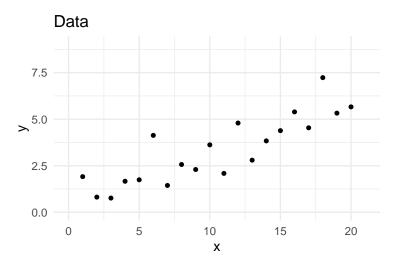
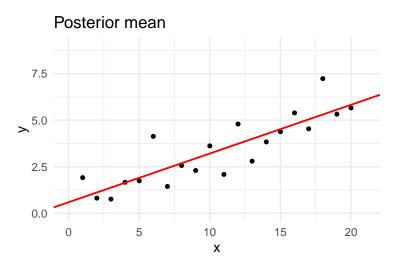
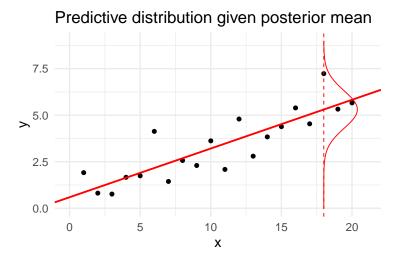
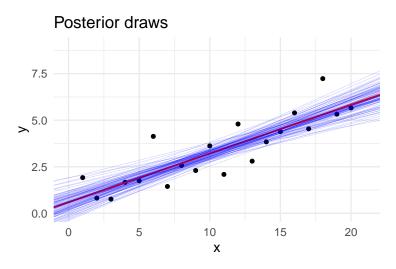
Chapter 3

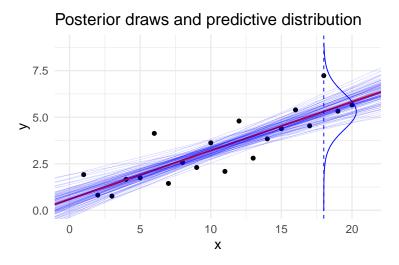
- 3.1 Marginalisation
- 3.2 Normal distribution with a noninformative prior (important)
- 3.3 Normal distribution with a conjugate prior (important)
- 3.4 Multinomial model (can be skipped)
- 3.5 Multivariate normal with known variance (useful for chapter 4)
- 3.6 Multivariate normal with unknown variance (glance through)
- 3.7 Bioassay example (very important, related to one of the exercises)
- 3.8 Summary (summary)

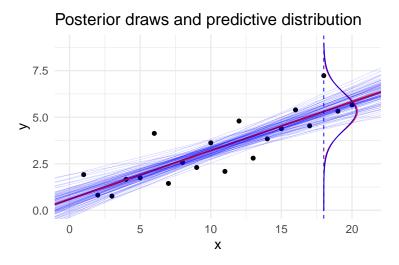












Marginalization

Joint distribution of parameters

$$p(\theta_1, \theta_2 \mid y) \propto p(y \mid \theta_1, \theta_2)p(\theta_1, \theta_2)$$

Marginalization

$$p(\theta_1 \mid y) = \int p(\theta_1, \theta_2 \mid y) d\theta_2$$

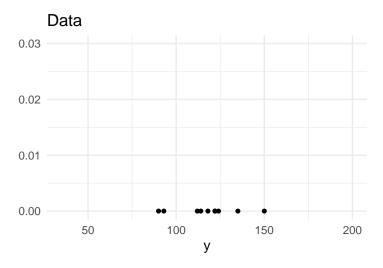
 $p(\theta_1 \mid y)$ is a marginal distribution

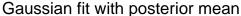
Marginalization - predictive distribution

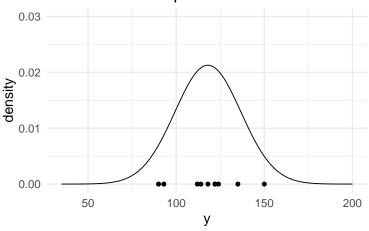
Marginalization over posterior distribution

$$p(\tilde{y} \mid y) = \int p(\tilde{y} \mid \theta) p(\theta \mid y) d\theta$$
$$= \int p(\tilde{y}, \theta \mid y) d\theta$$

 $p(\tilde{y} \mid y)$ is a predictive distribution

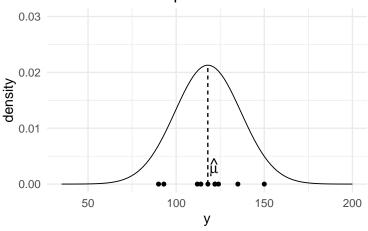




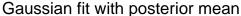


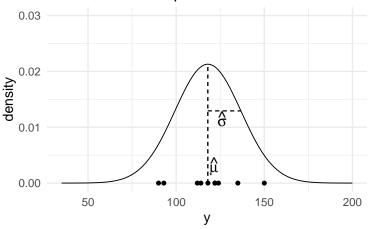
$$p(\mathbf{y} \mid \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2}(\mathbf{y} - \mu)^2\right)$$

Gaussian fit with posterior mean

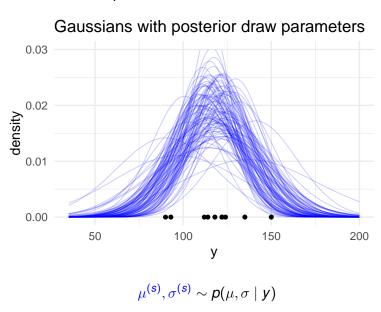


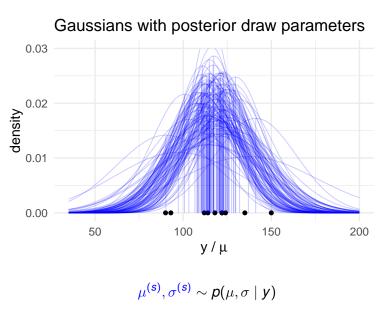
$$p(y \mid \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2}(y - \mu)^2\right)$$

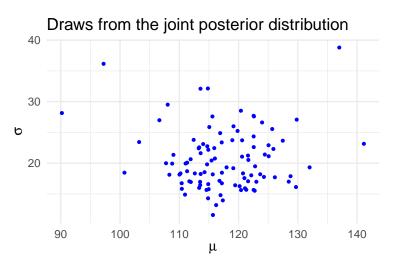




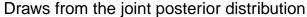
$$p(y \mid \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2}(y - \mu)^2\right)$$

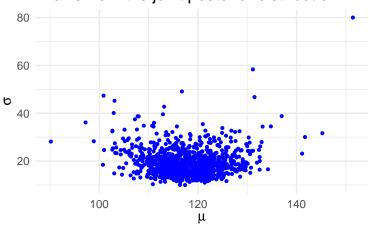






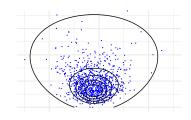
$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$



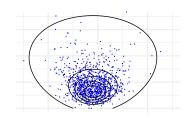


$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$

$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$

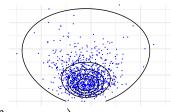


$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$
 with $p(\mu, \sigma^2) \propto \sigma^{-2}$



$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$
 with $p(\mu, \sigma^2) \propto \sigma^{-2}$

$$p(\mu, \sigma^2 \mid y) \propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2\right)$$



$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$
 with $p(\mu, \sigma^2) \propto \sigma^{-2}$

$$\sum_{i=1}^{n} (y_i - \mu)^2$$

$$p(\mu, \sigma^2 \mid y) \propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2\right)$$
$$= \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \left[\sum_{i=1}^n (y_i - \bar{y})^2 + n(\bar{y} - \mu)^2\right]\right)$$

where
$$\bar{y} = \frac{1}{n} \sum_{i=1}^{n} y_i$$

$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$
 with $p(\mu, \sigma^2) \propto \sigma^{-2}$

$$\begin{split} p(\mu, \sigma^2 \mid y) &\propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2\right) \\ &= \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \left[\sum_{i=1}^n (y_i - \bar{y})^2 + n(\bar{y} - \mu)^2\right]\right) \\ \text{where } \bar{y} &= \frac{1}{n} \sum_{i=1}^n y_i \\ &= \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \left[(n-1)s^2 + n(\bar{y} - \mu)^2\right]\right) \\ \text{where } s^2 &= \frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2 \end{split}$$

$$\sum_{i=1}^n (y_i - \mu)^2$$

$$\sum_{i=1}^{n} (y_i - \mu)^2$$

$$\sum_{i=1}^{n} (y_i^2 - 2y_i \mu + \mu^2)$$

$$\begin{split} &\sum_{i=1}^{n} (y_i - \mu)^2 \\ &\sum_{i=1}^{n} (y_i^2 - 2y_i \mu + \mu^2) \\ &\sum_{i=1}^{n} (y_i^2 - 2y_i \mu + \mu^2 - \bar{y}^2 + \bar{y}^2 - 2y_i \bar{y} + 2y_i \bar{y}) \end{split}$$

$$\begin{split} &\sum_{i=1}^{n} (y_i - \mu)^2 \\ &\sum_{i=1}^{n} (y_i^2 - 2y_i \mu + \mu^2) \\ &\sum_{i=1}^{n} (y_i^2 - 2y_i \mu + \mu^2 - \bar{y}^2 + \bar{y}^2 - 2y_i \bar{y} + 2y_i \bar{y}) \\ &\sum_{i=1}^{n} (y_i^2 - 2y_i \bar{y} + \bar{y}^2) + \sum_{i=1}^{n} (\mu^2 - 2y_i \mu - \bar{y}^2 + 2y_i \bar{y}) \end{split}$$

$$\begin{split} &\sum_{i=1}^{n} (y_{i} - \mu)^{2} \\ &\sum_{i=1}^{n} (y_{i}^{2} - 2y_{i}\mu + \mu^{2}) \\ &\sum_{i=1}^{n} (y_{i}^{2} - 2y_{i}\mu + \mu^{2} - \bar{y}^{2} + \bar{y}^{2} - 2y_{i}\bar{y} + 2y_{i}\bar{y}) \\ &\sum_{i=1}^{n} (y_{i}^{2} - 2y_{i}\bar{y} + \bar{y}^{2}) + \sum_{i=1}^{n} (\mu^{2} - 2y_{i}\mu - \bar{y}^{2} + 2y_{i}\bar{y}) \\ &\sum_{i=1}^{n} (y_{i} - \bar{y})^{2} + n(\mu^{2} - 2\bar{y}\mu - \bar{y}^{2} + 2\bar{y}\bar{y}) \end{split}$$

$$\sum_{i=1}^{n} (y_{i} - \mu)^{2}$$

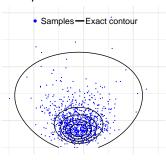
$$\sum_{i=1}^{n} (y_{i}^{2} - 2y_{i}\mu + \mu^{2})$$

$$\sum_{i=1}^{n} (y_{i}^{2} - 2y_{i}\mu + \mu^{2} - \bar{y}^{2} + \bar{y}^{2} - 2y_{i}\bar{y} + 2y_{i}\bar{y})$$

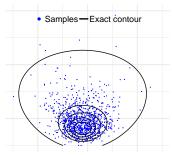
$$\sum_{i=1}^{n} (y_{i}^{2} - 2y_{i}\bar{y} + \bar{y}^{2}) + \sum_{i=1}^{n} (\mu^{2} - 2y_{i}\mu - \bar{y}^{2} + 2y_{i}\bar{y})$$

$$\sum_{i=1}^{n} (y_{i} - \bar{y})^{2} + n(\mu^{2} - 2\bar{y}\mu - \bar{y}^{2} + 2\bar{y}\bar{y})$$

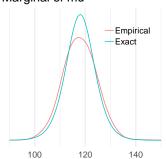
$$\sum_{i=1}^{n} (y_{i} - \bar{y})^{2} + n(\bar{y} - \mu)^{2}$$



$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$

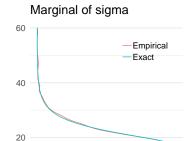


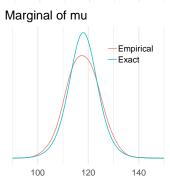
Marginal of mu



$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$
 marginals $p(\mu \mid y) = \int p(\mu, \sigma \mid y) d\sigma$

Joint posterior • Samples—Exact contour





$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$
 marginals $p(\mu \mid y) = \int p(\mu, \sigma \mid y) d\sigma$ $p(\sigma \mid y) = \int p(\mu, \sigma \mid y) d\mu$

$$p(\sigma^2 \mid y) \propto \int p(\mu, \sigma^2 \mid y) d\mu$$

$$p(\sigma^2 \mid y) \propto \int p(\mu, \sigma^2 \mid y) d\mu$$

$$\propto \int \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \left[(n-1)s^2 + n(\bar{y} - \mu)^2 \right] \right) d\mu$$

$$p(\sigma^{2} \mid y) \propto \int p(\mu, \sigma^{2} \mid y) d\mu$$

$$\propto \int \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^{2}} \left[(n-1)s^{2} + n(\bar{y} - \mu)^{2} \right] \right) d\mu$$

$$\propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^{2}} (n-1)s^{2}\right)$$

$$\int \exp\left(-\frac{n}{2\sigma^{2}} (\bar{y} - \mu)^{2}\right) d\mu$$

$$\begin{split} \rho(\sigma^2 \mid y) &\propto \int \rho(\mu, \sigma^2 \mid y) d\mu \\ &\propto \int \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \left[(n-1)s^2 + n(\bar{y} - \mu)^2\right]\right) d\mu \\ &\propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2}(n-1)s^2\right) \\ &\int \exp\left(-\frac{n}{2\sigma^2}(\bar{y} - \mu)^2\right) d\mu \\ &\int \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2}(y - \theta)^2\right) dy = 1 \end{split}$$

$$\begin{split} \rho(\sigma^2 \mid y) &\propto \int \rho(\mu, \sigma^2 \mid y) d\mu \\ &\propto \int \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \left[(n-1)s^2 + n(\bar{y} - \mu)^2\right]\right) d\mu \\ &\propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2}(n-1)s^2\right) \\ &\int \exp\left(-\frac{n}{2\sigma^2}(\bar{y} - \mu)^2\right) d\mu \\ &\int \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2}(y-\theta)^2\right) dy = 1 \\ &\propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2}(n-1)s^2\right) \sqrt{2\pi\sigma^2/n} \end{split}$$

Marginal posterior $p(\sigma^2 \mid y)$ (easier for σ^2 than σ)

$$\begin{split} \rho(\sigma^2 \mid y) & \propto & \int \rho(\mu, \sigma^2 \mid y) d\mu \\ & \propto & \int \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \left[(n-1)s^2 + n(\bar{y} - \mu)^2\right]\right) d\mu \\ & \propto & \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} (n-1)s^2\right) \\ & \int \exp\left(-\frac{n}{2\sigma^2} (\bar{y} - \mu)^2\right) d\mu \\ & \int \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2} (y - \theta)^2\right) dy = 1 \\ & \propto & \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} (n-1)s^2\right) \sqrt{2\pi\sigma^2/n} \\ & \propto & (\sigma^2)^{-(n+1)/2} \exp\left(-\frac{(n-1)s^2}{2\sigma^2}\right) \end{split}$$

Marginal posterior $p(\sigma^2 \mid y)$ (easier for σ^2 than σ)

$$\begin{split} \rho(\sigma^2 \mid y) & \propto & \int \rho(\mu, \sigma^2 \mid y) d\mu \\ & \propto & \int \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \left[(n-1)s^2 + n(\bar{y}-\mu)^2\right]\right) d\mu \\ & \propto & \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2}(n-1)s^2\right) \\ & \int \exp\left(-\frac{n}{2\sigma^2}(\bar{y}-\mu)^2\right) d\mu \\ & \int \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2}(y-\theta)^2\right) dy = 1 \\ & \propto & \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2}(n-1)s^2\right) \sqrt{2\pi\sigma^2/n} \\ & \propto & (\sigma^2)^{-(n+1)/2} \exp\left(-\frac{(n-1)s^2}{2\sigma^2}\right) \\ \rho(\sigma^2 \mid y) & = & \operatorname{Inv-}\chi^2(\sigma^2 \mid n-1, s^2) \end{split}$$

Gaussian - non-informative prior

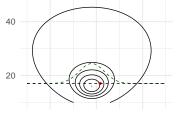
Known mean

$$\sigma^2 \mid y \sim \text{Inv-}\chi^2(n, v)$$
where $v = \frac{1}{n} \sum_{i=1}^{n} (y_i - \theta)^2$

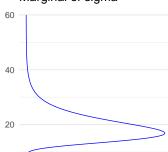
Unknown mean

$$\sigma^2 \mid y \sim \text{Inv-}\chi^2(n-1,s^2)$$
 where $s^2 = \frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2$

-Exact contour plot —Cond. distribution of mu Sample from joint post.—Sample from the marg.

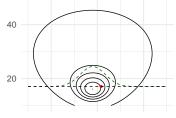


Marginal of sigma

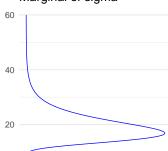


$$p(\mu, \sigma^2 \mid \mathbf{y}) = p(\mu \mid \sigma^2, \mathbf{y})p(\sigma^2 \mid \mathbf{y})$$

-Exact contour plot —Cond. distribution of mu Sample from joint post.—Sample from the marg.



Marginal of sigma

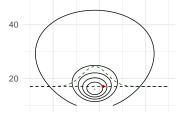


$$p(\mu, \sigma^2 \mid y) = p(\mu \mid \sigma^2, y)p(\sigma^2 \mid y)$$

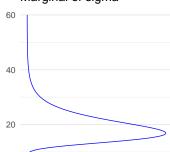
$$p(\sigma^2 \mid y) = \text{Inv-}\chi^2(\sigma^2 \mid n-1, s^2)$$

$$(\sigma^2)^{(s)} \sim p(\sigma^2 \mid y)$$

-Exact contour plot —Cond. distribution of mu Sample from joint post.—Sample from the marg.



Marginal of sigma



Factorization

60

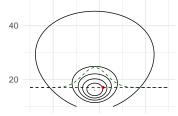
$$p(\mu, \sigma^2 \mid y) = p(\mu \mid \sigma^2, y)p(\sigma^2 \mid y)$$

$$p(\sigma^2 \mid y) = \text{Inv-}\chi^2(\sigma^2 \mid n - 1, s^2)$$

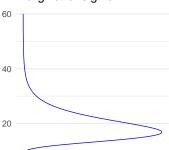
$$(\sigma^2)^{(s)} \sim p(\sigma^2 \mid y)$$

$$p(\mu \mid \sigma^2, y) = N(\mu \mid \bar{y}, \sigma^2/n)$$

-Exact contour plot —Cond. distribution of mu Sample from joint post.—Sample from the marg.



Marginal of sigma



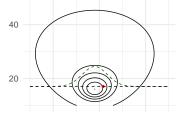
$$p(\mu, \sigma^2 \mid y) = p(\mu \mid \sigma^2, y)p(\sigma^2 \mid y)$$

$$p(\sigma^2 \mid y) = \text{Inv-}\chi^2(\sigma^2 \mid n-1, s^2)$$

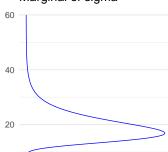
$$(\sigma^2)^{(s)} \sim p(\sigma^2 \mid y)$$

$$p(\mu \mid \sigma^2, y) = N(\mu \mid \bar{y}, \sigma^2/n) \propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \left[(n-1)s^2 + n(\bar{y} - \mu)^2 \right] \right)$$

-Exact contour plot —Cond. distribution of mu Sample from joint post.—Sample from the marg.



Marginal of sigma



$$p(\mu, \sigma^{2} \mid y) = p(\mu \mid \sigma^{2}, y)p(\sigma^{2} \mid y)$$

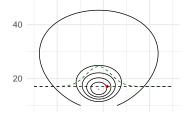
$$p(\sigma^{2} \mid y) = \text{Inv-}\chi^{2}(\sigma^{2} \mid n - 1, s^{2})$$

$$(\sigma^{2})^{(s)} \sim p(\sigma^{2} \mid y)$$

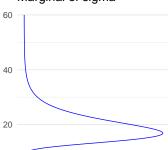
$$p(\mu \mid \sigma^{2}, y) = N(\mu \mid \bar{y}, \sigma^{2}/n)$$

$$\mu^{(s)} \sim p(\mu \mid \sigma^{2}, y)$$

-Exact contour plot —Cond. distribution of mu Sample from joint post. —Sample from the marg.



Marginal of sigma



Factorization

60

$$p(\mu, \sigma^{2} \mid y) = p(\mu \mid \sigma^{2}, y)p(\sigma^{2} \mid y)$$

$$p(\sigma^{2} \mid y) = \text{Inv-}\chi^{2}(\sigma^{2} \mid n - 1, s^{2})$$

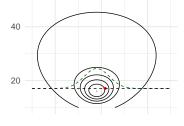
$$(\sigma^{2})^{(s)} \sim p(\sigma^{2} \mid y)$$

$$p(\mu \mid \sigma^{2}, y) = N(\mu \mid \bar{y}, \sigma^{2}/n)$$

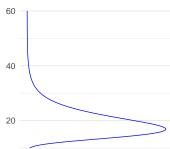
$$\mu^{(s)} \sim p(\mu \mid \sigma^{2}, y)$$

$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$

60
-Exact contour plot — Cond. distribution of mu Sample from joint post. — Sample from the marg.



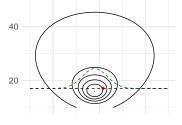
Marginal of sigma



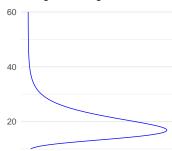
$$p(\mu, \sigma^2 \mid y) = p(\mu \mid \sigma^2, y) p(\sigma^2 \mid y)$$

60

-Exact contour plot —Cond. distribution of mu Sample from joint post.—Sample from the marg.



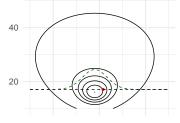
Marginal of sigma



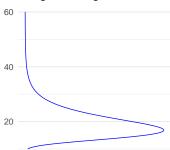
$$p(\mu, \sigma^2 \mid \mathbf{y}) = p(\mu \mid \sigma^2, \mathbf{y})p(\sigma^2 \mid \mathbf{y})$$
$$(\sigma^2)^{(s)} \sim p(\sigma^2 \mid \mathbf{y})$$

60

-Exact contour plot —Cond. distribution of mu Sample from joint post.—Sample from the marg.



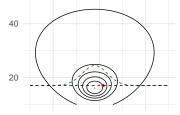
Marginal of sigma



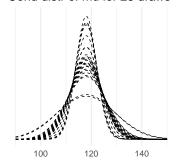
$$p(\mu, \sigma^2 \mid \mathbf{y}) = p(\mu \mid \sigma^2, \mathbf{y}) p(\sigma^2 \mid \mathbf{y})$$
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$$p(\mu \mid (\sigma^2)^{(s)}, \mathbf{y}) = N(\mu \mid \bar{\mathbf{y}}, (\sigma^2)^{(s)}/n)$$

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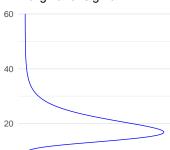
-Exact contour plot —Cond. distribution of mu Sample from joint post.—Sample from the marg.



Cond distr of mu for 25 draws



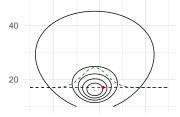
Marginal of sigma



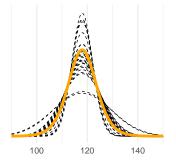
$$p(\mu, \sigma^2 \mid y) = p(\mu \mid \sigma^2, y)p(\sigma^2 \mid y)$$
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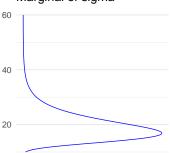
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Cond distr of mu for 25 draws



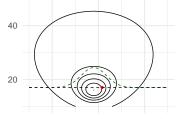
Marginal of sigma



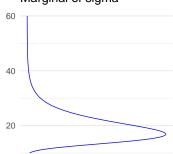
$$p(\mu, \sigma^2 \mid y) = p(\mu \mid \sigma^2, y)p(\sigma^2 \mid y)$$
$$(\sigma^2)^{(s)} \sim p(\sigma^2 \mid y)$$
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$$p(\mu \mid y) \approx \frac{1}{S} \sum_{s=1}^{S} N(\mu \mid \bar{y}, (\sigma^2)^{(s)}/n)$$

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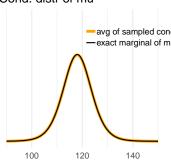
-Exact contour plot —Cond. distribution of mu Sample from joint post.—Sample from the marg.



Marginal of sigma



Cond. distr of mu



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$$p(\mu \mid y) = \int_0^\infty p(\mu, \sigma^2 \mid y) d\sigma^2$$

$$p(\mu \mid y) = \int_0^\infty p(\mu, \sigma^2 \mid y) d\sigma^2$$

$$\propto \int_0^\infty \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \left[(n-1)s^2 + n(\bar{y} - \mu)^2 \right] \right) d\sigma^2$$

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Transformation

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Recognize gamma integral $\Gamma(u) = \int_0^\infty x^{u-1} \exp(-x) dx$

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$$\propto [(n-1)s^{2} + n(\mu - \bar{y})^{2}]^{-n/2}$$

$$\propto \left[1 + \frac{n(\mu - \bar{y})^{2}}{(n-1)s^{2}}\right]^{-n/2}$$

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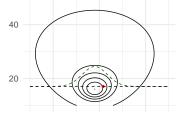
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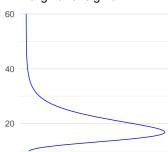
$$p(\mu \mid y) = t_{n-1}(\mu \mid \bar{y}, s^2/n) \quad \text{Student's } t$$

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-Exact contour plot —Cond. distribution of mu Sample from joint post.—Sample from the marg.



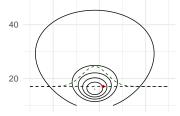
Marginal of sigma



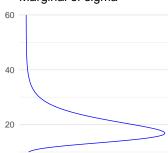
Predictive distribution for new \tilde{y}

$$p(\tilde{y} \mid y) = \int p(\tilde{y} \mid \mu, \sigma) p(\mu, \sigma \mid y) d\mu \sigma$$

-Exact contour plot —Cond. distribution of mu Sample from joint post.—Sample from the marg.



Marginal of sigma

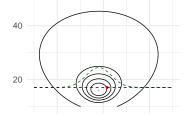


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Marginal of sigma



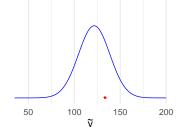
Predictive distribution for new \tilde{y}

 $\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$

$$ilde{m{y}^{(s)}} \sim m{p}(ilde{m{y}} \mid \mu^{(s)}, \sigma^{(s)})$$

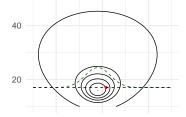
Posterior predictive distribution

 $p(\tilde{y} \mid y) = \int p(\tilde{y} \mid \mu, \sigma) p(\mu, \sigma \mid y) d\mu \dot{\underline{\sigma}}$ Sample from the predictive distribution given the posterior same



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-Exact contour plot —Cond. distribution of mu Sample from joint post. — Sample from the marg.



Marginal of sigma



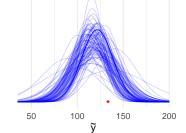
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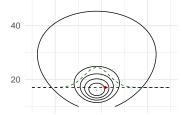
$$\tilde{y}^{(s)} \sim p(\tilde{y} \mid \mu^{(s)}, \sigma^{(s)})$$

Posterior predictive distribution

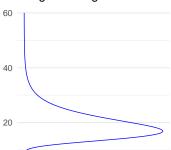


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-Exact contour plot —Cond. distribution of mu Sample from joint post. — Sample from the marg.



Marginal of sigma



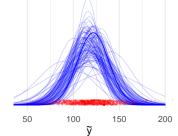
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 Sample from the predictive distribution given the posterior sample from the predictive distribution given the predictive distribution from the predictive distr

$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$

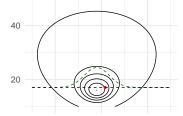
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Posterior predictive distribution

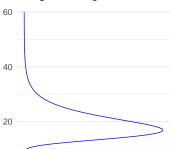


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-Exact contour plot —Cond. distribution of mu Sample from joint post. — Sample from the marg.



Marginal of sigma



Predictive distribution for new \tilde{y}

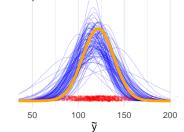
$$p(\tilde{y} \mid y) = \int p(\tilde{y} \mid \mu, \sigma) p(\mu, \sigma \mid y) d\mu \sigma$$
 • Sample from the predictive distribution given the posterior sample from the predictive distribution given the posterior sample from the predictive distribution given the posterior sample from the predictive distribution given given

$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$
$$\tilde{y}^{(s)} \sim p(\tilde{y} \mid \mu^{(s)}, \sigma^{(s)})$$

Posterior predictive distribution

· Sample from the predictive distribution

Exact predictive distribution



Posterior predictive distribution given known variance

$$p(\tilde{y} \mid \sigma^2, y) = \int p(\tilde{y} \mid \mu, \sigma^2) p(\mu \mid \sigma^2, y) d\mu$$

Posterior predictive distribution given known variance

$$p(\tilde{y} \mid \sigma^2, y) = \int p(\tilde{y} \mid \mu, \sigma^2) p(\mu \mid \sigma^2, y) d\mu$$
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this is up to scaling factor same as $p(\mu \mid \sigma^2, y)$

Posterior predictive distribution given known variance

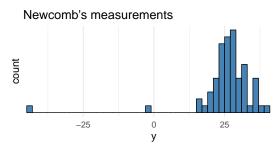
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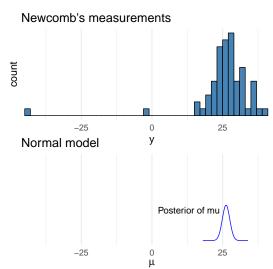
Simon Newcomb's light of speed experiment in 1882

Newcomb measured (n = 66) the time required for light to travel from his laboratory on the Potomac River to a mirror at the base of the Washington Monument and back, a total distance of 7422 meters.



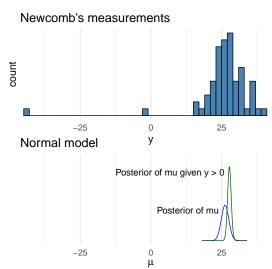
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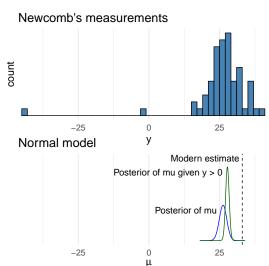
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- Conjugate prior has to have a form $p(\sigma^2)p(\mu \mid \sigma^2)$ (see the chapter notes)

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- Handy parametrization

$$\mu \mid \sigma^2 \sim N(\mu_0, \sigma^2/\kappa_0)$$

$$\sigma^2 \sim \text{Inv-}\chi^2(\nu_0, \sigma_0^2)$$

which can be written as

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which can be written as

$$p(\mu, \sigma^2) = \text{N-Inv-}\chi^2(\mu_0, \sigma_0^2/\kappa_0; \nu_0, \sigma_0^2)$$

- μ and σ^2 are a priori dependent
 - if σ^2 is large, then μ has wide prior

Joint posterior (exercise 3.9)

$$p(\mu, \sigma^2 \mid y) = \text{N-Inv-}\chi^2(\mu_n, \sigma_n^2/\kappa_n; \nu_n, \sigma_n^2)$$

where

$$\mu_{n} = \frac{\kappa_{0}}{\kappa_{0} + n} \mu_{0} + \frac{n}{\kappa_{0} + n} \bar{y}$$

$$\kappa_{n} = \kappa_{0} + n$$

$$\nu_{n} = \nu_{0} + n$$

$$\nu_{n} \sigma_{n}^{2} = \nu_{0} \sigma_{0}^{2} + (n - 1) s^{2} + \frac{\kappa_{0} n}{\kappa_{0} + n} (\bar{y} - \mu_{0})^{2}$$

Multinomial model for categorical data

- Extension of binomial
- Observation model

$$p(y \mid \theta) \propto \prod_{j=1}^k \theta_j^{y_j},$$

- BDA3 p. 69-

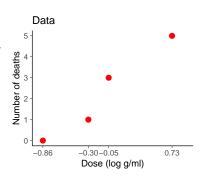
Multivariate Gaussian

Observation model

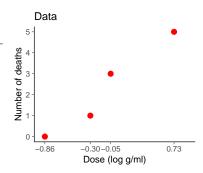
$$p(y \mid \mu, \Sigma) \propto \mid \Sigma \mid^{-1/2} \exp \left(-\frac{1}{2} (y - \mu)^T \Sigma^{-1} (y - \mu) \right),$$

- BDA3 p. 72-
- New recommended LKJ-prior mentioned in Appendix A, see more in Stan manual

Dose, x_i (log g/ml)	Number of animals, n_i	Number of deaths, <i>y_i</i>
-0.86	5	0
-0.30	5	1
-0.05	5	3
0.73	5	5



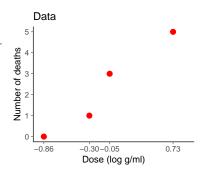
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Find out lethal dose 50% (LD50)

- used to classify how hazardous chemical is
- 1984 EEC directive has 4 levels (see the chapter notes)

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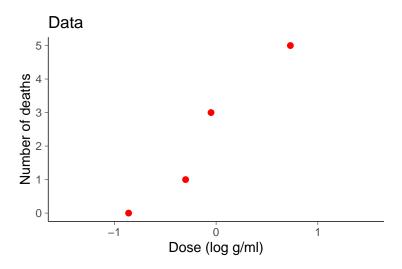


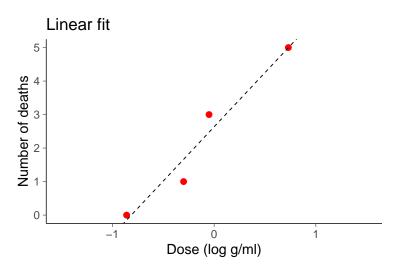
Find out lethal dose 50% (LD50)

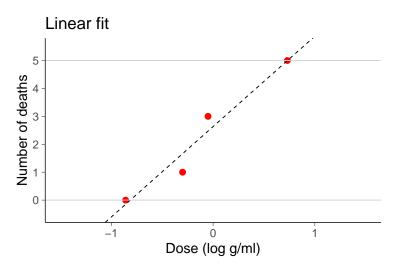
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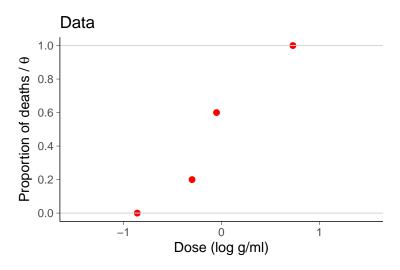
Bayesian methods help to

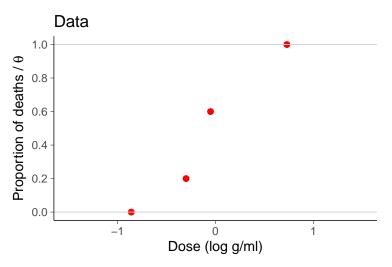
- reduce the number of animals needed
- easy to make sequential experiment and stop as soon as desired accuracy is obtained





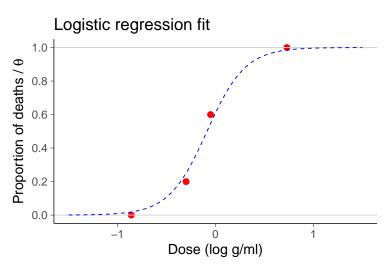






Binomial model

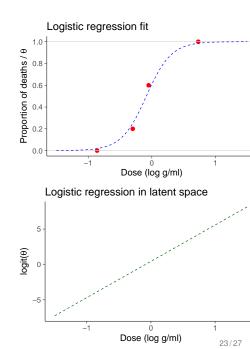
$$y_i \mid \theta_i \sim \text{Bin}(\theta_i, n_i)$$



Binomial model

$$y_i \mid \theta_i \sim \text{Bin}(\theta_i, n_i), \quad \text{logit}(\theta_i) = \log\left(\frac{\theta_i}{1 - \theta_i}\right) = \alpha + \beta x_i$$

$$y_i \mid \theta_i \sim \text{Bin}(\theta_i, n_i)$$
 $\text{logit}(\theta_i) = \log\left(\frac{\theta_i}{1 - \theta_i}\right)$
 $= \alpha + \beta x_i$



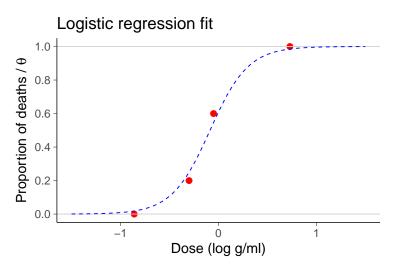
Logistic regression fit
$$y_i \mid \theta_i \sim \mathsf{Bin}(\theta_i, n_i)$$

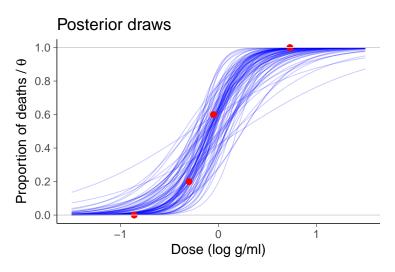
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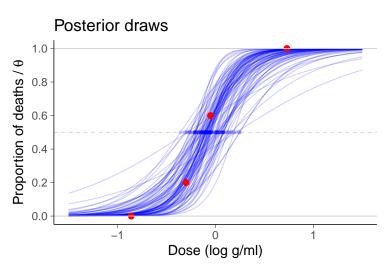
$$= \alpha + \beta x_i$$

$$\theta = \frac{1}{1 + \exp(-(\alpha + \beta x))}$$
Logistic regression fit
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Logistic regression in latent space

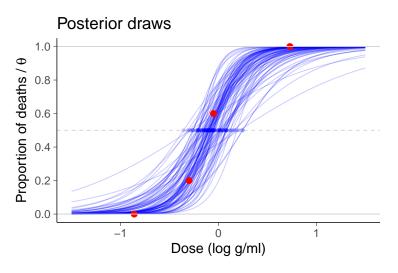
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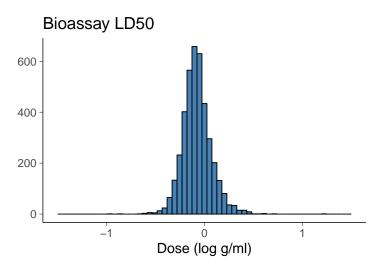




LD50:
$$E\left(\frac{y}{n}\right) = logit^{-1}(\alpha + \beta x) = 0.5$$



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Binomial model

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Likelihood

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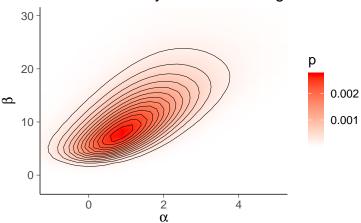
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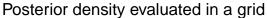
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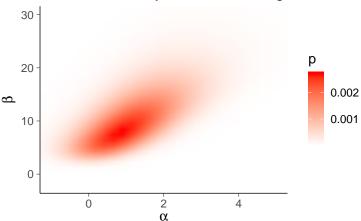
Posterior (with uniform prior on α, β)

$$p(\alpha, \beta \mid y, n, x) \propto p(\alpha, \beta) \prod_{i=1}^{n} p(y_i \mid \alpha, \beta, n_i, x_i)$$

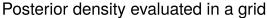
Posterior density evaluated in a grid

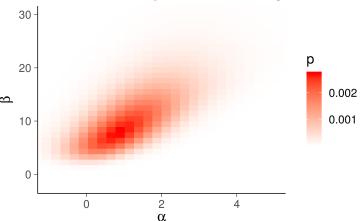






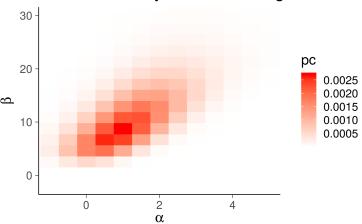
Density evaluated in grid, but plotted using interpolation



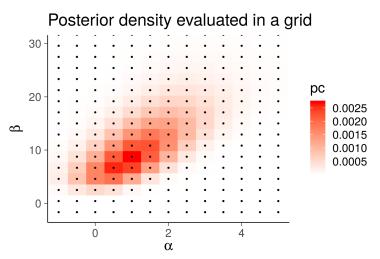


Density evaluated in grid, and plotted without interpolation

Posterior density evaluated in a grid

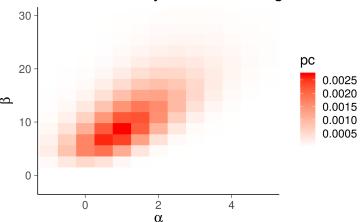


Density evaluated in a coarser grid

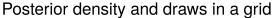


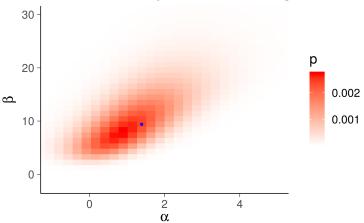
- Approximate the density as piecewise constant function
- Evaluate density in a grid over some finite region
- Density times cell area gives probability mass in each cell

Posterior density evaluated in a grid



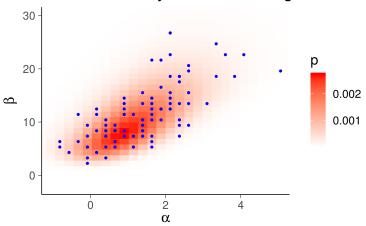
- Densities at 1, 2, and 3: 0.0027 0.0010 0.0001
- Probabilities of cells 1, 2, and 3: 0.0431 0.0166 0.0010
- Probabilities of cells sum to 1





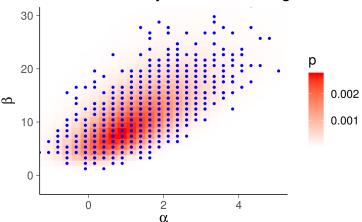
- Sample according to grid cell probabilities

Posterior density and draws in a grid



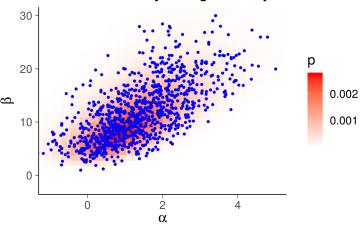
- Sample according to grid cell probabilities

Posterior density and draws in a grid



- Sample according to grid cell probabilities
- Several draws can be from the same grid cell

Posterior density in a grid and jittered draws



- Jitter can be added to improve visualization

Grid sampling

- Draws can be used to estimate expectations, for example

$$E[x_{\text{LD50}}] = E[-\alpha/\beta] \approx \frac{1}{S} \sum_{s=1}^{S} \frac{\alpha^{(s)}}{\beta^{(s)}}$$

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 Instead of sampling, grid could be used to evaluate functions directly, for example

$$\mathsf{E}[-\alpha/\beta] \approx \sum_{t=1}^T \mathbf{w}_{\mathrm{cell}}^{(t)} \frac{\alpha^{(t)}}{\beta^{(t)}},$$

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Grid sampling gets computationally too expensive in high dimensions