

## Assignment for Applied Regression Analysis

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Applied Statistics

### 3.1

a. From the table below we can conclude that

$$y = -1.80837 + 0.00360x_2 + 0.19386x_7 - 0.00482x_8.$$

But the coefficient of intercept is not significant.

Parameter Estimates					
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t
Intercept	1	-1.80837	7.90086	-0.23	0.8209
x2	1	0.00360	0.00069500	5.18	<.0001
x7	1	0.19396	0.08823	2.20	0.0378
x8	1	-0.00482	0.00128	-3.77	0.0009

b. We know that the  $p\_value < 0.05$ , so the regression is reasonable.

Analysis of Variance					
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	3	257.09428	85.69809	29.44	<.0001
Error	24	69.87000	2.91125		
Corrected Total	27	326.96429			

c. The last column of the first table shows the results of hypothesis testing for regression coefficients. We know that  $x_2$  contributes the most to the regression model, followed by  $x_8$ , and finally  $x_7$ .

d.  $R^2 = 0.7683$ ,  $R^2_{Adj} = 0.7596$ .

Root MSE	1.70624	R-Square	0.7683
Dependent Mean	6.96429	Adj R-Sq	0.7596
Coeff Var	24.49984		

e. From the table we can get the same conclusion as the  $t$  test for  $\beta_7$  calculated in part c above.

Summary of Stepwise Selection								
Step	Variable Entered	Variable Removed	Number Vars In	Partial R-Square	Model R-Square	C(p)	F Value	Pr > F
1	x8		1	0.5447	0.5447	27.1368	31.10	<.0001
2	x2		2	0.1986	0.7433	6.8324	19.34	0.0002
3	x7		3	0.0430	0.7863	4.0000	4.83	0.0378

### 3.2

From the result computed by SAS, we know that square of the simple correlation coefficient between the observed values  $y_i$  and the fitted values  $\hat{y}_i$  is 0.7924, which very nearly equals  $R^2$  (0.7863) calculated before.

Pearson Correlation Coefficients, N = 28 Prob >  r  under H0: Rho=0		
	y	y_fit
y	1.00000	0.05208 0.7924
y_fit	0.05208 0.7924	1.00000

### 3.3

a. From the result calculated by SAS we can know that a 95% CI on  $\beta_7$  is [0.01186, 0.37607].

Parameter Estimates							
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	95% Confidence Limits	
Intercept	1	-1.80837	7.90086	-0.23	0.8209	-18.11494	14.49820
x2	1	0.00360	0.00069500	5.18	<.0001	0.00216	0.00503
x7	1	0.19396	0.08823	2.20	0.0378	0.01186	0.37607
x8	1	-0.00482	0.00128	-3.77	0.0009	-0.00745	-0.00218

b. A 95% CI on the mean number of games won by a team when  $x_2 = 2300$ ,  $x_7 = 56.0$ ,  $x_8 = 2100$  is [6.4362, 7.9966].

Output Statistics						
Obs	Dependent Variable	Predicted Value	Std Error Mean Predict	95% CL Mean		Residual
1	.	7.2164	0.3780	6.4362	7.9966	.

### 3.4

a. From the table below we can know that the  $p$ -value calculated  $< 0.05$ , so the regression is significant.

Analysis of Variance					
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	2	179.06619	89.53309	15.13	<.0001
Error	25	147.89810	5.91592		
Corrected Total	27	326.96429			

b.  $R^2 = 0.5477$ ,  $R^2_{Adj} = 0.5115$ .

Root MSE	2.43227	R-Square	0.5477
Dependent Mean	6.96429	Adj R-Sq	0.5115
Coeff Var	34.92486		

These two figures are both smaller than those calculated in Problem 3.1, which indicates that the regressor in Problem 3.1 is better. So  $x_2$  can make a significant contribution to a better regression model.

c. From the result calculated by SAS we can know that a 95% CI on  $\beta_7$  is  $[-0.19716, 0.29391]$ .

Parameter Estimates							
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	95% Confidence Limits	
Intercept	1	17.94432	9.86248	1.82	0.0808	-2.36785	38.25649
x7	1	0.04837	0.11922	0.41	0.6884	-0.19716	0.29391
x8	1	-0.00654	0.00176	-3.72	0.0010	-0.01016	-0.00292

A 95% CI on the mean number of games won by a team when  $x_2 = 2300$ ,  $x_7 = 56.0$ ,  $x_8 = 2100$  is  $[5.8286, 8.0238]$ .

Output Statistics						
Obs	Dependent Variable	Predicted Value	Std Error Mean Predict	95% CL Mean		Residual
1	.	6.9262	0.5329	5.8286	8.0238	.

we know that

$$\mathbf{X} = \begin{bmatrix} 1 & x_{11} & x_{12} & \cdots & x_{1k} \\ 1 & x_{21} & x_{22} & \cdots & x_{2k} \\ \vdots & \vdots & \vdots & & \vdots \\ 1 & x_{n1} & x_{n2} & \cdots & x_{nk} \end{bmatrix}$$

$$(\mathbf{X}'\mathbf{X})^{-1} = \frac{1}{S_{xx}} \begin{pmatrix} \sum x_i^2/n & -\bar{x} \\ -\bar{x} & 1 \end{pmatrix}$$

so

$$\begin{aligned} h_{ii} &= (1 \quad x_i) \left[ \frac{1}{S_{xx}} \begin{pmatrix} \sum x_i^2/n & -\bar{x} \\ -\bar{x} & 1 \end{pmatrix} \right] \begin{pmatrix} 1 \\ x_i \end{pmatrix} \\ &= \left[ \frac{1}{S_{xx}} \right] \left[ \frac{\sum x_i^2}{n} - x_i \bar{x} - x_i \bar{x} + x_i^2 \right] \\ &= \left[ \frac{1}{S_{xx}} \right] \left[ \frac{\sum x_i^2}{n} + (x_i - \bar{x})^2 - \bar{x}^2 \right] \\ &= \left[ \frac{1}{S_{xx}} \right] \left[ \frac{\sum x_i^2 - n\bar{x}^2}{n} + (x_i - \bar{x})^2 \right] \\ &= \frac{1}{n} + \frac{(x_i - \bar{x})^2}{S_{xx}} \end{aligned}$$

and

$$\begin{aligned} h_{ij} &= (1 \quad x_i) \left[ \frac{1}{S_{xx}} \begin{pmatrix} \sum x_i^2/n & -\bar{x} \\ -\bar{x} & 1 \end{pmatrix} \right] \begin{pmatrix} 1 \\ x_j \end{pmatrix} \\ &= \left[ \frac{1}{S_{xx}} \right] \left[ \frac{\sum x_i^2}{n} - x_i \bar{x} - x_j \bar{x} + x_i x_j \right] \\ &= \left[ \frac{1}{S_{xx}} \right] \left[ \frac{\sum x_i^2}{n} + (x_i - \bar{x})(x_j - \bar{x}) - \bar{x}^2 \right] \\ &= \left[ \frac{1}{S_{xx}} \right] \left[ \frac{\sum x_i^2 - n\bar{x}^2}{n} + (x_i - \bar{x})(x_j - \bar{x}) \right] \\ &= \frac{1}{n} + \frac{(x_i - \bar{x})(x_j - \bar{x})}{S_{xx}} \end{aligned}$$

From the these expressions we can easily know that these quantities will increase as  $x_i$  moves farther from  $\bar{x}$ .

### 3.30

$$\begin{aligned} \hat{\beta} &= (\mathbf{X}'\mathbf{X})^{-1}\mathbf{y} = (\mathbf{X}'\mathbf{X})^{-1}(\mathbf{X}\beta + \epsilon) \\ &= (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{X}\beta + (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\epsilon \\ &= \beta + \mathbf{R}\epsilon \\ \mathbf{R} &= (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}' \end{aligned}$$

### 3.31

$$\begin{aligned}
e &= (\mathbf{I} - \mathbf{H})\mathbf{y} \\
&= (\mathbf{I} - \mathbf{H})(\mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}) \\
&= [\mathbf{I} - \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}']\mathbf{X}\boldsymbol{\beta} + (\mathbf{I} - \mathbf{H})\boldsymbol{\varepsilon} \\
&= (\mathbf{I} - \mathbf{H})\boldsymbol{\varepsilon}
\end{aligned}$$

3.39

$$\mathbf{W}'\mathbf{W} = \begin{bmatrix} 1 & r_{12} & r_{13} & \cdots & r_{1k} \\ r_{12} & 1 & r_{23} & \cdots & r_{2k} \\ r_{13} & r_{23} & 1 & \cdots & r_{3k} \\ \vdots & \vdots & \vdots & & \vdots \\ r_{1k} & r_{2k} & r_{3k} & \cdots & 1 \end{bmatrix}$$

The  $j^{th}$  VIF is the  $j^{th}$  diagonal element of  $(\mathbf{W}'\mathbf{W})^{-1}$ .

$$(\mathbf{W}'\mathbf{W})^{-1} = \frac{1}{1 - R_j^2}$$