

# Regressions and the SMATR Package in R

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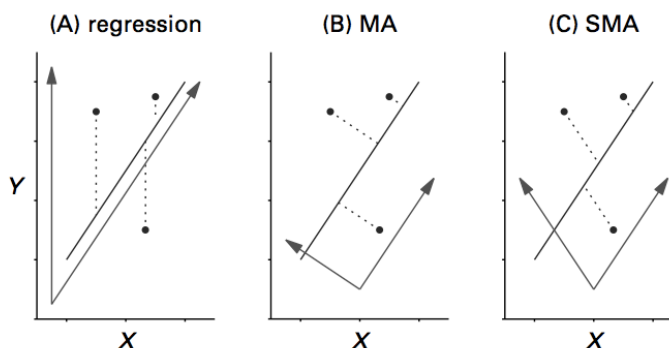
5-2-2012

## Roadmap...

- Types of 'regression'
- When to use, and not use, the different techniques (Warton et al. 2006)
- Example (Reich et al. 2010)
- SMATR tutorial

# Regression Types-Summary

*“Linear regression, MA and SMA are all least squares methods—the line is estimated by minimising the sum of squares of residuals from the line....The differences in methods of estimation of the lines are due to differences in the direction in which errors from the line are measured”*



**Fig. 4.** The direction in which residuals are measured is (A) vertical for linear regression (B) perpendicular to the line for major axis estimation (C) the fitted line reflected about the  $Y$  axis for standardised major axis estimation. Axes are plotted on the same scale. The broken lines indicate residuals, and the arrows represent the fitted and residual axes, which are useful for understanding methods of estimation and inference about these lines.

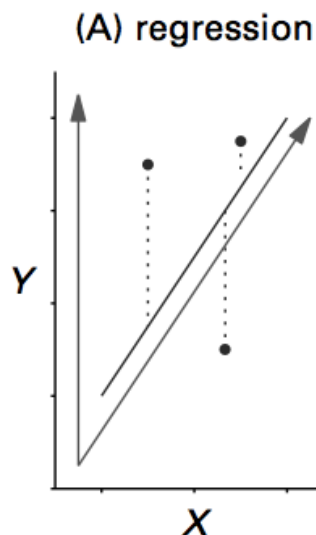
Warton et al. 2006

## Regression Types

Linear regression (Model I regression)

$$\sum_{i=1}^N (y_i - \hat{y}_i)^2$$

Minimize the residuals above, where  $\hat{y}_i$  is the fitted/predicted value of  $y_i$ .



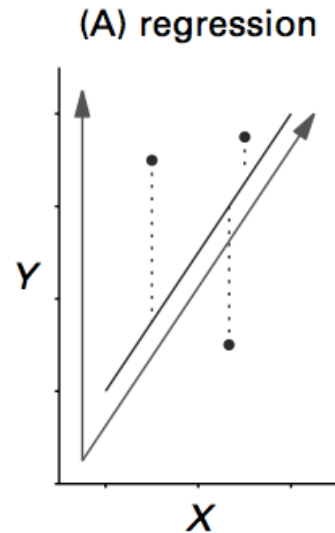
Warton et al. 2006

# Regression Types

## Linear regression (Model I regression)

*Regression is useful to...*

- Predict Y from X
- Determine if Y is associated with X (whether the slope is 0)
- Determine how strongly Y and X are associated (what is the  $r^2$ , the variation in Y that can be explained by linear regression on X)



Warton et al. 2006

# Regression Types

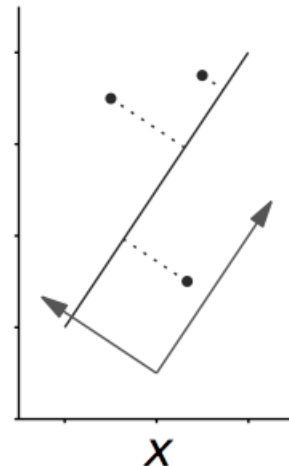
## Major Axis

(a.k.a. First principal component axis of the covariance matrix)

(a.k.a “model II regression”)

(B) MA

*“The major axis is the line that minimises the sum of squares of the shortest distances from the data points to the line. The shortest distance from a point to a line is perpendicular to it, so in this method residuals are measured perpendicular to the line.”*



Warton et al. 2006

# Regression Types

## Standardized Major Axis (MA)

(a.k.a Geometric Mean Functional Relationship)

(a.k.a Reduced Major Axis)

(Along with MA, “model II regression”)

*“The standardised major axis is the major axis calculated on standardised data, then rescaled to the original axes. This is typically done when two variables are not measured on comparable scales, in which case it might not seem reasonable to give the X and Y directions equal weight when measuring departures from the line.”*

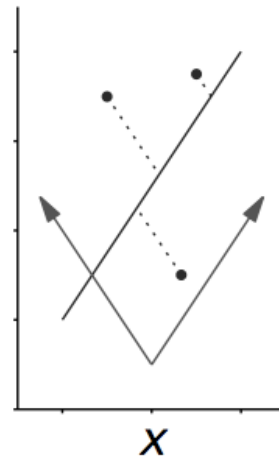
One way to think about it (from Smith 2009):

$$\text{Minimize } \Sigma(X - \bar{X})(Y - \bar{Y})$$

See Warton et al. 2006 Appendix for more derivations

Warton et al. 2006

(C) SMA



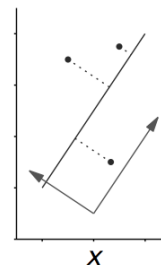
# Regression Types

## MA and SMA

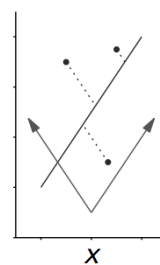
*MA and SMA are useful to...*

- Summarize the relationship between two variables (a single dimension describes the two dimensions of the data)
- Contexts:
  - Allometry
  - ‘law-like’ relationships
  - Testing agreement of two methods of measurement

(B) MA



(C) SMA



Warton et al. 2006

# Regression Types

## Summary of when to use what

Table 1. Which method of line-fitting should be used when

| Purpose  | Key statistic | Appropriate method           |
|--|---------------|------------------------------|
| Predict $Y$ from $X$ ( $X$ may even be random or may include measurement error)  | $\hat{y}$     | Linear regression            |
| Test for an association between $Y$ and $X$  | $P$           | Linear regression            |
| Estimate the line best describing the bivariate scatter of $Y$ and $X$   | $\hat{\beta}$ | MA or SMA                    |
| Test if the slope equals a specific value (1, or $\frac{3}{4}$ , etc.) for the line best describing the relationship between $Y$ and $X$ | $\hat{\beta}$ | MA or SMA                    |
| Estimate the strength of the linear relationship between $Y$ and $X$   | $r^2$         | Correlation                  |
| Predict $Y$ from some underlying $X$ that has been measured with error, so that only $(X + \delta)$ is observed                          | $\hat{y}$     | Method-of-moments regression |
| Estimate the line best describing the bivariate scatter of $Y$ and $X$ , when only $(X + \delta_X)$ and $(Y + \delta_Y)$ are observed    | $\hat{\beta}$ | Method-of-moments MA or SMA  |

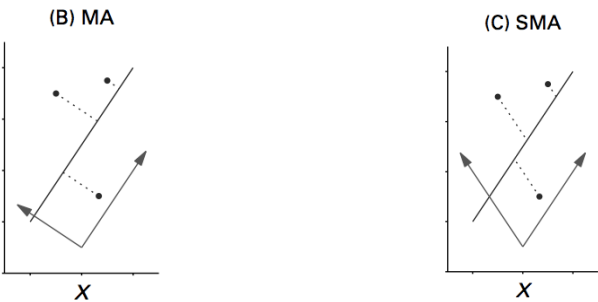
Abbreviations: MA, major axis; SMA, standardised major axis.

Warton et al. 2006

# Regression Types

## MA and SMA

*But when do you use MA versus SMA?*



Warton et al. 2006

## But when do you use MA versus SMA?

Table 2. Properties of the major axis (MA) and standardised major axis (SMA) that favour one or the other for line-fitting. The properties and recommendations listed here have a wide consensus or a strong logical basis. Some of the references given here relate to discussion of the equivalent question in the principal components literature – use of the covariance or correlation matrix for principal components analysis

| Property                      | Favours | Favoured in what situations   | Explanation   | References   |
|-------------------------------|---------|---|---|--|
| Efficiency                    | SMA     | All cases   | SMA lines are estimated with greater precision (standard error of the slope is smaller).  | Isobe <i>et al.</i> (1990); Jolicoeur (1990)                 |
| Scale dependence              | SMA     | When scale is arbitrary*  | The major axis is scale dependent – if all $Y$ values are doubled, the MA slope will not double.  | Harvey & Pagel (1991); Sokal & Rohlf (1995); Jolliffe (2002) |
| Inference in complex problems | MA      | When a method of inference for SMA is unavailable   | For some complex problems, procedures for analysis are currently available for MA but not for SMA.  | Anderson (1984); Jolliffe (2002)                             |
| Assumed error variances       | MA      | When there is no equation error and the measurement error variance is equal for $X$ and $Y$ | The major axis assumes the error variance is equal for $X$ and $Y$ , which is often a reasonable assumption when checking if two methods of measurement agree. Note that this argument does not hold if there is equation error (such as in allometry). | Sprent & Dolby (1980); Rayner (1985)                         |

\* Scale is arbitrary if the two variables are measured in qualitatively different units (e.g. kilograms and meters). Note that if both variables are log-transformed, units are no longer important and this consideration no longer applies, unless the power of  $X$  or  $Y$  is arbitrary (is there a reason for plotting  $Y$  versus  $X$  rather than  $Y^2$  versus  $X$ ?).

*Note: Also see table 3 of Warton et al. for more controversial relative merits of MA and SMA*

Warton et al. 2006

## Regression Types-Formulas

Table 4. Calculation formulae for estimation of bivariate lines for linear regression, the major axis and standardised major axis, and for inference about the slope ( $\beta$ ) or elevation ( $\alpha$ ) from one sample

|  | Linear regression   | Major axis  | Standardised major axis  |
|--|---|---|--|
| $\hat{\beta}$                          | $\frac{s_{xy}}{s_x^2}$  | $\frac{1}{2s_{xy}} \left( s_y^2 - s_x^2 + \sqrt{(s_y^2 - s_x^2)^2 + 4s_{xy}^2} \right)$   | $\text{Sign}(s_{xy}) \frac{s_y^2}{s_x}$  |
| $\hat{\alpha}$                         | $\bar{y} - \hat{\beta}\bar{x}$  | As for regression   | As for regression  |
| Residual axis (r)                      | $Y - \hat{\beta}X$  | As for regression   | As for regression  |
| Fitted axis (f)                        | $X$   | $\hat{\beta}Y + X$  | $Y + \hat{\beta}X$   |
| Test $H_0: \beta = b$                  | $(N-2) \frac{r_{\hat{\beta}}^2(b)}{1-r_{\hat{\beta}}^2(b)} \sim F_{1, N-2}$ | As for regression   | As for regression  |
| $s_{\hat{\beta}}^2$                    | $\frac{1}{N-2} \frac{s_y^2}{s_x^2} (1-r_{xy}^2)$                            | $\frac{(1+\hat{\beta}^2)}{N-2} \left( \frac{s_y^2}{s_x^2} + \frac{s_x^2}{s_y^2} - 2 \right)^{-1}$   | $\frac{1}{N-2} \frac{s_y^2}{s_x^2} (1-r_{xy}^2)$   |
| 100(1- $p$ )% CI for $\beta$ (primary) | $\hat{\beta} \pm s_{\hat{\beta}} t_{1-\frac{p}{2}, N-2}$                    | $\frac{1}{2(s_{xy} \pm \sqrt{Q})} \left( s_y^2 - s_x^2 + \sqrt{(s_y^2 - s_x^2)^2 + 4s_{xy}^2 - 4Q} \right)$<br>where $Q = \frac{1}{N-2} (s_x^2 s_y^2 - s_{xy}^2) f_{1-p, 1, N-2}$ | $\hat{\beta}(\sqrt{B+1} \pm \sqrt{B})$ , where<br>$B = \frac{1-r_{xy}^2}{N-2} f_{1-p, 1, N-2}$ |
| Secondary CI for $\beta$               | Not applicable  | $\frac{1}{2(s_{xy} \pm \sqrt{Q})} \left( s_y^2 - s_x^2 - \sqrt{(s_y^2 - s_x^2)^2 + 4s_{xy}^2 - 4Q} \right)$   | $-\hat{\beta}(\sqrt{B+1} \pm \sqrt{B})$  |
| $s_{\hat{\alpha}}^2$                   | $\frac{s_y^2}{N} + \bar{x}^2 s_{\hat{\beta}}^2$                             | As for regression   | As for regression  |
| Test $H_0: \alpha = a$                 | $\frac{\hat{\alpha} - a}{s_{\hat{\alpha}}} \sim t_{N-2}$                    | As for regression   | As for regression  |
| 100(1- $p$ )% CI for $\alpha$          | $\hat{\alpha} \pm s_{\hat{\alpha}} t_{1-\frac{p}{2}, N-2}$                  | As for regression   | As for regression  |

Notation: we wish to estimate the line  $Y = \alpha + \beta X$  from  $N$  pairs of observations of  $X$  and  $Y$ , as  $Y = \hat{\alpha} + \hat{\beta}X$ .  $\bar{x}$  and  $\bar{y}$  are the respective sample means of the observations of  $X$  and  $Y$ ,  $s_x^2$  is the sample estimate of the variance of  $X$ ,  $s_{xy}$  and  $r_{xy}$  are (respectively) the sample covariance and sample correlation coefficient of  $X$  and  $Y$ . The variables 'r' and 'f' represent residual and fitted axis scores, respectively, and  $r_{\hat{\beta}}(b)$  is the correlation between residual and axis scores, when these variables are calculated using a slope of  $b$  (not  $\hat{\beta}$ ). The terms  $t_{1-p, 1, N-2}$  and  $f_{1-p, 1, N-2}$  represent the 100 $p$ % critical values from the  $t_{N-2}$  and  $F_{1, N-2}$  distributions, respectively.  $H_0$  means 'null hypothesis'.

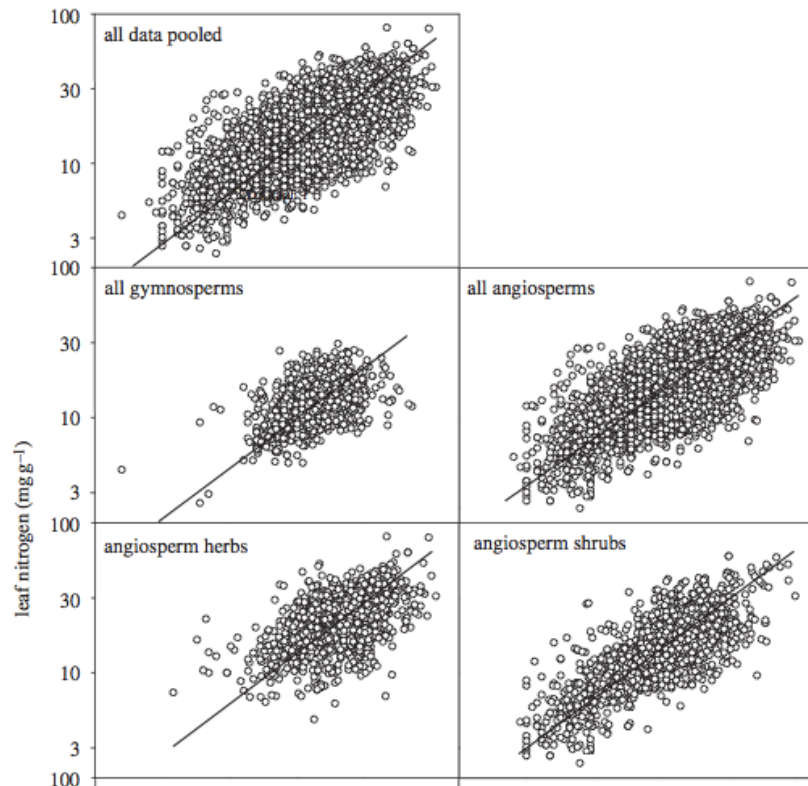
Warton et al. 2006

# Example: Reich et al. 2010 paper

Table 1. Scaling of leaf nitrogen concentration  $N_L$  in relation to leaf phosphorus concentration  $P_L$  for all data pooled, for plants grouped by phylogeny (angiosperm, gymnosperm), by four functional groups within the angiosperms and by three different biomes within the angiosperms. (All relations were significant ( $p < 0.0001$ ). All equations were fit using the log-log version of the equation:  $Y = \beta X^\alpha$ . Reduced major axis intercepts and slopes (exponents) are shown, as well as the lower and upper 95% CI of the exponent, and  $r^2$ .  $n$ , the number of observations, i.e. unique species-site combinations with data for  $N_L$  and  $P_L$  obtained from same individuals. Significant differences ( $p < 0.05$ ) in exponents among groups (for appropriate contrasts separated by blank lines) are shown by the lack of shared letters. Intercepts were not standardized to a common slope, and thus are not contrasted among groups. Biomes were broadly defined, such that temperate includes temperate and boreal; and moist tropical is both wet and moist tropical and subtropical.)

| plant group                         | $n$  | intercept | exponent | low CI | high CI | $r^2$ |
|-------------------------------------|------|-----------|----------|--------|---------|-------|
| all                                 | 9356 | 1.113     | 0.676    | 0.658  | 0.694   | 0.37  |
| <i>divisions</i>                    |      |           |          |        |         |       |
| angiosperm                          | 6466 | 1.166     | 0.637 a  | 0.621  | 0.653   | 0.48  |
| gymnosperm                          | 2890 | 1.002     | 0.696 a  | 0.650  | 0.746   | 0.22  |
| <i>angiosperm functional groups</i> |      |           |          |        |         |       |
| graminoid                           | 699  | 1.105     | 0.688 a  | 0.631  | 0.751   | 0.42  |
| forb                                | 1072 | 1.127     | 0.664 a  | 0.595  | 0.742   | 0.23  |
| shrub                               | 1518 | 1.155     | 0.652 a  | 0.624  | 0.682   | 0.56  |
| trees                               | 2878 | 1.195     | 0.633 a  | 0.610  | 0.658   | 0.48  |
| <i>biomes</i>                       |      |           |          |        |         |       |
| temperate                           | 3147 | 1.134     | 0.686 a  | 0.641  | 0.734   | 0.21  |
| Mediterranean                       | 714  | 1.143     | 0.655 a  | 0.623  | 0.689   | 0.68  |
| moist tropical                      | 1866 | 1.203     | 0.651 a  | 0.614  | 0.690   | 0.38  |

Reich et al. 2010



Reich et al. 2010



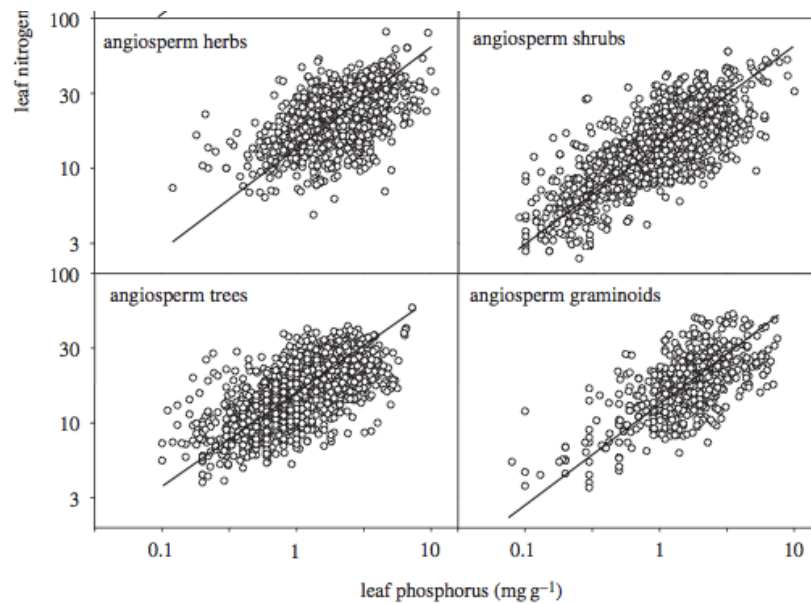


Figure 1. Relationships of leaf N ( $N_L$ ) to leaf P ( $P_L$ ) for all data pooled, and for plants grouped by both phylogeny (angiosperm, gymnosperm) and life form within the angiosperm group. The details of these relations using reduced major axis (RMA) regressions are presented in table 1.

Reich et al. 2010

## Example: Reich et al. 2010 paper

Table 2. Scaling relationships for angiosperms of leaf nitrogen concentration  $N_L$  and leaf phosphorus concentration  $P_L$  in relation to SLA for all data pooled, for plants grouped by life form and for three biomes. (All relations were significant ( $p < 0.0001$ ). All equations were fit using the log-log version of the equation:  $Y = \beta X^\alpha$ . Reduced major axis slopes (exponents) are shown, as well as the lower and upper 95% CI of the exponent, and  $r^2$ .  $n$ , number of observations. Significant differences ( $p < 0.05$ ) in exponents among groups (for appropriate contrasts separated by blank lines) are shown by the lack of shared letters.)

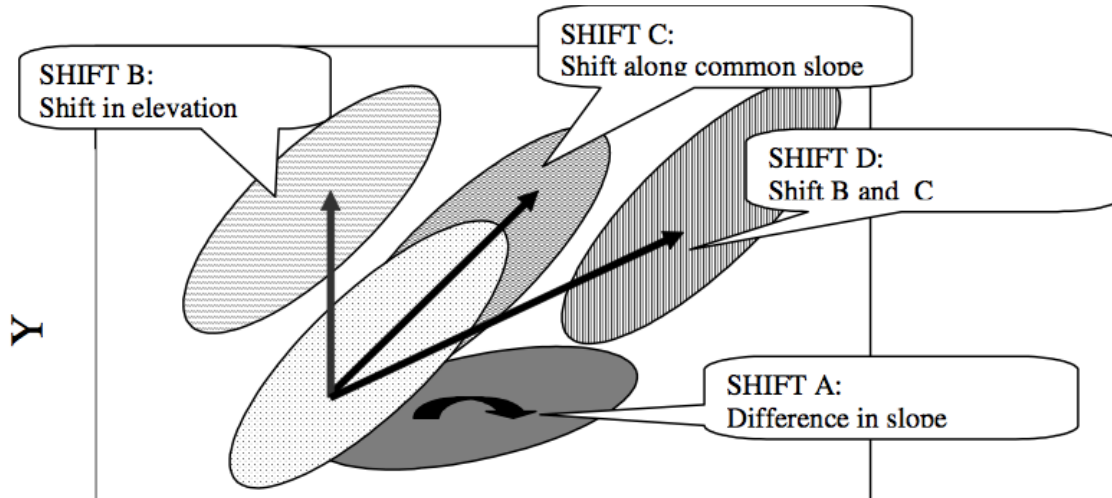
| plant group                         | $n$  | $N_L$    |                 |       | $P_L$    |                 |       |
|-------------------------------------|------|----------|-----------------|-------|----------|-----------------|-------|
|                                     |      | exponent | low CI, high CI | $r^2$ | exponent | low CI, high CI | $r^2$ |
| all                                 | 1819 | 0.76     | 0.73, 0.79      | 0.54  | 1.14     | 1.08, 1.12      | 0.45  |
| <i>angiosperm functional groups</i> |      |          |                 |       |          |                 |       |
| graminoid                           | 34   | 0.59 a   | 0.41, 0.85      | 0.52  | 1.17 a   | 0.81, 1.69      | 0.51  |
| forb                                | 22   | 0.72 a   | 0.33, 1.56      | 0.34  | 1.01     |                 | ns    |
| shrub                               | 535  | 0.83 a   | 0.76, 0.89      | 0.55  | 1.19 a   | 1.08, 1.31      | 0.45  |
| trees                               | 1112 | 0.75 a   | 0.70, 0.80      | 0.45  | 1.17 a   | 1.09, 1.26      | 0.39  |
| <i>biomes</i>                       |      |          |                 |       |          |                 |       |
| temperate                           | 449  | 0.67 a   | 0.59, 0.76      | 0.35  | 0.95 a   | 0.80, 1.12      | 0.24  |
| Mediterranean                       | 321  | 0.93 b   | 0.83, 1.03      | 0.54  | 1.41 b   | 1.26, 1.58      | 0.49  |
| moist tropical                      | 950  | 0.78 a   | 0.71, 0.85      | 0.34  | 1.21 ab  | 1.08, 1.35      | 0.24  |

Reich et al. 2010



# Possible tests with SMATR

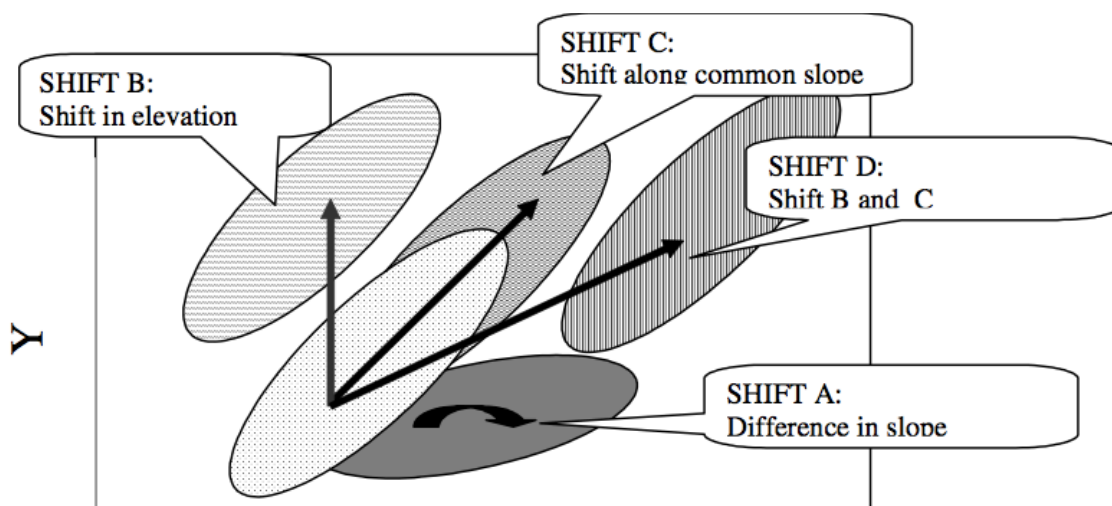
1. Test for differences among two or more fitted slopes (shift A)



From SMATR User's guide:[http://bio.mq.edu.au/ecology/SMATR/SMATR\\_users\\_guide.pdf](http://bio.mq.edu.au/ecology/SMATR/SMATR_users_guide.pdf)

# Possible tests with SMATR

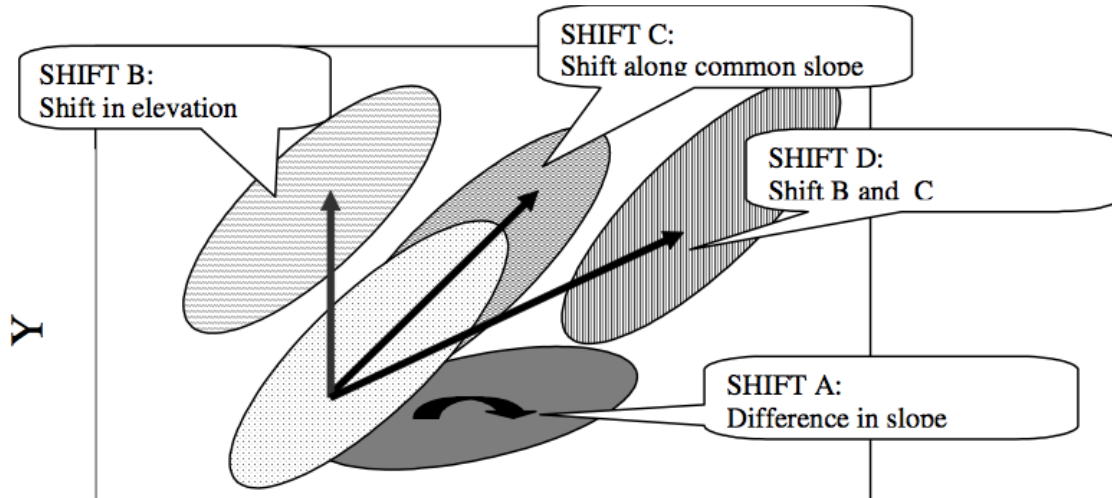
1. Test for differences among two or more fitted slopes (shift A)
2. If fitting a common slope can be justified (i.e., Test 1 is n.s.), test whether the fitted slopes share a common elevation (shift B)



From SMATR User's guide:[http://bio.mq.edu.au/ecology/SMATR/SMATR\\_users\\_guide.pdf](http://bio.mq.edu.au/ecology/SMATR/SMATR_users_guide.pdf)

# Possible tests with SMATR

1. Test for differences among two or more fitted slopes (shift A)
2. If fitting a common slope can be justified (i.e., Test 1 is n.s.), test whether the fitted slopes share a common elevation (shift B)
3. If fitting a common slope and elevation is justified, test whether the slopes fitted to each group are significantly separated along the common slope (shift C).



From SMATR User's guide:[http://bio.mq.edu.au/ecology/SMATR/SMATR\\_users\\_guide.pdf](http://bio.mq.edu.au/ecology/SMATR/SMATR_users_guide.pdf)

## Flow Chart for Comparing Bivariate Relationships

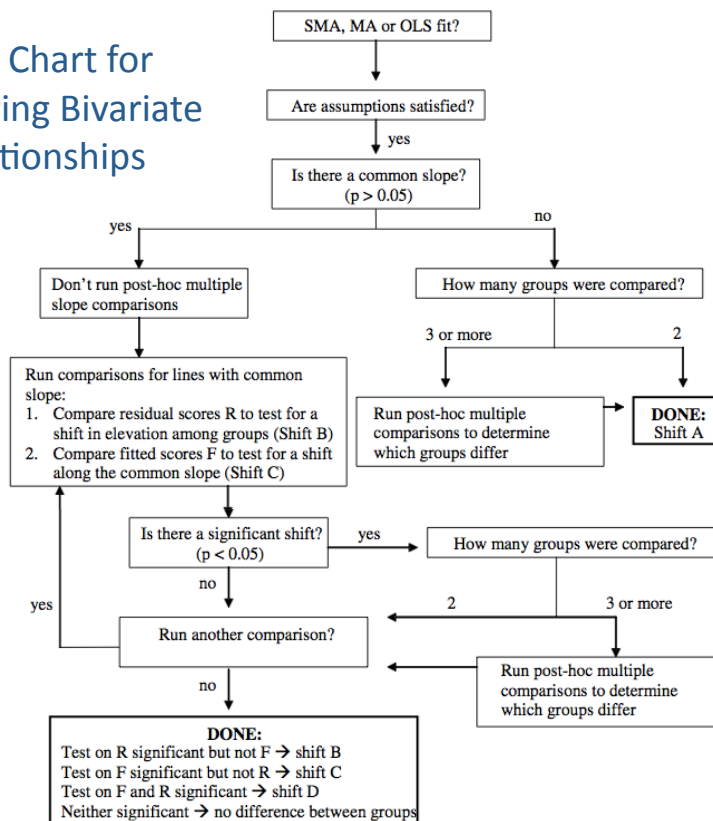


Figure from SMATR User's guide:[http://bio.mq.edu.au/ecology/SMATR/SMATR\\_users\\_guide.pdf](http://bio.mq.edu.au/ecology/SMATR/SMATR_users_guide.pdf)

# Relevant Resources

Papers that use the R SMATR package:

<http://bio.mq.edu.au/ecology/SMATR/pubs.html>

Another R package for model II regression:

<http://cran.r-project.org/web/packages/lmodel2/vignettes/mod2user.pdf>

SMATR documentation:

<http://cran.r-project.org/web/packages/smatr/smatr.pdf>

Plasticity and Allometry protocol:

<http://prometheuswiki.publish.csiro.au/tiki-index.php?page=Plasticity+and+Allometry+protocol>

# Works Cited

Reich et al. 2010. Evidence of a general 2/3-power law of scaling leaf nitrogen to phosphorus among major plant groups and biomes. *Proc. R. Soc. B.* 277: 877-883

Smith 2009. Use and misuse of the Reduced Major Axis for line-fitting. *Am. J. Phys. Anth.* 140: 476-486.

Warton et al. 2006. Bivariate line-fitting methods for allometry. *Biol. Rev.* 81: 259-291.

Warton website documents online: <http://bio.mq.edu.au/ecology/SMATR/>