#### Chapter 10

- 10.1 Numerical integration (overview)
- 10.2 Distributional approximations (overview, more in Chapter 4 and 13)
- 10.3 Direct simulation and rejection sampling (overview)
- 10.4 Importance sampling (used in PSIS-LOO discussed later)
- 10.5 How many simulation draws are needed? (Ex 10.1 and 10.2)
  - see chapter notes and extra slides for how many significant digits to report
- 10.6 Software (can be skipped)
- 10.7 Debugging (can be skipped)

#### **Notation**

- In this chapter, generic  $p(\theta)$  is used instead of  $p(\theta|y)$
- unnormalized distribution is denoted by  $q(\cdot)$ 
  - $\int q(\theta)d\theta \neq 1$ , but finite
  - $q(\cdot) \propto p(\cdot)$
- proposal distribution is denoted by  $g(\cdot)$

## Numerical accuracy – floating point

- Floating point presentation of numbers. e.g. with 64bits
  - closest value to zero is  $\approx 2.2 \cdot 10^{-308}$ 
    - generate sample of 600 from normal distribution: qr=rnorm(600)
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    - pbeta(0.5, 241945, 251527) → 1 (rounding)
    - pbeta(0.5, 241945, 251527, lower.tail=FALSE)  $\approx -1.2 \cdot 10^{-42}$  there is more accuracy near 0

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  - use log densities to avoid over- and underflows in floating point presentation
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    - e.g. in Metropolis-algorithm (ex5) compute the log of ratio of densities using the identity

$$E_{p(\theta|y)}[f(\theta)] = \int f(\theta)p(\theta|y)d\theta,$$
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• Monte Carlo methods which can sample from  $p(\theta^{(s)}|y)$  using only  $q(\theta^{(s)}|y)$ 

$$E_{p(\theta|y)}[f(\theta)] \approx \frac{1}{S} \sum_{s=1}^{S} f(\theta^{(s)})$$

$$E_{\theta}[f(\theta)] = \int f(\theta) p(\theta|y) d\theta$$

- Conjugate priors and analytic solutions (Ch 1-5)
- Grid integration and other quadrature rules (Ch 3, 10)
- Independent Monte Carlo, rejection and importance sampling (Ch 10)
- Markov Chain Monte Carlo (Ch 11-12)
- Distributional approximations (Laplace, VB, EP) (Ch 4, 13)

## Quadrature integration

- The simplest quadrature integration is grid integration
  - Evaluate function in a grid and compute

$$\mathsf{E}[-\alpha/\beta] \approx \sum_{t=1}^{T} w_{\mathrm{cell}}^{(t)} \frac{\alpha^{(t)}}{\beta^{(t)}}, \quad \begin{array}{c} 4 \\ 2 \\ 0 \\ -2 \end{array}$$

where  $\mathbf{w}_{\text{cell}}^{(t)}$  is the normalized probability of a grid cell t, and  $\alpha^{(t)}$  and  $\beta^{(t)}$  are center locations of grid cells

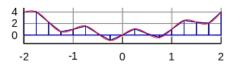
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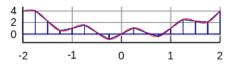
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- In 2D and higher
  - nested quadrature
  - product rules

## Monte Carlo - history

- Used already before computers
  - Buffon (18th century; needles)
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- Bayesians started to have enough cheap computation time in 1990s
  - BUGS project started 1989 (last OpenBUGS release 2014)
  - Gelfand & Smith, 1990
  - Stan initial release 2012

#### Monte Carlo

- Simulate draws from the target distribution
  - these draws can be treated as any observations
  - · a collection of draws is sample
- Use these draws, for example,
  - · to compute means, deviations, quantiles
  - to draw histograms
  - to marginalize
  - etc.

#### Monte Carlo vs. deterministic

- Monte Carlo = simulation methods
  - evaluation points are selected stochastically (randomly)
- Deterministic methods (e.g. grid)
  - evaluation points are selected by some deterministic rule
  - good deterministic methods converge faster (need less function evaluations)

- How many draws or how big sample size?
- If draws are independent
  - usual methods to estimate the uncertainty due to a finite number of observations (finite sample size)
- Markov chain Monte Carlo produces dependent draws
  - requires additional work to estimate the effective sample size

Expectation of unknown quantity

$$\mathsf{E}(\theta) \approx \frac{1}{L} \sum_{l} \theta^{(l)}$$

if L is big and  $\theta^{(I)}$  are independent, way may assume that the distribution of the expectation approaches normal distribution (see Ch 4) with variance  $\sigma_{\theta}^2/L$  (asymptotic normality)

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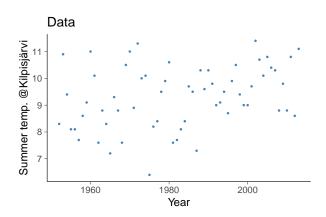
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- See Ch 4 for counter-examples for asymptotic normality

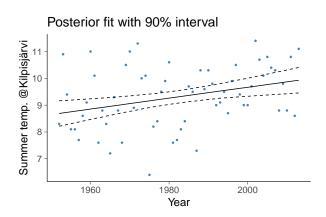
## Example: Kilpisjärvi summer temperature

Average temperature in June, July, and August at Kilpisjärvi, Finland

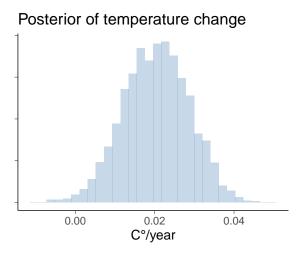


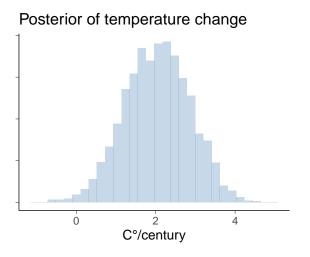
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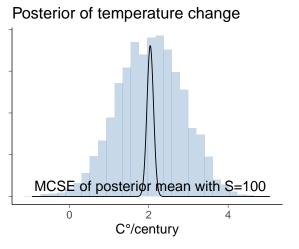
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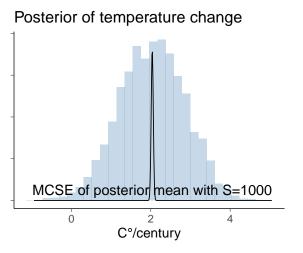
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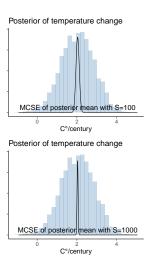


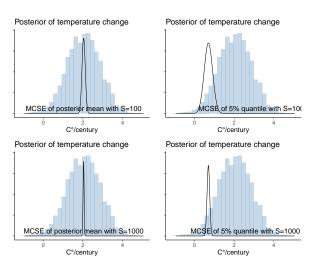


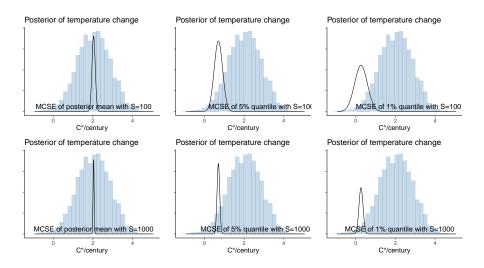
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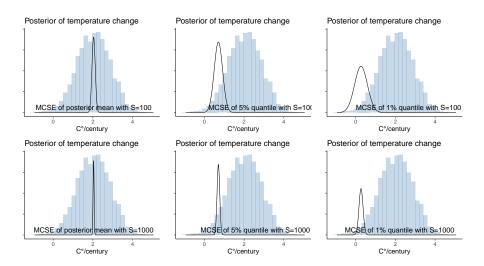


 $\sigma_{\theta} pprox$  0.827, MCSE pprox 0.0261, total deviation pprox 0.827 total deviation<sup>2</sup> =  $\sigma_{\theta}^2$  + MCSE<sup>2</sup>









Tail quantiles are more difficult to estimate

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where 
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 if  $\theta^{(l)} \in A$ 

- $I(\cdot)$  is binomially distributed as  $p(\theta \in A)$ 
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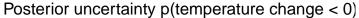
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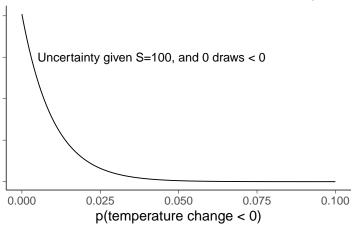
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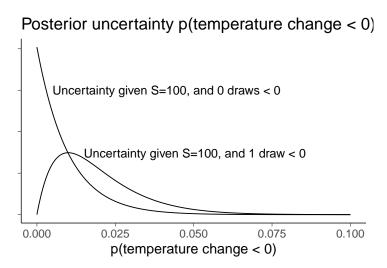
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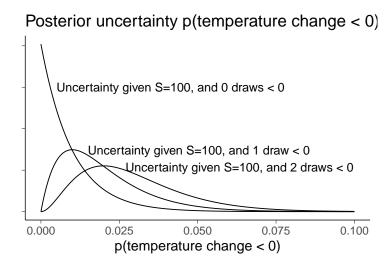
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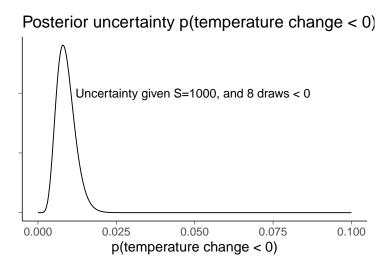
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- L = 2500 draws needed for 1% unit accuracy
- To estimate small probabilities, a large number of draws is needed
  - to be able to estimate p, need to get draws with  $\theta^{(l)} \in A$ , which in expectation requires  $L \gg 1/p$

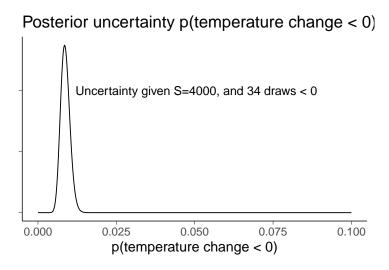












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  - 0.9960000 (NO!)
  - 1.00 (depends on the context)
  - With 4000 draws MCSE  $\approx$  0.002. We could report that probability is very likely larger than 0.99, or sample more to justify reporting three digits

- Too many digits make reading of the results slower and give false impression of the accuracy
- Don't show digits which are just random noise
  - check what is the Monte Carlo standard error
- Show meaningful digits given the posterior uncertainty
- Example: The mean and 90% central posterior interval for temperature increase C°/century based on posterior draws
  - 2.050774 and [0.7472868 3.3017524] (NO!)
  - 2.1 and [0.7 3.3]
  - 2 and [1 3] (depends on the context)
- Example: The probability that temp increase is positive
  - 0.9960000 (NO!)
  - 1.00 (depends on the context)
  - With 4000 draws MCSE  $\approx$  0.002. We could report that probability is very likely larger than 0.99, or sample more to justify reporting three digits
  - For probabilities close to 0 or 1, consider also when the model assumption justify certain accuracy

- Less draws needed with
  - · deterministic methods
  - marginalization (Rao-Blackwellization)
  - variance reduction methods, such, control variates

- Number of independent draws needed doesn't depend on the number of dimensions
  - but it may be difficult to obtain independent draws in high dimensional case

#### Direct simulation

- Produces independent draws
  - Using analytic transformations of unifrom random numbers (eg. appendix A)
  - factorization
  - numerical inverse-cdf
- Problem: restricted to limited set of models

#### Random number generators

- Good pseudo random number generators are sufficient for Bayesian inference
  - pseudo random generator uses deterministic algorithm to produce a sequence which is difficult to make difference from truly random sequence
  - modern software used for statistical analysis have good pseudo RNGs

## Direct simulation: Example

Box-Muller -method:
 If U<sub>1</sub> and U<sub>2</sub> are independent draws from distribution U(0,1), and

$$X_1 = \sqrt{-2\log(U_1)}\cos(2\pi U_2)$$
$$X_2 = \sqrt{-2\log(U_1)}\sin(2\pi U_2)$$

then  $X_1$  and  $X_2$  are independent draws from the distribution N(0,1)

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then  $X_1$  and  $X_2$  are independent draws from the distribution N(0,1)

- not the fastest method due to trigonometric computations
- for normal distrbution more than ten different methods
- · e.g. R uses inverse-CDF

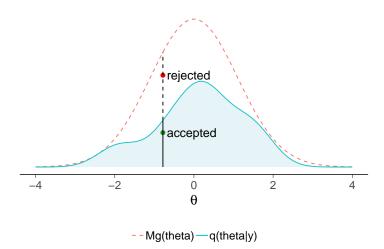
# Grid sampling and curse of dimensionality

- 10 parameters
- if we don't know beforehand where the posterior mass is
  - need to choose wide box for the grid
  - need to have enough grid points to get some of them where essential mass is
- e.g. 50 or 1000 grid points per dimension
  - $\rightarrow$  50<sup>10</sup>  $\approx$  1e17 grid points
  - $\rightarrow$  1000<sup>10</sup>  $\approx$  1e30 grid points
- R and my current laptop can compute density of normal distribution about 20 million times per second
  - → evaluation in 1e17 grid points would take 150 years
  - ightarrow evaluation in 1e30 grid points would take 1 500 billion years

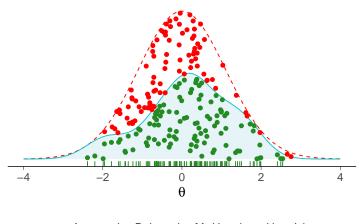
## Indirect sampling

- Rejection sampling
- Importance sampling
- Markov chain Monte Carlo (next week)

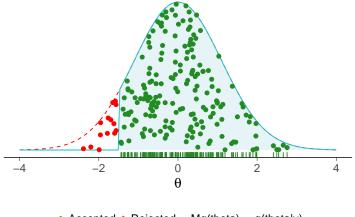
- Proposal forms envelope over the target distribution  $q(\theta|y)/Mg(\theta) \le 1$
- Draw from the proposal and accept with probability  $Mg(\theta)/q(\theta|y)$



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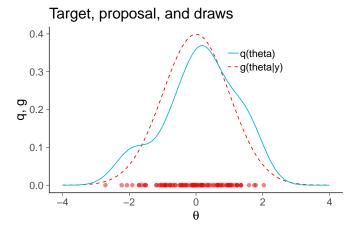


- Proposal forms envelope over the target distribution  $q(\theta|y)/Mg(\theta) \le 1$
- Draw from the proposal and accept with probability  $Mg(\theta)/q(\theta|y)$
- Common for truncated distributions

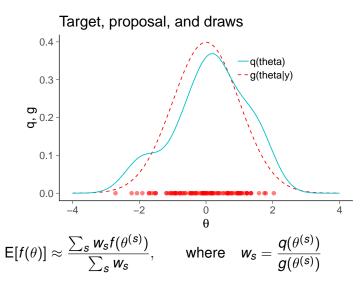


- The number of accepted draws is the effective sample size
  - with bad proposal distribution may require a lot of trials
  - selection of good proposal gets very difficult when the number of dimensions increase
  - reliable diagnostics and thus can be a useful part

- Proposal does not need to have a higher value everywhere

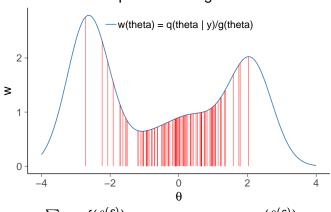


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- Proposal does not need to have a higher value everywhere

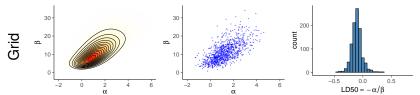
#### Draws and importance weights

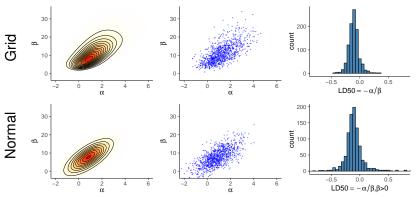


$$\mathsf{E}[f( heta)] pprox rac{\sum_{\mathcal{S}} w_{\mathcal{S}} f( heta^{(\mathcal{S})})}{\sum_{\mathcal{S}} w_{\mathcal{S}}}, \qquad ext{where} \quad w_{\mathcal{S}} = rac{q( heta^{(\mathcal{S})})}{g( heta^{(\mathcal{S})})}$$

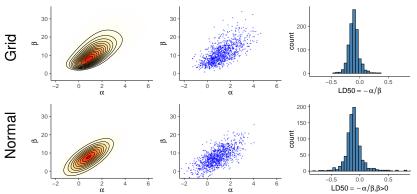
- Resampling using normalized importance weights can be used to pick a smaller number of draws with uniform weights
- Selection of good proposal gets more difficult when the number of dimensions increase
- Often used to correct distributional approximations

- Variation of the weights affect the effective sample size
  - if single weight dominates, we have effectively one sample
  - if weights are equal, we have effectively S draws
- Central limit theorem holds only if variance of the weight distribution is finite
- See Vehtari, Simpson, Gelman, Yuling and Gabry (2019).
  Pareto smoothed importance sampling. arXiv preprint arXiv:1507.02646, https://arxiv.org/abs/1507.02646 for improved diagnostics and stability.

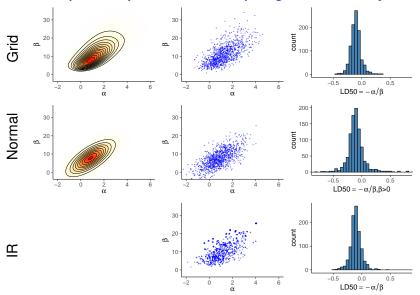


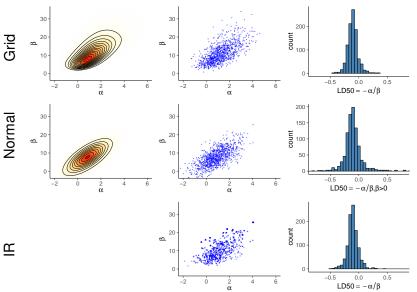


Normal approximation is discussed more in BDA3 Ch 4

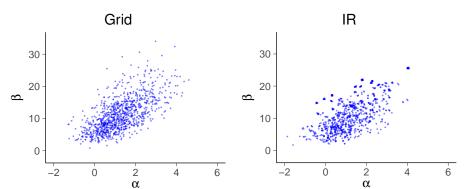


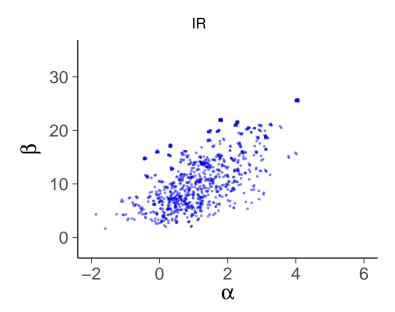
Normal approximation is discussed more in BDA3 Ch 4 But the normal approximation is not that good here: Grid  $sd(LD50) \approx 0.1$ , Normal  $sd(LD50) \approx .75!$ 

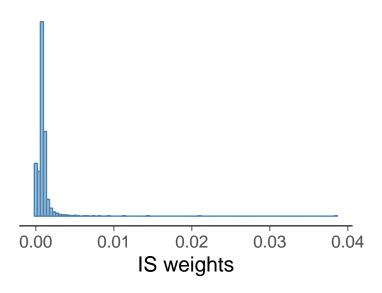


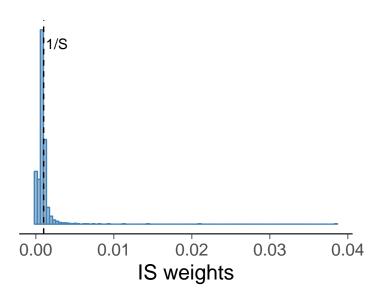


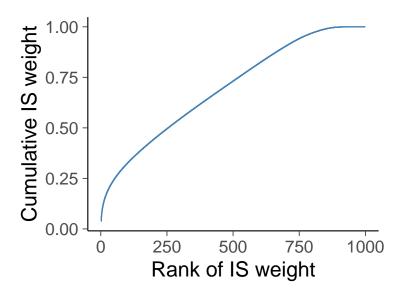
Grid sd(LD50)  $\approx$  0.1, IR sd(LD50)  $\approx$  0.1

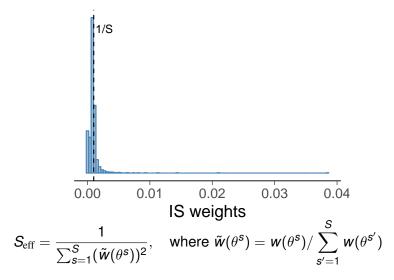


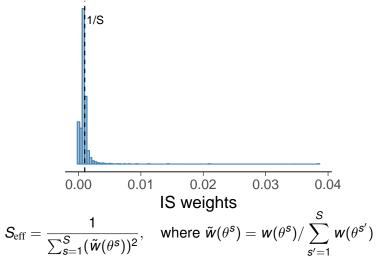




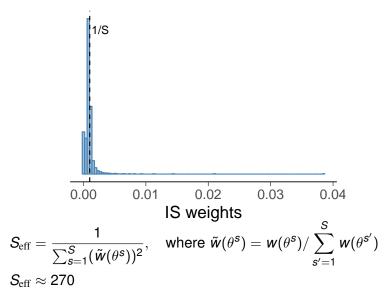


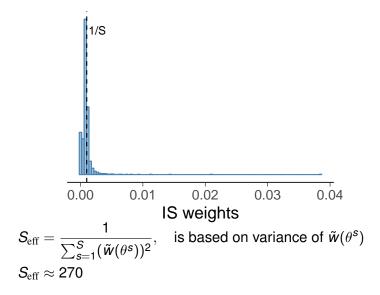


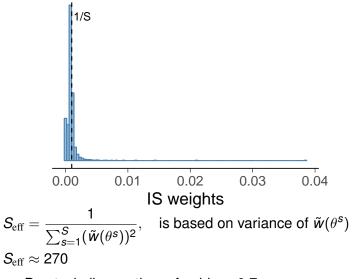




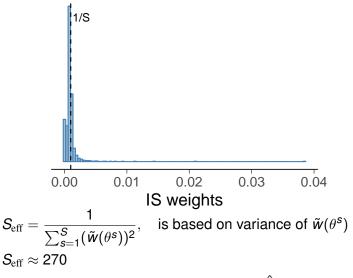
BDA3 1st (2013) and 2nd (2014) printing have an error for  $\tilde{w}(\theta^s)$ . The normalized weights equation should not have the multiplier S (the normalized weights should sum to one). Errata for the book http://www.stat.columbia.edu/~gelman/book/errata\_bda3.txt







Pareto-k diagnostic preferably < 0.7:



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 Pareto-k diagnostic estimate the number of existing moments (|1/k|)

- Pareto-k diagnostic estimate the number of existing moments ([1/k])
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# Importance sampling leave-one-out cross-validation

- Later in the course you will learn how  $p(\theta|y)$  can be used as a proposal distribution for  $p(\theta|y_{-i})$ 
  - which allows fast computation of leave-one-out cross-validation

$$p(y_i|y_{-i}) = \int p(y_i|\theta)p(\theta|y_{-i})d\theta$$

#### Curse of dimensionality

- Number of grid points increases exponentially
- Concentration of the measure, ie, where is the most of the mass?

### Markov chain Monte Carlo (MCMC)

- Pros
  - Markov chain goes where most of the posterior mass is
  - Certain MCMC methods scale well to high dimensions
- Cons
  - Draws are dependent (affects jow many draws are needed)
  - Convergence in practical time is not guaranteed
- MCMC methods in this course
  - Gibbs: "iterative conditional sampling"
  - Metropolis: "random walk in joint distribution"
  - Dynamic Hamiltonian Monte Carlo: "state-of-the-art" used in Stan