

Linear Dimensionality Reduction





- ***Introduction**
- Principal Component Analysis
- Factor Analysis
- Multidimensional Scaling
- Linear Discriminant Analysis





Introduction





Why Dimensionality Reduction?

- Scientific: understand structure of data (visualization)
- Statistical: fewer dimensions allows better generalization
- Computational: compress data for efficiency (both time/space)
- Direct: use as a model for anomaly detection



Feature Selection vs. Extraction

Feature selection:

- Choosing K < D important features and discarding the remaining D-K features.
- Subset selection algorithms

Feature extraction:

- Projecting the original D dimensions to $K(\langle D)$ new dimensions.
- Unsupervised methods (without using output information):
 - Principal component analysis (PCA)
 - Factor analysis (FA)
 - Multidimensional scaling (MDS)
- Supervised methods (using output information):
 - Linear discriminant analysis (LDA)
- The linear methods above also have nonlinear extensions.



预备知识——特征向量和特征矩阵

• 设A有n个特征值及特征向量,则

$$A * x_1 = \lambda_1 * x_1$$

$$A * x_2 = \lambda_2 * x_2$$

$$\vdots$$

$$A * x_n = \lambda_n * x_n$$

• 将上面的写到一起成矩阵形式:

$$A*(x_1 \ x_2 \ ... \ x_n) = \begin{pmatrix} x_1 \ x_2 \ ... \ x_n \end{pmatrix} * \begin{pmatrix} \lambda_1 \ 0 \ ... \ 0 \ \lambda_2 \ ... \ 0 \ ... \ \lambda_n \end{pmatrix}$$

• 若(x1, x2,..., xn)可逆,则左右两边都求逆,则方阵A可直接通过特征值和特征向量进行唯一的表示,令

$$Q=(x_{1,1},x_{2},...,x_{n,1})$$
 $\Sigma = diag(\lambda_{1},\lambda_{2},...,\lambda_{n})$
上述公式表示为: $A = Q\Sigma Q^{-1}$



协方差矩阵

• 样本X和样本Y的协方差(Covariance):

$$Cov(X,Y) = \frac{\sum_{i=1}^{n} (X_i - \overline{X})(Y_i - \overline{Y})}{(n-1)}$$

- · 当样本是n维数据时,它们的协方差实际上是协方差矩阵(对称方阵)
 - 比如对于3维数据(x, y, z), 计算它的协方差就是:

$$C = \begin{array}{cccc} cov(x,x) & cov(x,y) & cov(x,z) \\ cov(y,x) & cov(y,y) & cov(y,z) \\ cov(z,x) & cov(z,y) & cov(z,z) \end{array}$$



Principal Component Analysis



Principal Component Analysis

- PCA finds a linear mapping from the D-dimensional input space to a K-dimensional space (K < D) with minimum information loss according to some criterion.
- Projection of x on the direction of w: $z = \mathbf{w}^T \mathbf{x}$
- Finding the first principal component w1 s.t. var(z1) is maximized:

$$var(Z_1) = var(\mathbf{w}_1^T \mathbf{x}) = E[(\mathbf{w}_1^T \mathbf{x} - \mathbf{w}_1^T \mu)^2]$$

$$= E[\mathbf{w}_1^T (\mathbf{x} - \mu)(\mathbf{x} - \mu)^T \mathbf{w}_1]$$

$$= \mathbf{w}_1^T E[(\mathbf{x} - \mu)(\mathbf{x} - \mu)^T] \mathbf{w}_1 = \mathbf{w}_1^T \mathbf{\Sigma} \mathbf{w}_1$$

where

$$cov(\mathbf{x}) = E[(\mathbf{x} - \mu)(\mathbf{x} - \mu)^T] = \mathbf{\Sigma}$$



- Maximization of var(z1) subject to ||w1|| = 1 can be solved as a constrained optimization problem using a Lagrange multiplier.
- Maximization of Lagrangian:

$$\mathbf{w}_1^T \mathbf{\Sigma} \mathbf{w}_1 - \alpha (\mathbf{w}_1^T \mathbf{w}_1 - 1)$$

• Taking the derivative of the Lagrangian w.r.t. w1 and setting it to 0, we get an eigenvalue equation for the first principal component w1:

$$\Sigma \mathbf{w}_1 = \alpha \mathbf{w}_1$$

• Because we have

$$\mathbf{w}_1^T \mathbf{\Sigma} \mathbf{w}_1 = \alpha \mathbf{w}_1^T \mathbf{w}_1 = \alpha$$

• we choose the eigenvector with the largest eigenvalue for the variance to be maximum.



- The second principal component w2 should also maximize the variance var(z2), subject to the constraints that ||w2||=1 and that w2 is orthogonal to w1.
- Maximization of Lagrangian:

$$\mathbf{w}_2^T \mathbf{\Sigma} \mathbf{w}_2 - \alpha (\mathbf{w}_2^T \mathbf{w}_2 - 1) - \beta (\mathbf{w}_2^T \mathbf{w}_1 - 0)$$

• Taking the derivative of the Lagrangian w.r.t. w2 and setting it to 0, we get the following equation:

$$2\mathbf{\Sigma}\mathbf{w}_2 - 2\alpha\mathbf{w}_2 - \beta\mathbf{w}_1 = 0$$

• We can show that $\beta = 0$ and hence have this eigenvalue equation:

$$\Sigma \mathbf{w}_2 = \alpha \mathbf{w}_2$$

• implying that w2 is the eigenvector of Σ with the second largest eigenvalue.

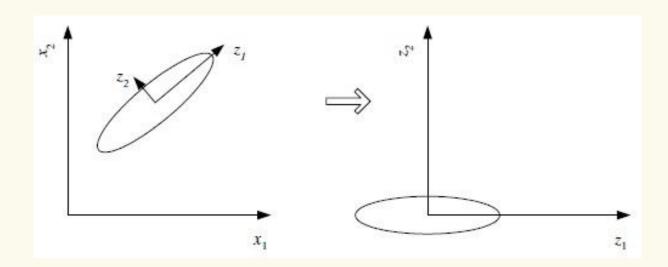


What PCA Does

• Transformation of data:

$$z = W^T (x - m)$$

- where the columns of W = [w1;w2; : :] are the eigenvectors of Σ and m is the sample mean.
- Centering the data at the origin and rotating the axes:



If the variance on z2 is too small, it can be ignored to reduce the dimensionality from 2 to 1.



How to Choose K

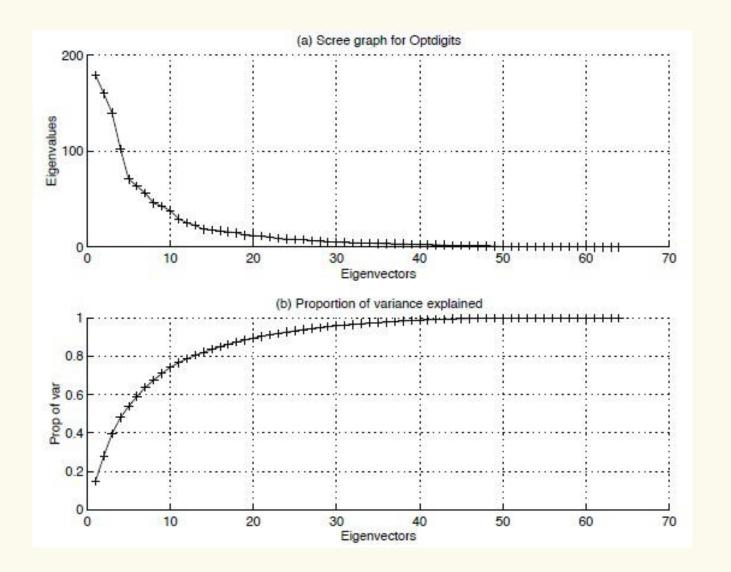
Proportion of variance (PoV) explained:

$$\frac{\lambda_1 + \lambda_2 + \ldots + \lambda_K}{\lambda_1 + \lambda_2 + \ldots + \lambda_D}$$

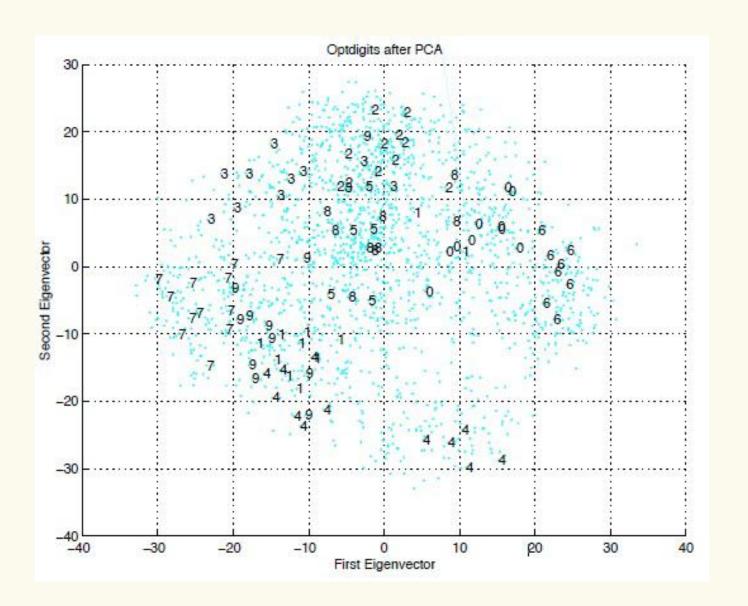
where λ i are sorted in descending order.

- Typically, stop at PoV > 0:9
- Scree graph plotting PoV against K; stop at "elbow".









Factor Analysis

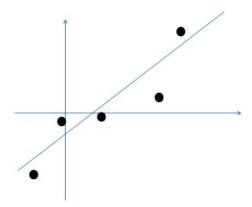
Definition

- FA assumes that there is a set of latent factors z_j which when acting in combination to generate the observed variables x.
- ► The goal of FA is to characterize the dependency among the observed variables by means of a smaller number of factors.
- $X = \mu + AF + \epsilon$
- ▶ What exactly that mean?

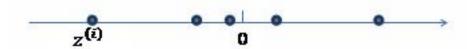
A simple Example

- $X = \mu + AF + \epsilon$
- It is easy to say whether a lady is beautiful or not, but the definition of beautiful can be very different. But we can attribute it to outer and inter beauty, which relates to the "F" above
- besides, the ε relates to some special taste of the individuals, which we don't care much in FA.

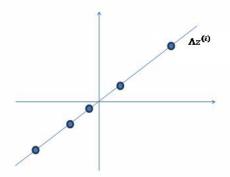
- Suppose we have 5 samples in a 2-D plane
- Lets prove it could be present with points from 1-D plane



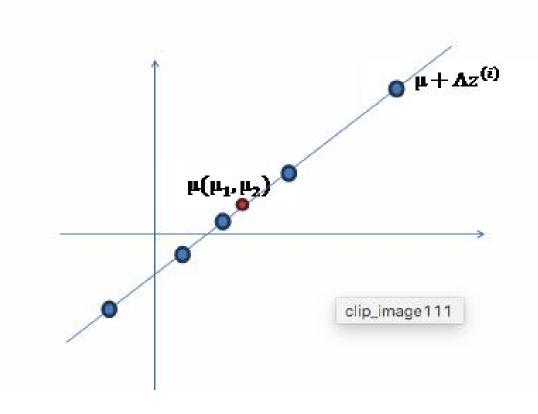
Now we have 5 points with Gaussian distribution in the 1-D plane



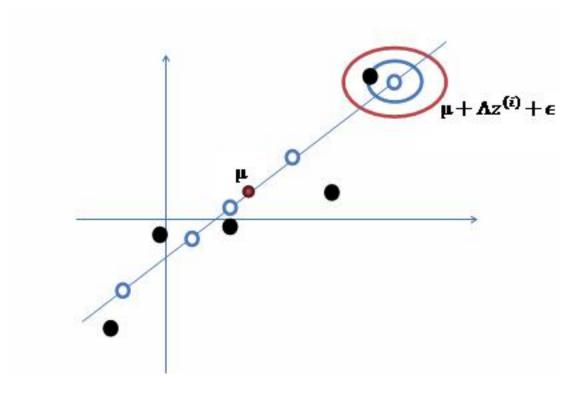
- ▶ The reason to choose Gaussian distribution is that its $\mu = 0$ and var = 1
- ▶ Now we project it to the 2-D plane with Covariance Matrix



Plus the μ



Plus the ε



Summary

- From the example above, we know that the samples of high dimensional planes can be reduced to lower dimensional.
- We have already known μ , and we don't care much of the special variation ϵj , so our main task in FA is to find A to project Z to X, that is, find the covariance Matrix "A"
- ► Unfortunately, usually Z could not be found easily, which means we cannot use a function to present Z
- ▶ How to solve it?

The solution of covariance Matrix

$$\Lambda = \left(\sum_{i=1}^m (x^{(i)} - \mu) \mu_{z^{(i)}|x^{(i)}}^T\right) \left(\sum_{i=1}^m \mu_{z^{(i)}|x^{(i)}} \mu_{z^{(i)}|x^{(i)}}^T + \sum_{z^{(i)}|x^{(i)}}\right)^{-1}.$$

$$\mu = \frac{1}{m} \sum_{i=1}^{m} x^{(i)}.$$

$$\Phi = \frac{1}{m} \sum_{i=1}^{m} x^{(i)} x^{(i)^T} - x^{(i)} \mu_{z^{(i)}|x^{(i)}}^T \Lambda^T - \Lambda \mu_{z^{(i)}|x^{(i)}} x^{(i)^T} + \Lambda (\mu_{z^{(i)}|x^{(i)}} \mu_{z^{(i)}|x^{(i)}}^T + \Sigma_{z^{(i)}|x^{(i)}}) \Lambda^T$$

Detailed solution

http://www.cnblogs.com/jerrylead/archive/2011/05/11/ 2043317.html

The difference with PCA

- ▶ PCA meanly to reduce the dimension, it involves extracting linear composites of observed variables.
- FA is based on a formal model predicting observed variables from theoretical latent factors
- ▶ The direction of FA is opposite to that of PCA (from courseware PPT):
 - ightharpoonup PCA (from x to z): z = WT (x μ)
 - FA (from z to x): $x \mu = Vz + ε$

Multidimensional Scaling

definition

- Problem formulation:
 - ► Given the pairwise distances between pairs of points in some space (but the exact coordinates of the points and their dimensionality are unknown).
 - ▶ We want to embed the points in a lower-dimensional space such that the pairwise distances in this space are as close as possible to those in the original space.
- ► The projection to the lower-dimensional space is not unique because the pairwise distances are invariant to such operations as translation, rotation and reflection.

task

► The I objects could have I^2 distances

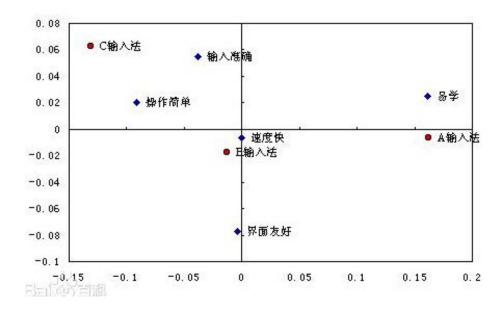
$$\Delta := \begin{pmatrix} \delta_{1,1} & \delta_{1,2} & \cdots & \delta_{1,I} \\ \delta_{2,1} & \delta_{2,2} & \cdots & \delta_{2,I} \\ \vdots & \vdots & & \vdots \\ \delta_{I,1} & \delta_{I,2} & \cdots & \delta_{I,I} \end{pmatrix}.$$

- lacksquare The task of MDS is to find $\|x_i x_j\| pprox \delta_{i,j}$
- And we can take it as an optimize problem:

$$\min_{x_1, \dots, x_I} \sum_{i < j} (\|x_i - x_j\| - \delta_{i,j})^2.$$

example

- ▶ It located multi variances into 2-D or 3-D planes and calculate their distances to reflect their similarities and difference.
- ▶ Perceptual Mapping (知觉图)



Fomula

$$\mathbf{B} = \mathbf{X} \mathbf{X}^T$$

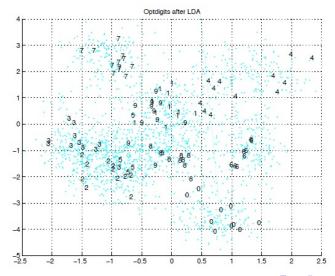
$$\mathbf{B} = \mathbf{C}\mathbf{D}\mathbf{C}^T = \mathbf{C}\mathbf{D}^{1/2}\mathbf{D}^{1/2}\mathbf{C}^T = (\mathbf{C}\mathbf{D}^{1/2})(\mathbf{C}\mathbf{D}^{1/2})^T$$

where C is the matrix whose columns are the eigenvectors of B and D1/2 is the diagonal matrix whose diagonal elements are the square roots of the eigenvalues.

Linear Discriminant Analysis

- Unlike PCA, FA and MDS, LDA is a supervised dimensionality reduction method.
- LDA is typically used with a classifier for classification problems.
- Goal: the classes are well-separated after projecting to a low-dimensional space by utilizing the label information (output information).

Example



2-Class Case

- Given sample $\mathcal{X} = \{(\mathbf{x}^{(i)}, y^{(i)})\}$, where $y^{(i)} = 1$ if $\mathbf{x}^{(i)} \in C_1$ and $y^{(i)} = 0$ if $\mathbf{x}^{(i)} \in C_2$.
- Find vector \mathbf{w} on which the data are projected such that the examples from C_1 and C_2 are as well separated as possible.
- Projection of x onto w (dimensionality reduced from D to 1):

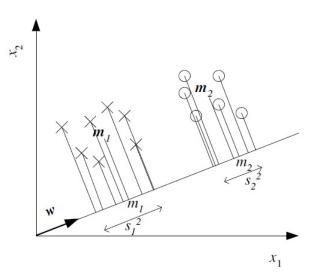
$$z = \mathbf{w}^T \mathbf{x}$$

• $\mathbf{m}_j \in \mathbb{R}^D$ and $m_j \in \mathbb{R}$ are sample means of C_j before and after projection:

$$m_1 = \frac{\sum_i \mathbf{w}^T \mathbf{x}^{(i)} y^{(i)}}{\sum_i y^{(i)}} = \mathbf{w}^T \mathbf{m}_1$$

$$m_2 = \frac{\sum_i \mathbf{w}^T \mathbf{x}^{(i)} (1 - y^{(i)})}{\sum_i (1 - y^{(i)})} = \mathbf{w}^T \mathbf{m}_2$$

Projection



Between-Class Scatter

Between-class scatter:

$$(m_1 - m_2)^2 = (\mathbf{w}^T \mathbf{m}_1 - \mathbf{w}^T \mathbf{m}_2)^2$$

= $\mathbf{w}^T (\mathbf{m}_1 - \mathbf{m}_2) (\mathbf{m}_1 - \mathbf{m}_2)^T \mathbf{w}$
= $\mathbf{w}^T \mathbf{S}_B \mathbf{w}$

where

$$\mathbf{S}_B = (\mathbf{m}_1 - \mathbf{m}_2)(\mathbf{m}_1 - \mathbf{m}_2)^T$$

Within-Class Scatter

Within-class scatter:

$$S_1^2 = \sum_{i} (\mathbf{w}^T \mathbf{x}^{(i)} - m_1)^2 y^{(i)}$$

$$= \sum_{i} \mathbf{w}^T (\mathbf{x}^{(i)} - \mathbf{m}_1) (\mathbf{x}^{(i)} - \mathbf{m}_1)^T \mathbf{w} y^{(i)}$$

$$= \mathbf{w}^T \mathbf{S}_1 \mathbf{w}$$

where
$$\mathbf{S}_1 = \sum_i (\mathbf{x}^{(i)} - \mathbf{m}_1) (\mathbf{x}^{(i)} - \mathbf{m}_1)^T y^{(i)}$$
. Similarly, $s_2^2 = \mathbf{w}^T \mathbf{S}_2 \mathbf{w}$ with $\mathbf{S}_2 = \sum_i (\mathbf{x}^{(i)} - \mathbf{m}_2) (\mathbf{x}^{(i)} - \mathbf{m}_2)^T (1 - y^{(i)})$. So

$$s_1^2 + s_2^2 = \mathbf{w}^T \mathbf{S}_W \mathbf{w}$$

where $S_W = S_1 + S_2$.



Fisher's Linear Discriminant

 Fisher's linear discriminant refers to the vector w that maximizes the Fisher criterion (a.k.a. generalized Rayleigh quotient):

$$J(\mathbf{w}) = \frac{(m_1 - m_2)^2}{s_1^2 + s_2^2} = \frac{\mathbf{w}^T \mathbf{S}_B \mathbf{w}}{\mathbf{w}^T \mathbf{S}_W \mathbf{w}}$$

 Taking the derivative of J w.r.t. w and setting it to 0, we obtain the following generalized eigenvalue problem:

$$S_B w = \lambda S_W w$$

or, if S_W is nonsingular, an equivalent eigenvalue problem:

$$\mathbf{S}_W^{-1}\mathbf{S}_B\mathbf{w} = \lambda\mathbf{w}$$

Fisher's Linear Discriminant (2)

Alternatively, for the 2-class case, we note that

$$S_B w = (m_1 - m_2)(m_1 - m_2)^T w = c(m_1 - m_2)$$

for some constant c and hence $S_B w$ is in the same direction of $m_1 - m_2$.

So we get

$$\mathbf{w} = \mathbf{S}_W^{-1}(\mathbf{m}_1 - \mathbf{m}_2) = (\mathbf{S}_1 + \mathbf{S}_2)^{-1}(\mathbf{m}_1 - \mathbf{m}_2)$$

The constant factor is irrelevant and hence is discarded.

K > 2 Classes

• Find the matrix $\mathbf{W} \in \mathbb{R}^{D \times K}$ such that

$$\mathbf{z} = \mathbf{W}^T \mathbf{x} \in \mathbb{R}^K$$

Within-class scatter matrix for class C_k:

$$\mathbf{S}_k = \sum_i y_k^{(i)} (\mathbf{x}^{(i)} - \mathbf{m}_k) (\mathbf{x}^{(i)} - \mathbf{m}_k)^T$$

where $y_k^{(i)} = 1$ if $\mathbf{x}^{(i)} \in C_k$ and 0 otherwise.

Total within-class scatter matrix:

$$\mathbf{S}_W = \sum_{k=1}^K \mathbf{S}_k$$



K > 2 Classes (2)

Between-class scatter matrix:

$$\mathbf{S}_B = \sum_{k=1}^K N_k (\mathbf{m}_k - \mathbf{m}) (\mathbf{m}_k - \mathbf{m})^T$$

where **m** is the overall mean and $N_k = \sum_i y_k^{(i)}$.

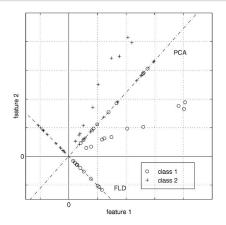
The optimal solution is the matrix W that maximizes

$$J(\mathbf{W}) = \frac{\mathrm{Tr}(\mathbf{W}^T \mathbf{S}_B \mathbf{W})}{\mathrm{Tr}(\mathbf{W}^T \mathbf{S}_W \mathbf{W})}$$

which corresponds to the eigenvectors of $S_W^{-1}S_B$ with the largest eigenvalues.

• Take new dimensionality $d \le K - 1$: since S_W is the sum of K rank-1 matrices and only K - 1 of them are independent, S_B has a maximum rank of K - 1.

Application in Face Recognition: PCA vs. LDA



- PCA (Eigenface) maps features to a subspace that contains most energy.
- FLD (Fisherface) maps features to a subspace that most separate the classes.

Application in Face Recognition: PCA vs. LDA (2)

- PCA is an unsupervised dimension reduction algorithm, while LDA is supervised.
- PCA is good at outlier cleaning, and LDA could learn the within-class deviation.
- These two methods only extract 1st and 2nd statistical moments.
- The combination of PCA and LDA could enhance the performance.
- PCA serves as the first-step processing of several kinds of face recognition technique.
- Techniques of dimension reduction are frequently used in face recognition.