

Time Series Analysis- GroupH - Assignment 2

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Question 1

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In [1]: %matplotlib inline
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```
import numpy as np
import matplotlib.pyplot as plt
from statsmodels.tsa.ar_model import AR
from statsmodels.tsa.arima_model import ARMA
import statsmodels as sm
import pandas as pd
```

```
In [2]: #Simulate AR(1)
```

```
alpha=0.1
beta = 0.3
sigma=0.005
T_list=[100,250,1250]
N = 2000

for T in T_list:
    alpha_sum = 0
    beta_sum = 0
    sigma_sum = 0
    for i in range(0,N):
        x0=alpha/(1-beta)
        x=np.zeros(T+1)
        x[0]=x0
        eps=np.random.normal(0.0,sigma,T)
        for i in range(1,T+1):
            x[i]=alpha+beta*x[i-1]+eps[i-1]

        y=x[0:T]
        yp=x[1:(T+1)]
        m=np.sum(y)/T
        mp=np.sum(yp)/T
        betaCMLE=np.inner(y-m,yp-mp)/np.inner(y-m,y-m)
        alphaCMLE=mp-betaCMLE*m
        sigmaCMLE=np.sqrt(np.inner(yp-betaCMLE*y-alphaCMLE,
                                   yp-betaCMLE*y-alphaCMLE)/T)

        alpha_sum += alphaCMLE
        beta_sum += betaCMLE
        sigma_sum += sigmaCMLE
```

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alpha_hat = alpha_sum / N
beta_hat = beta_sum / N
sigma_hat = sigma_sum / N
print("T = ", T, ", (alpha,beta,sigma) = " , [alpha_hat,beta_hat,sigma_hat])

T = 100 , (alpha,beta,sigma) = [0.1024153805831, 0.2831247621365, 0.004951490497068]
T = 250 , (alpha,beta,sigma) = [0.1009745294900, 0.2932032633954, 0.004971390214618]
T = 1250 , (alpha,beta,sigma) = [0.1002329906964, 0.2983269719304, 0.004993956810083]

```

Comment

- We run simulations $N(=2000)$ times to give $T(=100,250,1250)$ observations from the given AR(1) model and then fit the observations into AR(1) to give the average MLE estimates.
- The output is:

$$T = 100, [\hat{\alpha}, \hat{\beta}, \hat{\sigma}] = [0.102415, 0.283124, 0.00495149]$$

$$T = 250, [\hat{\alpha}, \hat{\beta}, \hat{\sigma}] = [0.100974, 0.293203, 0.00497139]$$

$$T = 1250, [\hat{\alpha}, \hat{\beta}, \hat{\sigma}] = [0.100232, 0.298326, 0.00499396]$$

we can see that the estimates are biased compare with the given parameters

$$[\alpha, \beta, \sigma] = [0.1, 0.3, 0.005]$$

- for these biased estimates, we can see that when T goes larger, the bias get smaller because the error is $O(\frac{1}{N})$, which means that they are consistent.

Question 2

```

In [7]: import numpy as np
import pandas as pd
import matplotlib.pyplot as plt
from statsmodels.tsa.stattools import adfuller

In [37]: def perform_adf_test(xt):
    """
    input: a timeseries
    output: outputs the statistic values from the ADF test
    """
    test = adfuller(xt)
    print('ADF Statistic: %f' % test[0])
    print('p-value: %f' % test[1])
    for key, value in test[4].items():
        print('\t%s: %.3f' % (key, value))

def plot_series(eur, diff_rates):
    plt.figure()

```

```

plt.plot(eur,label='EUR')
plt.plot(diff_rates,label='Rates differential')
plt.title('EUR/USD')
plt.xlabel('Date')
plt.ylabel('Last')
plt.legend()
plt.show()

def plot_coint_residual(u):
    plt.figure()
    plt.plot(u,label='Cointegrating residual')
    plt.title('EUR/USD - (FED-ECB)')
    plt.xlabel('Date')
    plt.ylabel('Last')
    plt.legend()
    plt.show()

In [54]: def perform_engle_granger_test(eur, diff_rates, coint_residual, confidence):
        """
        input: a cointegrating residual vector and the level of confidence
        output: True if the series are cointegrated
        """
        # run Dickey-Fuller on each series separately to check if they are I(d).
        print("EURUSD")
        eur_p = perform_adf_test(eur)
        print("Diff")
        diff_p = perform_adf_test(diff_rates)
        # this should be stationary
        print("delta EURUSD")
        lagged_eur = eur.diff().dropna()
        lag_eur_p = perform_adf_test(lagged_eur)

        # this too should be stationary
        print("delta rates")
        lagged_rates = diff_rates.diff().dropna()
        lag_rates_p = perform_adf_test(lagged_rates)

        # test unit root on the cointegration residual process
        # this would need to be stationary to say there is cointegration

        print("u")
        perform_adf_test(coint_residual)
        test = adfuller(coint_residual)
        p_value=test[1]
        if (p_value>confidence):
            print("The cointegrating residual is not stationary")

```

```

        return False

    print ("The series are cointegrated")
    return True

```

```

In [57]: def problem_2():
    # Section 2.1: Perform a cointegration test
    # Read data and generate time series
    eur = pd.read_excel('HW2_Data.xlsx', sheet_name = 'EUR',
                        index_col=0, parse_dates=True, infer_datetime_format=True)
    fed = pd.read_excel('HW2_Data.xlsx', sheet_name = 'FEDL01',
                        index_col=0, parse_dates=True, infer_datetime_format=True)
    ecb = pd.read_excel('HW2_Data.xlsx', sheet_name = 'EUORDEPOT',
                        index_col=0, parse_dates=True, infer_datetime_format=True)

    diff = fed - ecb

    plot_series(eur, diff)
    eur_rates = eur['Last Price'].dropna()
    diff_interest_rates = diff['Last Price'].dropna()
    merged_data = pd.concat([eur_rates.rename('eur'), diff_interest_rates.rename
    ('rates')], axis=1)
    print(merged_data.head())
    clean_data = merged_data.dropna()
    u = (clean_data['rates'] - clean_data['eur'])

    plot_coint_residual(u)

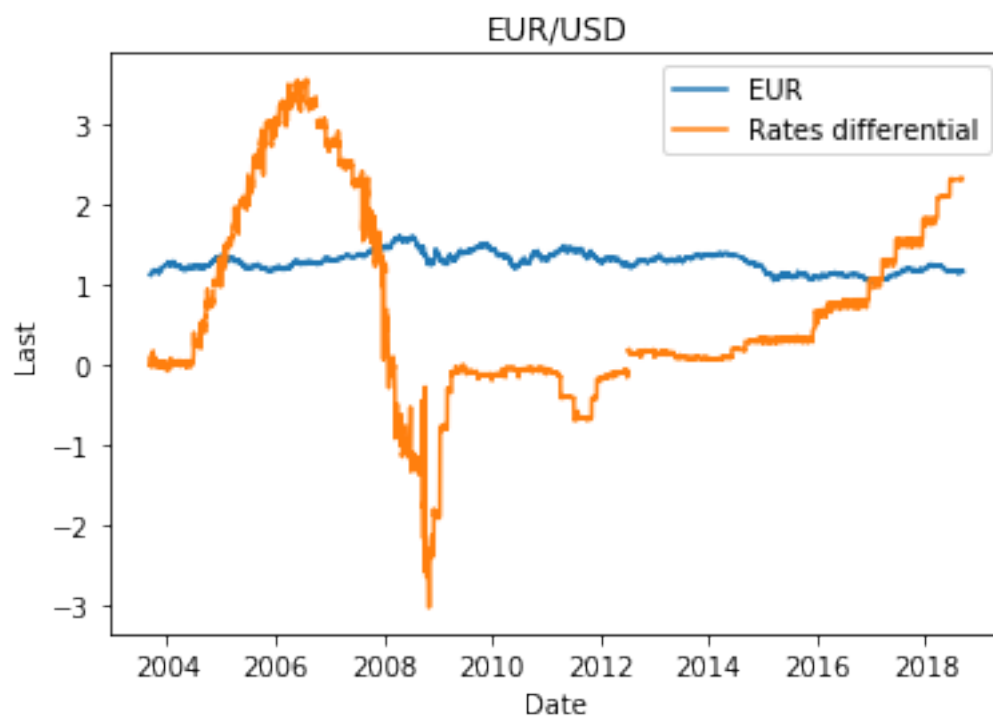
    perform_engle_granger_test(eur_rates, diff_interest_rates, u, confidence=0.05)

```

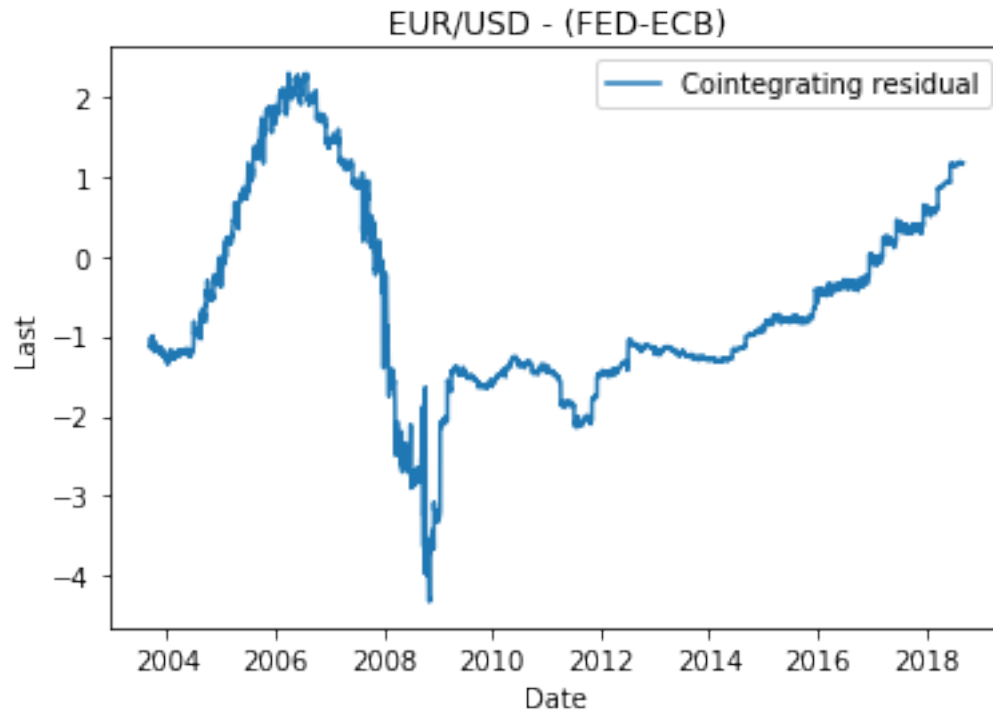
```

In [58]: problem_2()

```



Date	eur	rates
2003-09-11	1.1210	0.00
2003-09-12	1.1291	0.02
2003-09-15	1.1284	0.11
2003-09-16	1.1177	-0.03
2003-09-17	1.1285	-0.03



EURUSD

ADF Statistic: -1.966646

p-value: 0.301386

5%: -2.862

1%: -3.432

10%: -2.567

Diff

ADF Statistic: -1.016293

p-value: 0.747276

5%: -2.862

1%: -3.432

10%: -2.567

delta EURUSD

ADF Statistic: -62.279418

p-value: 0.000000

5%: -2.862

1%: -3.432

10%: -2.567

delta rates

ADF Statistic: -9.600190

p-value: 0.000000

5%: -2.862

1%: -3.432

10%: -2.567

u
ADF Statistic: -0.930077
p-value: 0.777906
5%: -2.862
1%: -3.432
10%: -2.567
The cointegrating residual is not stationary

Comment

First, we run ADF on EURUSD, the short interest difference between Fed and ECB. The test result shows that both series are not stationary at 5% significance level. And the first difference of both series are stationary at 5% significance level. Thus we can conclude that both series are $I(1)$.

Second, we use the given cointegration vector $\alpha = (1, -1)$ to calculate the residual of the cointegration process u_t . Then we test the stationarity of the residuals. The test shows that the residual is not stationary at 5% significance level. Thus the vector $\alpha = (1, -1)$ is not the cointegration vector of EURUSD exchange rate and the differential between Fed and ECB short interest rates in these currencies.

TS. HW2. Group H.

Q3: i we first calculate $p(x_1, x_2, \dots, x_T | \theta, x_0)$

$$p(x_1, \dots, x_T | x_0, \theta) = \frac{1}{(2\pi\sigma^2)^{T/2}} \cdot \exp\left(-\frac{1}{2\sigma^2} \sum_{t=1}^T \varepsilon_t^2\right).$$

(Note that $\hat{\varepsilon}_t = x_t - \alpha - \beta x_{t-1} - \delta t \sim N(0, \sigma^2)$).

$$\text{Thus. } p(x_1, \dots, x_T | x_0, \theta) = \frac{1}{(2\pi\sigma^2)^{T/2}} \cdot \exp\left(-\frac{1}{2\sigma^2} \sum_{t=1}^T (x_t - \alpha - \beta x_{t-1} - \delta t)^2\right).$$

ii draw log-likelihood estimation.

$$-\log L(\theta | y) = \frac{1}{2} T \log \sigma^2 + \frac{1}{2\sigma^2} \sum_{t=1}^T (x_t - \alpha - \beta x_{t-1} - \delta t)^2 + \text{constant}$$

Minimizing this function yields:

$$\begin{pmatrix} \hat{\alpha} \\ \hat{\beta} \\ \hat{\delta} \end{pmatrix} = \begin{pmatrix} T & \sum_{t=1}^T x_{t-1} & \sum_{t=1}^T t \\ \sum_{t=1}^T x_{t-1} & \sum_{t=1}^T x_{t-1}^2 & \sum_{t=1}^T t x_{t-1} \\ \sum_{t=1}^T t & \sum_{t=1}^T t x_{t-1} & \sum_{t=1}^T t^2 \end{pmatrix}^{-1} \begin{pmatrix} \sum_{t=1}^T x_t \\ \sum_{t=1}^T x_t x_{t-1} \\ \sum_{t=1}^T t x_t \end{pmatrix}.$$

$$\hat{\sigma}^2 = \frac{1}{T} \sum_{t=1}^T (x_t - \hat{\alpha} - \hat{\beta} x_{t-1} - \hat{\delta} t)^2.$$