

2017 秋概率论与数理统计 A 答案

一.选择题 (每道题 3 分, 共 15 分)

1.A 2.B 3.B 4.C 5.D

二.填空题 (每道题 3 分, 共 15 分)

1. 0.2; 2. $\frac{2e^y}{\pi(1+e^{2y})}$; 3. 10; 4. 0.95; 5. 32.917, 拒绝原假设, 认为各台机器生

产的薄板厚度有显著差异。

三 (6 分)

解: (1) 设 B = “主人回来树还活着”, 再设 A = “邻居记得浇水”, 则由全概率公式有

$$P(B) = P(A)P(B|A) + P(\bar{A})P(B|\bar{A}) = 0.9 \times (1 - 0.1) + 0.1 \times (1 - 0.8) = 0.83$$

—————3 分

$$(2) P(\bar{A}|\bar{B}) = \frac{P(\bar{A})P(\bar{B}|\bar{A})}{P(\bar{B})} = \frac{0.1 \times 0.8}{1 - 0.83} = \frac{8}{17} = 0.471$$

—————6 分

四. (9 分) 解: (1) $S(D) = \int_0^1 dx \int_{x^2}^{\sqrt{x}} dy = \frac{1}{3}$

$$f(x, y) = \begin{cases} 3, & (x, y) \in D \\ 0, & \text{其他} \end{cases}$$

—————2 分

(2)

$$f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \begin{cases} \int_{x^2}^{\sqrt{x}} 3 dx = 3(\sqrt{x} - x^2), & 0 < x < 1, \\ 0, & \text{其他} \end{cases}$$

$$f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \begin{cases} \int_{y^2}^{\sqrt{y}} 3 dx = 3(\sqrt{y} - y^2), & 0 < y < 1 \\ 0, & \text{其他} \end{cases}$$

由于 $f(x, y) \neq f_X(x)f_Y(y)$, 故 X 与 Y 不相互独立. —————5 分

(2) $Z = U + X$, 所以, $Z \in [0, 2)$,

所以, $z \leq 0$ 时, $F(z) = P(Z \leq z) = 0$;

$z \geq 2$ 时, $F(z) = P(Z \leq z) = 1$;

$$\begin{aligned} 0 < z < 1 \text{ 时, } F(z) &= P(Z \leq z) = P(U = 0, X \leq z) + P(U = 1, X \leq z - 1) \\ &= P(X > Y, X \leq z) + P(X \leq Y, X \leq z - 1) \\ &= \int_0^z dx \int_{x^2}^x 3 dy \\ &= \frac{3}{2} z^2 - z^3 \end{aligned}$$

$$\begin{aligned} 1 \leq z < 2 \text{ 时, } F(z) &= P(Z \leq z) = P(U = 0, X \leq z) + P(U = 1, X \leq z - 1) \\ &= P(X > Y, X \leq z) + P(X \leq Y, X \leq z - 1) \end{aligned}$$

$$= \frac{1}{2} + \int_0^{z-1} dx \int_x^{\sqrt{x}} 3dy = \frac{1}{2} + 2(z-1)^{\frac{3}{2}} - \frac{3}{2}(z-1)^2$$

—————9 分

五. (6 分)

解: (1) 由已知得 $f_{Y|X}(y|x) = \begin{cases} \frac{1}{x}, & 0 < y < x \\ 0, & \text{其他} \end{cases}$,

所以, $f(x, y) = f_X(x)f_{Y|X}(y|x) = \begin{cases} 5x^3, & 0 < y < x < 1 \\ 0, & \text{其他} \end{cases}$ ———2 分

(2), $E(X) = \int_0^1 x \cdot 5x^4 dx = \frac{5}{6}$,

$$f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \begin{cases} \int_y^1 5x^3 dx = \frac{5(1-y^4)}{4} & 0 < y < 1, \\ 0, & \text{其他} \end{cases}$$

$$E(Y) = \int_0^1 y \frac{5(1-y^4)}{4} dy = \frac{5}{12}$$

$$\text{或 } EY = \int_0^1 dx \int_0^x y 5x^3 dy = \frac{5}{12}$$

$$E(XY) = \int_0^1 dx \int_0^x xy 5x^3 dy = \frac{5}{14}$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = \frac{5}{14} - \frac{5}{6} \times \frac{5}{12} = \frac{5}{504}$$
 ———6 分

六. (9 分)

解: (1) 参数 σ 的矩估计:

$$\mu_1 = EX = \int_{-\infty}^{+\infty} x \frac{1}{2\sigma} e^{-\frac{|x|}{\sigma}} dx = 0,$$

$$\mu_2 = E(X^2) = \int_{-\infty}^{+\infty} x^2 \frac{1}{2\sigma} e^{-\frac{|x|}{\sigma}} dx = 2\sigma^2, \quad \sigma = \sqrt{\frac{\mu_2}{2}},$$

所以参数 σ 的矩估计 $\hat{\sigma}_1 = \sqrt{\frac{\frac{1}{n} \sum_{i=1}^n X_i^2}{2}}$ 。

参数 λ 的极大似然估计: 似然函数为

$$L(x_1, \dots, x_n; \sigma) = \prod_{i=1}^n \left(\frac{1}{2\sigma} e^{-\frac{|x_i|}{\sigma}} \right) = \frac{1}{(2\sigma)^n} \exp \left\{ -\frac{1}{\sigma} \sum_{i=1}^n |x_i| \right\}$$

求对数

$$\ln L(\sigma) = -n \ln(2\sigma) - \frac{1}{\sigma} \sum_{i=1}^n |x_i|$$

求导数, 令其为零, 得似然方程

$$\frac{d \ln L(\sigma)}{d \sigma} = -\frac{n}{\sigma} + \frac{1}{\sigma^2} \sum_{i=1}^n |x_i| \stackrel{!}{=} 0$$

解似然方程得

$$\sigma = \frac{1}{n} \sum_{i=1}^n |x_i|$$

故参数 σ 的极大似然估计为 $\hat{\sigma}_2 = \frac{1}{n} \sum_{i=1}^n |x_i|$. —————4 分

$$(2) \text{ 因为 } E\hat{\sigma}_2 = E\left(\frac{1}{n} \sum_{i=1}^n |X_i|\right) = E(|X|) = \int_{-\infty}^{+\infty} |x| \frac{1}{2\sigma} e^{-\frac{|x|}{\sigma}} dx = \sigma,$$

所以 $\hat{\sigma}_2 = \frac{1}{n} \sum_{i=1}^n |x_i|$ 是 σ 的无偏估计。 —————6 分

$$(3) \ln f(x, \sigma) = -\ln(2\sigma) - \frac{|x|}{\sigma}, \quad \frac{\partial \ln f(x, \sigma)}{\partial \sigma} = -\frac{1}{\sigma} + \frac{|x|}{\sigma^2} = \frac{|x| - \sigma}{\sigma^2},$$

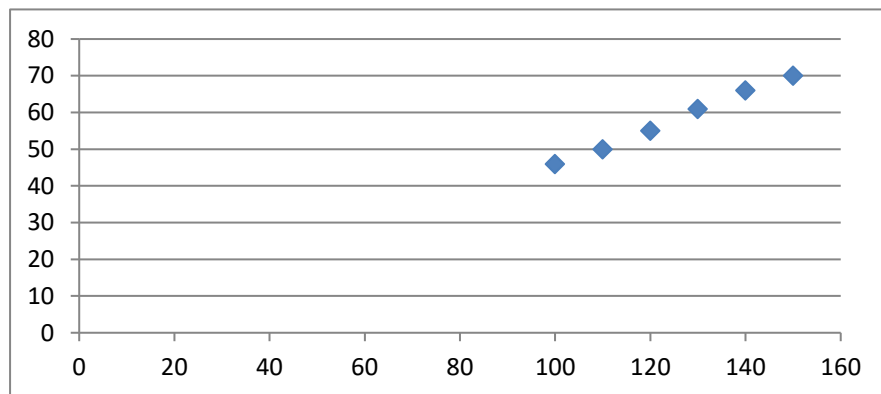
$$\text{Fisher 信息量为 } I(\sigma) = E\left(\frac{\partial \ln f(X, \sigma)}{\partial \sigma}\right)^2 = \frac{E(|X| - \sigma)^2}{\sigma^4} = \frac{D(|X|)}{\sigma^4},$$

$$D(|X|) = E(|X|^2) - (E|X|)^2 = E(X^2) - (E|X|)^2 = 2\sigma^2 - \sigma^2 = \sigma^2,$$

所以 σ 得 C—R 方差下界为 $L = \frac{1}{nI(\sigma)} = \frac{\sigma^2}{n}$ —————9 分

七 (10 分)

(1)



x 与 y 大致呈统计线性关系。 —————2 分

$$(2) \quad \bar{x} = 125, \bar{y} = 58, L_{xx} = \sum_{i=1}^n (x_i - \bar{x})^2 = 1750, \quad L_{yy} = \sum_{i=1}^n (y_i - \bar{y})^2 = 434,$$

$$L_{xy} = \sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y}) = 870, \quad \hat{b} = \frac{L_{xy}}{L_{xx}} = 0.4971, \quad \hat{a} = \bar{y} - \hat{b}\bar{x} = -4.1375$$

所以, 回归方程为 $\hat{y} = \hat{a} + \hat{b}x = -4.1375 + 0.4971x$ —————4 分

$$(3), U = \sum_{i=1}^n (\hat{y}_i - \bar{y})^2 = \hat{b}L_{xy} = 432.477, \quad Q = L_{yy} - U = 1.523, \quad n = 6$$

$$\text{所以, } \hat{\sigma} = \sqrt{\frac{Q}{n-2}} = \sqrt{\frac{1.523}{4}} = 0.617 \quad \text{-----6 分}$$

$$(4) \quad b \text{ 的置信度为 } 0.95 \text{ 的置信区间为 } (\hat{b} - t_{\frac{\alpha}{2}}(n-2)\hat{\sigma}\sqrt{\frac{1}{L_{xx}}}, \hat{b} + t_{\frac{\alpha}{2}}(n-2)\hat{\sigma}\sqrt{\frac{1}{L_{xx}}})$$

$$t_{\frac{\alpha}{2}}(n-2)\hat{\sigma}\sqrt{\frac{1}{L_{xx}}} = t_{0.025}(4) \times 0.617 \times \sqrt{\frac{1}{1750}} = 0.0409, \text{ 所以置信区间为 } (0.4562, 0.538) \quad \text{-----7 分}$$

(5) 检验 $H_0: b = 0, H_1: b \neq 0$

$$\text{检验统计量为: } F = \frac{U}{Q/(n-2)} = (n-2)\frac{U}{Q}, \text{ 假设 } H_0 \text{ 成立时, } F \sim F(1, n-2),$$

$$\text{拒绝域为 } K_0 = \{F \geq F_{\alpha}(1, n-2)\}, \quad \alpha = 0.05, F_{\alpha}(1, n-2) = F_{0.05}(1, 4) = 7.71,$$

$$\text{样本值代入得 } F = 4 \times \frac{432.477}{1.523} = 1135.856 > 7.71,$$

拒绝原假设 H_0 , 即回归方程回归显著。

-----10 分