

- Boosting (Freund et al., 1999) and decision tree algorithms (Schmid, 1994) sometimes do well on NLP tasks, but they are used less frequently these days, especially as the field increasingly emphasizes big data and simple classifiers.
- More recent work has shown how **deep learning** can perform non-linear classification. One way to use deep learning in NLP is by learning word representations while jointly learning how these representations combine to classify instances (Collobert and Weston, 2008). This approach is very hot at the moment, so I will discuss it towards the end of the semester.

5.6 Summary of classifiers

So now we've talked about four different classifiers. That's it! No more classifiers in this class. Yay? Anyway, let's review.

features?	probabilities?	easy?	complexity	tuning	estimation		Objective	
no	yes	very	$\mathcal{O}(NV)$	smoothing $lpha$	$ heta_{ij} = rac{c(x_i, y=j) + lpha}{c(y=j) + Vlpha}$	$\max \sum_i \log \mathbf{p}(m{x}_i, y_i)$	Joint likelihood	Naive Bayes
yes	yes	not really	$\mathcal{O}(NVT)$	regularizer $\lambda oldsymbol{ heta} _2^2$	$rac{\partial \mathcal{L}}{\partial oldsymbol{ heta}} = \sum_i oldsymbol{f}(oldsymbol{x}_i, y_i) - E[oldsymbol{f}(oldsymbol{x}_i, y)]$	$\max \sum_i \log \mathrm{p}(y_i m{x}_i)$	Conditional likelihood	Logistic Regression
yes	no	yes	$\mathcal{O}(NVT)$	weight averaging	$\boldsymbol{\theta}^{(t)} \leftarrow \boldsymbol{\theta}^{(t-1)} + \boldsymbol{f}(\boldsymbol{x}_i, y_i) - \boldsymbol{f}(\boldsymbol{x}_i, \hat{y})$	$\min \sum_i \delta(y_i, \hat{y})$	0-1 loss	Perceptron
yes	no	yes	$\mathcal{O}(NVT)$	slack penalty C	$\boldsymbol{\theta}^{(t)} \leftarrow \boldsymbol{\theta}^{(t-1)} + \tau_t(\boldsymbol{f}(\boldsymbol{x}_i, y_i) - \boldsymbol{f}(\boldsymbol{x}_i, \hat{y}))$	$\sum_i [1-\gamma(oldsymbol{ heta};oldsymbol{x}_i,y_i)]_+$	Hinge loss	PA

Table 5.1: Comparison of classifiers. N= number of examples, V= number of features, T= number of instances.

Chapter 6 Word-sense disambiguation

Todo: add notes about WSD here

Chapter 7

Learning without supervision

So far we've assumed the following setup:

- A training set where you get observations x_i and labels y_i
- A **test set** where you only get observations x_i

What if you never get labels y_i ?

For example, you get a bunch of text, and you suspect that there are at least two different meanings for the word *concern*.¹

The immediate context includes two groups of words:

- services, produces, banking, pharmaceutical, energy, electronics
- about, said, that, over, in, with, had

Suppose we plot each instance of concern on a graph

- x-axis is the density of words in group 1
- y-axis is the density of words in group 2

Two blobs might emerge. These blobs would correspond to two different sense of *concern*.

• But in reality, we don't know the word groupings in advance.

¹example from Pedersen and Bruce (1997)

- We have to try to apply the same idea in a very high dimensional space, where every word gets its own dimension (and most dimensions are irrelevant!)
- Or we have to automatically find a low-dimensional projection. More on that much later in the course.

Here's a related scenario:

- You look at thousands of news articles from today
- Plot them on a graph of Miley vs Syria
- Three clumps emerge (Miley, Syria, others)
- Those clumps correspond to natural document classes
- Again, in reality this is a hugely high-dimensional graph

So these examples show that we can find structure in data, even without labels.

7.1 K-means clustering

You might know about classic clustering algorithms like K-means. These algorithms are iterative:

- 1. Guess the location of cluster centers.
- 2. Assign each point to the nearest center.
- 3. Re-estimate the centers as the mean of the assigned points.
- 4. Goto 2.

This is an algorithm for finding coherent "blobs" of documents.

There is a variant called "soft k-means."

- Instead of assigning each point x_i to a specific cluster z_i
- You assign it a distribution over clusters $q_i(z_i)$

We're now going to explore a more principled, statistical version of soft K-means, called EM clustering.

By understanding the statistical principles underlying the algorithm, we can extend it in a number of cool ways.

The Expectation-Maximization Algorithm 7.2

Let's go back to the Naive Bayes model:

$$\log p(\boldsymbol{x}, \boldsymbol{y}; \boldsymbol{\phi}, \mu) = \sum_{i} \log p(\boldsymbol{x}_{i} | y_{i}; \boldsymbol{\phi}) P(y_{i}; \mu)$$

For example, x can describe the documents that we see today, and y can correspond to their labels. But suppose we never observe y_i ? Can we still do something?

Since we don't know y, let's marginalize it:

$$\log p(\boldsymbol{x}) = \log \sum_{\boldsymbol{y}} p(\boldsymbol{x}|\boldsymbol{y}; \boldsymbol{\phi}) p(\boldsymbol{y}; \boldsymbol{\mu})$$
 (7.1)

$$= \log \sum_{\mathbf{y}} \prod_{i} p(\mathbf{x}_{i}|y_{i}; \boldsymbol{\phi}) p(y_{i}; \boldsymbol{\mu})$$
 (7.2)

$$= \sum_{i} \log \sum_{y_i} p(\boldsymbol{x}_i|y_i; \boldsymbol{\phi}) p(y_i; \boldsymbol{\mu})$$
 (7.3)

Now we introduce an auxiliary variable q_i , for each y_i . We have the usual constraints: $\sum_{y} q_i(y) = 1$ and $\forall y, q_i(y) \geq 0$. In other words, q_i defines a probability distribution over Y, for each instance i. Now since $\frac{q_i(y)}{q_i(y)} = 1$,

$$\log p(\boldsymbol{x}) = \sum_{i} \log \sum_{y_i} p(\boldsymbol{x}_i | y_i; \boldsymbol{\phi}) p(y_i; \boldsymbol{\mu}) \frac{q_i(y)}{q_i(y)}$$
$$= \sum_{i} \log E_q[\frac{p(\boldsymbol{x}_i | y; \boldsymbol{\phi}) p(y; \boldsymbol{\mu})}{q_i(y)}],$$

by the definition of expectation. (Note that E_q just means the expectation under the distribution q.)

Now we apply Jensen's inequality. Jensen's equality says that because log is concave, we can push it inside the expectation, and obtain a lower bound.

$$\log p(\boldsymbol{x}) \ge \sum_{i} E_{q} \left[\log \frac{p(\boldsymbol{x}_{i}|y;\boldsymbol{\phi})p(y_{i};\boldsymbol{\mu})}{q_{i}(y)}\right]$$
$$\mathcal{J} = \sum_{i} E_{q} \left[\log p(\boldsymbol{x}_{i}|y;\boldsymbol{\phi})\right] + E_{q} \left[\log p(y;\boldsymbol{\mu})\right] - E_{q} \left[q_{i}(y)\right]$$

By maximizing \mathcal{J} , we are maximizing a lower bound on the joint log-likelihood $\log p(x)$.

Now, \mathcal{J} is a function of two arguments:

- the distributions $q_i(y)$ for each i
- the parameters μ and ϕ

We'll optimize with respect to each of these in turn, holding the other one fixed.

The E-step

First, we expand the expectation in the lower bound as:

$$\mathcal{J} = \sum_{i} E_q[\log p(\boldsymbol{x}_i|y;\boldsymbol{\phi})] + E_q[\log p(y;\mu)] - E_q[q_i(y)]$$
$$= \sum_{i} \sum_{y} q_i(y) \left(\log p(\boldsymbol{x}_i|Y_i=y;\boldsymbol{\phi}) + \log p(y;\mu) - \log q_i(y)\right)$$

As in relative frequency estimation of Naive Bayes, we need to add a Lagrange multiplier to ensure $\sum_y q_i(y)=1$, so

$$\mathcal{J} = \sum_{i} \sum_{y} q_{i}(y) \left(\log \mathbf{p}(\mathbf{x}_{i} | Y_{i} = y; \boldsymbol{\phi}) + \log \mathbf{p}(y; \boldsymbol{\mu}) - \log q_{i}(y) \right) + \lambda_{i} (1 - \sum_{y} q_{i}(y))$$

$$\frac{\partial \mathcal{J}}{\partial q_{i}(y)} = \log \mathbf{p}(\mathbf{x}_{i} | Y_{i} = y; \boldsymbol{\phi}) + \log \mathbf{p}(y; \boldsymbol{\theta}) - \log q_{i}(y) - 1 - \lambda_{i}$$

$$\log q_{i}(y) = \log \mathbf{p}(\mathbf{x}_{i} | Y_{i} = y; \boldsymbol{\phi}) + \log \mathbf{p}(y; \boldsymbol{\mu}) - 1 - \lambda_{i}$$

$$q_{i}(y) \propto \mathbf{p}(\mathbf{x}_{i} | Y_{i} = y; \boldsymbol{\phi}) \mathbf{p}(y; \boldsymbol{\mu})$$

$$\propto \mathbf{p}(\mathbf{x}_{i}, y; \boldsymbol{\phi}, \boldsymbol{\mu})$$

$$q_{i}(y) = \frac{\mathbf{p}(\mathbf{x}_{i}, y; \boldsymbol{\phi}, \boldsymbol{\mu})}{\sum_{y'} \mathbf{p}(\mathbf{x}_{i}, y'; \boldsymbol{\phi}, \boldsymbol{\mu})}$$

$$= P(Y_{i} = y | \mathbf{x}_{i}; \boldsymbol{\theta}, \boldsymbol{\phi})$$

After normalizing, each $q_i(y)$ – which is the soft distribution over clusters for data x_i – is set to the conditional probability $P(y_i|x_i)$ under the current parameters μ , ϕ .

This is called the E-step, or "expectation step," because it is derived from updating the expected likelihood under q(y).

The M-step

Next, we hold q(y) fixed and update the parameters. Let's do ϕ , which parametrizes p(x|y). Again, we start by adding Lagrange multipliers to the lower bound,

$$\begin{split} \mathcal{J} &= \sum_{i} \sum_{y} q_{i}(y) \left(\log \mathsf{p}(\boldsymbol{x}_{i} | Y_{i} = y; \boldsymbol{\phi}) + \log \mathsf{p}(y; \boldsymbol{\mu}) - \log q_{i}(y) \right) + \sum_{y} \lambda_{y} (1 - \sum_{j} \phi_{y,j}) \\ \frac{\partial \mathcal{J}}{\partial \phi_{y,j}} &= \sum_{i} q_{i}(y) \frac{x_{i,j}}{\phi_{y,j}} - \lambda_{y} \\ \lambda_{h} \phi_{y,j} &= \sum_{i} q_{i}(y) x_{i,j} \\ \phi_{y,j} &= \frac{\sum_{i} q_{i}(y) x_{i,j}}{\sum_{j'} \sum_{i} q_{i}(y) x_{i,j'}} = \frac{E_{q}[\mathsf{count}(y,j)]}{E_{q}[\mathsf{count}(y)]} \end{split}$$

So ϕ_y is now equal to the relative frequency estimate of the **expected counts** under the distribution q(y).

- As in supervised Naïve Bayes, we can apply smoothing to add α to all these counts
- The update for μ is identical: $\mu_y \propto \sum_i q_i(y)$, the expected proportion of cluster Y = y. If needed, we can add smoothing here too.
- So, everything in the M-step is just like Naive Bayes, except we used expected counts rather than observed counts.

Coordinate ascent

Algorithms that alternate between updating various subsets of the parameters are called "coordinate-ascent" algorithms.

The objective function \mathcal{J} is **biconvex**, meaning that it is separately convex in q(y) and $\langle \mu, \phi \rangle$, but it is not jointly convex.

- ullet Each step is guaranteed not to decrease ${\mathcal J}$
- This is called hill-climbing: you never go down.
- Specifically, EM is guaranteed to converge to a **local optima** a point which is as good or better than any of its immediate neighbors. But there may be many such points.

- But the overall procedure is **not** guaranteed to find a global maximum.
- This means that initialization is important: where you start can determine where you finish.
- This is not true in most of the supervised learning algorithms that we have considered, such as logistic regression; in that case, we are optimizing $\log p(y|x;\theta)$, which is defined so as to be convex with respect to the parameter θ . This means that for logistic regression (and many other supervised learning algorithms), we don't need to worry about initialization, because it won't affect our ultimate solution: we are guaranteed to reach the global minimum.

7.3 Applications of EM

EM is not really an "algorithm" like, say, quicksort. Rather, it's a framework for learning with missing data. The recipe for using EM on a problem of interest to you is something like this:

- Introduce latent variables z, such that it's easy to write the probability $P(\mathcal{D}, z)$, where \mathcal{D} is your observed data, and easy to estimate the associated parameters.
- Derive the E-step updates for q(z), which is typically factored as $q(z) = \prod_i q_{z_i}(z_i)$.

Some applications of this basic setup are presented here.

Word sense clustering

In the "demos" folder, you can find a demonstration of expectation-maximization for word sense clustering. I assume we know that there are two senses, and that the senses can be distinguished by the contextual information in the document. The basic framework is identical to the clustering model of EM as presented above.

Semi-supervised learning

Nigam et al. (2000) offer another application of EM: **semi-supervised learning**. They apply this idea to document classification in the classic "20 Newsgroup" dataset.

- In this setting, we have labels for some of the instances, $\langle \boldsymbol{x}^{(\ell)}, \boldsymbol{y}^{(\ell)} \rangle$, but not for others, $\langle \boldsymbol{x}^{(u)} \rangle$.
- Can unlabeled data improve learning?

We will choose parameters to maximize the joint likelihood,

$$\log p(x^{(\ell)}, x^{(u)}, y^{(\ell)}) = \log p(x^{(\ell)}, y^{(\ell)}) + \log p(x^{(u)})$$
(7.4)

- We treat the labels of the unlabeled documents as missing data. In the E-step we impute q(y) for the unlabeled documents only.
- The M-step computes estimates of μ and ϕ from the sum of the observed counts from $\langle \boldsymbol{x}^{(\ell)}, \boldsymbol{y}^{(\ell)} \rangle$ and the expected counts from $\langle \boldsymbol{x}^{(u)} \rangle$ and $q(\boldsymbol{y})$.
- We can further parametrize this approach by weighting the unlabeled documents by a scalar λ , which is a tuning parameter.

Multi-component modeling

- One of the classes in 20 newsgroups is comp.sys.mac.hardware.
- Suppose that there are two kinds of posts: reviews of new hardware, and question-answer posts about hardware problems.
- The language in these **components** of the mac.hardware class might have little in common.
- So we might do better if we model these components separately.

We can envision a new generative process here:

- For each document i,
 - draw the label $y_i \sim \text{Categorical}(\theta)$
 - draw the component $z_i|y_i \sim \text{Categorical}(\psi_{y_i})$
 - draw the vector of counts $x_i|z_i \sim \text{Multinomial}(\phi_{z_i})$

Our labeled data includes $\langle x_i, y_i \rangle$, but not z_i , so this is another case of missing data.

$$p(\boldsymbol{x}_i, y_i) = \sum_{z} p(\boldsymbol{x}_i, y_i, z)$$
$$= p(\boldsymbol{x}_i | z; \boldsymbol{\phi}) p(z | y_i; \boldsymbol{\psi}) p(y_i; \boldsymbol{\mu})$$

Again, we can apply EM

- We need a distribution over the missing data, $q_i(z)$. This is updated during the E-step.
- During the m-step, we compute:

$$\begin{split} \psi_{y,z} = & \frac{E_q[\mathsf{count}(y,z)]}{\sum_{z'} E_q[\mathsf{count}(y,z')]} \\ \phi_{j,y,z} = & \frac{E_q[\mathsf{count}(z,j)]}{\sum_{j'} E_q[\mathsf{count}(z,j')]} \end{split}$$

• Suppose we assume each class y is associated with K components, \mathcal{Z}_y . We can add a constraint to the E-step so that $q_i(z) = 0$ if $z \notin \mathcal{Z}_y \wedge Y_i = y$.

Chapter 8

Language models

A **language model** is used to compute the probability of a sequence of text. Why would we want to do this? Thus far, we have considered problems where text is the **input**, and we want to select an output, such as a document class or a word sense. But in many of the most prominent problems in language technology, text itself is the output:

- machine translation
- speech recognition
- summarization

As we will soon see, we can produce more **fluent** text output by computing the probability of the text.

Specifically, suppose we have a vocabulary of word types

$$V = \{aardvark, abacus, \dots, zither\}$$
 (8.1)

Given a sequence of word tokens w_1, w_2, \ldots, w_M , with $w_i \in \mathcal{V}$, we would like to compute the probability $p(w_1, w_2, \ldots, w_M)$. We will do this in a data-driven way, assuming we have a **corpus** of text.

- For now, we'll assume that the vocabulary V covers all the word tokens that we will ever see. Of course, we can enforce this by allocating a special token for unknown words. However, this might not be a great solution, as we will see later.
- Language models typically make an independence assumption across sentences, $p(s_1, s_2, ...) = \prod_j p(s_j)$, where each sentence $s_j = [w_1, w_2, ..., w_{N_j}]$.

So for our purposes, it is sufficient to compute the probability of sentences. The justification for this assumption is that the probability of words that are not in the same sentence don't depend on each other too much. Clearly this isn't true: once I mention *Manuel Noriega* once in a document, I'm far more likely to mention him again (Church, 2000). But the dependencies between words within a sentence are usually even stronger, and are more relevant to the fluency considerations inherent in applications such as translation and speech recognition (which are typically evaluated at the sentence level anyway).

So how can we compute the probability of a sentence? The simplest idea would be to apply a **relative frequency estimator**:

$$= \frac{\text{count}(Computers are useless, they can only give you answers)}}{\text{count}(\text{all sentences ever spoken})}$$
(8.3)

It's useful to think about this estimator in terms of bias and variance.

- In the theoretical limit of infinite data, it might work. But in practice, we are asking for accurate counts over an infinite number of events, since sentences can be arbitrarily long.
- Even if we set an aggressive upper bound of, say, n=20, the number of possible sentences is $\#|\mathcal{V}|^{20}$. A small vocabularly for English would have $\#|\mathcal{V}|=10^4$, so we would have 10^{80} possible sentences.
- Clearly, this estimator is extremely data-hungry. We need to introduce bias to have a chance of making reliable estimates.

Are language models meaningful? What are the probabilities of the following two sentences?

- Colorless green ideas sleep furiously
- Furiously sleep ideas green colorless

Noam Chomsky used this pair of examples to argue that the probability of a sentence is a meaningless concept:

- Any English speaker can tell that the first sentence is grammatical but the second sentence is not.
- Yet neither sentence, nor their substrings, had ever appeared at the time that Chomsky wrote this article (they have appeared lots since then).
- Thus, he argued, empirical probabilities can't distinguish grammatical from ungrammatical sentences.

Pereira (2000) showed that by identifying *classes* of words (e.g., noun, verb, adjective, adverb — but not necessarily these grammatical categories), it is easy to show that the first sentence is more probable than the second. We will talk about class-based language models later.

Are language models useful? Suppose we want to translate a sentence from Spanish:

- El cafe negro me gusta mucho.
- Word-for-word: *The coffee black me pleases much.*
- But a good language model of English will tell us:

$$P(The coffee black me pleases much) < P(I like black coffee a lot)$$
 (8.4)

How can we use this fact?

Warren Weaver on translation as decoding:

When I look at an article in Russian, I say: 'This is really written in English, but it has been coded in some strange symbols. I will now proceed to decode.'

This motivates a generative model (like Naive Bayes!):

- English sentence $w^{(e)}$ generated from language model $p_e(w^{(e)})$
- Spanish sentence $w^{(s)}$ generated from noisy channel $p_{s|e}(w^{(s)}|w^{(e)})$

(picture)

Then the **decoding** problem is: $\max_{\boldsymbol{w}^{(e)}} p(\boldsymbol{w}^{(e)}|\boldsymbol{w}^{(s)}) \propto p(\boldsymbol{w}^{(s)}, \boldsymbol{w}^{(e)}) = p(\boldsymbol{w}^{(e)})p(\boldsymbol{w}^{(s)}|\boldsymbol{w}^{(e)})$

- The **translation model** is $p(w^{(s)}|w^{(e)})$. This ensures the **adequacy** of the translation.
- The **language model** is $p(w^{(e)})$. This ensures the **fluency** of the translation.

What else can we model with a noisy channel?

- Speech recognition (original = words; encoded = sound)
- Spelling correction (original = well-spelled text; encoded = text with spelling mistakes)
- Part of speech tagging (original = tags; encoded = words)
- Parsing (original = parse tree; encoded = words)
- ..

The noisy channel model allows us to decompose NLP systems into two parts:

- The translation model, which we need labeled data to estimate.
- The language model, which we need only *unlabeled* data to estimate.

Since there is always more unlabeled data, this means we can improve NLP systems just by improving $p_e(w)$.

8.1 N-gram language models

We began with the relative frequency estimator,

$$p(Computers are useless, they can only give you answers)$$

$$= \frac{\text{count}(Computers are useless, they can only give you answers)}}{(8.5)}$$

count(all sentences ever spoken) (0.0)

We'll define the probability of a sentence as the probability of the words (in order): $p(w) = p(w_1, w_2, ..., w_M)$. We can apply the chain rule:

$$p(w) = p(w_1, w_2, \dots, w_M)$$

= $p(w_1)p(w_2 \mid w_1)p(w_3 \mid w_2, w_1) \dots p(w_M \mid w_{M-1}, \dots, w_1)$

Each element in the product is the probability of a word given all its predecessors. We can think of this as a *word prediction* task: *Computers are [BLANK]*. The relative frequency estimate:

$$p(\textit{useless}|\textit{computers are}) = \frac{\textit{count}(\textit{computers are useless})}{\sum_{x} \textit{count}(\textit{computers are } x)} = \frac{\textit{count}(\textit{computers are useless})}{\textit{count}(\textit{computers are})}$$

Note that we haven't made any approximations yet, and we could have applied the chain rule in reverse order, $p(w) = p(w_M)p(w_{M-1}|w_M)\dots$, or in any other order. But this means that we also haven't really improved anything either: to compute the conditional probability $P(W_M \mid W_{M-1}, W_{M-2}, \dots)$, we need to model $\#|\mathcal{V}|^{N-1}$, with $\#|\mathcal{V}|$ events. We can't even **store** this probability distribution, let alone reliably estimate it.

N-gram models

N-gram models make a simple approximation: condition on only the past n-1 words.

$$p(w_m \mid w_{m-1} \dots w_1) \approx P(w_m \mid w_{m-1}, \dots, w_{m-n+1})$$

This means that the probability of a sentence w can be computed as

$$p(w_1,\ldots,w_M)\approx\prod_m p(w_m\mid w_{m-1},\ldots,w_{m-n+1})$$

• To compute the probability of a whole sentence, it's convenient to pad the beginning and end with special symbols \Diamond and \Box . Then the bigram (n=2) approximation to the probability of *I like black coffee* is:

$$p(I \mid \Diamond)p(like \mid I)p(black \mid like)p(coffee \mid black)p(\Box \mid coffee)$$
 (8.7)

- In this model, we have to estimate and store the probability of only $\#|\mathcal{V}|^n$ events. A very common choice is a trigram model, in which n=3.
- The n-gram probabilities can be determined by relative frequency estimation,

$$p(w|u,v) = \frac{\text{count}(u,v,w)}{\text{count}(u,v)} = \frac{\text{count}(u,v,w)}{\sum_{w'} \text{count}(u,v,w')}$$
(8.8)

There could be too problems with an n-gram language model:

- *n* **is too small**. In this case, we are missing important linguistic context. Consider the following sentences:
 - Gorillas always like to groom **THEIR** friends.
 - The computer that's on the 3rd floor of our office building **CRASHED**.

The bolded words depend crucially on their predecessors in italics: *their* depends on knowing that *gorillas* is plural, and *crashed* depends on knowing that the subject is a *computer*. The resulting model would offer probabilities that are too low for these sentences, and too high for sentences that fail basic linguistic tests like number agreement.

- *n* **is too big**. In this case, we can't make good estimates of the n-gram parameters from our dataset. See the slides for some examples of this.
- These two problems point to another bias/variance tradeoff. Can you see how it works?
- In reality, we often have **both** problems! Language is full of long-range dependencies, and datasets are small.

We will seek approaches to keep n large, while still making low-variance estimates of the underlying parameters. To do this, we will introduce a different sort of bias: **smoothing**. But before we talk about that, let's consider how we can evaluate language models.

8.2 Evaluating language models

- Because language models are typically components of larger systems (language modeling is not really an application itself), we would prefer extrinsic evaluation: does the LM help the task (translation or whatever). But this is often hard to do, and depends on details of the overall system which may be irrelevant to language modeling.
- **Intrinsic evaluation** is task-neutral. Better performance on intrinsic metrics may be expected to improve extrinsic metrics across a variety of tasks (unless we are over-optimizing the intrinsic metric).

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Held-out likelihood

A popular intrinsic metric is the **held-out likelihood**.

- We obtain a test corpus, and compute the (log) probability according to our model. It is crucial that the words in this corpus were not used in estimating the model itself.
- A good model should assign high probability to this held-out data.
- Specifically, we compute

$$\ell(\boldsymbol{w}) = \sum_{i} \sum_{m} \log p(w_m^{(i)} | w_{m-1}^{(i)}, \dots, w_{m-n+1}^{(i)}), \tag{8.9}$$

for all sentences $w^{(i)}$ in the held-out corpus.

Perplexity

Perplexity is a transformation of the held-out likelihood, into an informationtheoretic quantity. Specifically, we compute

$$PP(\boldsymbol{w}) = 2^{-\frac{\ell(\boldsymbol{w})}{M}},\tag{8.10}$$

where M is the total number of tokens in the held-out corpus.

- After this transformation, we now prefer lower values. In the limit, we obtain probability 1 for our held-out corpus, with $PP = 2^{-\log 1} = 1$.
- Assume a uniform, unigram model in which $P(s_i) = \frac{1}{V}$ for all V words in the vocabulary. Then,

$$PP(\boldsymbol{w}) = \left[\left(\frac{1}{V} \right)^{M} \right]^{-\frac{1}{M}}$$
$$= \left(\frac{1}{V} \right)^{-1} = V$$

- We can think of perplexity as the *weighted branching factor* at each word in the sentence.
 - If we have solved the word prediction problem perfectly, PP(w) = 1, because there is only one possible choice.
 - (c) Jacob Eisenstein 2014-2015. Work in progress.

- If we have only a uniform model that assigns equal probability to every word, PP(w) = V.
- Most models fall somewhere in between.
- Here's how you remember: lower perplexity is better, because you are less perplexed.

Example On 38M tokens of WSJ, $V \approx 20K$, (Jurafsky and Martin, 2009, page 97) obtain these perplexities on a 1.5M token test set.

• Unigram: 962

• Bigram: 170

• Trigram: 109

Will it keep going down? See slides from (Manning and Schütze, 1999).

Information theory*

Perplexity is very closely related to the concept of entropy, the expected value of the information contained in each word.

$$H(P) = -\sum_{\boldsymbol{w}} p(\boldsymbol{w}) \log p(\boldsymbol{w})$$
(8.11)

The true entropy of English (or any real language) is unknown. Claude Shannon, one of the founders of information theory, wanted to compute upper and lower bounds. He would read passages of 15 characters to his wife, and ask her to guess the next character, recording the number of guesses it took for her to get the correct answer. As a fluent speaker of English, his wife could provide a reasonably tight bound on the number of guesses needed per character. **Question: is this an upper bound or a lower bound?**

Cross-entropy is a relationship between two probability distributions, the true one P(W) and an estimate Q(W).

$$H(P,Q) = E_P[\log Q]$$

$$= -\sum_{\boldsymbol{w}} p(\boldsymbol{w}) \log q(\boldsymbol{w})$$

$$= \sum_{\boldsymbol{w}} p(\boldsymbol{w}) \log \frac{p(\boldsymbol{w})}{q(\boldsymbol{w})} - p(\boldsymbol{w}) \log p(\boldsymbol{w})$$

$$= D_{KL}(P||Q) + H(Q)$$

So the cross-entropy is the KL-divergence between P and Q – a non-symmetric distance measure between distributions, which we will see again later in the course – plus the entropy of P. Since P is the language itself, we can only control Q, and minimizing the cross-entropy is equivalent to minimizing the KL-divergence.

We do not have access to the true P(W), just a sequence $\mathbf{w} = \{w_1, w_2, \dots, \}$, which is sampled from P(W). In the limit, the length of \mathbf{w} is infinite, so we have,

$$\begin{split} H(P,Q) &= -\sum_{\boldsymbol{w}} \mathbf{p}(\boldsymbol{w}) \log q(vw) \\ &= -\lim_{M \to \infty} \frac{1}{M} \log q(\boldsymbol{w}) \\ &\approx -\frac{1}{M} \log q(\boldsymbol{w}) \\ PP(S) &= 2^{-\frac{1}{M} \log q(vw)} \end{split}$$

A good language model has low cross-entropy with P(W), and thus low perplexity.

Further aside: A related topic in psycholinguistics is the "constant entropy rate hypothesis," also called the "uniform information density hypothesis." The hypothesis is that speakers should prefer linguistic choices that convey a uniform amount of information over time (Jaeger, 2010). Some evidence:

- Speakers shorten predictable words, lengthen unpredictable ones
- High-entropy sentences take longer to read
- Syntactic reductions (e.g., *I'm* versus *I am*) are more likely when the reducible word contains less information.

8.3 Smoothing and discounting

We want to estimate P(W) from sparse statistics, avoiding p(w) = 0.

Laplace/Lidstone smoothing

Simplest idea: just add "pseudo-counts"

$$p_{Laplace}(w \mid v) = \frac{\text{count}(v, w) + \alpha}{\sum_{w'} \text{count}(v, w') + V\alpha}$$
(8.12)

Anything that we add to the numerator (α) must also appear in the dominator ($V\alpha$). We can capture this with the concept of **effective counts**:

$$c_i^* = (c_i + \alpha) \frac{N}{N + V\alpha}$$

The **discount** for each n-gram is:

$$d_i = \frac{c_i^*}{c_i} = \frac{(c_i + \alpha)}{c_i} \frac{N}{(N + \alpha)}$$

- In general, this is called Lidstone smoothing
- When $\alpha = 1$, we are doing Laplace smoothing
- When $\alpha=0.5$, we are following Jeffreys-Perks law
- Manning and Schütze (1999) offer more insight on the justifications for Jeffreys-Perks smoothing

Discounting and backoff

Discounting "borrows" probability mass from observed n-grams and redistributes it.

- In Lidstone smoothing, we borrow probability mass by increasing the denominator of the relative frequency estimates, and redistribute it by increasing the numerator for all n-grams.
- Instead, we could borrow the same amount of probability mass from all observed counts, and redistribute it among only the unobserved counts. This is called absolute discounting.

• For example, if we set an absolute discount d = 0.1 in a trigram model, we get: $p(w|denied\ the) =$

word	$counts\ c$	effective counts c^*	unsmoothed probability	smoothed probability
allegations	3	2.9	0.429	0.414
reports	2	1.9	0.286	0.271
claims	1	0.9	0.143	0.129
request	1	0.9	0.143	0.129
charges	0	0.2	0.000	0.029
benefits	0	0.2	0.000	0.029

- . . .
- We need not redistribute the probability mass equally. Instead, we can backoff to a lower-order language model.
- In other words: if you have trigrams, use trigrams; if you don't have trigrams, use bigrams; if you don't even have bigrams, use unigrams. (And what if you don't even have unigrams?). This is called **Katz backoff**.

$$\begin{split} c^*(u,v) = & c(u,v) - d \\ \mathbf{p}_{\text{backoff}}(v \mid u) = \begin{cases} \frac{c^*(u,v)}{c(u)} & \text{if } c(u,v) > 0 \\ \alpha(u) \times \frac{\mathbf{p}_{\text{backoff}}(v)}{\sum_{v': c(u,v') = 0} \mathbf{p}_{\text{backoff}}(v')} & \text{if } c(u,v) = 0 \end{cases} \end{split}$$

Typically we can set d to minimize perplexity on a development set.

Interpolation

An alternative to this discounting scheme is to do interpolation: the probability of a word in context is a weighted sum of its probabilities across progressively shorter contexts.

Instead of choosing a single n-gram order, we can take the weighted average:

$$p_{Interpolation}(w|u, v) = \lambda_1 p_1^*(w|u, v)$$

$$+ \lambda_2 p_2^*(w|u)$$

$$+ \lambda_3 p_1^*(w)$$

- p_k^* is the maximum likelihood estimate (MLE) of a k-gram model
- Constraint: $\sum_{z} \lambda_{z} = 1$
- We can tune λ on heldout data...
- Or we can use **expectation maximization!**

EM for interpolation We can add a latent variable z_m , indicating the order of the n-gram that generated word w_m . Generative story:

- For each word m
 - Draw $z_m \sim \text{Categorical}(\lambda(w_m))$
 - Draw $w_m \sim p_{z_m}^*(w_m|s_{m-1}, \dots s_{m-z_m})$

As always we have two quantities of interest in our EM application:

- The parameters, λ .
- Our beliefs about the latent variables. Let $q_m(z)$ be our degree of belief that word token w_m was generated from a n-gram of order z.

Having defined these quantities, we can derive EM updates:

- E-step: $q_m(z) = p(z|w_{1:m}) = \frac{p_z^*(w_m|w_{m-1},...,w_{m-z+1})}{\sum_{z'} p_{z'}^*(w_m|w_{m-1},...w_{m-z'+1})} p(z'|\lambda(w_m))$
- M-step: $\lambda(w)_z = \frac{E_q[\operatorname{count}(W=w,Z=z)]}{\sum_{z'} E_q[\operatorname{count}(W=w,Z=z')]}$

By running the EM algorithm, we can obtain a good estimate of λ , which we can then use for unseen data. It should be clear how we can extend this approach to trigrams and beyond; Collins (2013) offers more details.

Kneser-ney smoothing

Kneser-ney smoothing also incorporates discounting, but redistributes the resulting probability mass in a different way. Consider the example:

I recently visited

- Francisco?
- Duluth?

Key idea: some words are more versatile than others.

- Suppose $p^*(Francisco) > p^*(Duluth)$, and $c(visited\ Francisco) = c(visited\ Duluth) = 0$.
- We would still guess that $p(visited\ Duluth) > P(visited\ Francisco)$, because Duluth is a more versatile word.

We define the Kneser-Ney bigram probability as

$$\begin{split} \mathbf{p}_{KN}(v|u) = &\begin{cases} \frac{\mathrm{count}(u,v) - d}{\mathrm{count}(u)}, & \mathrm{count}(u,v) > 0 \\ \alpha(u) \mathbf{p}_{\mathrm{continuation}}(v), & \mathrm{otherwise} \end{cases} \\ \mathbf{p}_{\mathrm{continuation}}(v) = &\frac{\#|u: \mathrm{count}(u,v) > 0|}{\sum_{v'} \#|u': \mathrm{count}(u',v') > 0|} \end{split}$$

- We reserve probability mass using absolute discounting *d*.
- The *continuation probability* $p_{continuation}(u)$ is proportional to the number of observed contexts in which u appears.
- As in Katz backoff, $\alpha(v)$ makes the probabilities sum to 1
- In practice, interpolation works a little better than backoff

$$p_{KN}(v|u) = \frac{\text{count}(u,v) - d}{\text{count}(u)} + \lambda(u)p_{\text{continuation}}(v)$$
(8.13)

• This idea of counting contexts may seem heuristic, but actually there is a cool justification from Bayesian nonparametrics (Teh, 2006).

8.4 Other types of Language Models

Interpolated Kneser-Ney is pretty close to state-of-the-art. But there are some interesting other types of language models, and they apply ideas that we have already learned.

Mixed-order n-gram models

Saul and Pereira (1997) described a "mixed-order" n-gram model, where you condition on multiple bigram contexts, skipping over intermediate words:

$$p(w_m|w_{m-1},\dots,w_{m-n+1}) = \sum_k \lambda_k(w_{m-k})\tilde{p}(w_m|w_m - k) \prod_j^{k-1} [1 - \lambda_j(w_{m-j})] \quad (8.14)$$

- This is an interpolated model, because we are taking the weighted average over a bunch of bigram probabilities.
- Note that the interpolation weight depends on the context word, $\lambda_k(w_{m-k})$. This means that some words can prefer certain dependency lengths for example, adjectives might prefer short dependencies, since they tend to affect adjacent nouns, while verbs might prefer longer dependencies, since they can affect indirect objects that are further away.
- The final product ensures that the weights in any particular context must add up to one: each λ_k is taking a slice of the probability mass that has already been used by the earlier contexts j < k.
- The parameters $\lambda_k(w)$ can be estimated by expectation maximization, just like in the interpolated N-gram model above.

Class-based language models

The reason we need smoothing is because the trigram probability model p(w|u,v) has a huge number of parameters. Let's simplify:

$$\mathbf{p}_{\text{class}}(w|v) = \sum_{z} P(w|z;\theta)P(z|v;\phi),$$

where $z \in [1, K], K \ll V$.

We get a bigram probability using 2VK parameters instead of V^2 . We could use EM to estimate θ and ϕ (Saul and Pereira, 1997).

- The latent variable is the class z, so the e-step updates $q_m(z)$
- The parameters are θ and ϕ , which can be updated in the M-step.

But this is usually too slow, so there are approximate algorithms, like "exchange clustering" (Brown et al 1992), which assigns each word type to a single class.

Discriminative language models

- Or we could just train a model to predict $p(w_m|w_{m-1},w_{m-2},...)$ directly.
- We might be able to use arbitrary features of the history to model long-range dependencies.
- Algorithms such as perceptron and logistic regression have been considered (Rosenfeld, 1996; Roark et al., 2007)
- Currently, "neural probabilistic language models" are attracting a lot of interest. The log-bilinear model (Mnih and Hinton, 2008) looks like this:

$$p_{\theta}^{h}(w) = \frac{\exp(s_{\theta}(w, h))}{\sum_{w'} \exp(s_{\theta}(w', h))}$$
$$s_{\theta}(w, h) = \hat{\boldsymbol{q}}_{h}^{\mathsf{T}} \boldsymbol{q}_{w} + b_{w},$$

where h is the history context, \hat{q}_h is a latent description of the history, q_w is a latent description of the word, and b_w is an offset. The history context can be computed from the words themselves, as $\hat{q}_h = \sum_i^{m-1} C_i q_i$, where the matrix C_i is applied to context position i. All parameters can be estimated to directly maximize the probability of a corpus, using gradient ascent.

Recent work has focused on efficiently training such models, with increasingly convincing results on large training sets (Mikolov et al., 2011).

Chapter 9

Morphology

¹ So far we have been focusing on NLP at the word level. Today we go **inside of words**.

We've already hinted at a morphological problem by introducing the idea of **lemmas**, where *serve/served/serving* all have the lemma *serve*.

From the perspective of document classification, these multiple forms may just seem like an annoyance, which we can get rid of by lemmatization or stemming (more on this later).

But morphology conveys information which might be crucial for some applications:

- Information retrieval
 - With a query like *bagel*, we want to get hits for *bagels*.
 - Same for *corpus/corpora*, *goose/geese*.
 - But we don't always want all the inflected forms. For example, a query for *Apple* may not want hits for *apples*
- Time. Morphology often indicates when events happen. For example, in French:

J'achete un velo
J'acheterai un velo
J'achetais un velo
J'ai acheté un velo
J'acheterais un velo

¹This chapter is pretty rough; better to see Chapter 2 of (Bender, 2013).

• Causality. Consider the difference between the Spanish examples:

Si tu vas a GT, tu seras rica If you go to GT, you will be rich Si tu vas a GT, tu eres rica. If you go to GT, you are rich

• Lexical semantics: suppose *antichrist* is not in your sentiment dictionary. Do you think it is a positive or negative word?

In addition to recognizing morphology, there are applications in which we need to produce it.

- Translation: you (pl) are smart \rightarrow Ustedes son inteligentes vs Tu eres inteligente
- Text generation: (has-property you-pl smart) \rightarrow ustedes son inteligentes

Morphology, Orthography, and Phonology

• **Morphology** describes how meaning is constructed from combining affixes. For example, it is a morphological fact of English that adding the affix +*S* to many nouns creates a plural.

$$berry + PLURAL \rightarrow berry + s$$

Orthography specifically relates to writing.
 For example,

$$berry+s \rightarrow berries$$

is an orthographic rule. We have lots of these in English, which is one reason English spelling is difficult.

- Morphological rules also include stem changes, such as $goose + PLURAL \rightarrow geese$.
- **Phonology** describes how sounds combine. For example, the different pronunciations of the final *s* in *cats* (s) and *dogs* (z) follow from a phonological rule (example (25) in the Bender text, page 30).
- In English, morphologically distinct words may be pronounced differently even when they are spelled the same, and this can reflect morphological differences. *read*+PRESENT vs. *read*+PAST.
- Conversely, morphological variants may be spelled differently even when they sound the same, like *The Champions' league* vs *The Champion's league* vs *The Champions league*.

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Productivity

One idea for dealing with morphology is to build a morphologically aware dictionary:

- Map each surface form to its underlying lemma
- Include meaning of morphology: tense, number, animacy, posession, etc.
- Then when you encounter a surface form, just look it up.

```
duck duck/N+SG
ducks duck/N+PL
duck duck/VB+PRESENT
ducks duck/VBZ+PRESENT
```

Will this work? Besides the problem of ambiguity, still another problem is that morphology is **productive**, meaning that it applies to new words. If you only know the words *Google* or *iPad*, you can immediately understand their inflected forms.

- Have you Googled that yet?
- I have owned three iPads.

Derivational morphology (more on this later) is productive in another way: you can produce new words by applying morphological changes to existing words. hyper+un+desire+able+ity

In some languages, derivational morphology can create extremely complicated words. The J&M textbook has a fun example from Turkish:

In the homework, you'll see examples from Swahili, which also has complex morphology. A dictionary of all possible surface forms in such languages would be gargantuan. So instead of building a static dictionary, we will try to model the underlying morphological and orthographic rules.

9.1 Morphemes

Two broad classes: **stems** and **affixes**.

• Intuitively, stems are the "main" part of meaning, affixes are the modifiers

A Turkish word

uygarlaştıramadıklarımızdanmışsınızcasına

uygar_laş_tır_ama_dık_lar_ımız_dan_mış_sınız_casına

"as if you are among those whom we were not able to civilize (=cause to become civilized)"

uygar: *civilized* _laş: *become*

_tır: cause somebody to do something

_ama: not able
_dik: past participle

_lar: plural

_ımız: 1st person plural possessive (our)

_dan: among (ablative case)

_mış: past

_sınız: 2nd person plural (you)

K. Oflazer pc to J&M

Figure 9.1: From (Jurafsky and Martin, 2009)

- Typically, **stems** can appear on their own (they are **free**) and affixes cannot (they are **bound**).
- Types of affixes:
 - **Prefixes**: *un+learn*, *pre+view*.
 - * These examples are derivational. English has few inflectional prefixes, but other languages have many.
 - * For example, in Swahili: *u-na-kata* versus *u-me-kata* distinguishes *you are cutting* versus *you have cut. na* and *me* are prefixes, *kata* is the root.
 - **Suffixes**: *I learn+ed*, *She learn+s*, *three apple+s*, *four fox+es*.
 - **Circumfixes** go around the stem.
 - * None in English.
 - * German has a circumfix for the past participle: sagen (say) $\rightarrow ge+sag+t$ (said)
 - * French negation can be seen as a circumfix: *Je mange+NEG\rightarrowJe ne mange pas.* (I do not eat).
 - (c) Jacob Eisenstein 2014-2015. Work in progress.

* More generally, morphemes can be non-continuous, as in the Hebrew example (7) in the Bender reading (page 12).

(7)	Root	Pattern	Part of	Phonological	Orthographic	Gloss	
			Speech	Form	Form		
	ktb	CaCaC	(v)	katav	כתב	'wrote'	
	ktb	hiCCiC	(v)	hixtiv	הכתיב	'dictated'	
	ktb	miCCaC	(n)	mixtav	מכתב	'a letter'	
	ktb	CCaC	(n)	ktav	כתב	writing,	
						alphabet'	
						-	[heb]

In this example, the root *ktb* (related to writing) is combined with patterns that indicate where to insert vowels to produce different parts-of-speech and meanings.

- Infixes go inside the stem.
 - * Tagalog: hingi+AGENT→h+um+ingi
 - * Lakhota: *m'ani* (he walks), *ma-wá-ni* (l walk). The *wá* marks agreement with a first-person singular subject; it is an infix for this word, although it is a prefix in other words.
 - * English: absolutely+fucking —absofuckinglutely, but *absfuckingsolutely arguably doesn't work.

9.2 Types of morphology

- **Inflection** creates different forms of a single word:
 - tense: to be, being, I am, you are, he is, I was
 - number: book, books
 - case: *he, his, her, they, them, their*
- **Derivation** creates new words: $grace \rightarrow disgrace \rightarrow disgracefully$
- **Cliticization** combines *Georgia+'s* into *Georgia's*; the possessive clitic 's is syntactically independent but phonologically dependent.
 - (c) Jacob Eisenstein 2014-2015. Work in progress.

- Pronouns appear as clitics in French, e.g., *j'accuse* (I accuse), as does negation *Je n'accuse personne* (I don't accuse anyone).
- Another example is from Hebrew: *l'shana tova* (literally for year good, meaning happy new year); the preposition for appears as a clitic.
- Compounding combines two words in a new word:
 cream → ice cream → ice cream cone → ice cream cone bakery
- **Portmanteaus** combine words, truncating one or both. $smoke + fog \rightarrow smog$ $glass + asshole \rightarrow glasshole$
- Word formation is *productive*: new words are subject to all of these processes

Inflectional morphology

Inflectional morphology adds information about words. English has a very simple system of inflectional morphology, compared to many languages.

Affix	Syntactic/semantic effect	Examples
-s	number: plural	cats
-'s	possessive	cat's
-s	TENSE: present, subj: 3sg	jumps
-ed	TENSE: past	jumped
-ed/-en	ASPECT: perfective	eaten
-ing	ASPECT: progressive	jumping
-er	comparative	smaller
-est	superlative	smallest

Figure 9.2: From (Bender, 2013)

• English nouns are marked for number and possession; many language also mark nouns for **case**, which is the syntactic role that the noun plays in the sentence.

- In English, we do distinguish the case of some pronouns:
 - * *He* (NOMINATIVE) gave her (OBLIQUE) his (GENITIVE) guitar.
 - * She gave him her guitar.
 - * I gave you our guitar.
 - * You gave me your guitar.

Specifically, we distinguish the **nominative** case of personal pronouns (except for 2nd-person), and the **genitive** case; all other uses are the **oblique** case.

- Other languages such as Latin, Russian, Sanskrit, and Tamil, mark the case of all nouns. These languages have additional cases, such as dative (indirect object), accusative (direct object), and vocative (address).
- In German, noun is not inflected for case, but the articles and adjectives are:
 - * Der alte Mann gab dem kleinen Affen die grosse Banane (based on example 49 from Bender)
 - * The old man (NOM) gave the little monkey (DATIVE) the big banana (ACCUSATIVE)
 - * Notice how *der*, *dem*, and *die* all mean the, but carry the case marking.
 - The adjectives are also marked for case.
- Many languages such as Romance languages mark the gender and number of nouns by inflecting the article and adjective. e.g., Spanish:
 - * El coche rojo pasó la luz roja: the red car ran the red light
 - * Los coches rojos pasó las luces rojas: the red cars ran the red lights
 - * Note that the article and adjective must **agree** for the sentence to be grammatical.
 - * In English, demonstrative determiners mark number, *this book* vs *these books*. In English, the determiner and noun must agree, e.g. **this books*.
- Gender is not necessarily binary.
 - * English pronouns include neuter *it*; German, Sanskrit, and Latin do this for all nouns.
 - * Danish and Dutch distinguish **neuter** from **common** gender
 - * Other languages distinguish animate and inanimate
 - (c) Jacob Eisenstein 2014-2015. Work in progress.

- Number if not necessarily binary.
 - * Many languages, such as Arabic and Sanskrit, include a special **dual** number for two. English has residual traces of the dual number, with *both* vs *all* and *either* vs *any*.
 - * Some Austronesian languages have a **trial** number, for groups of three.
 - * Some languages, including Arabic, have a **paucal** number, for small groups.
- English verbs are inflected for tense and number distinguishing past (*I ate*), present (*I eat*), and 3rd-person singular (*She eats*). They are also inflected for aspect, distinguishing perfective (*I had eaten*) and progressive (*I am eating*). Note that the perfective and the past tense are identical for regular verbs, e.g. we had talked, we talked.
 - Many languages (e.g., Chinese and Indonesian), do not mark tense with morphology. Indonesian uses time signals.

Saya makan apel
Saya sedang makan apel
Saya telah makan apel
Saya akan makan apel
I eat an apple
I already ate an apple
I will eat an apple

- Romance languages distinguish many more tenses than English with morphology.
 - * Spanish has multiple past tenses: **preterite** and **imperfect**, distinguishing, e.g. *I ate onions yesterday* from *I ate onions every day*. These are distinguished by morphology: *comí cebollas ayer*, *comía cebollas cada día*.
 - * Spanish and French have endings for conditional (*comería cebollas*) and future (*comeré cebollas*)
 - * All of these are marked with time signals in English; future can also be marked this way in French and Spanish, e.g. *voy a comer cebollas*.
- Romance also have separate forms for every combination of number and person.
 - * (yo hablo / tu hablas / ella habla / nosotros hablamos / vosotros hableis / ellas hablan)
 - * (je parle / tu parles / elle parle / nous parlons / vous parlez / ils parlent)

- * In Spanish, they they eliminate pronouns (pro-drop) in cases where the morphology makes it clear (unless they want to add emphasis). Chinese is also a pro-drop language?
- * This doesn't happen in French, maybe because many different spellings (parle/parles/parlent) sound the same.
- * In English, we only distiniguish 3rd-person singular.
- Adjectives in English mark comparative and superlative (*taller*, *tallest*). As we have seen, they can mark gender and number in other languages.
- Other things can be marked with affixes, such as evidentiality how the speaker came to know the information. In Eastern Pomo (a California language), there are verb suffixes for four evidential categories:

```
-ink'e nonvisual sensory-ine inferential-le hearsay-ya direct knowledge
```

Example (41) from Bender (2013) shows evidentiary marking in Turkish, *Ahmet geldi* (Ahmet came, witnessed by the speaker) vs *Ahmet gelmis* (not witnessed by the speaker)

The **index of synthesis** measures the ratio of the number of morphemes in a given text to the number of words. Languages with complex morphology are called **synthetic**; languages with simple morphology are called **isolating** or **analytic**. English is relatively, but not extremely, analytic.

An approximation of the index of synthesis is the type-token ratio. Can you see why? If you count the number of unique surface forms in 10K *parallel* sentences from Europarl, you get:

• English: 16k word types

• French: 22k

• German: 32k

• Finnish: 55k

Language	Index of synthesis
Vietnamese	1.06
Yoruba	1.09
English	1.68
Old English	2.12
Swahili	2.55
Turkish	2.86
Russian	3.33
Inuit (Eskimo)	3.72

Figure 9.3: From (Bender, 2013)

Derivational Morphology

Derivational morphology is a way to create new words and change part-of-speech.

• nominalization

-V + -ation: computerization

-V+-er: walker

- Adj + -ness: fussiness

- Adj + -ity: obesity

• negation: undo, unseen, misnomer

- **adjectivization**: V + -able : doable, thinkable, N + -al : tonal, national, N + -ous: famous, glamorous
- **abverbization**: *ADJ* + -*ily*: *clumsily*
- **lots more**: rewrite, phallocentrism, ...

You can create totally new words this way. $word \rightarrow wordify \rightarrow wordification \rightarrow wordificationism \rightarrow antiwordificationism \rightarrow hyperantiwordificationism$

Irregularities

English morphology contains a lot of irregularities: *know/knew/known*, *foot/feet*, *go/went*.. if you're not a native speaker, learning these was probably a pain in the neck.

- the good news is, there are fewer of these all the time! for example, the past tense of *show* used to be *shew*, like *know/knew* (the past participle is still *shown*).
- the bad news is, the most common words will be the last to change (if ever).

Attaching affixes can cause orthographic and phonological changes:

- walk + ed = walked, but frame+ed = framed, emit+ed = emitted, easy + ier = easier
- this is usually due to phonetic or orthographic *constraints*
- *optimality theory* is an approach to systematizing such interacting constraints. There's a lot of research on finite state models of optimality theory, but you'll have to take a linguistics course for that Karttunen and Beesley (2005).

Chapter 10

Finite-state automata

Finite-state automata are a powerful formalism for representing a subset of formal languages, the **regular** languages. As we will see, this formalism can also be used as a building block for an incredibly wide range of methods for manipulating natural language too (Mohri et al., 2002). This chapter will especially focus on **morphology**, which concerns how words are built out of smaller units. For a good reference on morphology for natural language processing, see (Bender, 2013).

Knight and May (2009) show how finite-state automata can be composed together to create impressive applications. They start with one such application — transliteration — and explain how it works. Here, we'll build the formalism from the ground up, starting with finite-state acceptors, then adding weights, and then adding transduction, finally arriving at the same sorts of applications.

10.1 Automata and languages

Basics of the formalism:

- An alphabet Σ is a set of symbols
- A string ω is a sequence of symbols. The empty string ϵ contains zero symbols.
- A language $L \subseteq \Sigma^*$ is a set of strings.

An automaton is an abstract model of a computer which reads an input string, and either accepts or rejects it.

Chomsky Hierarchy Every automaton defines a language. Different automata define different classes of languages. The Chomsky Hierarchy:

- Finite-state automata define regular languages
- Pushdown automata define context-free languages
- Turing machines define recursively-enumerable languages

Finite-state automata A finite-state automaton $M = \langle Q, \Sigma, q_0, F, \delta \rangle$ consists of:

- A finite set of states $Q = \{q_0, q_1, \dots, q_n\}$
- A finite alphabet Σ of input symbols
- A start state $q_0 \in Q$
- A set of final states $F \subseteq Q$
- A transition function δ

Determinism

- In a deterministic (D)FSA, $\delta: Q \times \Sigma \to Q$.
- In a nondeterministic (N)FSA, $\delta: Q \times \Sigma \rightarrow 2^Q$
- We can determinize any NFSA using the powerset construction, but the number of states in the resulting DFSA may be 2^n .
- Any regular expression can be converted into an NFSA, and thus into a DFSA.

The English Dictionary as an FSA We can build a simple "chain" FSA which accepts any single word. So, we can define the English dictionary with an FSA. However, we can make this FSA much more compact. (see slides)

Begin by taking the union of all of the chain FSAs by defining epsilon transitions (that is, transitions which do not consume an input symbol) from the start state to chain FSAs for each word (5303 states / 5302 arcs using a 850 word dictionary of "basic English")

- Eliminate the epsilon transitions by pushing the first letter to the front (4454 states / 4453 arcs)
- **Determinize** (2609 / 2608)
- **Minimize** (744 / 1535). The cost of minimizing an acyclic FSA is O(E). This data structure is called a trie.

Operations We've now talked about three operations: union, determinization and minimization. Other important operations are:

intersection: only accept strings in both FSAs **negation** only accept strings not accepted by FSA M**concatenation**: accept strings of the form $s=[s_1s_2]$, where $s_1\in M_1$ and $s_2\in M_2$

FSAs are closed under all these operations, meaning that resulting automaton is still an FSA (and therefore still defines a regular language).

10.2 FSAs for Morphology

Now for some morphology. Suppose that we want to write a program that accepts words that could **possibly** be constructed in accordance with English derivational morphology, but none of the impossible ones:

- grace, graceful, gracefully
- disgrace, disgraceful, disgracefully, ...
- Google, Googler, Googleology,...
- *gracelyful, *disungracefully, ...

We could just make a list, and then take the union of the list using ϵ -transitions. The list would get very long, and it would not account for productivity (our ability to make new words like *antiwordificationist*). So let's try to use finite state machines instead. Our FSA will have to encode rules about morpheme ordering, called *morphotactics*.

Let's start with some examples:

• grace: $q_0 \rightarrow_{\text{stem}} q_1$

- dis-grace: $q_0 \rightarrow_{\text{prefix}} q_1 \rightarrow_{\text{stem}} q_2$
- grace-ful: $q_0 \rightarrow_{\text{stem}} q_1 \rightarrow_{\text{suffix}} q_2$
- $\bullet \ \textit{dis-grace-ful:} \ q_0 \to_{\mathsf{prefix}} q_1 \to_{\mathsf{stem}} q_2 \to_{\mathsf{suffix}} q_3$

Can we generalize these examples?

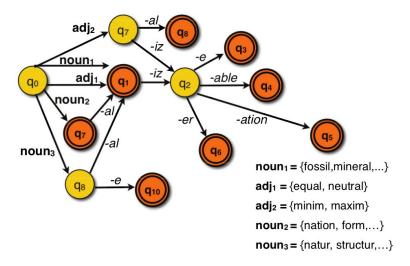


Figure 10.1: I can't find the attribution for this figure right now, sorry! I think it's either from Julia Hockenmaier's slides, or from Jurafsky and Martin (2009).

- This example abstracts away important details, like why *wordificate* is preferred to **wordifycate*. But this rule is part of English **orthography** (spelling), not **morphology**. "Two-level morphology" is an approach to integrating such orthographic transformations in a finite-state framework (Karttunen and Beesley, 2001).
- It also misses a key point: sometimes we have choices, and not all choices are considered to be equally good by fluent speakers.
 - Google counts:
 - * superfast: 70M; ultrafast: 16M; hyperfast: 350K; megafast: 87K
 - * suckitude: 426K; suckiness: 378K
 - * nonobvious: 1.1M; unobvious: 826K; disobvious: 5K

- Rather than asking whether a word is acceptable, we might like to ask how acceptable it is.
- But finite state acceptors gives us no way to express *preferences* among technically valid choices.
- We'll need to augment the formalism for this.

10.3 Weighted Finite State Automata

A weighted finite-state automaton $M = \langle Q, \Sigma, \pi, \xi, \delta \rangle$ consists of:

- A finite set of states $Q = \{q_0, q_1, \dots, q_n\}$
- A finite alphabet Σ of input symbols
- Initial weight function, $\pi: Q \to \mathbb{R}$
- Final weight function $\xi: Q \to \mathbb{R}$
- A transition function $\delta: Q \times \Sigma \times Q \to \mathbb{R}$

We have added a weight function that scores every possible transition.

- We can score any path through the WFSA by the sum of the weights.
- Arcs that we don't draw have infinite cost.
- The shortest-path algorithm can find the minimum-cost path for accepting a given string in $O(V \log V + E)$.

Applications of WFSAs

We can use WFSAs to score derivational morphology as suggested above. But let's start with a simpler example:

Edit distance . We can build an edit distance machine for any word. Here's one way to do this (there are others):

- Charge 0 for "correct" symbols and rightward moves
- Charge 1 for self-transitions (insertions)

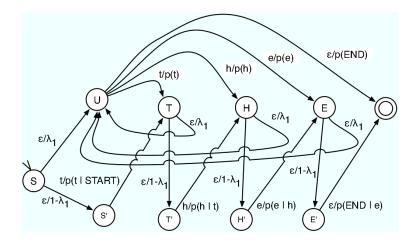


Figure 10.2: From (Knight and May, 2009)

- Charge 1 for rightward epsilon transitions (deletions)
- Charge 2 for "incorrect" symbols and rightward moves (substitutions)
- Charge ∞ for everything else

The total edit distance is the *sum* of costs across the best path through machine.

Probabilistic models For probabilistic models, we make the path costs equal to the likelihood:

$$\delta(q_1, s, q_2) = \mathbf{p}(s, q_2 \mid q_1) \tag{10.1}$$

This enables probabilistic models, such as N-gram language models.

- A unigram language model is just one state, with *V* edges.
- A bigram language model will have V states, with V^2 edges.

Knight and May (2009) show how to do an interpolated bigram/unigram language model using a WFSA. (Last year I wrote a note that I had found a better way, with only V+3 states rather than 2V+4. But now I can't find my solution!)

• Recall that an interpolated bigram language model is

$$\hat{p}(v|u) = \lambda p_2(v|u) + (1 - \lambda)p_1(v), \tag{10.2}$$

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with \hat{p} indicating the interpolated probability, p_2 indicating the bigram probability, and p_1 indicating the unigram probability.

- Unlike the basic n-gram language models, our interpolated model has nondeterminism: do we choose the bigram context or the unigram context?
- What should happen to the scores as we encounter a non-deterministic choice?
- For a sequence *a,b,a*, we want the final path score to be

$$\begin{split} \psi(a,\!b,\!a) = & (\lambda \mathbf{p}_2(a|*) + (1-\lambda)\mathbf{p}_1(a)) \\ & \times (\lambda \mathbf{p}_2(b|a) + (1-\lambda)\mathbf{p}_1(a)) \\ & \times (\lambda \mathbf{p}_2(b|a) + (1-\lambda)\mathbf{p}_1(b)) \end{split}$$

- So we could multiply along each step, and add probabilities across nondeterministic choices.
- With log-probabilities, we would add along each step, and use the log sum, $\log(e^a + e^b)$, to compute the score for non-deterministic branchings.

10.4 Semirings

We have now seen three examples: an acceptor for derivational morphology, and weighted acceptors for edit distance and language modeling. Several things are different across these examples.

- Scoring
 - In the derivational morphology FSA, we wanted a boolean "score": is the input a valid word or not?
 - In the edit distance WFSA, we wanted a numerical (integer) score, with lower being better.
 - In the interpolated language model, we wanted a numerical (real) score, with higher being better.
- Nondeterminism
 - In the derivational morphology FSA, we accept if there is any path to a terminating state.
 - (c) Jacob Eisenstein 2014-2015. Work in progress.

- In the edit distance WFSA, we want the score of the single best path.
- In the interpolated language model, we want to sum over non-deterministic choices.
- How can we combine all of these possibilities into a single formalism? The answer is semiring notation.

Formal definition

A semiring is a system $(\mathbb{K}, \oplus, \otimes, \overline{0}, \overline{1})$

- \mathbb{K} is the set of possible values, e.g. $\{\mathbb{R}_+ \cup \infty\}$, the non-negative reals union with infinity
- ullet \oplus is an addition operator
- ⊗ is a multiplication operator
- $\overline{0}$ is the additive identity
- $\overline{1}$ is the multiplicative identity

A semiring must meet the following requirements:

- $(a \oplus b) \oplus c = a \oplus (b \oplus c), (\overline{0} \oplus a) = a, a \oplus b = b \oplus a$
- $(a \otimes b) \otimes c = a \otimes (b \otimes c)$, $a \otimes \overline{1} = \overline{1} \otimes a = a$
- $\bullet \ \ a\otimes (b\oplus c)=(a\otimes b)\oplus (a\otimes c)\text{, }(a\oplus b)\otimes c=(a\otimes c)\oplus (b\otimes c)$
- $a \otimes \overline{0} = 0 \otimes \overline{a} = \overline{0}$

Semirings of interest :

where $\bigoplus_{\log}(a,b)$ is defined as $\log(e^a + e^b)$.

Semirings allow us to compute a more general notion of the "shortest path" for a WFSA.

- Our initial score is $\overline{1}$
- ullet When we take a step, we use \otimes to combine the score for the step with the running total.
- When nondeterminism lets us take multiple possible steps, we combine their scores using \oplus .

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Name	\mathbb{K}	\oplus	\otimes	$\overline{0}$	1	Applications
Boolean	{0,1}	V	\wedge	0	1	identical to an unweighted
						FSA
Probability	\mathbb{R}_{+}	+	×	0	1	sum of probabilities of all
						paths
Log-probability	$\mathbb{R}\cup -\infty\cup\infty$	\oplus_{\log}	+	$-\infty$	0	log marginal probability
Tropical	$\mathbb{R}\cup -\infty\cup\infty$	min	+	∞	0	best single path

Example Let's see how this works out for our language model example.

$$score(\{a,b,a\}) = \overline{1} \otimes \left(\lambda \otimes \mathsf{p}_2(a|*) \oplus (1-\lambda) \otimes \mathsf{p}_1(a)\right)$$
$$\otimes \left(\lambda \otimes \mathsf{p}_2(b|a) \oplus (1-\lambda) \otimes \mathsf{p}_1(b)\right)$$
$$\otimes \left(\lambda \otimes \mathsf{p}_2(a|b) \oplus (1-\lambda) \otimes \mathsf{p}_1(a)\right)$$

Now if we plug in the **probability semiring**, we get

$$score(\{a, b, a\}) = 1 \times (\lambda p_2(a|*) + (1 - \lambda)p_1(a))$$
$$\times (\lambda p_2(b|a) + (1 - \lambda)p_1(b))$$
$$\times (\lambda p_2(a|b) + (1 - \lambda)p_1(a))$$

But if we plug in the **log probability semiring**, we need the edge weights to be equal to $\log p_1$, $\log p_2$, $\log \lambda$, and $\log (1 - \lambda)$. Then we get:

$$\begin{split} score(\{a,b,a\}) &= 0 + \log \left(\exp(\log \lambda + \log \mathbf{p}_2(a|*)) + \exp(\log (1-\lambda) + \log \mathbf{p}_1(a)) \right) \\ &\quad + \log \left(\exp(\log \lambda + \log \mathbf{p}_2(b|a)) + \exp(\log (1-\lambda) + \log \mathbf{p}_1(b)) \right) \\ &\quad + \log \left(\exp(\log \lambda + \log \mathbf{p}_2(a|b)) + \exp(\log (1-\lambda) + \log \mathbf{p}_1(a)) \right) \\ &= 0 + \log (\lambda \mathbf{p}_2(a|*) + (1-\lambda) \mathbf{p}_1(a)) \\ &\quad + \log (\lambda \mathbf{p}_2(b|a) + (1-\lambda) \mathbf{p}_1(b)) \\ &\quad + \log (\lambda \mathbf{p}_2(a|b) + (1-\lambda) \mathbf{p}_1(a)), \end{split}$$

which is exactly equal to the log of the score from the probability semiring.

• The score on any specific path will be the semiring **product** of all steps along the path.

- The score of any input will be the semiring sum of the scores of all paths that successfully process the input.
- What happens if we use the tropical semiring?

10.5 Finite state transducers

FSAs and WFSAs apply to single strings. FSTs and WFSTs apply to pairs of string. FSTs define **regular relations** over pairs of strings. We can think of them in a few different ways:

- Recognizer: accepts string pairs iff they are in the relation
- Translator: reads an input, produces an output

Formally, a finite-state transducer $M = \langle Q, \Sigma, \Delta, q_0, F, \delta, \sigma \rangle$ consists of:

- A finite set of states $Q = \{q_0, q_1, \dots, q_n\}$
- Finite alphabets Σ for input symbols and Δ for output symbols
- Initial state $q_0 \in Q$ and final states $F \subseteq Q$
- A state transition function $\delta: \langle Q \times \Sigma^* \rangle \to 2^Q$
- A string transition function $\sigma: \langle Q \times \Sigma^* \rangle \to 2^{\Delta^*}$

Unlike NFSAs, not all NFSTs can be determinized. However, special subsets of NFSTs called **subsequential** transducers can be determinized efficiently (see 3.4.1 in Jurafsky and Martin (2009)).

We can build some simple NLP systems directly from FSTs.

- A zeroth-order translation system could be made from a single state and a set of self-transitions: $Q_0 \xrightarrow[el]{the} Q_0$, $Q_0 \xrightarrow[los]{the} Q_0$, $Q_0 \xrightarrow[los]{book} Q_0$, $Q_0 \xrightarrow[los]{book} Q_0$,...
- First-order translation would require a state per word in the "input" vocabulary: $Q_0 \xrightarrow[\epsilon]{the} Q_{the} \xrightarrow[los\ libros]{books} Q_0$, $Q_{the} \xrightarrow[el\ libro]{book} Q_0$.
- Inflectional morphology and orthography:

$$Q_0 \xrightarrow[wit]{wit} Q_{\text{regular}} \xrightarrow[+s]{+\text{PL}} Q_1$$

$$Q_0 \xrightarrow[wish]{wish} Q_{\text{needs-e}} \xrightarrow[+es]{+\text{PL}} Q_1$$

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10.6 Weighted FSTs

Weights can be added to FSTs in much the same way as they are added to FSAs.

- For any pair $\langle q \in Q, s \in \Sigma^* \rangle$, we have a set of possible transitions, $\langle q \in Q, t \in \Delta^*, \omega \in \mathbb{K} \rangle$, with a weight ω in the domain defined by the semiring.
- For example, we could augment the translation transducers defined above to allow alternative possible translations for a single word.
- The same semiring operations in WFSAs apply here too.

	acceptor	transducer
unweighted weighted	$FSA: \Sigma^* \to \{0, 1\}$ $FST: \Sigma^* \to \Sigma^*$	WFSA: $\Sigma^* \to \mathbb{R}$ WFST: $\Sigma^* \to \langle \Sigma^*, \mathbb{R} \rangle$

Example: General edit distance computer.

- $Q_0 \xrightarrow{a} Q_0 : 0$
- $Q_0 \xrightarrow{a} Q_0 : 1$
- $Q_0 \xrightarrow{\epsilon} Q_0 : 1$

The shortest path for a pair of strings $\langle s, t \rangle$ in this transducer has a score equal to the minimum edit distance between the strings (in the tropical semiring).

We can think of each path as defining a potential **alignment** between s and t.

Operations on FSTs

- Closed under union
- Closed under **inversion**, which switches input and output labels.
- Closed under **projection**, because FSAs are a special case of FSTs
- Not closed under difference, complementation, and intersection.

• Closed under **composition**.

FST composition is the basis for implementing the noisy channel model in FSTs, and can be used to support dozens of cool applications.

Finite state composition

Suppose we have a transducer T_1 from language I_1 to O_1 , and another transducer T_2 from O_1 to O_2 . The composition $T_1 \circ T_2$ is an FST from I_1 to O_2 .

- Unweighted definition: iff $\langle x, z \rangle \in T_1$ and $\langle z, y \rangle \in T_2$, then $\langle x, y \rangle \in T_1 \circ T_2$.
- Weighted definition:

$$(T_1 \circ T_2)(x,y) = \bigoplus_{z \in \Sigma^*} T_1(x,z) \otimes T_2(z,y)$$
(10.3)

Designing algorithms for automatic FST composition is relatively straightforward if there are no epsilon transitions; otherwise it's more challenging (Allauzen et al., 2009).

The simplest example

- $T_1: Q_0 \xrightarrow{x} Q_0, Q_0 \xrightarrow{y} Q_0$
- $T_2: Q_1 \xrightarrow{a} Q_1, Q_1 \xrightarrow{b} Q_2, Q_2 \xrightarrow{b} Q_2$
- $T_1 \circ T_2 : Q_1 \xrightarrow{x} Q_1, Q_1 \xrightarrow{y} Q_2, Q_2 \xrightarrow{y} Q_2$

For simplicity T_2 is written as a finite-state acceptor, not a transducer. Acceptors are a special case of transducers.

If we had weights, they would be combined through \otimes .

10.7 Applications of composition

Edit distance

Consider the general edit distance computer developed in section 10.6. It assigns scores to pairs of strings. If we compose it with an FSA for a given string (e.g., *tech*), we get a WFSA, who assigns score equal to the minimum edit distance from *tech* for the input string.

- Composing an FST with a FSA yields a FSA.
- A very useful design pattern is to build a **decoding** WFSA by composing a general-purpose WFST with an unweighted FSA representing the input.
- The best path through the resulting WFSA will be the minimum cost / maximum likelihood decoding.

Transliteration

English is written in a Roman script, but many languages are not. **Transliteration** is the problem of converting strings between scripts. It is especially important for names, which don't have agreed-upon translations.

A simple transliteration system can be implemented through the noisy-channel model.

- T_1 is an English character model, implemented as a transducer so that strings are scored as $\log p_r(c_1, c_2, \dots, c_M)$.
- T_2 is a character-to-character transliteration model. This can be based on explicit rules, 1 or on conditional probabilities $\log p_t(c^{(f)}|c^{(r)})$.
- T_3 is an acceptor for a given string that is to be transliterated.

The machine $T_1 \circ T_2 \circ T_3$ scores English character strings based on their orthographic fluency (T_1) and adequacy (T_2) .

Suppose you were given an Roman-script character model and a set of foreign-script strings, but no equivalent Roman-script strings. How would you use EM to learn a transliteration model?

Knight and May (2009) provide a more complex transliteration model, which transliterates between Roman and Katakana scripts, using a deep cascade that includes models of the underlying phonology. In their model,

Word-based translation

As we have seen, simple models of machine translation can be implemented as finite-state transducers.

• Recall the example: *the books / los libros*. Here we are interested in modeling English-to-Spanish translation, P(S|E).

¹http://en.wikipedia.org/wiki/Romanization_of_Russian

- Earlier we proposed a multi-state translation model to deal with the impact of pluralization on the Spanish article los
- If we build our translator by composing a Spanish language model P(S) with an Spanish-to-English transducer P(E|S), this is not necessary. This is an example of the **noisy channel model**.

Here are the specifics:

- T1 is a language model, implemented as a transducer, where every path inputs and outputs the same string, with a score equal to $\log p(w_1, w_2, \dots, w_M)$. This model's responsibility is to tell us that $p(el\ libros) \ll p(los\ libros)$.
 - For example, a bigram language model would be implemented with V states, with a score of $\log P_S(W_{n+1}=i|W_n=j)$ for the edge from the state representing word j to the state representing word i.
 - We are free to use higher-order language models. Would we need a trigram language model to correctly identify *los grandes libros* as the right translation of *the great books*?
- T2 is a word-to-word translation model.
 - It could be a bilingual dictionary, with edge weights taking the value
 In this case, the resulting translations will be scored only on boolean acceptability in the translation model, and on fluency according to the language model.
 - It could a probabilistic translation model, with each edge having a score equal to $\log p(w^{(e)}|w^{(s)})$, the log probability of the English word $w^{(e)}$ given the Spanish word $w^{(s)}$.
- T3 is an acceptor for a given foreign string $w^{(f)}$

So the translation model is $T_1 \circ T_2 \circ T_3$, which produces a WFSA, in which path are scored according to the joint log probability $\log p_{e|s}(\boldsymbol{w}^{(s)}, \boldsymbol{w}^{(e)})$. Using semiring notation, here's what happens in our example:

```
p(\textit{los libros}, \textit{the books}) = p_S(\textit{los}|\star) \otimes p_{E|S}(\textit{the}|\textit{los}) \otimes p_S(\textit{libros}|\textit{los}) \otimes p_{E|S}(\textit{books}|\textit{libros})
```

The composed FST can thus overcome its simplistic model of translation by having a more intelligent language model. This is useful, because language models can be trained without labeled data, while translation models cannot.

Structure prediction Transliteration and translation are examples of **structure prediction**, which will be a dominant theme in the remainder this course.

- Our goal for English-to-Spanish translation is to predict a sequence $w^{(s)}$, given linguistic input $w^{(e)}$.
- The set of possible sequences is very large in word-to-word translation, it is V^M . Therefore, we need to score these structures in a decomposable way.
- Finite-state composition give us a way to do that. We are implicitly decomposing the score for the tuple $\langle \boldsymbol{w}^{(e)}, \boldsymbol{w}^{(s)} \rangle$ into scores for adjacent words in $\boldsymbol{w}^{(s)}$ and aligned word pairs.
- This scoring function can be written as $\theta f(w^{(s)}, w^{(e)})$, where the FST defines a specific, decomposable feature function. More on this later.

Nondeterminism While T_2 might be non-deterministic, there is no non-determinacy about **paths**: a given pair $\langle \boldsymbol{w}^{(e)}, \boldsymbol{w}^{(f)} \rangle$ can only be transduced in one way.

Knight and May (2009) introduce a more complex translation model. They relax the assumption that there is a monotonic word-to-word alignment, allowing reorderings and multiword translations.

Word segmentation

Word segmentation is a challenging problem for speech, and also for languages like Chinese, which don't explicitly segment adjacent words in orthography. We can again use a finite-state approach.

- *T*₁ is again a language model. We can use a simplistic model representing a dictionary, or we can use a probabilistic model.
- T_2 is a sequence of unsegmented symbols.

Suppose you are given dictionary of permissible words. How would you use EM to learn the language model?

Stemming

As discussed on Tuesday, information retrieval systems that only return exact matches aren't very useful.

- Suppose you query: is it medically safe to kiss my cat on the lips
- You want to get a hit even for documents like: on the medical safety of kissing cats.
- In morphologically complex languages like Hebrew, these differences could cause early IR systems to miss more than 90% of relevant documents Choueka (1989)!
- Stemming improves the **recall** of information retrieval systems by converting all tokens of *kissing* to *kiss*, etc.

The **Porter Stemmer** is a very popular stemming program. It is written as a set of rules, which are applied in stages, e.g.

- if the word ends in *-ing* and the preceding part contains a vowel, delete the ending: *hopping* → *hopp*
- if the word ends in -ed, and the preceding part contains a vowel, delete the ending: hopped → hopp
- Next, if the remaining part contains double letters (besides ss, ll, or zz), remove one (hopp → hop)

We can think of these rules as a sequence of deterministic finite state transducers:

- We can build a transducer to strip off endings like -ing, using two states to indicate whether we have yet seen a vowel
- Next we can build a transducer to strip off double letter endings, with exceptions for s, l, and z.
- We can **compose** these transducers into a single machine.

Morphological analysis

Recall that when we talked about morphology, there were several types of interacting rules:

1. To pluralize regular words, add an *s*; but if the word ends in *sh*, you have to add *es*.

2. To conjugate 3rd-person singular, add an *s*; but if the word ends in *sh*, you have to add *es*.

Pluralization and conjugation are different morphological systems, but once they have decided to append an s, the subsequent orthographical rules are the same.

This suggests an intermediate representation:

- $dish+PL \rightarrow dish \land s\# \rightarrow dishes$
- $fist+PL \rightarrow fist \land s\# \rightarrow fists$
- $fish+PL \rightarrow fish\# \rightarrow fish$
- $fish+3S \rightarrow fish \land s\# \rightarrow fishes$
- Special symbols for morpheme and word boundaries
- The conjugation and pluralization FSTs only need to know what affix to add, and where to add it.
- Then an orthography FST takes over, and figures out how the morphemes should be combined.
- Finite-state composition allows us to automatically build a single machine for conjugation and pluralization, incorporating both the selection of affixes and orthographic constraints.

We have only described how to *generate* the English text. But if we compose this machine with a chain FSA representing an observed string, then we obtain an FSA where the set of accepted strings reveal the acceptable morphological analyses.

- For example, an utterance like *wishes* could have been produced by wish/N+3S or by wish/V+PL; both will be accepted by the resulting FSA.
- Suppose we want to distinguish the most likely morphological analysis given the context.
 - First, we add a loop back to the beginning in the morphological FST, so that we can accept multiple words.
 - (c) Jacob Eisenstein 2014-2015. Work in progress.

 Now we can build something like a language model WFST, but here we need probabilities of morphological analyses rather than just words. We might want to decompose this like

```
P(stem, affixes|context) \approx P(stem|context)P(affix|context) (10.4)
```

- If we have a morphologically annotated corpus, we can make smoothed relative frequency estimates of these probabilities. If we don't, what do we do?
- The morphological analyses are missing data, so we might be able to do EM:
 - * **E-step**: Decode all observed string sequences in the corpus, given the current probabilities.
 - * **M-step**: Treat decoding as ground truth, update probabilities
 - * Note: this is "Viterbi" EM, (a.k.a. "Hard" EM), because we are not computing probabilities over all possible decodings. We could do that using something called the expectation semiring Eisner (2002).

Context-sensitive spelling correction

Swype text entry in my old android phone does not consider context, much to my annoyance.

- I mean: *Prepare lecture for my class*
- It says: Prepare lecture foie my class

That's not smart. The bigram probabilities $P(lecture\ foie)$ and $P(foie\ my)$ should be ridiculously low. Once again, we can apply the noisy channel model, starting with T_1 as a language model and T_2 as a spelling model. If T_1 is a bigram language model, then the resulting machine is an FST with one state per word type.

- The cost of the transition from Q_a to Q_b on input s is the semiring product \otimes of two costs:
 - The transition cost from a to b in the language model
 - The "emission" cost from s to a in the edit distance machine. Depending on exactly how we're getting our input, we could memorize these costs (for all pairs of words) and work at the token level, rather than character level. But there's no conceptual difference; alternatively you could think about character-level edit distance FSTs sitting at each state.

 Given an input sequence s, we compose select the spelling-corrected string t with the greatest value,

$$\hat{t} = \max_{t} \bigoplus_{\pi} s \leadsto_{\pi} t, \tag{10.5}$$

where π is a path over input s and output t.

- Why are there multiple paths from s to t? In the edit distance machine, we can always model matching input/output symbols as an insertion and a deletion.
- Whether we care about these alternative paths depends on the semiring. In the tropical semiring, the solution is simply the minimum-cost path.

Norvig's spelling corrector

We haven't said much about where the weights come from. The bigram language model is probabilistic, and we can compute $\delta(q_s, q_t, t, t)$ as $-\log P(t|s)$. What about the edit distance model?

Before we get into that, here's an alternative approach, from Peter Norvig http://norvig.com/spell-correct.html (highly recommended). It may help explain how google is able to quickly and accurately spell-check your queries.

The approach can't easily be framed in exactly terms of the FST/semiring formalism, but it's close:

- Check if the word itself is in the dictionary. If so, return it.
- Else, check if any edit-distance=1 corrections is in the dictionary. If so, return the one with the highest unigram count.
- Else try corrections with edit-distance = 2.

The essay discusses extensions to bigrams a probabilistic model of errors. One way to build such a model is to look at data.

- Norvig suggests the Birbeck spelling error corpus, http://norvig.com/spell-correct.html.
 - Such a resource would tell us the likelihood of a word s being mispelled as t.

- But new mispellings are always possible (even inevitable)
- An alternative would be to parametrize the edit distance model with more specific probabilities: P(insertion), P(deletion), etc.
 - Maximum-likelihood estimation would compute

$$P(\text{insertion}) = \frac{\text{count}(\text{insertion})}{\text{count}(\text{all-characters})}$$
(10.6)

- But we will probably never get a corpus that gives us these counts.
- Can we bootstrap our way to success? Suppose we could just guess the probabilities.
 - * We could find the best path for each example in the training set, and keep track of the counts of insertions and deletions in these paths.

$$\begin{split} \hat{\pi}_{s,t} =& \arg\min_{\pi} \mathrm{cost}(s \leadsto_{\pi} t) \\ \mathrm{count}(\mathrm{insertion}) =& \sum_{\langle s,t \rangle \in \mathcal{D}} \mathrm{count}(\mathrm{insertion}, \hat{\pi}_{s,t}) \end{split}$$

* More probabilistically, we could compute the likelihood of each path, and use these likelihood to find the *expected* counts:

$$\mathrm{count}(\mathrm{insertion}) = \sum_{\langle s,t \rangle \in \mathcal{D}} \sum_{\pi} \frac{P(s \leadsto_{\pi} t)}{\sum_{\pi'} P(s \leadsto_{\pi'} t)} \mathrm{count}(\mathrm{insertion}, \pi_{s,t})$$

- Once we have the counts, we can go back and re-estimate the insertion probability (and all the other probabilities).
- This type of bootstrapping is another example of **expectation-maximization**.
 - We introduced EM in the simpler case of clustering.
 - * Each document had a distribution over clusters q(z)
 - * Each cluster had parameters θ .
 - * The E-step computed q(z), the M-step optimized θ .
 - Here, the Q distribution is over paths $Q(\pi)$.
 - In the E-step, we can compute $Q(\pi)$ given estimates of the parameters P(insertion).
 - (c) Jacob Eisenstein 2014-2015. Work in progress.

- In the M-step, we can update the parameters P(insertion) from expected counts under $Q(\pi)$.
- The number of paths can be very large, often too large to store.
 - In Viterbi EM, we just use the best path. This can work very well.
 - Alternatively, we can use the expectation semiring so that the "score" of the weighted path sum includes the expected counts Eisner (2001).

Each edge is a tuple of a probability and a vector of expected counts, $\langle p_i, p_i v_i \rangle$

$$\langle p_i, p_i v_i \rangle \otimes \langle p_j, p_j v_j \rangle = \langle p_i p_j, p_j p_i v_i + p_i p_j v_j \rangle = \langle p_k, p_k v_k \rangle$$
$$\langle p_i, v_i \rangle \oplus \langle p_j, v_j \rangle = \langle p_i + p_j, v_i + v_j \rangle$$

– If we can compute the semiring shortest path ($\mathcal{O}(n^3)$ at worst), we can compute the expected counts!

Speech Recognition

Speech recognition is yet another application of finite-state methods in NLP. It takes the idea of intermediate representations even further.

We want a mapping from intended words to observed acoustics. This can be seen as a multilevel process:

- G: WFST which generates English sentences: *I went to the bank*
- L: WFST which converts each word to context-indendent phonemes (WFST. Pronunciation dictionaries have this information): *Ay w eh n t t uw ...*
- C: WFST which converts context-independent phonemes to context-dependent phonemes: *Ay w eh n t uw ...*
- A: WFST which converts context-dependent phonemes to acoustic observations

By composing the FST $G \circ L \circ C \circ A$ with an FSA representing the observed acoustics O, we obtain a single WFSA which scores proposed English sentences for the observations O.

Inference The composed FSA would be ridiculously huge. Beam pruning is a technique for pruning away paths which are extremely unlikely, resulting in a faster, smaller FST.

Estimation It's really hard to get labeled data for all of these levels, for example context-dependent phones. We can treat this as a hidden variable and estimate it using EM.

Software There are mature software toolkits for working with finite state machines. OpenFST is a C++ package which I have had some experience with; it's fast and relatively well-documented. XFST and Carmel are other options.

10.8 Recap

We saw how finite-state composition can create powerful NLP systems out of simple, modular components.

- For example, we can compose a translation model (one state!) and a bigram language model (*V* states) to create a finite-state translation machine.
- If we compose the translation machine with a chain FSA representing the "input", we get a WFSA whose shortest path is the best translation of the input.
- We didn't talk about algorithms for composition, but the formal definition is

$$(T_1 \circ T_2)(x,y) = \bigoplus_{z} T_1(x,z) \otimes T_2(z,y). \tag{10.7}$$

In other words, to compute the score of $\langle x, y \rangle$ in the composed machine, we take the semiring addition over all strings z, for which we compute the extension of the score of $\langle x, z \rangle$ in T_1 with the score of $\langle z, y \rangle$ in T_2 .

- For example, the language model only has edges for $\langle s, s \rangle$, with score $\log P(s_i|s_{i-1})$. The translation model has edges for all $\langle s, t \rangle$, with score $\log P(t|s)$. So the composed machine must have edges for all $\langle s, t \rangle$, with score $\log P(t|s) \otimes \log P(s_i|s_{i-1})$.
- In the chain FSA, each edge takes exactly one string t_i , with score $\overline{1}$. So the composed machine is a WFSA, with edges for each possible s_i , each having score $\log P(t_i|s_i) \otimes \log P(s_i|s_{i-1}) \otimes \overline{1}$. This structure is known as a **trellis**.

Chapter 11

Part-of-speech tagging

Words can be grouped into rough classes based on syntax.

- Why is colorless green ideas sleep furiously more acceptable than ideas colorless furiously green sleep?
- Why is teacher strikes idle children ambiguous?

In both examples, word classes can provide an explanation.

- Word classes have strong ordering constraints:
 - JJNVR is likely
 - NJRJV is unlikely (why?)
- Ambiguity about word class leads to very different interpretations:
 - NNVN
 - NVJN (ouch!)

So clearly we have intuitions about a few parts-of-speech already: noun, verb, adjective, adverb. Jurafsky and Martin (2009) describe these as the four major **open** word classes, although apparently not all languages have all of them.

What other word classes are there?

• The Penn Treebank defined a set of 45 POS tags.¹

¹http://www.comp.leeds.ac.uk/ccalas/tagsets/upenn.html

- The Brown corpus defined a set of 87 POS tags.²
- Petrov et al. (2012) define a "universal" set of 12 tags.

Which is right?

Example Let's look at some data.

My name is Ozymandias, king of kings: Look on my works, ye Mighty, and despair!

The part-of-speech tags for this couplet from Ozymmandias are shown in Table 11.1.

- All tagsets distinguish basic categories like nouns, pronouns, verbs, adjectives, and punctuation.
- The Brown tagset includes a number of fine-grained distinctions:
 - specific tags for the *be, do,* and *have* verbs, which the other two tagsets just lump in with other verbs.
 - distinct tags for possessive determiners (*my name*) and possessive pronouns (*mine*)
 - distinct tags for the third-person singular pronouns (e.g., *it*, *he*) and other pronouns (e.g., *they*, *we*, *I*)
- The Universal tagset aggressively groups categories that are distinguished in the other tagsets:
 - all nouns are grouped, ignoring number and the proper/common distinction (see below)
 - all verbs are grouped, ignoring inflection
 - preposition and postpositions are grouped as adpositions
 - all punctuation is grouped
 - coordinating and subordinating conjunctions (e.g. and versus that) are grouped

 $^{^2 \}verb|http://www.comp.leeds.ac.uk/ccalas/tagsets/brown.html|$

⁽c) Jacob Eisenstein 2014-2015. Work in progress.

	Brown	PTB	Universal
My	possessive determiner (DD\$)	possessive pronoun (PRP\$)	pronoun (PRON)
name	noun, singular, common (NN)	NN	NOUN
is	verb "to be" 3rd person, singular (BEZ)	verb 3rd person, singular (VBZ)	VERB
Ozymandias	proper noun, singular (NP)	proper noun, singular (NNP)	NOUN
,	comma (,)	comma (,)	punctuation (.)
king	NN	NN	NOUN
of	preposition (IN)	preposition (IN)	adposition (ADP)
kings	noun, plural, common (NNS)	NNS	NOUN
:	colon (:)	mid-sentence punc (:)	
Look	verb, base: uninflected present, imperative, or infinite (VB)	VB	VERB
on	IN	IN	ADP
my	DD\$	PRP\$	PRON
works	NNS	NNS	NOUN
ye	personal pronoun, nominative, non 3S (PPSS)	personal pronoun, nominative (PRP)	PRON
mighty	adjective (JJ)	JJ	adjective (ADJ)
,	comma (,)	comma (,)	punctuation (.)
and	coordinating conjunction (CC)	CC	conjunction (CONJ)
despair	VB	VB	VERB

Table 11.1: Part-of-speech annotations from three tagsets.

So which is "right"? It depends. The Brown tags can be useful for certain applications, and they may have strong tag-to-tag relations that make tagging

easier (see next chapter). But they are more expensive to annotate. The Universal tags are intended to generalize across many types of text, and should be easier to annotate.



Figure 11.1: needs attribution

11.1 Details about parts-of-speech

As usual, Bender (2013) provides a useful linguistic perspective.

- Nouns describe entities and concepts
 - Proper nouns name specific people and entities: Georgia Tech, Janet, Buddhism. In English, they're usually capitalized. PTB tags: NNP (singular), NNPS (plural)
 - Common nouns cover everything else. In English, they're often preceded by articles, e.g. the book, the university. Common nouns decompose into
 - * **Count nouns** have a plural and need an article in the singular, *dogs*, *the dog*.
 - (c) Jacob Eisenstein 2014-2015. Work in progress.

- * **Mass nouns** don't have a plural and don't an article in the singular, snow is cold, gas is expensive
- **Pronouns** refer to specific noun phrases or entities or events.
 - * **Personal pronouns** refer to people or entities: *you, she, I, it, me*. PTB tag: PRP
 - * **Possessive pronouns** are pronouns that indicate possession: *your, her, my, its, one's, our.* PTB tag: PRP\$
 - Wh-pronouns are used in question forms (Where are you going?, WP) and as relative pronouns in forms like (The girl who played with fire)

Unlike other nouns, the set of possible pronouns cannot be expanded! It is a **closed class**.

- Verbs describe activities, processes, and events, e.g. eat, write, sleep
 - The Penn Treebank differentiates verbs by morphology: VB (infinitive),
 VBD (past), VBG (present participle), VBN (past participle), VBP (present, non 3rd person-singular), VBZ (present 3rd person singular)
 - modals are a closed subclasses of verbs, such as (should, can, will, must).
 They get PTB tag MD
 - **copula** is *be* with a predicate, e.g. *she* is *hungry*. The Brown Tagset distinguishes copula, but PTB doesn't.
 - auxiliary verbs include be, have, will to form complex tenses, e.g. we will have done it twice.
 - * Also includes *do* in questions and negation, e.g. *did you eat yet?*. Apparently this is from Welsh, which was spoken in England before the Anglo-Saxons invaded; *do* doesn't function this way in German.
 - * The Brown corpus has special tags for HAVE and DO, but the PTB doesn't.
- Adjectives describe properties of entities: antique, vast, trunkless
 - Attributive use: an antique land
 - **Predicative use**: the land was antique
 - Gradable adjectives (big) have comparative form (bigger) and superlative form (biggest)
 - (c) Jacob Eisenstein 2014-2015. Work in progress.

- Can you think of an adjective that is not gradable?
- With big, we can move to comparative form by adding the suffix -est. This is an example of agglutinative morphology. Can you think of an adjective in English where the relationship is not agglutinative? (rather, it's fusional). How about good, better, best?
- PTB tags: JJ, JJR, JJS
- Adverbs describe properties of events.
 - Manner: *slowly, slower, fast, hesitantly*
 - Degree: extremely, very, highly
 - Directional and locative: here, downstairs, near
 - Temporal: yesterday, Monday
 - Besides verbs, adverbs may also modify sentences, adjectives, or other adverbs. Apparently, the very ill man walks extremely slowly
 - Adverbs may also be gradable. Tags: RB, RBR, RBS
- Prepositions are a closed class of words that can come before noun phrases, forming a prepositional phrase that relates the noun phrase to something else in the sentence.
 - I eat sushi WITH soy sauce
 - I eat sushi WITH chopsticks
 - To gets its own tag TO, because it forms the infinitive with bare form verbs (VB), e.g. I want to eat
 - Everything else is tagged IN in the PTB.
- Coordinating conjunctions join two elements,
 - vast and trunkless legs
 - She eats burgers but she drinks soda
 - PTB tag: CC
- **Subordinating conjunctions** introduce a subordinate clause, e.g. *She thinks THAT Chomsky is wrong*. PTB tag: annotIN
 - (c) Jacob Eisenstein 2014-2015. Work in progress.

- **Particles** are oddball words that come with verbs and can change their meaning to a new **phrasal verb**, e.g., *come ON*, *he brushed himself OFF*, *let's check OUT that new restaurant*. They are a closed class, PTB tag RP.
- **Determiners** are a closed class of words that precede noun phrases.
 - Articles: the, an, a
 - Demonstratives: this, these, that
 - Quantifiers: some, every, few
 - Wh-determiners, WHICH burger should I choose?
 - PTB tag: DT

Oddballs

- **Existential there**, e.g. *There is no way out of here*, gets its own tag, EX.
- So does the possessive ending 's, POS
- So do numbers (CD), list items (LS), commas, and other non-alphabetic symbols.

11.2 Part of speech tagging

Part of speech tags relate to a number of other linguistic phenomena:

- Lexical semantics: *can*/V vs *can*/N, *teacher strikes children*, etc
- Pronunciation: *inSULT* vs *INsult*, *conTENT* vs *CONtent*
- Translation: $park/v \rightarrow garer$, $park/N \rightarrow parque$
- NP chunking: grep {JJ | NN}* {NN | NNS}

POS tagging is a useful preprocessing step for downstream applications. So how can we build an automatic POS tagger?

- Observation 1: it's "easy."
 - 60% of word types have only one possible POS tag (in English).
 - If you choose the majority POS tag for each token, you get 90% right.
- Observation 2: it's not easy: a few words have a lot of possible POS tags
 - (c) Jacob Eisenstein 2014-2015. Work in progress.

- We're taking it **back**/RB
- The shirt off my back/NN
- Go back/RP where you belong
- If you challenge him, I'll back/VBP you
- The back/]] roads are safer
- Observation 3: 90% is not actually very good. $0.9^{10} \approx .3$, so you will only get 30% of ten-word sentences correct. Sentences have exponentially many possible POS sequences:

Anyway, let's look at a tougher poem, Jabberwocky:

'Twas brillig, and the slithy toves Did gyre and gimble in the wabe: All mimsy were the borogoves, And the mome raths outgrabe

Forget *twas*. What about *slithy*? Can you guess the POS? What about *toves*? You don't know these words. What information are you using to guess?

Word identity: you do know that and is CC and the is DET

Context

- JJ NN is likely
- NN JJ is unlikely

Morphology

- -s → noun or verb
- -able → adjective (98% of the time!)
- -ly → adverb
- *un* → adjective or verb

- (not rules, just hints)

Let's put morphology on hold for a minute.

- Suppose we have an annotated corpus, with tagged sentences, $\langle w_{1:N_t}, y_{1:N_t} \rangle_{1:T}$.
- Then we could estimate the likelihood of a word given a tag,

$$P(w|y) = \frac{\text{count}(w,y)}{\text{count}(y)}$$
 (11.1)

As always, smoothing is possible...

• Given this same annotated corpus, we could also compute $P(y_n|y_{n-1})$, a sort of language model over tags.

$$P(y_n|y_{n-1}) = \frac{\text{count}(y_{n-1}, y_n)}{\text{count}(y_{n-1})}$$
(11.2)

- Let's combine these ideas via a generative story
 - For word n, draw tag $y_n \sim \text{Categorical}(\theta_{y_{n-1}})$
 - Then draw word $w_n \sim \text{Categorical}(\phi_{y_n})$

We've built a generative model that explains our observations w through a bigram generative model over the tags.

Under this model, we can compute

$$P(\boldsymbol{y}|\boldsymbol{w}) \propto P(\boldsymbol{w}, \boldsymbol{y}) \tag{11.3}$$

$$P(\boldsymbol{w}|\boldsymbol{y})P(\boldsymbol{y}) \tag{11.4}$$

$$\prod_{n=1}^{N} P(w_n|y_n)P(y_n|y_{n-1})$$
 (11.5)

- This is a **hidden Markov model**. It's Markov because the probability of y_n depends only on y_{n-1} and not any of the previous history. It's hidden because y_n is unknown when we decode.
- We can treat this as a special case of finite state transduction. Can you see how?

Chapter 12

Sequence Labeling

In sequence labeling, we want to assign tags to words. There are many applications:

- Part-of-speech tagging: Go/V to/P Georgia/N Tech/N next/J year/N ./.
- Named entity recognition: Go/O to/O Georgia/B-ORG Tech/I-ORG next/B-DATE year/I-DATE ./O
- Phrase chunking: Go/B-VP to/B-PP Georgia/B-NP Tech/I-NP next/B-NP year/I-NP ./O

In classification, we would choose each tag independently:

$$p(y_m|w_m) \perp p(y_n|w_n), \forall m \neq n$$
(12.1)

In sequence labeling, we choose the sequence of tags **jointly**. Probabilistically, we might try to choose $\hat{y} = \arg\max_{y \in \mathcal{T}^M} p(y|w)$. As we will see later, we can also write this in the form of a linear predictor:

$$\hat{\boldsymbol{y}} = \arg\max_{\boldsymbol{y} \in \mathcal{T}^M} \boldsymbol{\theta}^\mathsf{T} \boldsymbol{f}(\boldsymbol{x}, \boldsymbol{y}) \tag{12.2}$$

In either case, we have an immediate problem: finding the best scoring tag sequence in the set \mathcal{T}^M . As the notation suggests, the number of possible tag sequences is exponential in the length of the sequence. Consider part-of-speech tagging:

Even after using a tag dictionary to restrict the set of possible tags for each word, there are thirty-six possible tag sequences for this six-word sentence. This means we will need clever algorithms to compute $\arg \max_{\boldsymbol{v} \in \mathcal{T}^M}$. We cannot enumerate all possibilities.

Hidden Markov Models 12.1

Let's first think about tagging as a probabilistic model. Specifically, we want to maximize $p(y|w) \propto p(y,w)$, where w are words and y are tags. This is equivalent to Naive Bayes, but for sequence labeling.

As in Naive Bayes, we define the probability distribution p(w, y) through a generative story,

- For word m, draw tag $y_m \sim \text{Categorical}(\lambda_{y_{m-1}})$
- Then draw word $w_m \sim \text{Categorical}(\phi_{y_m})$

Under this model, we can compute

$$p(y \mid w) \propto p(w, y) \tag{12.3}$$

$$p_e(\boldsymbol{w} \mid \boldsymbol{y}; \phi) p_t(\boldsymbol{y}; \lambda) \tag{12.4}$$

$$\mathbf{p}_{e}(\boldsymbol{w} \mid \boldsymbol{y}; \phi) \mathbf{p}_{t}(\boldsymbol{y}; \lambda) \qquad (12.4)$$

$$\prod_{m}^{M} \mathbf{p}_{e}(w_{m} \mid y_{m}; \phi) \mathbf{p}_{t}(y_{m} \mid y_{m-1}; \lambda) \qquad (12.5)$$

- This is a **hidden Markov model**. It's Markov because the probability of y_m depends only on y_{m-1} and not any of the previous history. It's hidden because y_m is unknown when we decode.
 - The probability $p_e(w_m \mid y_m; \phi)$ is the **emission probability**, since the words are treated as emissions from the tags.
 - The probability $p_t(y_m \mid y_{m-1}; \lambda)$ is the **transition probability**, since it assigns probability to each possible tag-to-tag transition.

- We can describe this generative story as a graphical model. Note that although graphical models and finite-state models both use circles and arrows, the meaning is completely different.
- Our generative story assumes that the words are conditionally independent, given the tags, so

$$w_n \perp \{ \boldsymbol{w}_{m \neq n} \} \mid y_n.$$

Conditional independence is not the same as independence. We do **not** have $p(w_n, w_m) = p(w_n)p(w_m)$ because the tags are related to each other. For example, suppose that (a) nouns always follow determiners, (b) *the* is always a determiner and (c) *bike* is always a noun. Then

$$\begin{split} P(W_{m} = \textit{the}, W_{m+1} = \textit{bike}) &= \sum_{y_{m+1}, y_{m}} P(W_{m} = \textit{the}, W_{m+1} = \textit{bike}, y_{m+1}, y_{m}) \\ &= \sum_{y_{m+1}, y_{m}} P(W_{m+1} = \textit{bike} \mid y_{m+1}, y_{m}, W_{m} = \textit{the}) \\ &\times P(y_{m+1} \mid y_{m}, W_{m} = \textit{the}) P(y_{m} \mid W_{m} = \textit{the}) P(W_{m} = \textit{the}) \\ &= \sum_{y_{m+1}} P(W_{m+1} = \textit{bike} \mid y_{m+1}) \\ &\times \sum_{y_{m}} P(y_{m+1} \mid y_{m}) P(y_{m} \mid W_{m} = \textit{the}) P(W_{m} = \textit{the}) \\ &= P(W_{m+1} = \textit{bike} \mid y_{m+1} = \text{NOUN}) \times 1 \times 1 \times P(W_{m} = \textit{the}) \\ &> P(W_{m+1} = \textit{bike}) P(W_{m} = \textit{the}) \end{split}$$

- Another way to think about independence is that if we are told one tag, it affects all of our other tagging decisions.
 - For example, in the sentence teacher strikes idle children, we might choose tag sequence NN VBZ JJ NNS.
 - But if are given $y_3 = VBP$, then suddenly $y_2 = VBZ$ looks like a bad choice because $p_T(VBZ, VBP)$ is very small.
 - So we might now choose $y_2 = NNS$.
 - This change might cascade back to y_1 , etc (not in this case, but it could happen in theory)

- A classifier-based tagger, which treated the tags as IID, might ignore these dependencies, and produce a tag sequence that contained unlikely transitions like VBZ, VBP.
- A better alternative might be to tag the text from left-to-right; we could then condition on the previous tag, choosing

$$y_m = \arg \max_{y} \mathbf{p}_e(w_m | y_m) \mathbf{p}_t(y_m | y_{m-1})$$
 (12.6)

But this approach is "greedy," and can mistakenly commit to bad tagging decisions. For example, in *teacher strikes strand children*, we might initially choose $y_2 = \text{VBZ}$, because this is more common than the noun sense of *strikes*. However, we are then stuck, because *strand* has low probability as anything but a verb, yet the verb-verb transition also has low probability. The greedy tagger is unable to recover the globally optimal sequence, NN NNS VBP NNS, without backtracking.

• This is why we need **joint inference** over $y_{1:M}$ to find $\hat{y} = \arg \max_{y} p(w, y)$. The key challenge is to search over the exponential number of tag sequences efficiently.

12.2 Sequence labeling as finite-state transduction

To see whether efficient joint inference is possible, we first formulate the problem in terms of finite-state transduction.

- Transducer E has one state, and transduces from tags to words, with $\delta_{w,t}^{(e)} = \mathsf{p}_e(w|y)$.
- Transducer T has $\#|\mathcal{T}|$ states (if it's a bigram model), and transduces tags to tags, with $\delta^{(t)}_{y_m,y_{m-1}} = \mathsf{p}_t(y_m|y_{m-1})$.
- Recall the definition of finite state composition,

$$(T \circ E)(y, x) = \bigoplus_{z} T(y, z) \otimes E(z, x). \tag{12.7}$$

Since T only accepts identical tag pairs $\langle y, y \rangle$, we can ignore \bigoplus ; there's only one possible z = y. The result of $T \circ E$ transduces tags to words, with edge

weights

$$\begin{split} \delta_{w,y_{m},y_{m-1}}^{(toe)} &= & \delta_{w,y_{m}}^{(e)} \otimes \delta_{y_{m-1},y_{m}}^{(t)} \\ &= & p(w \mid y_{m}) \otimes p(y_{m} \mid y_{m-1}) \\ &= & p(w \mid y_{m}) p(y_{m} \mid y_{m-1}) \\ &= & p(w,y_{m} \mid y_{m-1}) \end{split}$$

Suppose we wanted to work with log probabilities instead. Then

$$\delta_{w,t} = \log p(x \mid y)$$

$$\delta_{y_m,y_{m-1}} = \log p(y_m \mid y_{m-1})$$

$$a \otimes b := a + b$$

$$\delta_{x,y_m,y_{m-1}} = \log p(x \mid y_m) \otimes \log p(y_m \mid y_{m-1})$$

$$= \log p(x \mid y_m) + \log p(y_m \mid y_{m-1})$$

$$= \log p(x, y_m \mid y_{m-1})$$

Can you see how many states the resulting FST will have?

- Finally, we compose with an acceptor S, which forms a chain for a sentence $w_1, \ldots w_M$.
- This composition $T \circ E \circ S$ yields a **trellis**-shaped weighted finite state acceptor (WFSA).
 - Number of columns = M, length of input.
 - Number of rows = T, number of tags.
 - Edges from states $\langle m, t_1 \rangle$ to $\langle m+1, t_2 \rangle$ have the score

$$\delta_{w_{m+1},t_2,t_1}^{(t\circ e)} = \delta_{t_2,t_1}^{(t)} \otimes \delta_{t_2,w_{m+1}}^{(e)} = P(Y_{m+1} = t_2 \mid Y_m = t_1) P(W_{m+1} = w_{m+1} \mid Y_{m+1} = t_2)$$
(12.8)

- Each path in the trellis corresponds to a unique sequence of tags, $y_{1:M}$, and every sequence of tags has a unique path. The score of the path is equal to $p(w_{1:M}, y_{1:M})$ by construction.
- If we define $\bigoplus = \max$ (as in the tropical semiring), then the score of the semiring shortest path is equal to $\max_{\boldsymbol{y}} p(\boldsymbol{w}_{1:M}, \boldsymbol{y}_{1:M})$.
- So, can we find this score (and therefore the best path) in polynomial time?
 - (c) Jacob Eisenstein 2014-2015. Work in progress.

- How expensive is it to construct the trellis?
 - Generic composition is polynomial but slower than we would like
 it depends on the vocabulary size.
 - * But since we know what the trellis is supposed to look like, we can just build it directly. This requires constant time per edge.
 - * How big is the trellis? $\mathcal{O}(MT)$ states, $\mathcal{O}(MT^2)$ edges.
- How expensive is it find the shortest path in the trellis?: Generic shorest path has a time cost of $\mathcal{O}(V \log V + E) = \mathcal{O}(MT \log MT + MT^2)$ and a space cost of $\mathcal{O}(V) = \mathcal{O}(MT^2)$.
- So:
 - * Building the trellis is polynomial.
 - * Shortest path is polynomial.
 - * Therefore, there must be a poly-time algorithm to find the best tag sequence, despite the apparently exponential number of paths.

12.3 The Viterbi algorithm

The Viterbi algorithm is a special-purpose best-path algorithm for FSTs in the shape of a trellis. It has a time cost of $\mathcal{O}(MT^2)$ and a space cost of $\mathcal{O}(MT)$. (This time cost improvement is important, because it is linear in the length of the sequence M, unlike the generic shortest-path algorithm, which is $M \log M$.

Based on the Markov assumption, we can decompose the likelihood recursively.

$$p(w_{1:M}, y_{1:M}) = p(w_M | y_M)p(y_M | y_{M-1})p(w_{1:M-1}, y_{1:M-1})$$

- Given y_{m-1} , we can choose y_m without considering any other element of the history.
- Suppose we know the best path to $y_m = k$. The best path to $y_{m+1} = k'$ through $y_m = k$ must include the best path to $y_m = k$.
- Suppose we know the score (probability) of the best path to each $y_m = k$, which we write $v_m(k) = \max_{y_1...y_{m-1}} p(\boldsymbol{w}_{1:m}, \boldsymbol{y}_{1:m-1}, y_m = k)$.
 - (c) Jacob Eisenstein 2014-2015. Work in progress.

What is the score of the best path to $y_{m+1} = k'$?

$$v_{m+1}(k') = \max_{\mathbf{y}_{1:m}} p(\mathbf{w}_{1:m+1}, \mathbf{y}_{1:m}, y_{m+1} = k')$$

$$= p_{e}(w_{m+1} \mid y_{m+1} = k') \max_{\mathbf{y}_{1:m}} P_{t}(Y_{m+1} = k' \mid y_{m}) p(\mathbf{w}_{1:m}, \mathbf{y}_{1:m})$$

$$= p_{e}(w_{m+1} \mid y_{m+1} = k') \max_{\mathbf{y}_{m} = k} P_{t}(Y_{m+1} = k' \mid Y_{m} = k) \max_{\mathbf{y}_{1:m-1}} p(\mathbf{w}_{1:m}, \mathbf{y}_{1:m-1}, y_{m} = k)$$

$$= p_{e}(w_{m+1} \mid y_{m+1} = k') \max_{\mathbf{y}_{m} = k} P_{t}(Y_{m+1} = k' \mid Y_{m} = k) v_{m}(k)$$

$$= p_{e}(w_{m+1} \mid y_{m+1} = k') \max_{\mathbf{y}_{m} = k} P_{t}(Y_{m+1} = k' \mid Y_{m} = k) v_{m}(k)$$

$$= p_{e}(w_{m+1} \mid y_{m+1} = k') \max_{\mathbf{y}_{m} = k} P_{t}(Y_{m+1} = k' \mid Y_{m} = k) v_{m}(k)$$

$$= p_{e}(w_{m+1} \mid y_{m+1} = k') \max_{\mathbf{y}_{m} = k} P_{t}(Y_{m+1} = k' \mid Y_{m} = k) v_{m}(k)$$

$$= p_{e}(w_{m+1} \mid y_{m+1} = k') \max_{\mathbf{y}_{m} = k} P_{t}(Y_{m+1} = k' \mid Y_{m} = k) v_{m}(k)$$

$$= p_{e}(w_{m+1} \mid y_{m+1} = k') \max_{\mathbf{y}_{m} = k} P_{t}(Y_{m+1} = k' \mid Y_{m} = k) v_{m}(k)$$

$$= p_{e}(w_{m+1} \mid y_{m+1} = k') \max_{\mathbf{y}_{m} = k} P_{t}(Y_{m+1} = k' \mid Y_{m} = k) v_{m}(k)$$

$$= p_{e}(w_{m+1} \mid y_{m+1} = k') \max_{\mathbf{y}_{m} = k} P_{t}(Y_{m+1} = k' \mid Y_{m} = k) v_{m}(k)$$

$$= p_{e}(w_{m+1} \mid y_{m+1} = k') \max_{\mathbf{y}_{m} = k} P_{t}(Y_{m+1} = k' \mid Y_{m} = k) v_{m}(k)$$

$$= p_{e}(w_{m+1} \mid y_{m+1} = k') \max_{\mathbf{y}_{m} = k} P_{t}(Y_{m+1} = k' \mid Y_{m} = k) v_{m}(k)$$

$$= p_{e}(w_{m+1} \mid y_{m+1} = k') \max_{\mathbf{y}_{m} = k} P_{t}(Y_{m+1} = k' \mid Y_{m} = k) v_{m}(k)$$

The base case is $v_0(\lozenge) = 1$, with zero probability for everything else.

We can generalize this recurrence using semiring notation:

$$v_{m+1}(k') = \delta_{w_{m+1}, y_{m+1} = k'}^{(e)} \otimes \bigoplus_{k} \delta_{k \to k'}^{(t)} \otimes v_m(k)$$
 (12.13)

Then if we want to move to log-probabilities, we have

$$v_{m+1}(k') = \log p_E(w_{m+1} \mid y_{m+1} = k') \otimes \bigoplus_k \log p_T(k \to k') \otimes v_m(k)$$
(12.14)
$$= \log p_E(w_{m+1} \mid y_{m+1} = k') + \max_k \log p_T(k \to k') + v_m(k)$$
(12.15)

- We will frequently use a semiring in which the edge weights are log probabilities and ⊗ is addition. This is partly because addition is notationally clearer than multiplication, and because in practical settings, you will use the log probabilities to avoid underflow.
- Note that we are setting $\oplus = \max$, as in the tropical semiring. This means that the score of the best tag sequence overall is $v_M(\square)$.
- To find the best tag sequence, we just need to keep back-pointers, from $v_m(k)$ to $v_{m-1}(k')$:

$$v_{m+1}(k') = \max_{k} \log \mathsf{p}_E(w_{m+1} \mid Y_{m+1} = k') + \log P_T(Y_{m+1} = k' \mid Y_m = k) + v_m(k) \tag{12.16}$$

$$= \log \mathsf{p}_E(w_{m+1} \mid y_{m+1} = k') + \left(\max_{k} \log P_T(Y_{m+1} = k' \mid Y_m = k) + v_m(k)\right) \tag{12.17}$$

$$b_{m+1}(k') = \arg \max_{k} \log \mathsf{p}_E(w_{m+1} \mid Y_{m+1} = k') + \log P_T(Y_{m+1} = k' \mid Y_m = k) + v_m(k) \tag{12.18}$$

$$= \arg \max_{k} \log P_T(Y_{m+1} = k' \mid Y_m = k) + v_m(k) \tag{12.19}$$

Note that the computation of the back-pointer doesn't depend on the emission probability $p_E(w_{m+1} \mid Y_{m+1} = k')$, since Y_m is conditionally independent from w_{m+1} given Y_{m+1} .

 In the probability semiring, we had ⊕ as addition; in the log-probability semiring, it was log addition. What happens if we try these addition operators? We'll see in a moment.

Example

Table 12.1: $\log p_e(w \mid y)$

	they	can	fish
N	-2	-3	-3
V	-10	-1	-3

Table 12.2: $\log p_t(y_m \mid y_{m-1})$

	N	V	END
START	-1	-2	$-\infty$
N	-3	-1	-2
V	-1	-3	-2

See the slides for how the Viterbi algorithm works in this example.

12.4 The forward algorithm

In an influential survey on HMMs, Rabiner (1989) defines three problems:

- **Decoding**: find the best tags *y* for a sequence *w*.
- **Likelihood**: compute the probability $p(w) = \sum_{y} p(w, y)$
- **Learning**: given only unlabeled data $\{w_1, w_2, \dots, w_D\}$, estimate the transition and emission distributions.

The Viterbi algorithm solves the decoding problem. We'll talk about the learning problem later. Let's now consider how to compute the likelihood $p(w) = \sum_{y} p(w, y)$.

• First, we move to a semiring where $a \oplus b = \log(e^a + e^b)$ instead of max. Then

$$\alpha_{m+1}(k') = \bigoplus_{k} \log p_{E}(w_{m+1} \mid Y_{m+1} = k') \otimes \log P_{T}(Y_{m+1} = k' \mid Y_{m} = k) \otimes \alpha_{m}(k)$$

$$(12.20)$$

$$= \log p_{E}(w_{m+1} \mid Y_{m+1} = k') \otimes \bigoplus_{k} \log P_{T}(Y_{m+1} = k' \mid Y_{m} = k) \otimes \alpha_{m}(k)$$

$$(12.21)$$

$$= \log p_{E}(w_{m+1} \mid Y_{m+1} = k') + \log \sum_{k} P_{T}(Y_{m+1} = k' \mid Y_{m} = k) \times e^{\alpha_{m}(k)}$$

$$(12.22)$$

$$= \log p_{E}(w_{m+1} \mid Y_{m+1} = k') + \log \sum_{k} P_{T}(Y_{m+1} = k' \mid Y_{m} = k) p(\mathbf{w}_{1:m}, Y_{m} = k)$$

$$(12.23)$$

$$= \log p_{E}(\mathbf{w}_{m+1} \mid Y_{m+1} = k') + \log P_{T}(Y_{m+1} = k', \mathbf{w}_{1:m})$$

$$= \log p(\mathbf{w}_{1:m+1}, Y_{m+1} = k')$$

$$(12.24)$$

$$= \log p(\mathbf{w}_{1:m+1}, Y_{m+1} = k')$$

$$(12.25)$$

- We used the inductive hypothesis in (12.23), and we used the HMM conditional independence assumptions $W_{m+1} \perp W_{1:m} \mid Y_{m+1}$ and $Y_{m+1} \perp W_{1:m} \mid Y_m$ in the following two steps.
- We can formalize this as an inductive proof by stating the base case,

$$\alpha_1(k) = \log p_e(w_1 \mid y_1) \otimes \log P_t(Y_1 = k \mid \Diamond) = \log p(w_1, Y_1 = k \mid Y_0 = \Diamond).$$
(12.26)

This is called the **forward** algorithm. The total probability of a sequence is $p(\mathbf{w}_{1:M}) = \alpha_M(\square)$. For a demo, see the slides.

Why solve the likelihood problem?

Why would we want to compute $p(w_{1:M})$?

Word class language models

- Remember p(colorless green ideas sleep furiously)
- We don't care about the specific tags, we just want to know the probability of the utterance, so we can compare it with p(*Furiously sleep ideas green colorless*).

Comparing HMMs

- Suppose we have a few HMMs, each of which could have generated the observations.
- We want to compute the marginal likelihood of the observations given each HMM, regardless of the path.
- This approach is sometimes used in gesture recognition.

Computing marginals: we'll soon be very interested in **marginal** probabilities $p(y_m | w_{1:M})$, for all $m \le M$. The likelihood $p(w_{1:M})$ is part of this computation.

12.5 HMM Details

Trigrams

- Can we use trigrams instead of bigrams for an HMM?
- How do we change the trellis? How big is the new trellis?
- Each cell represents a pair of tags, $\langle y_m, y_{m-1} \rangle$.
- The trellis still needs N columns, but now need T^2 rows.
- Each node can only connect to T neighbors in the next column, based on the trigram transition constraint. Number of edges = NT^3 .
 - (c) Jacob Eisenstein 2014-2015. Work in progress.

• We can prune very low probability edges for speed.

Estimation

In principle, we can use relative frequency estimation.

$$\lambda_{k,k'} \triangleq P_T(Y_m = k' \mid Y_{m-1} = k) = \frac{\text{count}(Y_m = k', Y_{m-1} = k)}{\text{count}(Y_{m-1} = k)}$$
$$\phi_{k,i} \triangleq P_E(W_m = i \mid Y_m = k) = \frac{\text{count}(W_m = i, Y_m = k)}{\text{count}(Y_m = k)}$$

In practice, we need smoothing and other tricks.

- The same smoothing ideas from language modeling can be applied.
 - interpolate between trigrams, bigrams, and unigrams

$$P(y_m \mid y_{m-1}, y_{m-2}) = \lambda_2 \hat{P}_2(y_m \mid y_{m-1}, y_{m-2}) + \lambda_1 \hat{P}_1(y_m \mid y_{m-1}) + (1 - \lambda_2 - \lambda_1) \hat{P}_0(y_m)$$

- reserve probability mass for unseen words

12.6 Tagging with features

Let's consider an example:

'Twas brillig, and the slithy toves Did gyre and gimble in the wabe: All mimsy were the borogoves, And the mome raths outgrabe

You probably don't know many of these words, yet it's not so hard to see what some of their tags should be. How do you do it?

Recall that the HMM can incorporate two sources of information:

- Word-tag probabilities, via $p_E(w_m|y_n)$.
- Local context, via $p_T(y_m|y_{m-1})$.

Local context is helpful, but the word-tag probabilities will be worthless for words like *brillig*, *slithy*, *toves*, *gyre*, etc.

But there are a lot of things that we're missing here. Crucially, unseen words have internal structure which can be exploited: *mimsy*, 3.1415, *Ke\$ha*

• Orthography and morphology. *Slithy toves* just kind of looks like JJ NNS, because of the apparent suffixes.

```
- -s \rightarrow NNS, VBZ

- -able \rightarrow JJ

- -ly \rightarrow RB

- un- \rightarrow JJ, RB, V

- ...
```

But morphological features are difficult to incorporate in a generative model, because they break the Naive Bayes assumption:

$$p(mimsy, -sy \mid JJ) \neq p(mimsy \mid JJ)p(-sy \mid JJ)$$

- **Capitalization**. This is especially relevant for named entity recognition, e.g., *I bought an apple* and *I bought an Apple phone*.
- More advanced HMMs incorporate morphological, orthographic, and typographic features by creating a more complex $p_E(w \mid y)$ emission probability. For example, the TNT Tagger took this approach, and is one of the best generative taggers (Brants, 2000).

In addition to word-internal features, we might want more fine-grained context. For example, in the PTB, *this* and *these* are both tagged DT. But *this* is likely to be followed by a singular noun NN, and *these* is likely to be followed by a plural noun NNS. So we might like to add word-context features to the probability $p(y_m \mid y_{m-1}, w_{m-1})$.

How can we incorporate these overlapping features? The solution is to build sequence labeling models based on the perceptron and logistic regression classifiers. The first model is called **structured perceptron**, since the label space consists of structures rather than individual labels (Collins, 2002). The second model is called a **conditional random field (CRF)**, due to its relation to Markov random fields (Lafferty et al., 2001). In this model, we explicitly compute $p(y \mid w)$.

In addition to incorporating overlapping features, these models have another advantage: they are discriminative, directly maximizing the conditional probability $p(\boldsymbol{y} \mid \boldsymbol{w})$, or minimizing the perceptron loss. As in standard classification, this criterion is more closely connected to the accuracy metrics that we usually care about.

12.7 Structured perceptron

Remember the perceptron update:

$$\hat{y} = \arg \max_{y} \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{w}, y)$$

$$\boldsymbol{\theta}^{(t+1)} \leftarrow \boldsymbol{\theta}^{(t)} + \boldsymbol{f}(\boldsymbol{w}, y^{*}) - \boldsymbol{f}(\boldsymbol{w}, \hat{y})$$
(12.28)

$$\boldsymbol{\theta}^{(t+1)} \leftarrow \boldsymbol{\theta}^{(t)} + \boldsymbol{f}(\boldsymbol{w}, y^*) - \boldsymbol{f}(\boldsymbol{w}, \hat{y})$$
 (12.28)

In sequence labeling, we have a **structured output** $y \in \mathcal{T}(w)$. Can we still apply the perceptron rule?

$$\hat{\boldsymbol{y}} = \arg\max_{\boldsymbol{y} \in \mathcal{T}(\boldsymbol{w})} \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{w}, \boldsymbol{y})$$
 (12.29)

$$\boldsymbol{\theta}^{(t+1)} \leftarrow \boldsymbol{\theta}^{(t)} + \boldsymbol{f}(\boldsymbol{w}, \boldsymbol{y}^*) - \boldsymbol{f}(\boldsymbol{w}, \hat{\boldsymbol{y}})$$
 (12.30)

This is called structured perceptron, because it learns to predict structured output y. The problem is that $\arg \max_{y}$ must search over the entire set of structures \mathcal{T} . The set of permissible outputs depends on the input w, and it is very large: $\mathcal{O}(K^M)$. What can we do?

Viterbi inference for structured perceptron

We will place a restriction on the scoring function $\theta^T f(w, y)$, which allows us to apply the Viterbi algorithm to find $\arg \max_{\boldsymbol{y}} \boldsymbol{\theta}^{\mathsf{T}} f(\boldsymbol{w}, \boldsymbol{y})!$

- Specifically, we require $\theta^T f(w, y) = \sum_m \theta^T f(w, y_m, y_{m-1}, m)$
- That is, the global score must be a **sum of local scores**.
- The local scores can consider any part of the observation, but must only consider adjacent elements in the label.

To apply Viterbi to structured perceptron, we have

$$v_m(k) = \bigoplus_{k'} \boldsymbol{\theta}^\mathsf{T} \boldsymbol{f}_m(\boldsymbol{w}, k, k', m) \otimes v_{m-1}(k')$$
$$= \max_{k'} \boldsymbol{\theta}^\mathsf{T} \boldsymbol{f}_m(\boldsymbol{w}, k, k', m) + v_{m-1}(k')$$
$$b_m(k) = \arg \max_{k'} \boldsymbol{\theta}^\mathsf{T} \boldsymbol{f}_m(\boldsymbol{w}, k, k', m) + v_{m-1}(k')$$

Suppose we want to apply this to POS tagging? What features might we want? Here are some:

- Word-tag features, e.g. $\langle W : slithy, JJ \rangle$
- Adjacent tag-tag features, e.g. $\langle T : JJ, NNS \rangle$
- Suffix-tag features, e.g., $\langle M : -es, NNS \rangle$
- Previous-word features, e.g., $\langle P_1 : the, JJ \rangle$
- Next-word features, e.g., $\langle N_1 : slithy, DT \rangle$
- Note that we can consider arbitrarily distant words, e.g. $\langle Y_m, W_{m-15} \rangle$, because this still fits in the constraint, $\boldsymbol{\theta}^\mathsf{T} \boldsymbol{f}(\boldsymbol{w}, \boldsymbol{y}) = \sum_m \boldsymbol{\theta}^\mathsf{T} \boldsymbol{f}(\boldsymbol{w}, y_m, y_{m-1}, m)$.

So to compute $v_m(k)$, we have to iterate over all $y_{m-1} = k'$,

- find the features $f(x_{1:M}, y_m = k, y_{m-1} = k', m)$,
- compute the inner product $\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{x}_{1:M}, y_m = k, y_{m-1} = k', m)$,
- add it to $v_{m-1}(k')$
- take the max over all k'

This only works because of the assumption that the feature function decomposes over local parts of the sequence! If we wanted a feature that considered arbitrary parts of tag sequence, there would be no way to incorporate it into the recurrence relation.

Example

$$w = \dots$$
 and the slithy toves $y = \dots$ CC DT IJ NNS

Then we can characterize the tagging DT JJ NNS of the text the slithy toves (from Jabberwocky) in terms of the following features:

```
 \begin{split} \textbf{\textit{f}} (\text{the slithy toves}, \operatorname{DT} \operatorname{JJ} \operatorname{NNS}) = & \{ \langle W: \operatorname{the}, \operatorname{DT} \rangle, \langle M:\varnothing, \operatorname{DT} \rangle, \langle T:\diamondsuit, \operatorname{DT} \rangle \\ & \langle W: \operatorname{slithy}, \operatorname{JJ} \rangle, \langle M: \operatorname{-thy}, \operatorname{JJ} \rangle, \langle T: \operatorname{DT}, \operatorname{JJ} \rangle \\ & \langle W: \operatorname{toves}, \operatorname{NNS} \rangle, \langle M: \operatorname{-es}, \operatorname{NNS} \rangle, \langle T: \operatorname{JJ}, \operatorname{NNS} \rangle \\ & \langle T: \operatorname{NNS}, \square \rangle \}  \end{split}
```

If this is the correct tagging, then we hope to learn a set of weights w such that $\theta^T f$ (the slithy toves, DT JJ NNS) is larger than the scores for other tag sequences, such as $\theta^T f$ (the slithy toves, DT NN VBZ).

Learning

If we define $f(w_{1:M}, y_m, y_{m-1}, m) = \{\langle W : w_m, y_m \rangle, \langle T : y_{m-1}, y_m \rangle\}$, then our model is identical to the HMM. If we set the weights of these features to the log of their maximum-likelihood estimates,

$$w_{\langle W:w_m,y_m\rangle} = \log \operatorname{count}(w_m,y_m) - \log \operatorname{count}(y_m)$$

$$w_{\langle T:y_{m-1},y_m\rangle} = \log \operatorname{count}(y_{m-1},y_m) - \log \operatorname{count}(y_{m-1}),$$

then we exactly recover the HMM.

But to use more overlapping features and to get the advantages of error-driven learning, we're going to do perceptron updates. It's exactly the same as the non-structured perceptron:

- $\boldsymbol{\theta}^{(t+1)} \leftarrow \boldsymbol{\theta}^{(t)} + \boldsymbol{f}(\boldsymbol{x}, \boldsymbol{y}) \boldsymbol{f}(\boldsymbol{x}, \hat{\boldsymbol{y}})$ is the standard update, using Viterbi to find $\hat{\boldsymbol{y}}$.
- As before, weight averaging is crucial to get good performance (Collins, 2002).
- As before, we can use Passive-Aggressive to do large-margin training (Crammer et al., 2006), computing the step size by dividing a non-negative **loss** $\ell(\boldsymbol{y}_i, \hat{\boldsymbol{y}})$ by the squared norm of the difference in the feature vectors, $||\boldsymbol{f}(\boldsymbol{y}_i, \boldsymbol{w}_i) \boldsymbol{f}(\hat{\boldsymbol{y}}, \boldsymbol{w}_i)||^2$. A reasonable choice of loss function is the Hamming loss, which is the number of incorrect tag predictions (Taskar et al., 2003; Tsochantaridis et al., 2004).

12.8 Conditional random fields

Structured perceptron works well in practice, and you will implement in your project 2, where it should work much better than the Hidden Markov Model.

- But sometimes we need probabilities, and SP doesn't give us that.
- The Conditional Random Field (CRF) is a probabilistic conditional model for sequence labeling.

 Just as structured perceptron is built on the perceptron classifier, conditional random fields are built on the logistic regression classifier.

$$p(y \mid \boldsymbol{x}) = \frac{e^{\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y, \boldsymbol{x})}}{\sum_{y' \in \mathcal{Y}} e^{\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y', \boldsymbol{x})}}$$
(12.31)

We can again move to structured prediction,

$$p(\boldsymbol{y} \mid \boldsymbol{w}) = \frac{e^{\boldsymbol{\theta}^\mathsf{T} \boldsymbol{f}(\boldsymbol{y}, \boldsymbol{w})}}{\sum_{\boldsymbol{y}' \in \mathcal{T}(\boldsymbol{w})} e^{\boldsymbol{\theta}^\mathsf{T} \boldsymbol{f}(\boldsymbol{y}', \boldsymbol{w})}}.$$

This is called a Conditional Random Field, because it models the sequence labeling task as a Markov random field, and estimates the probability of a set of variables **conditioned** on the others (as opposed to jointly, which the HMM does). We will need the same restriction on the scoring function as in Structured Perceptron: $\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{w}, \boldsymbol{y}) = \sum_{m} \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{w}, y_m, y_{m-1}, m)$.

Decoding in CRFs

Decoding in the CRF does not depend on the denominator of $p(y \mid w)$

$$\begin{split} \hat{\boldsymbol{y}} &= \arg\max_{\boldsymbol{y}} p(\boldsymbol{y} \mid \boldsymbol{w}) \\ &= \arg\max_{\boldsymbol{y}} \log p(\boldsymbol{y} \mid \boldsymbol{w}) \\ &= \arg\max_{\boldsymbol{y}} \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{y}, \boldsymbol{w}) - \log \sum_{\boldsymbol{y}' \in \mathcal{T}(\boldsymbol{w})} e^{\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{y}', \boldsymbol{w})} \\ &= \arg\max_{\boldsymbol{y}} \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{y}, \boldsymbol{w}) \\ &= \arg\max_{\boldsymbol{y}} \sum_{m=0}^{M} \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}_{m}(\boldsymbol{w}, y_{m}, y_{m-1}, m) \end{split}$$

So we can just apply Viterbi directly, as defined in the Structured Perceptron.

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Learning in CRFs

Learning is a little more complicated. As with logistic regression, we need to learn weights to minimize the regularized negative log conditional probability,

$$\begin{split} \ell &= \sum_{i=1}^{N} -\log p(\boldsymbol{y}_i \mid \boldsymbol{w}_i; \boldsymbol{\theta}) + \lambda ||\boldsymbol{\theta}||^2, \\ &= -\sum_{i} \boldsymbol{\theta}^\mathsf{T} \boldsymbol{f}(\boldsymbol{w}_i, \boldsymbol{y}_i) + \log \sum_{\boldsymbol{y}' \in \mathcal{T}(\boldsymbol{w}_i)} \exp \left(\boldsymbol{\theta}^\mathsf{T} \boldsymbol{f}(\boldsymbol{w}_i, \boldsymbol{y}')\right) + \lambda ||\boldsymbol{\theta}||^2, \end{split}$$

where λ controls the amount of regularization. As in logistic regression, the gradient includes is a difference between observed and expected counts:

$$\frac{d\ell}{d\theta_j} = \sum_{i} \text{count}(\boldsymbol{w}_i, \boldsymbol{y}_i)_j - E_{\boldsymbol{y}|\boldsymbol{w}_i;\boldsymbol{\theta}}[\text{count}(\boldsymbol{w}_i, \boldsymbol{y})_j] + \lambda \boldsymbol{\theta}$$

$$\text{count}(\boldsymbol{w}_i, \boldsymbol{y}_i)_j = \sum_{m}^{M} f_j(\boldsymbol{w}_i, y_{i,m}, y_{i,m-1}, m)$$

For example:

- If feature j is $\langle T : CC, DT \rangle$, then $count(\boldsymbol{w}_i, \boldsymbol{y}_i)_j$ is the count of times DT follows CC in the sequence \boldsymbol{y}_i .
- If feature j is $\langle M : -thy, JJ \rangle$, then $count(w_i, y_i)_j$ is the count of words ending in -thy in w_i that are tagged JJ in y_i .

The expected feature counts are more difficult to compute.

$$E_{\boldsymbol{y}|\boldsymbol{w};\boldsymbol{\theta}}[\operatorname{count}(\boldsymbol{w}_i,\boldsymbol{y})_j] = \sum_{\boldsymbol{y}\in\mathcal{T}(\boldsymbol{w}_i)} P(\boldsymbol{y}\mid\boldsymbol{w}_i;\boldsymbol{\theta}) f_j(\boldsymbol{w}_i,\boldsymbol{y})$$
(12.32)

- This looks bad: we have to sum over an exponential number of labelings again.
- But remember that the feature function decomposition implies that

$$f_j(\boldsymbol{w}, \boldsymbol{y}) = \sum_m f_j(\boldsymbol{w}, y_m, y_{m-1}, m)$$
(12.33)

This means we can compute the expectation as,

$$E_{\boldsymbol{y}|\boldsymbol{w};\boldsymbol{\theta}}[\operatorname{count}(\boldsymbol{w},\boldsymbol{y})_j] = \sum_{\boldsymbol{y}\in\mathcal{T}(\boldsymbol{w})} P(\boldsymbol{y}\mid\boldsymbol{w};\boldsymbol{\theta}) f_j(\boldsymbol{w},\boldsymbol{y})$$
(12.34)

$$= \sum_{\boldsymbol{y} \in \mathcal{T}(\boldsymbol{w})} P(\boldsymbol{y} \mid \boldsymbol{w}; \boldsymbol{\theta}) \sum_{m}^{M} f_{j}(\boldsymbol{w}, y_{m}, y_{m-1}, m)$$
(12.35)

$$= \sum_{m}^{M} \sum_{\boldsymbol{y} \in \mathcal{T}(\boldsymbol{w})} P(\boldsymbol{y} \mid \boldsymbol{w}; \boldsymbol{\theta}) f_{j}(\boldsymbol{w}, y_{m}, y_{m-1}, m)$$
(12.36)

$$= \sum_{m}^{M} \sum_{k,k' \in \mathcal{T}} \sum_{\boldsymbol{y} \in \mathcal{T}(\boldsymbol{w}): y_{m-1} = k, y_{m} = k'} P(\boldsymbol{y} \mid \boldsymbol{w}; \boldsymbol{\theta}) f_{j}(\boldsymbol{w}, k', k, m)$$
(12.37)

$$= \sum_{m}^{M} \sum_{k,k' \in \mathcal{T}} f_j(\boldsymbol{w}, k', k, m) \sum_{\boldsymbol{y} \in \mathcal{T}(\boldsymbol{w}): y_{m-1} = k, y_m = k'} P(\boldsymbol{y} \mid \boldsymbol{w}; \boldsymbol{\theta})$$
(12.38)

$$= \sum_{m}^{M} \sum_{k,k' \in \mathcal{T}} f_j(\boldsymbol{w}, k, k', m) P(y_{m-1} = k, y_m = k' \mid \boldsymbol{w}; \boldsymbol{\theta})$$
(12.39)

- The expected feature counts can be computed efficiently if we know the marginal probabilities $P(Y_m = k', Y_{m-1} = k \mid \boldsymbol{w}; \boldsymbol{\theta})$.
- This is the probability of traversing the trellis edge $\langle m-1,k\rangle \to \langle m,k'\rangle$, conditioned on the entire observation $w_{1:M}$. [Draw this in trellis]
- To compute this marginal probability, we will apply the forward-backward algorithm.

12.9 The forward-backward algorithm

Here we require a marginal probability, e.g. $P(Y_m = \text{NNP}, Y_{m-1} = \text{DT} \mid \boldsymbol{w}_{1:M})$. We can rewrite the marginal probability of a single tag transition $P(Y_m = k', Y_{m-1} = k \mid \boldsymbol{w}_{1:M})$ as a ratio:

$$P(Y_m = k', Y_{m-1} = k | \mathbf{w}_{1:M}) = \frac{P(Y_m = k', Y_{m-1} = k, \mathbf{w}_{1:M})}{P(\mathbf{w}_{1:M})}$$
(12.40)

Recall the Forward Algorithm, which defines the forward probabilities $\alpha_n(k) \triangleq P(Y_m = k, \boldsymbol{w}_{1:m})$. This allows us to compute the denominator $P(w_{1:M}) = \alpha_M(\square)$, with \square as the special final tag.

Forward algorithm

First we need to show how to use the forward algorithm in a CRF. We're going to switch to a semiring in which we work with probabilities rather than log-probabilities, so that we can use simple addition rather than log-addition. But in practice we'd work in a log-probability semiring, where $a \oplus b = \log(e^a + e^b)$. For a practical implementation, see <code>numpy.logaddexp</code> and <code>scipy.misc.logsumexp</code>.

$$\alpha_{m}(k) \triangleq P(Y_{m} = k, w_{1:m}) = \sum_{\boldsymbol{y}_{1:m}:y_{m}=k} p(\boldsymbol{y}_{1:m}, \boldsymbol{w}_{1:m})$$

$$= \sum_{\boldsymbol{y}_{1:m}:y_{m}=k} \exp(\sum_{n=0}^{m} \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{w}, y_{n}, y_{n-1}, n))$$

$$= \sum_{\boldsymbol{y}_{1:m}:y_{m}=k} \exp(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{w}, y_{m} = k, y_{m-1}, m)) \exp(\sum_{n=0}^{m-1} \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{n}, y_{n-1}, \boldsymbol{w}, n))$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{w}, y_{m} = k, y_{m-1} = k', m))$$

$$\times \sum_{\boldsymbol{y}_{1:m-1}:y_{m-1}=k'} \exp(\sum_{n=0}^{m-1} \boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{w}, y_{n}, y_{n-1}, n))$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{m} = k, y_{m-1} = k', \boldsymbol{w}, m)\right) \alpha_{m-1}(k')$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{m} = k, y_{m-1} = k', \boldsymbol{w}, m)\right) \alpha_{m-1}(k')$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{m} = k, y_{m-1} = k', \boldsymbol{w}, m)\right) \alpha_{m-1}(k')$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{m} = k, y_{m-1} = k', \boldsymbol{w}, m)\right) \alpha_{m-1}(k')$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{m} = k, y_{m-1} = k', \boldsymbol{w}, m)\right) \alpha_{m-1}(k')$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{m} = k, y_{m-1} = k', \boldsymbol{w}, m)\right) \alpha_{m-1}(k')$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{m} = k, y_{m-1} = k', \boldsymbol{w}, m)\right) \alpha_{m-1}(k')$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{m} = k, y_{m-1} = k', \boldsymbol{w}, m)\right) \alpha_{m-1}(k')$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{m} = k, y_{m-1} = k', \boldsymbol{w}, m)\right) \alpha_{m-1}(k')$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{m} = k, y_{m-1} = k', \boldsymbol{w}, m)\right) \alpha_{m-1}(k')$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{m} = k, y_{m-1} = k', \boldsymbol{w}, m)\right) \alpha_{m-1}(k')$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{m} = k, y_{m-1} = k', \boldsymbol{w}, m)\right) \alpha_{m-1}(k')$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{m} = k, y_{m-1} = k', \boldsymbol{w}, m)\right) \alpha_{m-1}(k')$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{m} = k, y_{m-1} = k', \boldsymbol{w}, m)\right) \alpha_{m-1}(k')$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{m} = k, y_{m-1} = k', \boldsymbol{w}, m)\right) \alpha_{m-1}(k')$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{m} = k, y_{m-1} = k', \boldsymbol{w}, m)\right) \alpha_{m-1}(k')$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'} \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(y_{m} = k, y_{m-1} = k', \boldsymbol{w}, m)\right) \alpha_{m-1}(k')$$

$$= \sum_{\boldsymbol{y}_{m-1}=k'}$$

Here's what's happening:

- In the first step (12.41), we introduce $y_{1:m-1}$ and then marginalize it out.
- In (12.42), we use the fact that the conditional probability $p(y \mid w) = \frac{p(y,w)}{\sum_{y'} p(y',w)}$, so that the joint probability p(y,w) can be computed without the denominator.

- We then make use of the feature decomposition to isolate the contribution of $\theta^T f(w, y_m, y_{m-1}, m)$. This is sometimes called a **potential**.
- The rest of the derivation follows the playback for the Forward algorithm in Hidden Markov Models: introduce $Y_{m-1} = k'$, and obtain the expression for $\alpha_{m-1}(k')$, which we compute recursively.

An example of this is found in the slides. The feature weights are shown in the slide. So to compute the forward variable for *can* tagged as N, we have,

$$\alpha_2(N) = (-4 \otimes -3 \otimes -3) \oplus (-5 \otimes -1 \otimes -3) \tag{12.48}$$

$$=(-10) \oplus (-9) \tag{12.49}$$

$$=\log\left(e^{-10} + e^{-9}\right) \tag{12.50}$$

$$\approx -8.7\tag{12.51}$$

Note that this is greater than the Viterbi score $v_2(N) = -9$, because it is the log of the sum of probabilities which **include** the Viterbi path.

Now we return to the numerator of $P(Y_m = k', Y_{m-1} = k \mid \boldsymbol{w}_{1:M}) = \frac{P(Y_m = k', Y_{m-1} = k, \boldsymbol{w}_{1:M})}{p(\boldsymbol{w}_{1:M})}$. We can factor this probability as,

$$P(Y_{m} = k', Y_{m-1} = k, \mathbf{w}_{1:M}) = p(\mathbf{w}_{m+1:M} \mid Y_{m} = k', Y_{m-1} = k, \mathbf{w}_{1:M})$$

$$\times P(Y_{m} = k', Y_{m-1} = k, \mathbf{w}_{1:M}) \qquad (12.52)$$

$$= p(\mathbf{w}_{m+1:M} \mid Y_{m} = k') P(Y_{m} = k', \mathbf{w}_{m} \mid Y_{m-1} = k)$$

$$\times P(Y_{m-1} = k, \mathbf{w}_{1:m-1}) \qquad (12.53)$$

$$\propto p(\mathbf{w}_{m+1:M} \mid Y_{m} = k') \exp\left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(k', k, \mathbf{w}, m)\right)$$

$$\times \alpha_{m-1}(k). \qquad (12.54)$$

Here's what's going on:

- We have recursively plugged in the forward variable $\alpha_{m-1}(k) = p(\boldsymbol{w}_{1:m-1}, Y_{m-1} = k)$.
- We have computed

$$P(Y_m = k', w_m \mid Y_{m-1} = k) \propto \exp\left(\boldsymbol{\theta}^\mathsf{T} \boldsymbol{f}(k', k, \boldsymbol{w}, m)\right), \tag{12.55}$$

• Now we define a **backward variable**, $b_m(k') = P(w_{m+1:M} \mid y_m = k')$, yielding

$$P(Y_m = k', Y_{m-1} = k \mid w_{1:M}) \propto b_m(k') \exp\left(\boldsymbol{\theta}^\mathsf{T} \boldsymbol{f}(\boldsymbol{w}, k', k, m)\right) \alpha_{m-1}(k).$$

If we want to compute the marginal for a single tag $y_m = k$, we have

$$P(Y_m = k \mid \mathbf{w}_{1:M}) \propto b_m(k)\alpha_m(k) = \mathbf{p}(\mathbf{w}_{m+1:M} \mid y_m)P(Y_m = k, \mathbf{w}_{1:m}) = P(Y_m = k, \mathbf{w}_{1:M})$$
(12.56)

This marginal probability is useful for applying expectation-maximization to unsupervised learning of HMMs. Specifically, we can compute the expected count

$$E[\text{count}(w=i, y=k)] = \sum_{m} P(Y_m = k \mid \mathbf{w}_{1:M}) \delta(w_m = i),$$
 (12.57)

which we then use in the M-step to estimate $\phi_{k,i} = \frac{E[\text{count}(w=i,y=k)]}{E[\text{count}(y=k)]}$.

The backwards algorithm

We still need to compute the backward variables. Fortunately, the backwards probability also decomposes into a recurrence relation,

$$b_{m}(k) \triangleq p(\boldsymbol{w}_{m+1:M} \mid Y_{m} = k)$$

$$= \sum_{k'} P(\boldsymbol{w}_{m+1:M}, Y_{m+1} = k' \mid Y_{m} = k)$$

$$= \sum_{k'} P(\boldsymbol{w}_{m+2:M} \mid Y_{m+1} = k', w_{m+1}) P(w_{n+1}, Y_{m+1} = k' \mid y_{m} = k)$$

$$\propto \sum_{k'} b_{m+1}(k') \exp \left(\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{f}(\boldsymbol{w}_{1:M}, y_{m+1}, y_{m}, m)\right)$$

$$b_{M}(\square) = 1$$

The steps are as follows:

- introduce $Y_{m+1} = k'$ and sum over all possible values
- apply the chain rule to extract Y_{m+1} and w_{m+1}
- $\bullet\,$ use the definition of the backwards probability and the model assumptions

The backwards algorithm moves from right to left, starting with $B_N(y_M) = P(\text{END} \mid y_M)$, the score for transitioning to the end state given y_N .

We can see an example in the slides.

Learning in CRFs: wrapup

The overall procedure looks like logistic regression:

• Use forward-backward to compute expected feature counts under $P(y \mid w; \theta)$:

$$E[f_j(\mathbf{y}, \mathbf{w})] = \sum_{m} \sum_{k,k'} P(Y_m = k, Y_{m-1} = k' | \mathbf{w}_{1:M}) f_j(\mathbf{w}, k, k', m)$$
(12.58)

$$= \sum_{m} \sum_{k,k'} \frac{\alpha_m(k)\beta_{m+1}(k')}{\alpha_M \square} f_j(\boldsymbol{w}, k, k', m).$$
 (12.59)

- Compute gradients as difference between feature counts and expected counts, f(y, vw) E[f(y, w)].
- Update θ , using quasinewton optimization, stochastic gradient descent, or AdaGrad.
- Iterate, recomputing the expected feature counts.

12.10 Unsupervised sequence labeling

In unsupervised sequence labeling, we want to induce a Hidden Markov Model (HMM) from a corpus of unannotated text $w_1, w_2, ..., w_N$. The resulting tags might be useful for some downstream task, or for better understanding the language's inherent structure; or, we might want to do probability density estimation for sequences. Another reason would be to do semi-supervised learning, imputing tag sequences for unlabeled data. For part-of-speech tagging, often we use a tag dictionary which lists the allowed tags for each word, simplifying the problem (Christodoulopoulos et al., 2010).

In any case, we can perform unsupervised learning by applying expectation-maximization. In the M-step, we compute the HMM parameters from expected counts:

$$\begin{split} \mathbf{p}_{E}(w = i | y = k) &= \phi_{k,i} = \frac{E[\mathsf{count}(w = i, y = k)]}{E[\mathsf{count}(y = k)]} \\ \mathbf{p}_{T}(y_{m} = k | y_{m-1} = k') &= \lambda_{k',k} = \frac{E[\mathsf{count}(y_{m} = k, y_{m-1} = k')]}{E[\mathsf{count}(y_{m-1} = k')]} \end{split}$$

The expected counts are computed in the E-step, using the forward and backward variables.

$$E[\operatorname{count}(W=i,Y=k)] = \sum_{m} P(Y_{m}=k|\boldsymbol{w})\delta(W_{m}=i)$$

$$= \sum_{m} \frac{P(Y_{m}=k,\boldsymbol{w}_{1:m})p(\boldsymbol{w}_{m+1:M}|Y_{m}=k)}{p(\boldsymbol{w}_{1:M})}\delta(w_{m}=i)$$

$$= \frac{1}{\alpha_{M}(\square)} \sum_{m} \alpha_{m}(k)\beta_{m}(k)\delta(w_{m}=i)$$

$$(12.62)$$

We use the chain rule to separate $w_{1:m}$ and $w_{m+1:M}$, and then use the definitions of the forward and backward variables. In the final step, we normalize by $p(w_{1:M}) = \alpha_M(\square) = \beta_0(\lozenge)$.

$$E[\operatorname{count}(Y_{m} = k, Y_{m-1} = k')] = \sum_{m} P(Y_{m} = k, Y_{m-1} = k' | \boldsymbol{w})$$

$$\propto \sum_{m} P(Y_{m-1} = k', \boldsymbol{w}_{1:m-1}) P(w_{m+1:M} | Y_{m} = k)$$

$$\times P(w_{m}, Y_{m} = k | Y_{m-1} = k')$$

$$= \sum_{m} P(Y_{m-1} = k', \boldsymbol{w}_{1:m-1}) P(w_{m+1:M} | Y_{m} = k)$$

$$\times P(w_{m} | Y_{m} = k) P(Y_{m} = k | Y_{m-1} = k')$$

$$\times P(w_{m} | Y_{m} = k) P(Y_{m} = k | Y_{m-1} = k')$$

$$= \sum_{m} \alpha_{m-1}(k') \beta_{m}(k) \phi_{k,w_{m}} \lambda_{k' \to k}$$

$$(12.68)$$

Again, we use the chain rule to separate out $w_{1:m-1}$ and $w_{m+1:M}$, and use the definitions of the forward and backward variables. The final computation also includes the parameters ϕ and λ , which govern (respectively) the emission and transition properties between w_m, y_m , and y_{m-1} . Note that the derivation only shows how to compute this to a constant of proportionality; we would divide by $p(w_{1:M})$ to go from the joint probability $P(Y_{m-1} = k', Y_m = k, \mathbf{w}_{1:M})$ to the desired conditional $P(Y_{m-1} = k', Y_m = k \mid \mathbf{w}_{1:M})$.

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