

MTH 9831. Solutions to Quiz 8.

- (1) (4 points) Let $S(t)$, $t \geq 0$, be a geometric Brownian motion $dS(t) = rS(t)dt + \sigma S(t)dW(t)$, where $W(t)$, $t \geq 0$, is a standard Brownian motion with respect to \mathbb{P} , $\sigma > 0$, and let $Y(t) = \max_{0 \leq u \leq t} S(u)$. Show that the process $(S(t), Y(t))$, $t \geq 0$, is Markov.

Solution. Let T be fixed, $t \in [0, T]$, and $f(x, y)$ be a non-negative Borel-measurable function. Filtration $\mathcal{F}(t)$ is the filtration associated with $W(t)$. We need to show that there exists a function $g(x, y)$ such that

$$\mathbb{E}[f(S(T), Y(T)) | \mathcal{F}(t)] = g(S(t), Y(t)).$$

Notice that

$$\begin{aligned} S(T) &= S(t) \frac{S(T)}{S(t)} = S(t) \exp(\sigma(W(T) - W(t)) + (r - \sigma^2/2)(T - t)) \\ &= S(t) \exp\left(\sigma\left(\hat{W}(T) - \hat{W}(t)\right)\right), \end{aligned}$$

where $\hat{W}(u) = W(u) + \alpha u$, $\alpha = \sigma^{-1}(r - \sigma^2/2)$, $u \geq 0$, is a drifted Brownian motion under \mathbb{P} . Moreover,

$$\begin{aligned} Y(T) &= Y(t) + Y(T) - Y(t) = Y(t) + \left(\max_{t \leq u \leq T} S(u) - Y(t)\right)_+ \\ &= Y(t) + S(t) \left(\max_{t \leq u \leq T} \frac{S(u)}{S(t)} - \frac{Y(t)}{S(t)}\right)_+ \\ &= Y(t) + S(t) \left(\exp\left(\sigma \max_{t \leq u \leq T} (\hat{W}(u) - \hat{W}(t))\right) - \frac{Y(t)}{S(t)}\right)_+ \end{aligned}$$

Since $\hat{W}(u) - \hat{W}(t)$ is independent of $\mathcal{F}(t)$ for all $u \in [t, T]$ and has the same distribution as $\hat{W}(u - t)$, $u \geq t$, by the independence lemma

$$\mathbb{E}[f(S(T), Y(T)) | \mathcal{F}(t)] = g(S(t), Y(t)),$$

where

$$g(x, y) = \mathbb{E}\left(f\left(xe^{\sigma\hat{W}(T-t)}, y + x\left(e^{\sigma\hat{M}(T-t)} - \frac{y}{x}\right)_+\right)\right),$$

and $\hat{M}(u) = \max_{0 \leq s \leq u} \hat{W}(s)$, $u \geq 0$. This finishes the proof of the Markov property.

Just to elaborate a bit further,¹ recall that the joint density of $(\hat{M}(u), \hat{W}(u))$ is given by

$$f_{\hat{M}(u), \hat{W}(u)}(m, w) = \frac{2(2m - w)}{u\sqrt{2\pi u}} e^{\alpha w - \frac{1}{2}\alpha^2 u - \frac{1}{2u}(2m - w)^2} \mathbf{1}_{\{w \leq m, m \geq 0\}}.$$

Thus, we can express $g(x, y)$ as the following integral:

$$\begin{aligned} g(x, y) &= \int_0^\infty \int_{-\infty}^m f\left(xe^{\sigma w}, y + x\left(e^{\sigma m} - \frac{y}{x}\right)_+\right) \\ &\quad \times \frac{2(2m - w)}{(T - t)\sqrt{2\pi(T - t)}} e^{\alpha w - \frac{1}{2}\alpha^2(T - t) - \frac{1}{2(T - t)}(2m - w)^2} dw dm. \end{aligned}$$

Clearly, function g depends on $T - t$, r , and σ but these dependencies are not reflected in the notation.

Remark. Observe that (up to the discounting factor) we obtained a formula for the time t value of any derivative security with expiration T whose payoff is a function of $S(T)$ and $Y(T)$.

¹this is not a part of the problem.

- (2) (2 points) Let $S(t)$, $Y(t)$, $t \geq 0$, be as above and Π be a partition of $[0, t]$ with $t_0 = 0 < t_1 < \dots < t_m = t$. Show that $\lim_{\|\Pi\| \rightarrow 0} \mathbb{E} \left[\left(\sum_{i=1}^m (Y(t_i) - Y(t_{i-1}))(S(t_i) - S(t_{i-1})) \right)^2 \right] = 0$. You may assume that as $\|\Pi\| \rightarrow 0$
- $$\mathbb{E} \left(\sum_{i=1}^m (Y(t_i) - Y(t_{i-1}))^2 \right)^2 \rightarrow 0 \quad \text{and} \quad \mathbb{E} \left(\sum_{i=1}^m (S(t_i) - S(t_{i-1}))^2 \right)^2 \rightarrow \sigma^2 \int_0^t \mathbb{E}(S^2(u)) du. \quad (1)$$

Solution. Let $t_0 = 0 < t_1 < \dots < t_m = t$ be a partition of $[0, t]$. Then by the Cauchy-Schwarz inequality we get

$$\left| \sum_{i=1}^m (Y(t_i) - Y(t_{i-1}))(S(t_i) - S(t_{i-1})) \right| \leq \left(\sum_{i=1}^m (Y(t_i) - Y(t_{i-1}))^2 \right)^{1/2} \left(\sum_{i=1}^m (S(t_i) - S(t_{i-1}))^2 \right)^{1/2}.$$

Squaring both sides, taking expectations, and again using the Cauchy-Schwarz inequality (now for expectations) we see that

$$\mathbb{E} \left(\sum_{i=1}^m (Y(t_i) - Y(t_{i-1}))(S(t_i) - S(t_{i-1})) \right)^2 \leq \sqrt{\mathbb{E} \left(\sum_{i=1}^m (Y(t_i) - Y(t_{i-1}))^2 \right)^2} \sqrt{\mathbb{E} \left(\sum_{i=1}^m (S(t_i) - S(t_{i-1}))^2 \right)^2}.$$

By (1) the right hand side converges to 0 as $\|\Pi\| \rightarrow 0$.

- (3) (4 points) Assume Black-Scholes model and consider a zero-strike Asian call option, i.e. the option with the payoff

$$V(T) = \frac{1}{T} \int_0^T S(u) du.$$

Suppose that at time t we have $S(t) = x \geq 0$ and $\int_0^t S(u) du = y \geq 0$. Use the fact that $e^{-rt}S(t)$ is a martingale under $\tilde{\mathbb{P}}$ to compute

$$e^{-r(T-t)} \tilde{\mathbb{E}} \left(\frac{1}{T} \int_0^T S(u) du \mid \mathcal{F}(t) \right).$$

Call your answer $v(t, x, y)$. Then find explicitly $\Delta(t)$ for all $0 \leq t \leq T$.

Solution. Let $Y(t) = \int_0^t S(u) du$. By Fubini's theorem for conditional expectations and the fact that $e^{-rt}S(t)$ is a martingale under $\tilde{\mathbb{P}}$ we have

$$\begin{aligned} \tilde{\mathbb{E}} \left(\int_0^T S(u) du \mid \mathcal{F}(t) \right) &= \tilde{\mathbb{E}} \left(\int_0^t S(u) du + \int_t^T S(u) du \mid \mathcal{F}(t) \right) \\ &= Y(t) + \int_t^T \tilde{\mathbb{E}}(S(u) \mid \mathcal{F}(t)) du = Y(t) + \int_t^T e^{ru} \tilde{\mathbb{E}}(e^{-ru} S(u) \mid \mathcal{F}(t)) du \\ &= Y(t) + \int_t^T e^{ru} e^{-rt} S(t) du = Y(t) + S(t) \int_t^T e^{r(u-t)} du \\ &= Y(t) + S(t) \frac{1}{r} (e^{r(T-t)} - 1). \end{aligned}$$

Multiplying by $e^{-r(T-t)}$ and dividing by T we get

$$v(t, S(t), Y(t)) = \frac{1}{T} e^{-r(T-t)} Y(t) + \frac{1}{rT} (1 - e^{-r(T-t)}) S(t).$$

Therefore,

$$v(t, x, y) = \frac{1}{T} e^{-r(T-t)} y + \frac{1}{rT} (1 - e^{-r(T-t)}) x, \quad v_x(t, x, y) = \frac{1}{rT} (1 - e^{-r(T-t)}),$$

and $\Delta(t) = \frac{1}{rT} (1 - e^{-r(T-t)})$, $0 \leq t \leq T$, a non-random smooth function.