

HW4

April 3, 2019

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2 Problem 4.1

```
In [1]: import pandas as pd
import numpy as np
from sklearn.linear_model import LinearRegression
import statsmodels.api as sm
from scipy.stats import pearsonr
```

```
In [2]: data = pd.read_csv('BetaExample.txt', sep="|")
data.DATE=pd.to_datetime(data.DATE)
data=data.dropna()
```

```
In [3]: data1=data.loc[data.TICKER=="IBM"]
data2=data.loc[data.TICKER=="AAPL"]
data3=data.loc[data.TICKER=="TSLA"]
```

2.1 (a)

2.1.1 IBM

```
In [4]: d=data1[data1.DATE<pd.to_datetime("2015-01-01")]
X=np.array(d[["RM"]])
y=np.array(d[["R"]])
```

```
In [5]: est = sm.OLS(y, X)
est2 = est.fit()
print(est2.summary())
```

OLS Regression Results

```
=====
Dep. Variable:          y      R-squared:          0.332
Model:                  OLS    Adj. R-squared:      0.332
Method:                 Least Squares    F-statistic:    3133.
Date:                   Wed, 03 Apr 2019    Prob (F-statistic):    0.00
Time:                   10:09:54    Log-Likelihood:    17657.
No. Observations:      6301    AIC:              -3.531e+04
```

```

Df Residuals:          6300    BIC:          -3.531e+04
Df Model:              1
Covariance Type:      nonrobust

```

```

=====
              coef      std err          t      P>|t|      [0.025      0.975]
-----
x1              0.9067      0.016     55.970      0.000      0.875      0.938
=====
Omnibus:              1227.530    Durbin-Watson:              1.964
Prob(Omnibus):         0.000    Jarque-Bera (JB):          34867.449
Skew:                  0.179    Prob(JB):              0.00
Kurtosis:              14.519    Cond. No.              1.00
=====

```

Warnings:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

$\beta_{IBM} = 0.91$, P is low, we reject the null hypothesis.

2.1.2 AAPL

```

In [6]: d=data2[data2.DATE<pd.to_datetime("2015-01-01")]
        X=np.array(d[["RM"]])
        y=np.array(d["R"])

In [7]: est = sm.OLS(y, X)
        est2 = est.fit()
        print(est2.summary())

```

OLS Regression Results

```

=====
Dep. Variable:          y    R-squared:          0.192
Model:                OLS    Adj. R-squared:      0.192
Method:             Least Squares    F-statistic:      1500.
Date:                Wed, 03 Apr 2019    Prob (F-statistic):  1.61e-294
Time:                10:09:54    Log-Likelihood:    13915.
No. Observations:      6301    AIC:              -2.783e+04
Df Residuals:          6300    BIC:              -2.782e+04
Df Model:              1
Covariance Type:      nonrobust
=====
              coef      std err          t      P>|t|      [0.025      0.975]
-----
x1              1.1363      0.029     38.728      0.000      1.079      1.194
=====
Omnibus:              1964.903    Durbin-Watson:              1.980
Prob(Omnibus):         0.000    Jarque-Bera (JB):          232261.428
Skew:                  -0.417    Prob(JB):              0.00

```

Kurtosis: 32.732 Cond. No. 1.00
=====

Warnings:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

$\beta_{AAPL} = 1.14$, **P is low, we reject the null hypothesis.**

2.1.3 TSLA

```
In [8]: d=data3[data3.DATE<pd.to_datetime("2015-01-01")]
        X=np.array(d[["RM"]])
        y=np.array(d[["R"]])
```

```
In [9]: est = sm.OLS(y, X)
        est2 = est.fit()
        print(est2.summary())
```

```

                                OLS Regression Results
=====
Dep. Variable:                  y      R-squared:                0.122
Model:                            OLS      Adj. R-squared:           0.121
Method:                 Least Squares      F-statistic:            157.4
Date:                Wed, 03 Apr 2019      Prob (F-statistic):       6.51e-34
Time:                10:09:54      Log-Likelihood:           2219.2
No. Observations:          1135      AIC:                     -4436.
Df Residuals:              1134      BIC:                     -4431.
Df Model:                   1
Covariance Type:            nonrobust
=====

```

	coef	std err	t	P> t	[0.025	0.975]
x1	1.3101	0.104	12.548	0.000	1.105	1.515

```

=====
Omnibus:                 262.175      Durbin-Watson:           2.005
Prob(Omnibus):            0.000      Jarque-Bera (JB):        2947.095
Skew:                     0.726      Prob(JB):                 0.00
Kurtosis:                 10.760      Cond. No.                 1.00
=====

```

Warnings:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

$\beta_{TSLA} = 1.14$, **P is low, we reject the null hypothesis.**

2.2 (b)

let $h = (a, b, c)^T$

from self-financing condition, we get $a + b + c = 0$

from unit exposure to AAPL, we get $b = 1$

from zero exposure to beta, we get $a\beta_{IBM} + b\beta_{AAPL} + c\beta_{TSLA} = 0$

we use following code to solve this linear problem

```
In [33]: a = np.array([[1,1,1], [0,1,0],[0.91,1.14,1.31]])
         b = np.array([0.0,1.0,0.0])
         x = np.linalg.solve(a, b)
         x
```

```
Out[33]: array([-0.425,  1.    , -0.575])
```

$$\therefore h = (-0.425, 1, -0.575)^T$$

2.3 (C)

```
In [11]: d1=data1[data1.DATE>=pd.to_datetime("2015-01-01")]
         d2=data2[data2.DATE>=pd.to_datetime("2015-01-01")]
         d3=data3[data3.DATE>=pd.to_datetime("2015-01-01")]
```

```
In [17]: sum(d1.R*(-0.425))+sum(d2.R*1)+sum(d3.R*(-0.575))
```

```
Out[17]: -0.021396825000000078
```

cumulative sum of return in 2015 is -0.0214

2.4 (d)

```
In [13]: ret=np.array(d1.R*(-0.425))+np.array(d2.R*1)+np.array(d3.R*(-0.575))
         ret_m=np.array(data1[data1.DATE>=pd.to_datetime("2015-01-01")].RM)
```

```
In [14]: pearsonr(ret,ret_m)
```

```
Out[14]: (0.02866801914913187, 0.6512598504889181)
```

P-value is 0.65, not greater than 0.95. So it is not significantly different from 0.

3 Problem 4.2

3.1 (a)

$$\begin{aligned} E[h'r] &= \frac{1}{n} \sum_i E[r_i] = \beta E[r_m] + \frac{1}{n} \sum_i E[\epsilon_i] \\ V[h'r] &= \frac{1}{n^2} \sum_i (\beta^2 \sigma_m^2 + \sigma_i^2) + \frac{1}{n^2} \sum_{i,j,i \neq j} \beta^2 \sigma_m^2 = \beta^2 \sigma_m^2 + \sum_i \frac{1}{n^2} \sigma_i^2 \\ \therefore f(\beta, \sigma_m^2) &= \beta^2 \sigma_m^2 \quad g(\sigma_1^2, \dots, \sigma_n^2) = \frac{1}{n^2} \sum_i \sigma_i^2 \end{aligned}$$

3.2 (b)

```
In [2]: beta=0.5
        sigmaM=0.2
        sigmaI=0.03

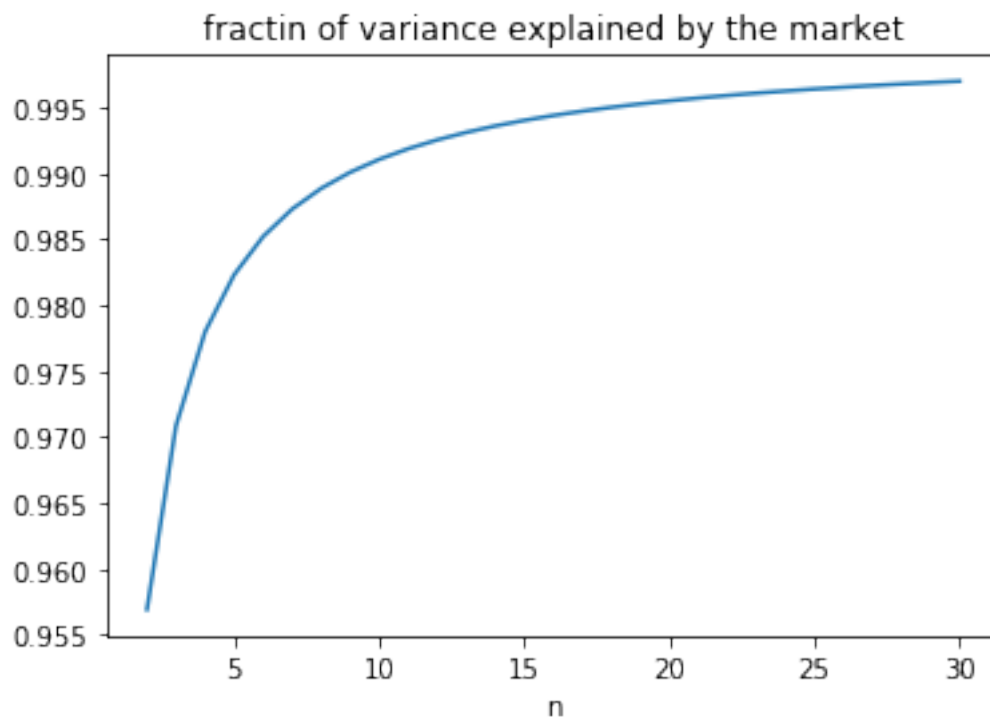
        def f(n):
            return beta**2*sigmaM**2

        def g(n):
            return sigmaI**2/n

        def fraction(n):
            return f(n)/(f(n)+g(n))

        x=[]
        y=[]
        for i in range(2,31):
            x.append(i)
            y.append(fraction(i))

        import matplotlib.pyplot as plt
        plt.plot(x,y)
        plt.title("fractin of variance explained by the market")
        plt.xlabel('n')
        plt.show()
```



3.3 (c)

In this case,

$$E[h'r] = \frac{1}{n} \sum_i \beta E[r_M] + E[\epsilon_i]$$

```
In [3]: Epsilon=1.5*sigmaI
        ErM=0.07

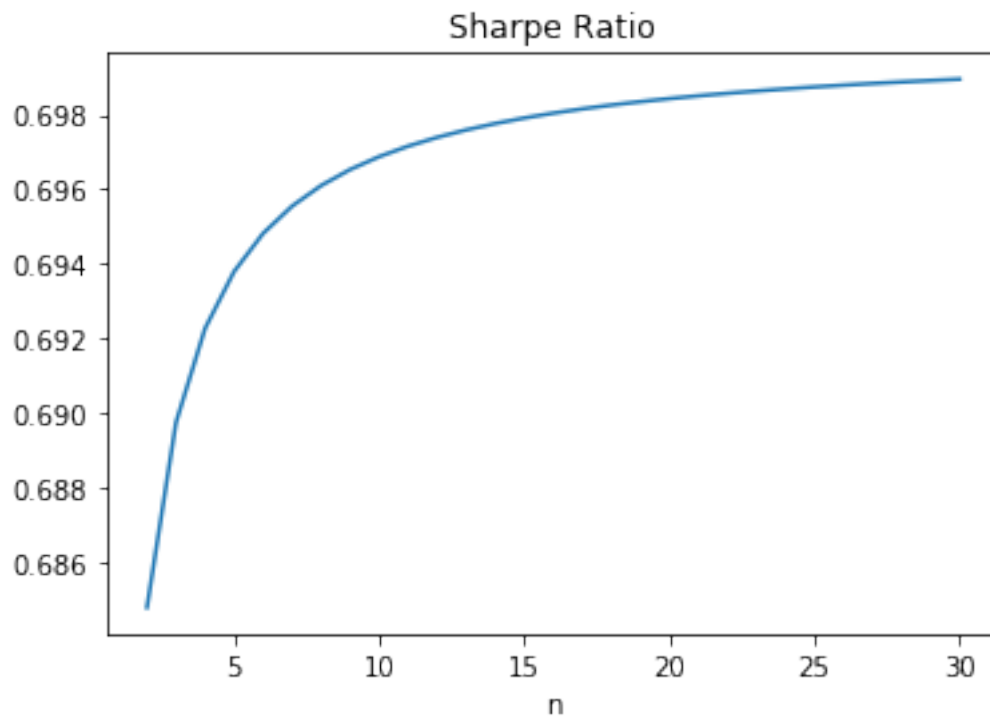
        def E(n):
            return beta*ErM+Epsilon

        def V(n):
            return f(n)+g(n)

        def Sharpe(n):
            return (E(n)-0.01)/np.sqrt(V(n))

        x=[]
        y=[]
        for i in range(2,31):
            x.append(i)
            y.append(Sharpe(i))

        plt.plot(x,y)
        plt.title("Sharpe Ratio")
        plt.xlabel('n')
        plt.show()
```

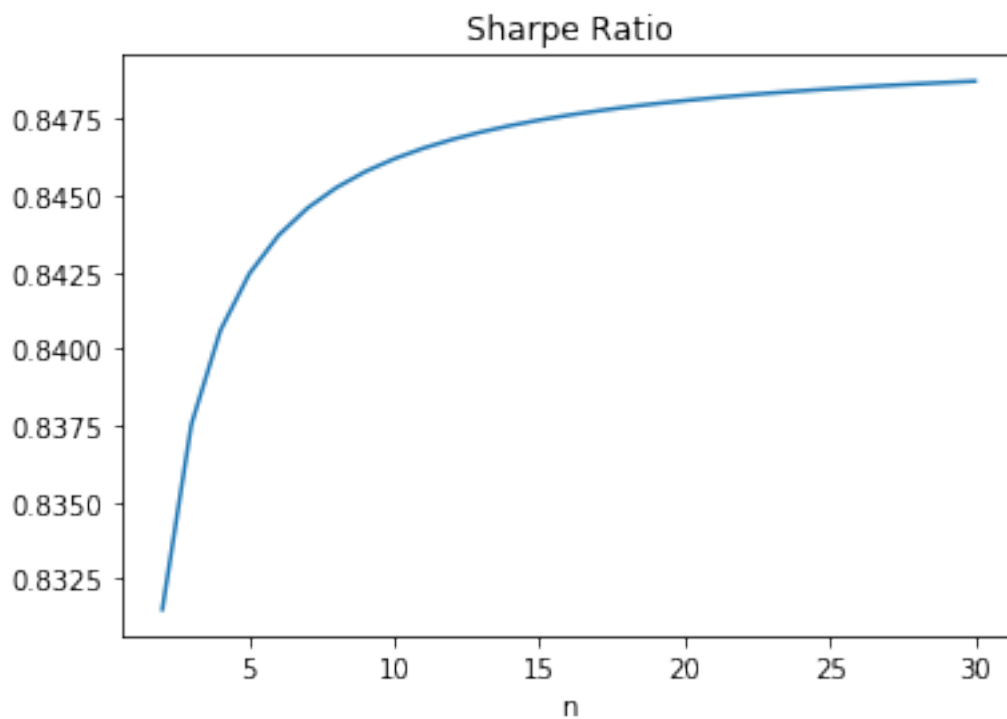


If we change Sharpe of ϵ_i from 1.5 to 2.0

```
In [4]: Epsilon=2.0*sigmaI

x=[]
y=[]
for i in range(2,31):
    x.append(i)
    y.append(Sharpe(i))

plt.plot(x,y)
plt.title("Sharpe Ratio")
plt.xlabel('n')
plt.show()
```



3.4 (d)

Suppose Sharpe of ϵ_i is 1.5

$$E[h'r] = E[\epsilon_i] = 1.5\sigma_i$$

$$V[h'r] = \sigma_i^2$$

$$\therefore \text{Sharpe} = \frac{1.5\sigma_i - 0.01}{\sigma_i} = 1.4$$

Compare to above scenario, this is much better.