

MA3676 - Stochastic Models

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1 Week 1

1.1 Axioms of Probability and conditional probabilities

1.1.1 Axioms of probability

- **Axiom 1:** To each event E there corresponds a number $\mathbb{P}[E] \geq 0$, where $\mathbb{P}[E]$ is called the probability of E .
- **Axiom 2:** $\mathbb{P}[\mathbb{S}] = 1$.
- **Axiom 3:** If events A_1, A_2, A_3, \dots are mutually exclusive then

$$\mathbb{P}\left[\bigcup_{i=1}^{\infty} A_i\right] = \sum_{i=1}^{\infty} \mathbb{P}[A_i]. \quad (1)$$

1.1.2 Conditional probability

Definition 1

$$\mathbb{P}[A|B] = \frac{\mathbb{P}[A \cap B]}{\mathbb{P}[B]}. \quad (2)$$

Definition 2 Two events A and B are independent if $\mathbb{P}[A \cap B] = \mathbb{P}[A]\mathbb{P}[B]$.

Definition 3

$$\mathbb{P}[A_i|B] = \frac{\mathbb{P}[A_i]\mathbb{P}[B|A_i]}{\sum_{j=1}^n \mathbb{P}[A_j]\mathbb{P}[B|A_j]}. \quad (3)$$

2 Week 2

2.1 Discrete random variables

2.1.1 Bernoulli distribution

Definition 4 If the probability of event E is p , then the variable I_E is an indicator variable, which is said to follow the Bernoulli distribution with parameter P

$$\mathbb{P}[I_E = 1] = p, \mathbb{P}[I_E = 0] = 1 - p, \quad (4)$$

or in shorthand, $\mathbb{P}[1] = p, \mathbb{P}[0] = 1 - p$.

2.1.2 Independence of random variables

Definition 5 Two random variables X and Y are independent if for any possible value x of X and any possible value y of Y the events $\{X = x\}$ and $\{Y = y\}$ are independent.

2.1.3 Binomial distribution

Definition 6

$$\mathbb{P}[\{Y_n = m\}] = \binom{n}{m} p^m 1^{n-m}. \quad (5)$$

2.1.4 Poisson distribution

Definition 7

$$P(m) = \frac{\lambda^m}{m!} e^{-\lambda}. \quad (6)$$

2.1.5 Geometric Distribution

Definition 8

$$P(n) = p^{n-1}(1 - p). \quad (7)$$

2.2 Sums and Series

2.2.1 Finite geometric sum

Definition 9

$$S_n = \sum_{k=0}^n a^k = 1 + a + a^2 + \cdots + a^n, \quad (8)$$

hence

$$S_n = \frac{1 - a^{n+1}}{1 - a}. \quad (9)$$

2.2.2 Convergence of series

In order for the series $\sum_{n=1}^{\infty} b_n$ to converge, the limit $\lim_{A \rightarrow \infty} (b_1 + b_2 + b_3 + \cdots + b_A)$ must exist.

A necessary condition for the convergence of this series is that the limit of the sequence b_n exists and is equal to 0:

$$\lim_{n \rightarrow \infty} b_n = 0 \quad (10)$$

2.2.3 Series examples

Definition 10 *Geometric series:*

$$\sum_{n=0}^{\infty} a^n = \frac{1}{1 - a}. \quad (11)$$

Definition 11 *Power series:*

$$\sum_{n=0}^{\infty} \frac{x^n}{n!} = e^x. \quad (12)$$

Definition 12 *Logarithmic series:*

$$\sum_{n=1}^{\infty} \frac{x^n}{n} = -\ln(1 - x). \quad (13)$$

Definition 13 *Trigonometric series:*

$$\sum_{n=0}^{\infty} \frac{(-1)^n x^{2n+1}}{(2n+1)!} = \sin x, \quad (14)$$

$$\sum_{n=0}^{\infty} \frac{(-1)^n x^{2n}}{(2n)!} = \cos x. \quad (15)$$

3 Week 3

3.1 Expectation value

3.1.1 Expectation of a random variable

Definition 14 Consider a discrete random variable X which takes its values in a set $\{x_1, x_2, \dots\}$ with probabilities $\mathbb{P}[\{X = x_i\}] = p_i$. Then, the expectation value of X is defined as

$$\mathbb{E}[X] = \sum_i x_i p_i. \quad (16)$$

3.1.2 Expectation of functions of random variables

Definition 15

$$\mathbb{E}[g(X)] = \sum_i g(x_i) p_i. \quad (17)$$

3.1.3 Linearity of expectation

Definition 16

$$\mathbb{E}[aX + bY] = a\mathbb{E}[X] + b\mathbb{E}[Y]. \quad (18)$$

3.2 Expectation and variance values

3.2.1 Variance of a random variable

Definition 17

$$\text{Var}[X] = \mathbb{E}[(X - \mathbb{E}[X])^2]. \quad (19)$$

3.2.2 Variance of a binomial random variable

Definition 18

$$\text{Var}\left[\sum_{i=1}^n X_i\right] = n \text{Var}[X]. \quad (20)$$

4 Week 4

4.1 Generating functions

4.1.1 The concept of a stochastic process: terminology

- The collection of all indices marking the random variables: the index set.
- All possible values of each X_i : the state space.
- Particular values x_i taken by X_i are the states of the process.
- The elements of the index set are conventionally called times.
- The whole collection of values $\{x_i\}_{i=1}^N$ is a realisation of the process run for N steps.
- A stochastic process can be viewed as a single, multi-component random variable.

4.1.2 Definitions and basic properties

Definition 19 A (discrete) “Random Walk” is a stochastic process S_n which can be represented as a sum of random variables (“steps”):

$$S_n = S_0 + \sum_{k=1}^n X_k. \quad (21)$$

Definition 20 A “Simple Random Walk” is obtained if X_k are independently identically distributed (i.i.d.) modified Bernoulli random variables taking values in the set $\{-1, 1\}$, i.e.

$$\mathbb{P}[X_k = 1] = p, \quad \mathbb{P}[X_k = -1] = 1, \quad p + q = 1, \quad \forall k \in \mathbb{N}.$$

4.1.3 Generating functions: definition

Definition 21 Consider a discrete random variable X , and denote $\mathbb{P}[X = n] = p_n$. We now define the corresponding generating function $G(s)$ as

$$G(s) = \mathbb{E}[s^X]. \quad (22)$$

Definition 22

$$G(s) = \sum_n p_n s^n. \quad (23)$$

4.1.4 Expectation of a generating function

Definition 23

$$\mathbb{E}[X] = G'(1) \quad (24)$$

4.1.5 Variance of a generating function

Definition 24

$$\text{Var}[X] = G''(1) + G'(1) - G'(1)^2 \quad (25)$$

5 Week 5

5.1 Random Walks

5.1.1 Probability mass function

Definition 25

$$G_{S_n} = \sum_{m=0}^n \binom{n}{m} p^{n-m} q^m s^{n-2m}, \quad (26)$$

where the powers of s correspond to values of S_n .

Definition 26 *The probability of displacement by k is denoted as*

$$P_k^{(n)} \equiv \mathbb{P}[S_n - S_0 = k] = \binom{n}{\frac{n-k}{2}} p^{\frac{n+k}{2}} q^{\frac{n-k}{2}}. \quad (27)$$

5.1.2 General random walks

Theorem 1 *Suppose $S_n = \sum_{k=1}^n X_k$ is a random walk such that X_k 's are i.i.d. random variables, characterised by a generating function $G_X(s) = \mathbb{E}[s^X]$. Then, the generating function of S_n is*

$$G_{S_n}(s) = G_X(s)^n \quad (28)$$

5.1.3 Returning to the initial position

Theorem 2 *For an unrestricted unbiased one-dimensional random walk the probability of return $\mathcal{R} = 1$.*

Here, the number of returns N to the state i is represented by $\mathbb{E}[N]$.

Lemma 1

$$\mathbb{E}[N] = \sum_{n=1}^{\infty} P_0^{(n)}. \quad (29)$$

Lemma 2 *If $\mathcal{R} < 1$, then*

$$\mathbb{E}[N] = \frac{\mathcal{R}}{1 - \mathcal{R}}. \quad (30)$$

6 Week 6

6.1 Gambler's ruin

Denote events such that:

- R_n : player A is eventually ruined starting from initial capital n .
- W_n : player A wins the first game starting from initial capital n .
- $L_n = \overline{W}_n$: player A loses the first game starting from initial capital n .

6.1.1 Probability of ruin

Definition 27 The probability of ruin of a player from initial capital n is

$$\mathbb{P}[R_n] = \mathbb{P}[R_n|W_n]\mathbb{P}[W_n] + \mathbb{P}[R_n|L_n]\mathbb{P}[L_n], \quad (31)$$

which simplifies to

$$P_n = pP_{n+1} + qP_{n-1}, \quad (32)$$

where $\mathbb{P}[R_n|L_n] = P_{n-1}$, $\mathbb{P}[R_n|W_n] = P_{n+1}$, $\mathbb{P}[W_n] = p$ and $\mathbb{P}[L_n] = q = 1 - p$.

Definition 28 The probability of ruin from initial capital a with total pool $a + b$ is given by

$$P_a = \frac{1 - (1/p)^b}{(p/q)^a - (1/p)^b}, \quad (33)$$

where $p \neq q$. If $p = q = 1/2$, then

$$P_a = \frac{b}{a + b}. \quad (34)$$

7 Week 7

7.1 First step decomposition

7.1.1 Time to absorption

Definition 29 Let T_n be the random variable equal to the number of steps, starting from n , which the random walker will make before reaching either of the boundaries. Denote as

$$\mathcal{T} = \mathbb{E}[T_n]$$

the expected number of steps until the walker reaches either of the boundaries.

Definition 30

$$\mathcal{T}_n = A + B(q/p)^n + \frac{n}{q - p}, \quad p \neq q, \quad (35)$$

$$\mathcal{T}_n = A + Bn - n^2, \quad q = p = 1/2. \quad (36)$$

8 Week 8

8.1 Boundary Conditions

Use first step decomposition on the boundary conditions to solve for A and B. Use examples to understand.

9 Week 9

9.1 Introduction to Markov chains

Definition 31 A stochastic process S_n with discrete state space and discrete index set is a Markov chain if $\mathbb{P}[S_{n+1} = j | S_0 = i_0, S_1 = i_1, \dots, S_n = i_n] = \mathbb{P}[S_{n+1} = j | S_n = i_n]$, $\forall j, i_0, i_1, \dots, i_n$ and $\forall n \in \mathbb{N}$.

This says that the future (S_{n+1}) is conditionally independent of the past $(\{S_i\}_{i=1}^{n-1})$, given the knowledge of the present S_n .

9.2 Markov chains

9.2.1 Random walk on a circle

Definition 32 Treat p_{ij} as elements of an $(n \times n)$ -matrix and $\mathbf{p}_i(n)$ as elements of an n -component row vector $\underline{\pi}(n)$ where the matrix \mathbf{p} provides the full description of transitions and $\underline{\pi}$ the current states of each position. The total probability decomposition is a product of these such that:

$$\underline{\pi}(n+1) = \underline{\pi}(n)\mathbf{p}. \quad (37)$$

9.2.2 Chapman-Kolmogorov equation

Theorem 3

$$\pi(n) = \pi(0)\mathbf{p}^n. \quad (38)$$

9.2.3 Multiplying stochastic matrices

Theorem 4 *A product of two (or more, by induction) stochastic matrices is again a stochastic matrix.*

10 Week 10

10.1 Markov chains

10.1.1 Equilibrium distribution

Theorem 5 *If such a distribution exists, then it means formally the existence of the limit*

$$\lim_{n \rightarrow \infty} \pi(n) = \pi, \quad (39)$$

and therefore

$$\pi = \pi\mathbf{p}. \quad (40)$$

10.1.2 Matrix analysis of Markov chains

Theorem 6 $\lambda = 1$ *is an eigenvalue of any stochastic matrix* \mathbf{p} .

11 Week 17

11.1 Markov chains

11.1.1 Revisiting a position

In a random walk with only ± 1 steps allowed, even positions can be reached with an even number of steps and vice versa. Thus a position i can be revisited only after an even number of steps:

$$P_{ii}^{(2m)} = \frac{(2m)!}{(m!)^2} p^m q^m \quad \text{and} \quad P_{ii}^{(2m-1)} = 0 \quad (41)$$

for any initial position.

Definition 33 *A state, return to which is only possible after $m, 2m, 3m, \dots$ steps is called periodic if $m \geq 2$. Any other state is called aperiodic.*

Definition 34 *If the probability of return to a given state is equal to one, such a state is called recurrent. If this probability is less than one, the state is called transient.*

Note that this definition refers to the eventual probability of return, after an arbitrary number of steps.

11.1.2 Classification of states

These concepts apply to general discrete stochastic processes; in Markov chain applications $P_{ij}^{(m)}$ is $(p^{(m)})_{ij}$.

Definition 35 *A state i is said to communicate with a state j if it is possible eventually to reach j starting from i , i.e., if $P_{ij}^{(m)} > 0$ for some m .*

Definition 36 *States i and j intercommunicate if, in addition, $\exists m' : P_{ji}^{(m')} > 0$.*

This property is denoted as $i \leftrightarrow j$, and it is transitive:

$$\text{If } i \leftrightarrow j \text{ and } j \leftrightarrow k, \text{ then } i \leftrightarrow k. \quad (42)$$

11.1.3 Recurrence time

Starting from a state i the probability that the chain will next visit the same state after m steps can be formally expressed as

$$f_{ii}^{(m)} = \mathbb{P}[(X_{n+m} = i \cap (X_{n+k} \neq i \forall k : 1 \leq k \leq m-1) | X_n = i)]. \quad (43)$$

The sum

$$R_i = \sum_{m=1}^{\infty} f_{ii}^{(m)} \quad (44)$$

is the probability of return, so $R_i = 1$ for a recurrent state i . This means that $f_{ii}^{(m)}$ has the meaning of a discrete probability distribution of recurrence “times” $T_i : f_{ii}^{(m)} = \mathbb{P}[T_i = m]$.

Thus $\mu_i = \mathbb{E}[T_i] = \sum_{m=1}^{\infty} m f_{ii}^{(m)}$ is the expectation of the recurrence time - often known as the mean recurrence time for a state i .

Definition 37 Depending on whether μ_i is finite or infinite, all recurrent states are classified as positive-recurrent, or null-recurrent, respectively.

Definition 38 A positive-recurrent aperiodic state is ergodic.

Two more special types of states are:

- **Ephemeral** states: those states i for which $p_{ki} = 1$, for any k , and
- **Absorbing** states: those states i for which $p_{ii} = 1$.

11.1.4 States are “contagious”

Theorem 7 Two intercommunicating states must be of the same type.

Theorem 8 Let i be a recurrent state which communicates with j ; then i and j intercommunicate, and j is also a recurrent state of the same type and period as i .

Corollary 1 A pair of recurrent states of the same type either intercommunicate, or neither of them communicates with the other.

Corollary 2 A recurrent state of one type cannot communicate with a recurrent state of another type, nor with a transient state.

Note, a transient state can communicate with other transient states, as well as with both types of recurrent states.

11.1.5 Closed sets of states

Definition 39 A set C of states is closed if no state in C communicates with any state outside of C . A sufficient condition for this is

$$p_{ij} = 0 \forall i \in C \text{ and } \forall j \notin C. \quad (45)$$

Definition 40 A closed set C of states is irreducible if every pair of states in C intercommunicate.

- Since a recurrent state cannot communicate with a transient state, the set of all recurrent states is closed, but not necessarily irreducible.
- If a closed set contains only one state then that state is absorbing.
- An irreducible closed set contains no smaller closed set.

Theorem 9 The state space \mathbb{S} of a Markov chain can be uniquely partitioned as

$$\mathbb{S} = T \cup C_1 \cup C_2 \cup \dots$$

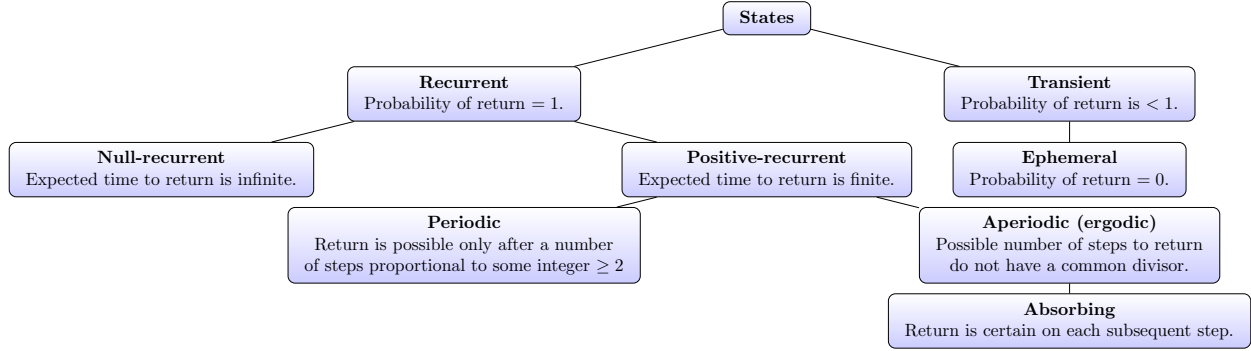
where T is the set of all transient states, and C_1, C_2, \dots are irreducible closed sets of recurrent states, each set containing states of the same type.

Theorem 10 If the state space is finite, then at least one state is recurrent, and, assuming the chain is time-homogeneous, all recurrent states are positive recurrent.

11.1.6 Corollaries of the decomposition theorem

Corollary 3 *If the chain starts in, or eventually reaches one of the irreducible closed sets C_r , then it never leaves it. C_r thus becomes the state space for the subsequent part of the realisation of the chain, and every state in C_r is visited infinitely often. Thus the subset C_r , together with the corresponding transition probabilities, forms a Markov chain in it's own right.*

Corollary 4 *If the chain starts in the set T of transient states, then it either moves between states within T forever, visiting any one state only a finite number of times, or it eventually enters one of the closed sets C_r of recurrent states, and it remains in C_r forever after.*



11.1.7 Block matrices

$$M = \left(\begin{array}{ccc|ccc|ccc} 1 & 2 & 3 & 4 & 5 & 6 & 7 \\ 8 & 9 & 10 & 11 & 12 & 13 & 14 \\ \hline 15 & 16 & 17 & 18 & 19 & 20 & 21 \\ 22 & 23 & 24 & 25 & 26 & 27 & 28 \\ 29 & 30 & 31 & 32 & 33 & 34 & 35 \end{array} \right) = \begin{pmatrix} A & B & C \\ D & E & F \end{pmatrix}.$$

Each of the blocks is defined in the obvious way:

$$A = \begin{pmatrix} 1 & 2 & 3 \\ 8 & 9 & 10 \end{pmatrix}, \quad B = \begin{pmatrix} 4 & 5 \\ 11 & 12 \end{pmatrix} \quad \text{etc.}$$

If we were to then have a matrix N such that

$$N = \begin{pmatrix} G & H \\ J & K \\ L & S \end{pmatrix}$$

that is the correct size so MN exists, then

$$MN = \begin{pmatrix} A & B & C \\ D & E & F \end{pmatrix} \begin{pmatrix} G & H \\ J & K \\ L & S \end{pmatrix} = \begin{pmatrix} AG + BJ + CL & AH + BK + CS \\ DG + EJ + FL & DH + EK + FS \end{pmatrix}. \quad (46)$$

11.1.8 Permutations of rows and columns

Consider a 3×3 -transition matrix, corresponding to states 1, 2 and 3:

$$\mathbf{p} = \begin{pmatrix} .1 & .7 & .2 \\ .3 & .5 & .2 \\ .2 & .4 & .4 \end{pmatrix}.$$

If we now decide to list the states in the order 3, 2, 1, then we need to swap the first and third row, **and** the first and third column, so that we get

$$\begin{pmatrix} .4 & .4 & .2 \\ .2 & .5 & .3 \\ .2 & .7 & .1 \end{pmatrix}.$$

11.1.9 Reducibility

Theorem 11 *If columns and rows of a square matrix A can be permuted so that it acquires the form*

$$A = \begin{pmatrix} B & 0 \\ C & D \end{pmatrix} \quad (47)$$

where B (and hence D) are square matrices, and 0 is a matrix consisting of all zeros, then A is called reducible.

Theorem 12 *Transition matrix \mathbf{p} of a decomposable (reducible) Markov chain is reducible.*

Suppose $\mathbb{S} = C \cup T$, with k_C states in the closed set, and k_T transient states. By definition of reducibility, the transition matrix \mathbf{p} can be rearranged into the form

$$\mathbf{p} = \begin{pmatrix} P & 0 \\ R & Q \end{pmatrix}. \quad (48)$$

Here, P corresponds to transitions among recurrent states of the closed set C , elements of Q are probabilities of transitions among transient states, and elements of R give the probabilities of transitions from transient to recurrent states. This form is retained after an arbitrary number of steps.

12 Week 18

12.1 Markov Chains

12.1.1 Perron-Frobenius theorem

Definition 41 *If all other eigenvalues λ satisfy the strict inequality $|\lambda| < r$, then A is called primitive.*

Theorem 13 *Suppose a matrix A is irreducible, and all elements of A are non-negative. Then the following statements are true:*

1. *A has at least one real positive eigenvalue.*
2. *The largest (if there is more than one) such real positive eigenvalue r dominates all other eigenvalues: If α is any other eigenvalue of A , then $|\alpha| \leq r$.*
3. *r is non-degenerate, i.e., it is a simple root of the characteristic equations*

$$\det(A_\lambda \mathbb{I}) = 0. \quad (49)$$

4. *Corresponding to r are the right and left eigenvectors all of whose components are positive. Conversely, the only eigenvectors whose components are positive are those corresponding to r .*
5. *r increases when any element of A increases. Correspondingly, r decreases when any element of A decreases.*
6. *If A is not primitive, i.e., there are $d > 1$ eigenvalues, such that $|\lambda| = r$, then they are all simple and different, and are located symmetrically on the unit circle in the complex plane: if they are numbered from 0 to $d - 1$, they can be written as*

$$\lambda_k = r e^{2\pi i k / d}, \quad (49)$$

with $k = 0, 1, \dots, d - 1$. If there are two such roots, they are equal to r and $-r$.

Further, if $d > 1$, the rows and columns of the matrix can be permuted to bring it to the block form

$$\begin{pmatrix} 0 & A_1 & 0 & \cdots & 0 \\ 0 & 0 & A_2 & \cdots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \cdots & A_{d-1} \\ A_d & 0 & 0 & \cdots & 0 \end{pmatrix}.$$

12.1.2 Limiting behaviour

Corollary 5 *If A is irreducible and primitive, and x and y are the right (column) and left (row) eigenvectors of A respectively, corresponding to the dominant eigenvalue r , and normalised so that $yx = 1$, then*

$$\lim_{n \rightarrow \infty} (1/r)^n A^n = x \otimes y, \quad (49)$$

where \otimes is used to denote the “outer” product of vectors: column times row.

Such products are known as dyadic products and the limits are uniform for all elements, so

$$\lim_{n \rightarrow \infty} (1/r)^n (A^n)_{ij} = x_i y_j.$$

Theorem 14 *For irreducible and primitive \mathbf{p} ,*

$$\lim_{n \rightarrow \infty} \mathbf{p}^n = e \otimes \pi. \quad (50)$$

It follows that

Corollary 6 *If \mathbf{p} is irreducible and primitive, there exists an equilibrium distribution.*

12.1.3 Reducible matrices

Given a reducible matrix A , its rows and columns can be permuted to the canonical form:

$$\begin{pmatrix} B & 0 \\ C & D \end{pmatrix}.$$

If B is reducible in its turn, its rows and columns can be permuted further to exhibit the structure:

$$\begin{pmatrix} B_{11} & 0 & 0 \\ B_{21} & B_{22} & 0 \\ C_1 & C_2 & D \end{pmatrix}, \quad (51)$$

where $(C_1 \ C_2) = C$.

For the application to Markov chains, only the following special case of the canonical form is relevant:

$$p = \begin{pmatrix} P_1 & 0 & \cdots & 0 & 0 \\ 0 & P_2 & \cdots & 0 & \vdots \\ \vdots & \vdots & \ddots & \vdots & \vdots \\ 0 & 0 & \cdots & P_n & 0 \\ & & & R & Q \end{pmatrix}. \quad (52)$$

Each block P_k contains probabilities of transitions inside k -th closed set, the Q block describes transitions among transient states, and the R block describes transitions from transient to recurrent states.

Theorem 15 *If the chain has transient states and/or more than one closed set of recurrent states, then p is reducible, and vice versa. The number of closed sets of recurrent states is equal to the multiplicity of the dominant eigenvalue $r = 1$.*

12.1.4 Classification of chains

Considering only chains with a finite overall number of states:

- **Reducible:**

- One closed set of recurrent states:
Single eigenvalue equal to 1.
Equilibrium distribution exists.
- Two or more closed sets of recurrent states:
Two or more eigenvalues equal to 1.
Equilibrium distribution does not exist.

- **Irreducible:**

- **Periodic:**

Single eigenvalue equal to 1. Two or more eigenvalues with absolute value equal to 1. Period equal to the number of such eigenvalues.

Equilibrium distribution does not exist.

- **Ergodic:**

Single eigenvalue equal to 1. No other eigenvalues with absolute value equal to 1.

Equilibrium distribution exists.

12.1.5 Algebraic test of reducibility

Definition 42 A matrix M is said to be greater than 0 ($M > 0$) if each element of M is greater than 0 ($M_{ij} > 0 \forall i, j$).

Theorem 16 An $n \times n$ -matrix p is reducible iff $(\mathbb{I}_n + p)^{n-1} > 0$.

12.1.6 Long-time limit

$$p^n = \begin{pmatrix} P^n & 0 \\ V_n & Q^n \end{pmatrix}, \quad (53)$$

where $V_n = RP^{n-1} + QRP^{n-2} + Q^2RP^{n-3} + \dots + Q^{n-2}RP + Q^{n-1}R$. This sum can be written in compact notation as:

$$V_n = \sum_{j=0}^{n-1} Q^j RP^{n-1-j}. \quad (54)$$

V_n satisfies a recursion:

$$V_n = QV_{n-1} + RP^{n-1}. \quad (55)$$

12.1.7 Evaluating the limits

- $\lim_{n \rightarrow \infty} P^n = \begin{pmatrix} \Pi & 0 \\ V & 0 \end{pmatrix}$, where $V = \lim_{n \rightarrow \infty} V_n$
- $\lim_{n \rightarrow \infty} Q^n = 0$.

Theorem 17

$$R\Pi + QV = V \quad (56)$$

$$V = (\mathbb{I} - Q)^{-1}R\Pi, \quad (57)$$

where \mathbb{I} is the unit matrix of the same dimension as Q .

12.1.8 Special case

Theorem 18 When the chain has a single closed set of recurrent states:

$$\lim_{n \rightarrow \infty} p^n = e_{r+t} \otimes (\pi, 0_t). \quad (58)$$

Here, t is the number of transient states and r is the number of columns in R .

12.1.9 Absorption problems in reducible chains

Theorem 19 A Markov chain with a single ergodic closed set possesses an equilibrium distribution.

13 Week 19

13.1 Markov chains

13.1.1 Absorption probabilities

Let the transition matrix p be of the form

$$p = \begin{pmatrix} P & 0 \\ R & Q \end{pmatrix},$$

where P has a block-diagonal structure with each square block containing elements encoding transitions among the states of the n -th closed set.

The structure of R is

$$R = \begin{pmatrix} R_{11} & R_{12} & \cdots & R_{1r} \\ R_{21} & R_{22} & \cdots & R_{2r} \\ \vdots & \vdots & \ddots & \vdots \\ R_{t1} & R_{t2} & \cdots & R_{tr} \end{pmatrix}.$$

Each $R_{\alpha n}$ is a row of k_n elements $R_{\alpha j}$ describing the transitions between the α -th transient state and the j -th state in the closed set C_n .

If we are interested only in the properties of the chain up to the time of absorption, the subsequent transitions among the states of a closed set are irrelevant. Thus, we can replace our original chain with an equivalent one which:

Each of the closed sets is replaced by a single compound absorbing state. (59)

Correspondingly, each row $R_{\alpha j}$ is replaced with the number

$$\tilde{R}_{\alpha n} = \sum_{j \in C_n} R_{\alpha j}. \quad (60)$$

The probabilities of absorption are given by the elements of the matrix

$$\tilde{V} = (\mathbb{I} - Q)^{-1} \tilde{R}. \quad (61)$$

13.1.2 First-step decomposition in absorption problems

Let $V_{\alpha j}$ be the eventual probability to be in the j -th recurrent state starting from the transient state α . Then:

Theorem 20

$$V_{\alpha j} = \sum_{r \in C_j} \Pi_{rj} R_{\alpha r} + \sum_{\beta} V_{\beta j} Q_{\alpha \beta}. \quad (62)$$

In matrix form, this is the same as eq.(57).

13.1.3 Recurrence time

Theorem 21 *The mean recurrence time for an ergodic state j is equal to the inverse of the equilibrium probability in the state j .*

Theorem 22 *The mean recurrence time μ for a state j is given by:*

$$\mu_j = \sum_{m=1}^{\infty} m f_{jj}^{(m)} = \frac{1}{\pi_j}. \quad (63)$$

14 Week 20

14.1 Markov chains

14.1.1 Mean time to absorption

Theorem 23

$$\mu = (\mathbb{I} - Q)^{-1} e^{(t)}, \quad (64)$$

where μ is the column of mean absorption times and $(\mathbb{I} - Q)^{-1} e^{(t)}$ is the mean number of visits.

15 Week 21

15.1 Branching Processes

15.1.1 Population dynamics

- **Deterministic:** In deterministic models the rules of the evolution of a population size from one generation to the next are fixed, and complexity arises, e.g., from non-linear interaction between different subspecies of the population (e.g., predators eat prey).
- **Stochastic:** In stochastic models one takes into account the fact that each member of the population may have a varying number of immediate descendants, so that this number is modeled as a random variable subject to a certain probability law.

A deterministic model often arises as the average behaviour of a stochastic model when the number of individuals is large.

15.1.2 Galton-Watson branching process

Theorem 24 *The total population size of the $n + 1$ -st generation is*

$$Z_{n+1} = \sum_{i=1}^{Z_n} X_i. \quad (65)$$

The X_i are taken to be i.i.d. random variables.

Define two types of generating functions. $\mathbb{G}_n(s) = \mathbb{E}[s^{Z_n}]$ for the population in the n -th generation. and $G(s) = \mathbb{E}[s^X]$ for family size.

Theorem 25

$$\mathbb{G}_{n+1}(s) = \mathbb{G}_n[G(s)] \quad (66)$$

and

$$\mathbb{G}_n(s) = G(\mathbb{G}_{n-1}(s)). \quad (67)$$

Theorem 26 *The probability of eventual extinction ξ is given by*

$$\xi = \lim_{n \rightarrow \infty} \mathbb{G}_n(0). \quad (68)$$

This also means

$$\xi = G(\xi). \quad (69)$$

Corollary 7 *The extinction probability ξ of the branching process is given by the smallest of the two non-negative roots of the equation*

$$\xi = G(\xi),$$

where $G(s)$ is the generating function for the distribution of the number of descendants of a single individual.

Corollary 8 *Extinction is certain iff mean “family” size is less than or equal to one.*

16 Week 23

16.1 Introduction to Martingales

16.1.1 Martingales

Definition 43 *A martingale is a stochastic process which possesses the property of being a “fair game”, and does not allow for infinite expectations.*

Formally: Z_n is a martingale with respect to itself if:

1. $\mathbb{E}[Z_{n+1} | \{Z_i\}_{i=1}^n] = Z_n$ and
2. $\mathbb{E}[|Z_n|] < \infty$.

A simple unbiased random walk without boundaries is a martingale.

16.1.2 Biased random walk

1. A more general definition of martingale with respect to another process (as opposed to with respect to itself) could be Y_n is a martingale with respect to S_n .
2. Since martingales have many very useful properties, constructing a martingale on the basis of the process one wants to study is a powerful mathematical tool.

16.1.3 Properties of martingales

Theorem 27 *If Z_n is a martingale, then*

$$\mathbb{E}[Z_n] = \mathbb{E}[Z_0]. \quad (70)$$

16.1.4 De Moivre's martingale

Definition 44 *De Moivre's martingale is defined as*

$$Y_n = (q/p)^{S_n}, \quad (71)$$

where $p \neq q$.

16.1.5 Stopping times

Definition 45 *A random variable T is a stopping time with respect to a stochastic process Y_n if $\{T = n\} \in \mathcal{F}_n$ where \mathcal{F}_n is the powerset associated with the sample space of all possible realisations of the sequence*

$$Y_0, Y_1, Y_2, \dots, Y_{n-1}, Y_n.$$

\mathcal{F}_n is known as a filtration. Martingales are properly defined with respect to filtrations, rather than the underlying stochastic processes:

$$\mathbb{E}[Y_{n+1} | \mathcal{F}_n] = Y_n \quad (72)$$

defines a martingale.

17 Week 24

17.1 Introduction to martingales

17.1.1 Optional stopping theorem

Theorem 28 *Let Y_n be a martingale with respect to a filtration \mathcal{F}_n and let T be a stopping time. Then, $\mathbb{E}[Y_T] = \mathbb{E}[Y_0]$ if the following conditions are satisfied:*

1. $\mathbb{P}[T < \infty] = 1$,
2. $\mathbb{E}[T] < \infty$,
3. \exists a constant c such that for all $n < T$ the following condition is satisfied:

$$\mathbb{E} \left[|Y_{n+1} - Y_n| \middle| \mathcal{F}_n \right] \leq c. \quad (73)$$

17.1.2 Gambling strategies

Consider a martingale $W_n = \sum_{k=1}^n Z_k$. Suppose now the gambler attempts to execute a strategy by varying the amount of individual bets depending on the outcome(s) preceding game(s). Let us denote these bets as α_n , so that the capital of the gambler after n games is

$$C_n = \sum_{k=1}^n \alpha_k Z_k. \quad (74)$$

Theorem 29 *If each bet α_k depends only on the outcomes of the preceding $k - 1$ games, then C_n is a martingale.*

17.1.3 Wald's identity

Assuming the validity of the optional stopping theorem, we have for a suitable stopping time T ,

$$\mathbb{E}[s^{S_T} G(S)^{-T}] = 1. \quad (75)$$