

第X章习题一



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$$1. \bar{x} = np$$

$$\Rightarrow \hat{p} = \frac{\bar{x}}{n}$$

$$2. \bar{x} = \lambda$$

$$\Rightarrow \hat{\lambda} = \bar{x}$$

$$3. \begin{cases} \bar{x} = \frac{a+b}{2} \\ \frac{1}{n} \sum_{i=1}^n x_i^2 = \left(\frac{a+b}{2}\right)^2 + \frac{(b-a)^2}{12} \end{cases}$$

$$\Rightarrow \begin{cases} a+b = 2\bar{x} \\ b-a = \sqrt{12\left(\frac{1}{n} \sum_{i=1}^n x_i^2 - \bar{x}^2\right)} \end{cases}$$

$$\Rightarrow \begin{cases} \hat{a} = \bar{x} - \sqrt{3\left(\frac{1}{n} \sum_{i=1}^n x_i^2 - \bar{x}^2\right)} \\ \hat{b} = \bar{x} + \sqrt{3\left(\frac{1}{n} \sum_{i=1}^n x_i^2 - \bar{x}^2\right)} \end{cases}$$

$$4. \bar{x} = \theta$$

$$\Rightarrow \hat{\theta} = \bar{x}$$

$$5. \begin{cases} \bar{x} = \mu \\ \frac{1}{n} \sum_{i=1}^n x_i^2 = \mu^2 + \sigma^2 \end{cases}$$

$$\Rightarrow \begin{cases} \hat{\mu} = \bar{x} \\ \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n x_i^2 - \bar{x}^2 \end{cases}$$

$$6. \bar{x} = \frac{1}{p}$$

$$\Rightarrow \hat{p} = \frac{1}{\bar{x}}$$

第七章 习题 2.



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离散: $L(\theta) = \prod_{i=1}^n P(X_i = x_i)$ 连续: $L(\theta) = \prod_{i=1}^n f(x_i, \theta)$

1. ① $L(p) = \prod_{i=1}^m P(X_i = x_i) = \prod_{i=1}^m \left[\binom{n}{x_i} p^{x_i} (1-p)^{n-x_i} \right]$

② $\ln L(p) = \ln \prod_{i=1}^m \left[\binom{n}{x_i} p^{x_i} (1-p)^{n-x_i} \right] = \sum_{i=1}^m x_i \ln p + \sum_{i=1}^m (n-x_i) \ln(1-p)$

$\frac{d \ln L(p)}{dp} = \frac{\sum_{i=1}^m x_i}{p} - \frac{\sum_{i=1}^m (n-x_i)}{1-p} = 0$

$\sum_{i=1}^m x_i \cdot (1-p) - \sum_{i=1}^m (n-x_i) p = 0$

$\sum_{i=1}^m x_i - \sum_{i=1}^m x_i \cdot p - mn p + \sum_{i=1}^m x_i \cdot p = 0$

$\hat{p} = \frac{\sum_{i=1}^m x_i}{mn} = \frac{\bar{x}}{n}$

2. $L(\lambda) = \prod_{i=1}^n \frac{\lambda^{x_i}}{x_i!} e^{-\lambda} = \frac{\lambda^{\sum_{i=1}^n x_i}}{\prod_{i=1}^n (x_i!)} e^{-n\lambda}$

$\ln L(\lambda) = \sum_{i=1}^n x_i \cdot \ln \lambda - \ln \prod_{i=1}^n (x_i!) + (-n\lambda)$

$\frac{d \ln L(\lambda)}{d\lambda} = \frac{\sum_{i=1}^n x_i}{\lambda} - n = 0 \Rightarrow \lambda = \frac{\sum_{i=1}^n x_i}{n} = \bar{x}$

3. $L(a, b) = \prod_{i=1}^n f(x_i, a, b) = \begin{cases} \frac{1}{(b-a)^n} & a \leq \min\{x_i\} \leq \max\{x_i\} \leq b \\ 0 & \text{其它} \end{cases}$

$L(a, b)$ 关于 b 是单调减, 故 b 取 b 参数空间最小值, $\hat{b} = \max\{x_i\}$
 $L(a, b)$ 关于 a 是单调增, 故 a 取 a 参数空间最大值, $\hat{a} = \min\{x_i\}$



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$$4. L(\theta) = \prod_{i=1}^n f(x_i, \theta) = \begin{cases} \frac{1}{\theta^n} e^{-\frac{1}{\theta} \sum_{i=1}^n x_i}, & \min\{x_i\} > 0 \\ 0, & \text{其它} \end{cases}$$

$$\ln L(\theta) = \begin{cases} -\frac{1}{\theta} \sum_{i=1}^n x_i - n \ln \theta, & \min\{x_i\} > 0 \\ 0, & \text{其它} \end{cases}$$

$$\frac{d \ln L(\theta)}{d\theta} = \begin{cases} \frac{1}{\theta^2} \sum_{i=1}^n x_i - \frac{n}{\theta}, & \min\{x_i\} > 0 \\ 0, & \text{其它} \end{cases}$$

$$\hat{\theta} = \bar{x}, \quad (\min\{x_i\} > 0)$$

$$5. L(\mu, \sigma^2) = \prod_{i=1}^n f(x_i, \mu, \sigma^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x_i - \mu)^2}{2\sigma^2}} \\ = \left(\frac{1}{\sqrt{2\pi}}\right)^n (\sigma^2)^{-\frac{n}{2}} e^{-\frac{\sum_{i=1}^n (x_i - \mu)^2}{2\sigma^2}}$$

$$\ln L(\mu, \sigma^2) = \ln \left(\frac{1}{\sqrt{2\pi}}\right)^n - \frac{n}{2} \ln \sigma^2 - \frac{\sum_{i=1}^n (x_i - \mu)^2}{2\sigma^2}$$

$$\frac{\partial \ln L(\mu, \sigma^2)}{\partial \mu} = \frac{n}{\sum_{i=1}^n \frac{1}{\sigma^2}} \sum_{i=1}^n (x_i - \mu) = 0$$

$$\frac{\partial \ln L(\mu, \sigma^2)}{\partial \sigma^2} = -\frac{n}{2} \frac{1}{\sigma^2} + \frac{\sum_{i=1}^n (x_i - \mu)^2}{\sigma^4} = 0$$

$$\Rightarrow \begin{cases} \hat{\mu} = \bar{x} \\ \hat{\sigma}^2 = s_n^2 \end{cases}$$



$$6. L(p) = \prod_{i=1}^n P(X_i = x_i) = \prod_{i=1}^n (1-p)^{x_i-1} p = (1-p)^{\sum_{i=1}^n x_i - n} p^n$$

$$\ln L(p) = (\sum_{i=1}^n x_i - n) \ln(1-p) + n \ln p$$

$$\frac{d \ln L(p)}{dp} = - \frac{\sum_{i=1}^n x_i - n}{1-p} + \frac{n}{p} = 0$$

$$n(1-p) = p \cdot \sum_{i=1}^n x_i - np$$

$$\hat{p} = \frac{\sum_{i=1}^n x_i}{n} = \bar{x}$$

二、 $\frac{1}{n} \sum_{i=1}^n x_i = \frac{\theta}{2} \Rightarrow \hat{\theta} = 2\bar{x}$, $E(\hat{\theta}) \neq \theta$, 不是无偏估计

矩估计: $f(x) = \begin{cases} \frac{1}{\theta} & 0 \leq x \leq \theta \\ 0 & \text{其它} \end{cases}$

最大似然估计: $L(\theta) = \prod_{i=1}^n f(x_i, \theta) = \begin{cases} \frac{1}{\theta^n} & 0 \leq \min\{x_i\} \leq \max\{x_i\} \leq \theta \\ 0 & \text{其它} \end{cases}$

由上式知, $L(\theta)$ 关于 θ 单调递减, 故 $\hat{\theta} = \max\{x_i\}$, 令 $Y = \max\{x_i\}$

$E(\hat{\theta}) = E(\max\{x_i\}) = EY$, $F_Y(y) = P(X_1 \leq y, X_2 \leq y, \dots, X_n \leq y)$

$= P(X_1 \leq y) \cdot P(X_2 \leq y) \cdots P(X_n \leq y) = \prod_{i=1}^n F_{X_i}(y) = \begin{cases} 0 & y < 0 \\ \frac{y^n}{\theta^n} & 0 \leq y \leq \theta \\ 1 & y > \theta \end{cases}$

$f_Y(y) = \begin{cases} n\theta^{-n} y^{n-1} & 0 \leq y \leq \theta \\ 0 & \text{其它} \end{cases}$

$E(\hat{\theta}) = E(Y) = \int_{-\infty}^{+\infty} y f_Y(y) dy = \int_0^{\theta} y \cdot n \cdot \theta^{-n} \cdot y^{n-1} dy = n \cdot \theta^{-n} \int_0^{\theta} y^n dy$

$= \frac{n}{n+1} \theta^{-n} y^{n+1} \Big|_0^{\theta} = \frac{n}{n+1} \theta^{-n} \cdot \theta^{n+1} = \frac{n}{n+1} \theta \neq \theta$

故不是无偏估计



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$$\text{三. (1)} \quad \frac{1}{n} \sum_{i=1}^n X_i = EX \Rightarrow \bar{x} = \int_{-\infty}^{+\infty} x f(x) dx = \int_0^1 (\theta+1)x^{\theta+1} dx$$

$$= \frac{\theta+1}{\theta+2} x^{\theta+2} \Big|_0^1 = \frac{\theta+1}{\theta+2} \Rightarrow \hat{\theta} = \frac{2\bar{x}-1}{1-\bar{x}}$$

$$(2) \quad L(\theta) = \prod_{i=1}^n f(x_i, \theta) = \begin{cases} (\theta+1)^n \left(\prod_{i=1}^n x_i \right)^{\theta}, & 0 < \min\{x_i\} \leq \max\{x_i\} < 1 \\ 0, & \text{其它} \end{cases}$$

当 $0 < \min\{x_i\} \leq \max\{x_i\} < 1$ 时:

$$\ln L(\theta) = n \ln(\theta+1) + \theta \ln \prod_{i=1}^n x_i$$

$$\frac{d \ln L(\theta)}{d\theta} = \frac{n}{\theta+1} + \frac{n}{\sum_{i=1}^n} (\ln x_i) = 0$$

$$\hat{\theta} = -1 - \frac{n}{\sum_{i=1}^n \ln x_i}$$

$$\text{五、} \quad EX = 2\theta(1-\theta) + 2\theta^2 + 3(1-2\theta) = 2\theta - 2\theta^2 + 2\theta^2 + 3 - 6\theta + 3 - 4\theta$$

$$\bar{x} = 2$$

$$\text{矩估计: } \bar{x} = EX \Rightarrow 3 - 4\theta = 2 \Rightarrow \hat{\theta} = \frac{1}{4}$$

$$\text{最大似然估计: } L(\theta) = \prod_{i=1}^n P(X_i = x_i)$$

$$= \theta^2 \cdot [2\theta(1-\theta)]^2 \cdot \theta^2 \cdot (1-2\theta)^4 = 4\theta^6 (1-\theta)^2 (1-2\theta)^4$$

$$\ln L(\theta) = 6 \ln 4\theta + 2 \ln(1-\theta) + 4 \ln(1-2\theta)$$

$$\frac{d \ln L(\theta)}{d\theta} = 6 \frac{1}{4\theta} \cdot 4 - 2 \frac{1}{1-\theta} + 4 \frac{1}{1-2\theta} (-2)$$

$$= \frac{24\theta^2 - 28\theta + 6}{\theta(\theta-1)(2\theta-1)} = 0 \Rightarrow \hat{\theta} = \frac{7 \pm \sqrt{13}}{12}$$

$$\text{又} \because \frac{7+\sqrt{13}}{12} > \frac{1}{2} \quad \therefore \hat{\theta} = \frac{7-\sqrt{13}}{12}$$

地址: 青岛市松岭路238号

邮编: 266100

电话: 0532-66782730

传真: 0532-66782799

网址: <http://www.ouc.edu.cn>



一. 1. 最有效

2. $an + bm = 1$ $\frac{4}{4n+m}$ $\frac{1}{4n+m}$

3. $\hat{\mu}_3$

二. $E\left(C \sum_{i=1}^{n-1} (X_{i+1} - X_i)^2\right) = 6^2$

$X_{i+1} - X_i \sim N(0, 26^2)$

$\Rightarrow E(X_{i+1} - X_i)^2 = (E(X_{i+1} - X_i))^2 + D(X_{i+1} - X_i) = 26^2$

$E\left(C \sum_{i=1}^{n-1} (X_{i+1} - X_i)^2\right) = C \sum_{i=1}^{n-1} E(X_{i+1} - X_i)^2 = C(n-1)26^2 = 6^2$

$\Rightarrow C = \frac{1}{2(n-1)}$

三. $L(\theta) = \prod_{i=1}^n f(X_i, \theta) = \begin{cases} 2^n e^{-2 \sum_{i=1}^n (X_i - \theta)}, & \min\{X_i\} > \theta \\ 0, & \text{其它} \end{cases}$

当 $\min\{X_i\} > \theta$ 时: $\ln L(\theta) = n \ln 2 - 2 \sum_{i=1}^n (X_i - \theta)$
 $= n \ln 2 - 2 \sum_{i=1}^n X_i + 2n\theta$

$\frac{d \ln L(\theta)}{d\theta} = 2n > 0$, θ 最大而增大, $\hat{\theta} = \min\{X_i\}$

令 $Y = \min\{X_i\}$, $F_Y(y) = 1 - (1 - F_X(y))^n \Rightarrow f_Y(y) = -n(1 - F_X(y))^{n-1} [-f_X(y)]$

$F_X(x) = \int_{-\infty}^x f(t) dt = \begin{cases} 0, & x \leq \theta \\ \int_{\theta}^x 2e^{-2(t-\theta)} dt = 1 - e^{-2(x-\theta)}, & x > \theta \end{cases}$

$f_Y(y) = \begin{cases} 0, & y \leq \theta \\ 2ne^{-2n(y-\theta)}, & y > \theta \end{cases}$

$EY = \int_{-\infty}^{+\infty} y f_Y(y) dy = \int_{\theta}^{+\infty} 2ny e^{-2n(y-\theta)} dy$
 $= \int_0^{+\infty} 2ny e^{-2ny+2n\theta} dy$



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$$\begin{aligned}
 &= e^{2n\theta} \int_0^{+\infty} 2ny e^{-2ny} dy = -e^{2n\theta} \int_0^{+\infty} y d e^{-2ny} = e^{2n\theta} [y e^{-2ny}]_0^{+\infty} - \int_0^{+\infty} e^{-2ny} dy \\
 &= -e^{2n\theta} [0 - \theta e^{-2n\theta} + \frac{1}{2n} \int_0^{+\infty} e^{-2ny} d(-2ny)] \\
 &= -e^{2n\theta} [-\theta e^{-2n\theta} + 0 - \frac{1}{2n} e^{-2n\theta}] = \theta + \frac{1}{2n} \neq \theta
 \end{aligned}$$

不是无偏估计

四. 令 $Y = \max\{X, 0\}$, $F_Y(y) = \{F_X(y)\}^3$

$$F_X(x) = \begin{cases} 0 & x < 0 \\ \frac{x}{\theta} & 0 \leq x \leq \theta \\ 1 & x > \theta \end{cases} \quad F_Y(y) = \begin{cases} 0 & y < 0 \\ (\frac{y}{\theta})^3 & 0 \leq y \leq \theta \\ 1 & y > \theta \end{cases} \quad f_Y(y) = \begin{cases} 3\theta^{-3}y^2 & 0 \leq y \leq \theta \\ 0 & \text{其它} \end{cases}$$

$$EY = \int_{-\infty}^{+\infty} y f_Y(y) dy = \int_0^{\theta} 3\theta^{-3} y^3 dy = \frac{3}{4}\theta$$

$$E(\hat{\theta}) = \frac{4}{3} EY = \theta \quad \text{故 } \hat{\theta} \text{ 是 } \theta \text{ 无偏估计}$$

五. $E\hat{\mu} = (a_1 + \dots + a_n)\mu = \mu \Rightarrow a_1 + \dots + a_n = 1$

$$D\hat{\mu} = (a_1^2 + \dots + a_n^2)\sigma^2 \geq \frac{(a_1 + \dots + a_n)^2}{n} \sigma^2 = \frac{1}{n} \sigma^2$$

上式取等条件为 $a_1 = a_2 = \dots = a_n$

$$\because a_1 + \dots + a_n = 1$$

$\therefore a_i = \frac{1}{n}$ 时, $D\hat{\mu}$ 最小, $\hat{\mu}$ 最有效