

# Time Series Analysis- GroupH - Assignment 2

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### Question 1

```
In [1]: %matplotlib inline
```

```
import numpy as np
import matplotlib.pyplot as plt
from statsmodels.tsa.ar_model import AR
from statsmodels.tsa.arima_model import ARMA
import statsmodels as sm
import pandas as pd
```

```
In [2]: #Simulate AR(1)
```

```
alpha=0.1
beta = 0.3
sigma=0.005
T_list=[100,250,1250]
N = 2000

for T in T_list:
    alpha_sum = 0
    beta_sum = 0
    sigma_sum = 0
    for i in range(0,N):
        x0=alpha/(1-beta)
        x=np.zeros(T+1)
        x[0]=x0
        eps=np.random.normal(0.0,sigma,T)
        for i in range(1,T+1):
            x[i]=alpha+beta*x[i-1]+eps[i-1]

        y=x[0:T]
        yp=x[1:(T+1)]
        m=np.sum(y)/T
        mp=np.sum(yp)/T
        betaCMLE=np.inner(y-m,yp-mp)/np.inner(y-m,y-m)
        alphaCMLE=mp-betaCMLE*m
        sigmaCMLE=np.sqrt(np.inner(yp-betaCMLE*y-alphaCMLE,
                                   yp-betaCMLE*y-alphaCMLE)/T)

        alpha_sum += alphaCMLE
        beta_sum += betaCMLE
        sigma_sum += sigmaCMLE
```

```

alpha_hat = alpha_sum / N
beta_hat = beta_sum / N
sigma_hat = sigma_sum / N
print("T = ", T, ", (alpha,beta,sigma) = " , [alpha_hat,beta_hat,sigma_hat])

T = 100 , (alpha,beta,sigma) = [0.1024153805831, 0.2831247621365, 0.004951490497068]
T = 250 , (alpha,beta,sigma) = [0.1009745294900, 0.2932032633954, 0.004971390214618]
T = 1250 , (alpha,beta,sigma) = [0.1002329906964, 0.2983269719304, 0.004993956810083]

```

- \$ Comment\$
- We run simulations  $N(=2000)$  times to give  $T(=100,250,1250)$  observations from the given AR(1) model and then fit the observations into AR(1) to give the average MLE estimates.
- The output is:

$$T = 100, [\hat{\alpha}, \hat{\beta}, \hat{\sigma}] = [0.102415, 0.283124, 0.00495149]$$

$$T = 250, [\hat{\alpha}, \hat{\beta}, \hat{\sigma}] = [0.100974, 0.293203, 0.00497139]$$

$$T = 1250, [\hat{\alpha}, \hat{\beta}, \hat{\sigma}] = [0.100232, 0.298326, 0.00499396]$$

we can see that the estimates are biased compare with the given parameters

$$[\alpha, \beta, \sigma] = [0.1, 0.3, 0.005]$$

- for these biased estimates, we can see that when  $T$  goes larger, the bias get smaller because the error is  $O(\frac{1}{N})$ , which means that they are consistent.

## Question 2

```

In [3]: import numpy as np
import pandas as pd
import matplotlib.pyplot as plt
from statsmodels.tsa.stattools import adfuller

```

```

In [4]: def perform_adf_test(xt):
        """
        input:  a timeseries representing the daily difference between SPY and IWV returns
        output: outputs the statistic values from the ADF test
        """
        test = adfuller(xt)
        print('ADF Statistic: %f' % test[0])
        print('p-value: %f' % test[1])
        for key, value in test[4].items():
            print('\t%s: %.3f' % (key, value))

```

```

In [5]: def problem_2():
    # Section 2.1: Generate time series
    eur = pd.read_excel('HW2_Data.xlsx', sheet_name = 'EUR',
                        index_col=0, parse_dates=True, infer_datetime_format=True)
    fed = pd.read_excel('HW2_Data.xlsx', sheet_name = 'FEDL01',
                        index_col=0, parse_dates=True, infer_datetime_format=True)
    ecb = pd.read_excel('HW2_Data.xlsx', sheet_name = 'EUORDEPOT',
                        index_col=0, parse_dates=True, infer_datetime_format=True)

    diff = fed - ecb

    plt.figure()
    plt.plot(eur,label='EUR')
    plt.title('EUR/USD')
    plt.xlabel('Date')
    plt.ylabel('Last')
    plt.legend()
    plt.show()

    plt.plot(diff,label='Rates differential')
    plt.title('Interest Rates Differential')
    plt.xlabel('Date')
    plt.ylabel('Last')
    plt.legend()
    plt.show()

    # run Dickey-Fuller on each series separately.
    eur_rates = eur['Last Price'].dropna()
    diff_interest_rates = diff['Last Price'].dropna()
    print("EURUSD")
    perform_adf_test(eur_rates)
    print("Diff")
    perform_adf_test(diff_interest_rates)

    # this should be stationary
    print("delta EURUSD")
    lagged_eur = eur['Last Price'].pct_change(1).dropna()
    perform_adf_test(lagged_eur)

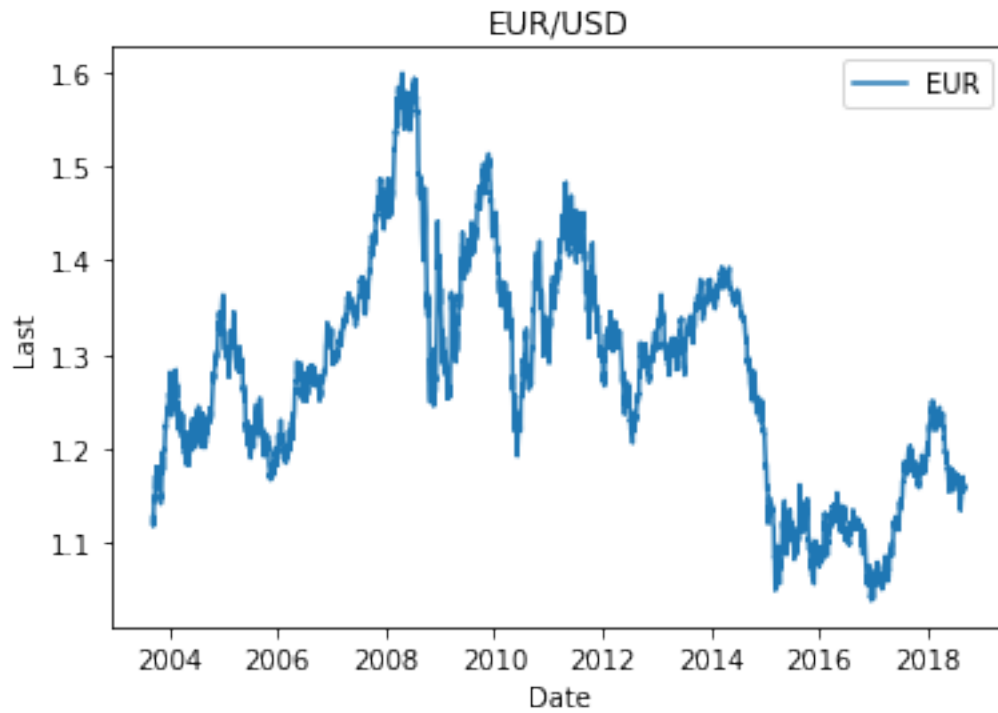
    # this too should be stationary
    print("delta rates")
    lagged_rates = diff['Last Price'].pct_change(1).replace([np.inf, -np.inf], np.nan).dropna()
    perform_adf_test(lagged_rates)

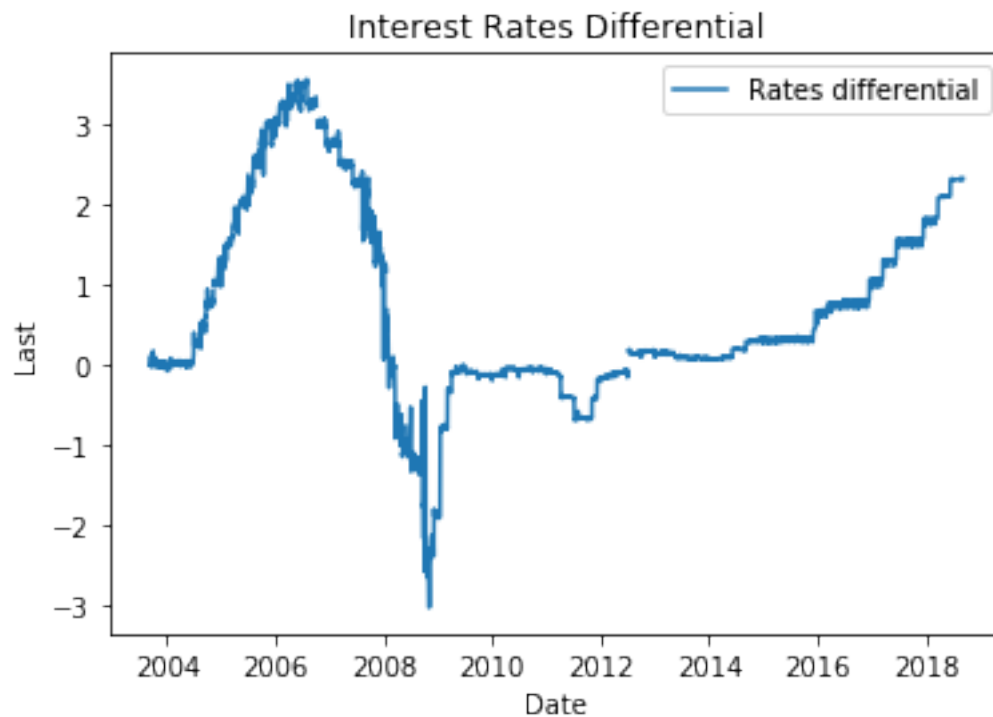
    # form vector  $a = (1, -1)'$  ==>  $u_t = x - y$ 
    merged_data = pd.concat([eur_rates.rename('eur'), diff_interest_rates.rename('rates')])
    clean_data = merged_data.dropna()
    u = (clean_data['eur'] - clean_data['rates'])

```

```
# test unit root on u, this would have to be stationary  
print("u")  
perform_adf_test(u)
```

In [9]: problem\_2()





EURUSD

ADF Statistic: -1.966646

p-value: 0.301386

5%: -2.862

1%: -3.432

10%: -2.567

Diff

ADF Statistic: -1.016293

p-value: 0.747276

5%: -2.862

1%: -3.432

10%: -2.567

delta EURUSD

ADF Statistic: -62.486528

p-value: 0.000000

5%: -2.862

1%: -3.432

10%: -2.567

delta rates

ADF Statistic: -9.344125

p-value: 0.000000

5%: -2.862

1%: -3.432

10%: -2.567

u

ADF Statistic: -0.930077

p-value: 0.777906

5%: -2.862

1%: -3.432

10%: -2.567

#### *Comment*

First, we run ADF on EURUSD, the short interest difference between Fed and ECB. The test result shows that both series are not stationary at 5% significance level. And the first difference of both series are stationary at 5% significance level. Thus we can conclude that both series are I(1).

Second, we use the given cointegration vector  $\alpha = (1, -1)$  to calculate the residual of the cointegration process  $u_t$ , which  $u_t = EURUSD - Diff$ . Then we test the stationarity of the residuals. The test shows that the residual is not stationary at 5% significance level. Thus the vector  $\alpha = (1, -1)$  is not the cointegration vector of EURUSD exchange rate and the differential between Fed and ECB short interest rates in these currencies.

TS. HW2. Group H.

Q3: i we first calculate  $p(x_1, x_2, \dots, x_T | \theta, x_0)$

$$p(x_1, \dots, x_T | x_0, \theta) = \frac{1}{(2\pi\sigma^2)^{T/2}} \cdot \exp\left(-\frac{1}{2\sigma^2} \sum_{t=1}^T \varepsilon_t^2\right).$$

(Note that  $\hat{\varepsilon}_t = x_t - \alpha - \beta x_{t-1} - \delta t \sim N(0, \sigma^2)$ ).

$$\text{Thus. } p(x_1, \dots, x_T | x_0, \theta) = \frac{1}{(2\pi\sigma^2)^{T/2}} \cdot \exp\left(-\frac{1}{2\sigma^2} \sum_{t=1}^T (x_t - \alpha - \beta x_{t-1} - \delta t)^2\right).$$

ii draw log-likelihood estimation.

$$-\log L(\theta | y) = \frac{1}{2} T \log \sigma^2 + \frac{1}{2\sigma^2} \sum_{t=1}^T (x_t - \alpha - \beta x_{t-1} - \delta t)^2 + \text{constant}$$

Minimizing this function yields:

$$\begin{pmatrix} \hat{\alpha} \\ \hat{\beta} \\ \hat{\delta} \end{pmatrix} = \begin{pmatrix} T & \sum_{t=1}^T x_{t-1} & \sum_{t=1}^T t \\ \sum_{t=1}^T x_{t-1} & \sum_{t=1}^T x_{t-1}^2 & \sum_{t=1}^T t x_{t-1} \\ \sum_{t=1}^T t & \sum_{t=1}^T t x_{t-1} & \sum_{t=1}^T t^2 \end{pmatrix}^{-1} \begin{pmatrix} \sum_{t=1}^T x_t \\ \sum_{t=1}^T x_t x_{t-1} \\ \sum_{t=1}^T t x_t \end{pmatrix}.$$

$$\hat{\sigma}^2 = \frac{1}{T} \sum_{t=1}^T (x_t - \hat{\alpha} - \hat{\beta} x_{t-1} - \hat{\delta} t)^2.$$