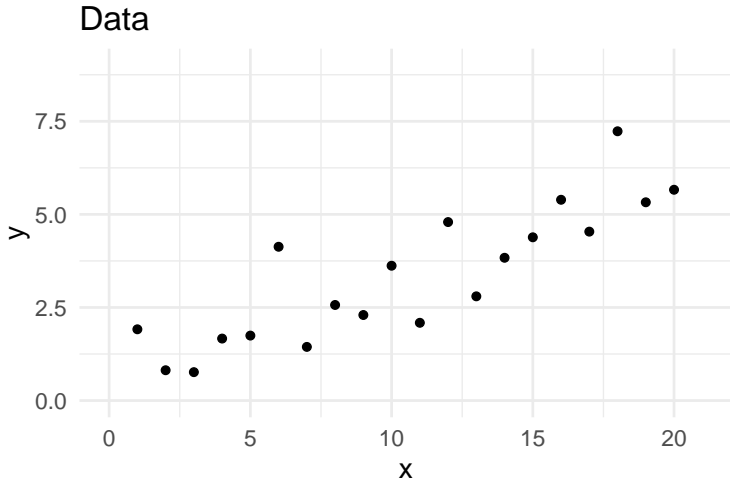


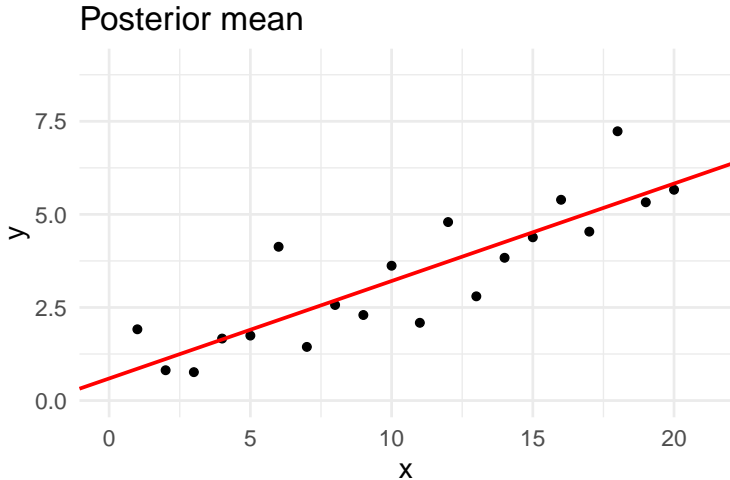
Chapter 3

- 3.1 Marginalisation
- 3.2 Normal distribution with a noninformative prior (important)
- 3.3 Normal distribution with a conjugate prior (important)
- 3.4 Multinomial model (can be skipped)
- 3.5 Multivariate normal with known variance (useful for chapter 4)
- 3.6 Multivariate normal with unknown variance (glance through)
- 3.7 Bioassay example (very important, related to one of the exercises)
- 3.8 Summary (summary)

Example of uncertainty in modeling

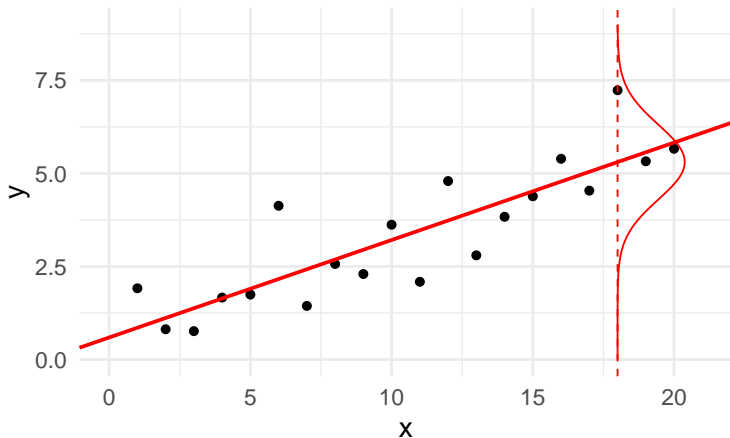


Example of uncertainty in modeling



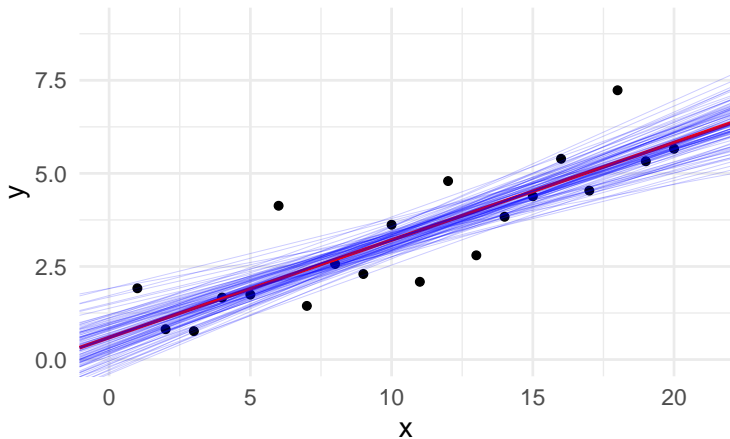
Example of uncertainty in modeling

Predictive distribution given posterior mean



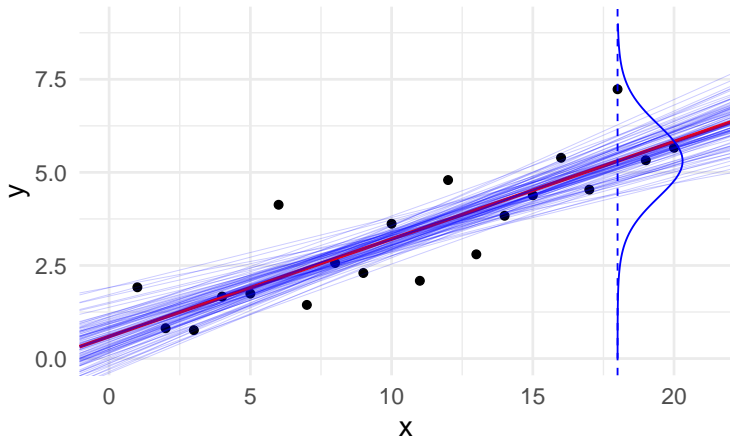
Example of uncertainty in modeling

Posterior draws



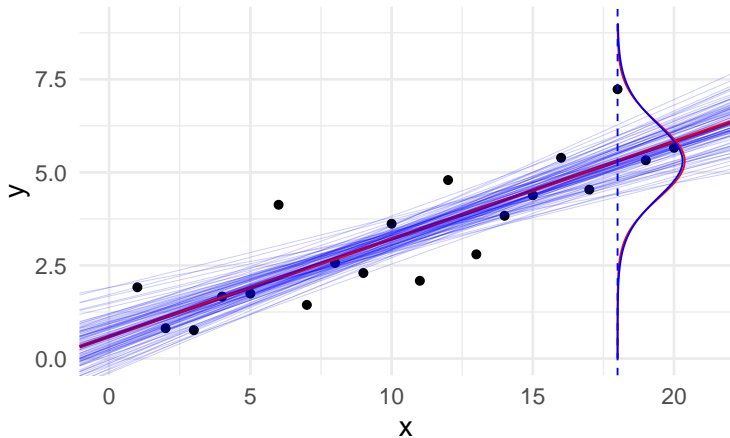
Example of uncertainty in modeling

Posterior draws and predictive distribution



Example of uncertainty in modeling

Posterior draws and predictive distribution



Marginalization

- Joint distribution of parameters

$$p(\theta_1, \theta_2 \mid y) \propto p(y \mid \theta_1, \theta_2)p(\theta_1, \theta_2)$$

- Marginalization

$$p(\theta_1 \mid y) = \int p(\theta_1, \theta_2 \mid y) d\theta_2$$

$p(\theta_1 \mid y)$ is a marginal distribution

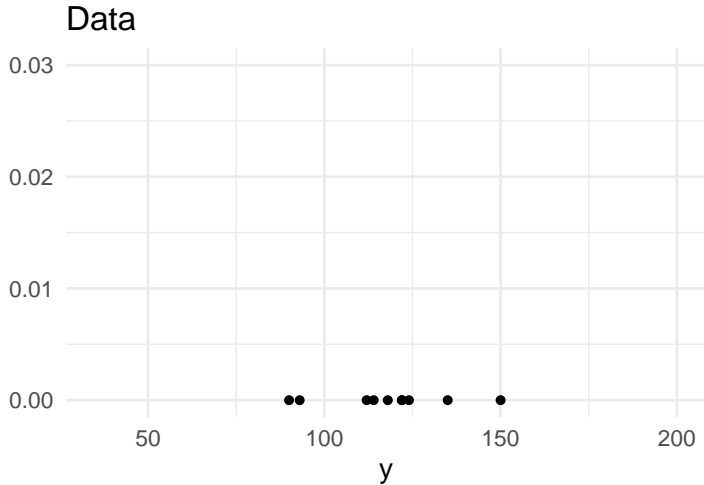
Marginalization - predictive distribution

- Marginalization over posterior distribution

$$\begin{aligned} p(\tilde{y} | y) &= \int p(\tilde{y} | \theta) p(\theta | y) d\theta \\ &= \int p(\tilde{y}, \theta | y) d\theta \end{aligned}$$

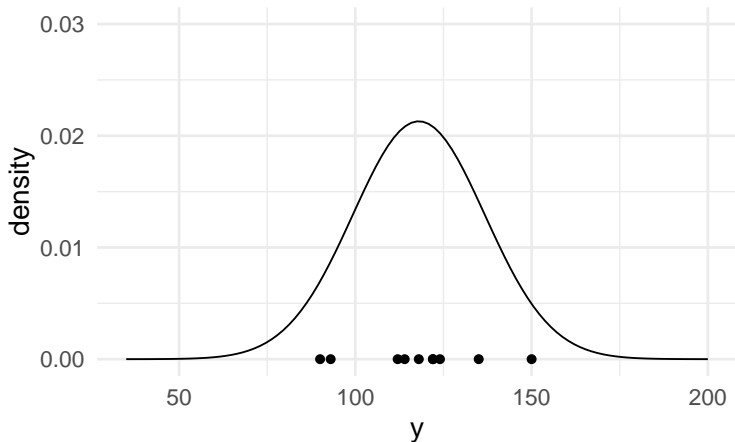
$p(\tilde{y} | y)$ is a predictive distribution

Gaussian example



Gaussian example

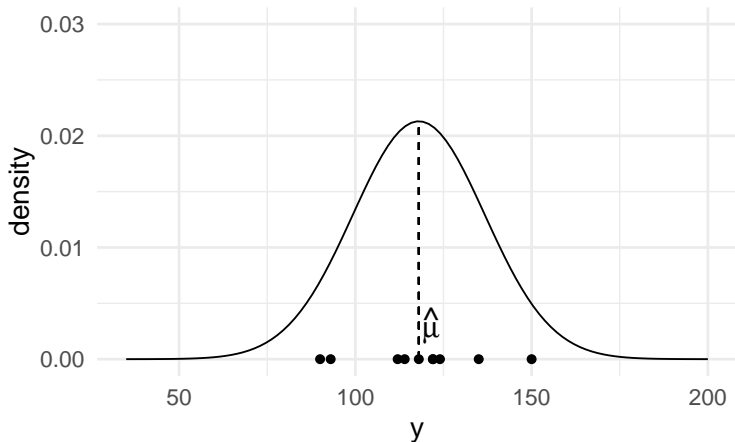
Gaussian fit with posterior mean



$$p(\textcolor{red}{y} \mid \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2}(\textcolor{red}{y} - \mu)^2\right)$$

Gaussian example

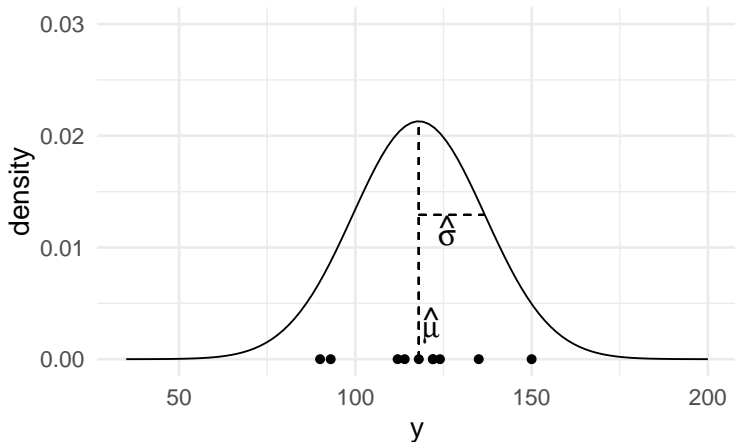
Gaussian fit with posterior mean



$$p(y \mid \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2}(y - \mu)^2\right)$$

Gaussian example

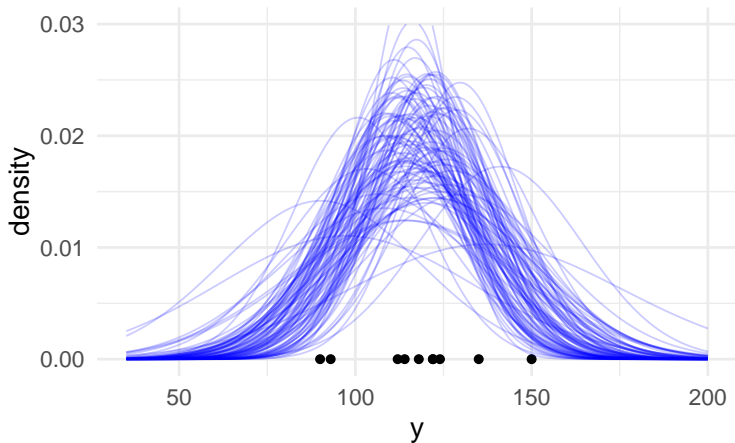
Gaussian fit with posterior mean



$$p(y \mid \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2}(y - \mu)^2\right)$$

Gaussian example

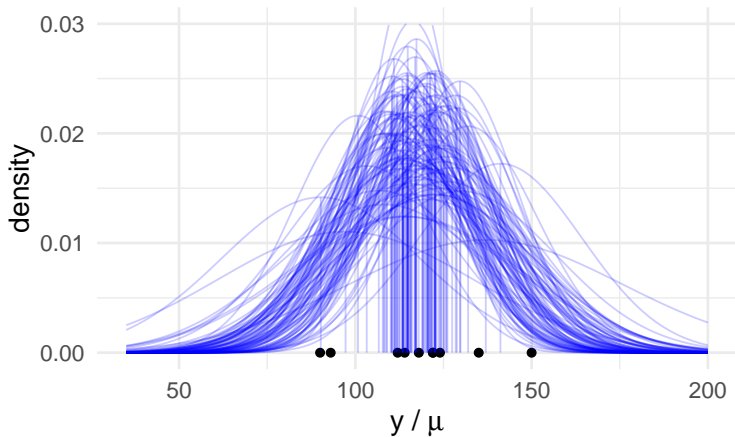
Gaussians with posterior draw parameters



$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$

Gaussian example

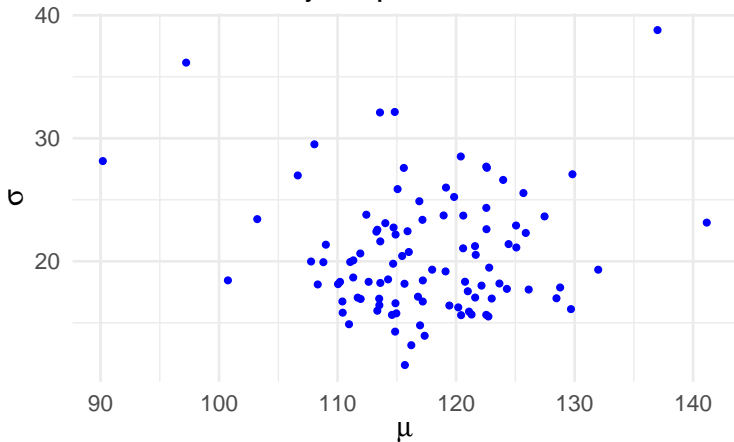
Gaussians with posterior draw parameters



$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$

Gaussian example

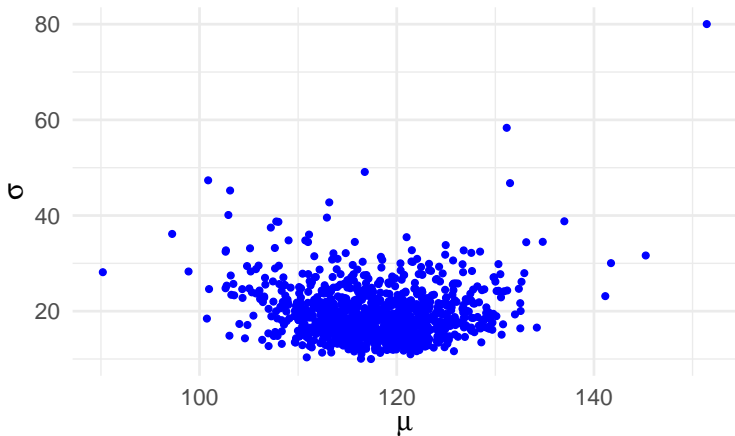
Draws from the joint posterior distribution



$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$

Gaussian example

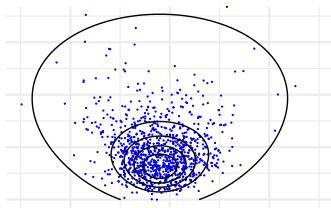
Draws from the joint posterior distribution



$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$

Joint posterior

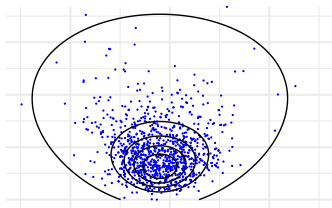
$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$



Joint posterior

$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$

with $p(\mu, \sigma^2) \propto \sigma^{-2}$

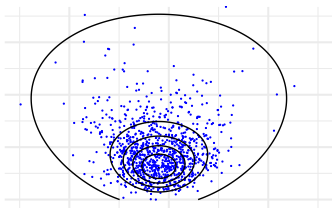


Joint posterior

$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$

with $p(\mu, \sigma^2) \propto \sigma^{-2}$

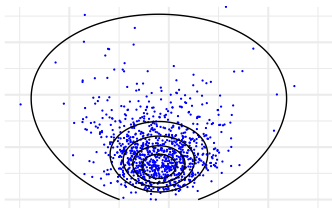
$$p(\mu, \sigma^2 \mid y) \propto \sigma^{-n-2} \exp \left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2 \right)$$



Joint posterior

$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$

with $p(\mu, \sigma^2) \propto \sigma^{-2}$



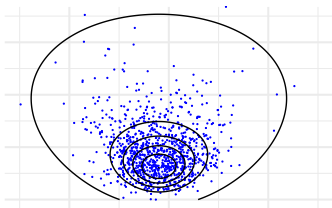
$$\begin{aligned} p(\mu, \sigma^2 \mid y) &\propto \sigma^{-n-2} \exp \left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2 \right) \\ &= \sigma^{-n-2} \exp \left(-\frac{1}{2\sigma^2} \left[\sum_{i=1}^n (y_i - \bar{y})^2 + n(\bar{y} - \mu)^2 \right] \right) \end{aligned}$$

$$\text{where } \bar{y} = \frac{1}{n} \sum_{i=1}^n y_i$$

Joint posterior

$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$

with $p(\mu, \sigma^2) \propto \sigma^{-2}$



$$\begin{aligned} p(\mu, \sigma^2 \mid y) &\propto \sigma^{-n-2} \exp \left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2 \right) \\ &= \sigma^{-n-2} \exp \left(-\frac{1}{2\sigma^2} \left[\sum_{i=1}^n (y_i - \bar{y})^2 + n(\bar{y} - \mu)^2 \right] \right) \end{aligned}$$

$$\text{where } \bar{y} = \frac{1}{n} \sum_{i=1}^n y_i$$

$$= \sigma^{-n-2} \exp \left(-\frac{1}{2\sigma^2} \left[(n-1)s^2 + n(\bar{y} - \mu)^2 \right] \right)$$

$$\text{where } s^2 = \frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2$$

Gaussian - non-informative prior

$$\sum_{i=1}^n (y_i - \mu)^2$$

Gaussian - non-informative prior

$$\sum_{i=1}^n (y_i - \mu)^2$$

$$\sum_{i=1}^n (y_i^2 - 2y_i\mu + \mu^2)$$

Gaussian - non-informative prior

$$\sum_{i=1}^n (y_i - \mu)^2$$

$$\sum_{i=1}^n (y_i^2 - 2y_i\mu + \mu^2)$$

$$\sum_{i=1}^n (y_i^2 - 2y_i\mu + \mu^2 - \bar{y}^2 + \bar{y}^2 - 2y_i\bar{y} + 2y_i\bar{y})$$

Gaussian - non-informative prior

$$\sum_{i=1}^n (y_i - \mu)^2$$

$$\sum_{i=1}^n (y_i^2 - 2y_i\mu + \mu^2)$$

$$\sum_{i=1}^n (y_i^2 - 2y_i\mu + \mu^2 - \bar{y}^2 + \bar{y}^2 - 2y_i\bar{y} + 2y_i\bar{y})$$

$$\sum_{i=1}^n (y_i^2 - 2y_i\bar{y} + \bar{y}^2) + \sum_{i=1}^n (\mu^2 - 2y_i\mu - \bar{y}^2 + 2y_i\bar{y})$$

Gaussian - non-informative prior

$$\sum_{i=1}^n (y_i - \mu)^2$$

$$\sum_{i=1}^n (y_i^2 - 2y_i\mu + \mu^2)$$

$$\sum_{i=1}^n (y_i^2 - 2y_i\mu + \mu^2 - \bar{y}^2 + \bar{y}^2 - 2y_i\bar{y} + 2y_i\bar{y})$$

$$\sum_{i=1}^n (y_i^2 - 2y_i\bar{y} + \bar{y}^2) + \sum_{i=1}^n (\mu^2 - 2y_i\mu - \bar{y}^2 + 2y_i\bar{y})$$

$$\sum_{i=1}^n (y_i - \bar{y})^2 + n(\mu^2 - 2\bar{y}\mu - \bar{y}^2 + 2\bar{y}\bar{y})$$

Gaussian - non-informative prior

$$\sum_{i=1}^n (y_i - \mu)^2$$

$$\sum_{i=1}^n (y_i^2 - 2y_i\mu + \mu^2)$$

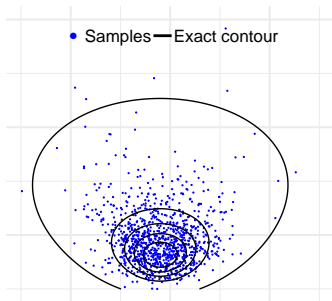
$$\sum_{i=1}^n (y_i^2 - 2y_i\mu + \mu^2 - \bar{y}^2 + \bar{y}^2 - 2y_i\bar{y} + 2y_i\bar{y})$$

$$\sum_{i=1}^n (y_i^2 - 2y_i\bar{y} + \bar{y}^2) + \sum_{i=1}^n (\mu^2 - 2y_i\mu - \bar{y}^2 + 2y_i\bar{y})$$

$$\sum_{i=1}^n (y_i - \bar{y})^2 + n(\mu^2 - 2\bar{y}\mu - \bar{y}^2 + 2\bar{y}\bar{y})$$

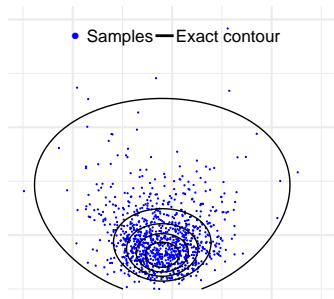
$$\sum_{i=1}^n (y_i - \bar{y})^2 + n(\bar{y} - \mu)^2$$

Joint posterior

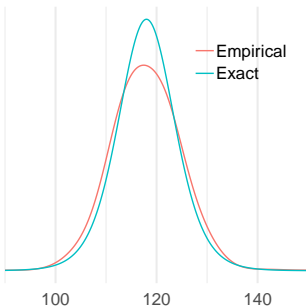


$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma \mid y)$$

Joint posterior



Marginal of mu

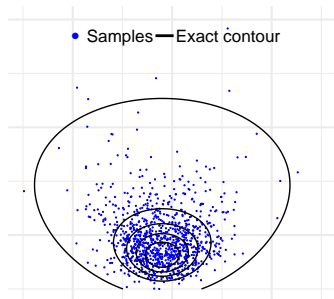


$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma | y)$$

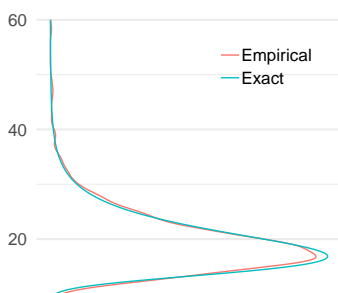
marginals

$$p(\mu | y) = \int p(\mu, \sigma | y) d\sigma$$

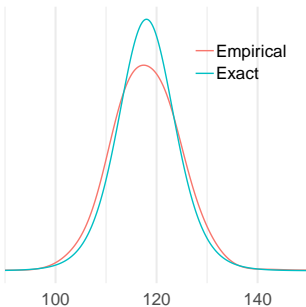
Joint posterior



Marginal of sigma



Marginal of mu



$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma | y)$$

marginals

$$p(\mu | y) = \int p(\mu, \sigma | y) d\sigma$$

$$p(\sigma | y) = \int p(\mu, \sigma | y) d\mu$$

Marginal posterior $p(\sigma^2 \mid y)$ (easier for σ^2 than σ)

$$p(\sigma^2 \mid y) \propto \int p(\mu, \sigma^2 \mid y) d\mu$$

Marginal posterior $p(\sigma^2 \mid y)$ (easier for σ^2 than σ)

$$\begin{aligned} p(\sigma^2 \mid y) &\propto \int p(\mu, \sigma^2 \mid y) d\mu \\ &\propto \int \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \left[(n-1)s^2 + n(\bar{y} - \mu)^2\right]\right) d\mu \end{aligned}$$

Marginal posterior $p(\sigma^2 \mid y)$ (easier for σ^2 than σ)

$$\begin{aligned} p(\sigma^2 \mid y) &\propto \int p(\mu, \sigma^2 \mid y) d\mu \\ &\propto \int \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \left[(n-1)s^2 + n(\bar{y} - \mu)^2\right]\right) d\mu \\ &\propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} (n-1)s^2\right) \\ &\quad \int \exp\left(-\frac{n}{2\sigma^2} (\bar{y} - \mu)^2\right) d\mu \end{aligned}$$

Marginal posterior $p(\sigma^2 | y)$ (easier for σ^2 than σ)

$$\begin{aligned} p(\sigma^2 | y) &\propto \int p(\mu, \sigma^2 | y) d\mu \\ &\propto \int \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \left[(n-1)s^2 + n(\bar{y} - \mu)^2\right]\right) d\mu \\ &\propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} (n-1)s^2\right) \\ &\quad \int \exp\left(-\frac{n}{2\sigma^2} (\bar{y} - \mu)^2\right) d\mu \\ &\quad \int \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2} (y - \theta)^2\right) dy = 1 \end{aligned}$$

Marginal posterior $p(\sigma^2 | y)$ (easier for σ^2 than σ)

$$\begin{aligned} p(\sigma^2 | y) &\propto \int p(\mu, \sigma^2 | y) d\mu \\ &\propto \int \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} [(n-1)s^2 + n(\bar{y} - \mu)^2]\right) d\mu \\ &\propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} (n-1)s^2\right) \\ &\quad \int \exp\left(-\frac{n}{2\sigma^2} (\bar{y} - \mu)^2\right) d\mu \\ &\quad \int \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2} (y - \theta)^2\right) dy = 1 \\ &\propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} (n-1)s^2\right) \sqrt{2\pi\sigma^2/n} \end{aligned}$$

Marginal posterior $p(\sigma^2 | y)$ (easier for σ^2 than σ)

$$\begin{aligned} p(\sigma^2 | y) &\propto \int p(\mu, \sigma^2 | y) d\mu \\ &\propto \int \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} [(n-1)s^2 + n(\bar{y} - \mu)^2]\right) d\mu \\ &\propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} (n-1)s^2\right) \\ &\quad \int \exp\left(-\frac{n}{2\sigma^2} (\bar{y} - \mu)^2\right) d\mu \\ &\quad \int \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2} (y - \theta)^2\right) dy = 1 \\ &\propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} (n-1)s^2\right) \sqrt{2\pi\sigma^2/n} \\ &\propto (\sigma^2)^{-(n+1)/2} \exp\left(-\frac{(n-1)s^2}{2\sigma^2}\right) \end{aligned}$$

Marginal posterior $p(\sigma^2 | y)$ (easier for σ^2 than σ)

$$\begin{aligned} p(\sigma^2 | y) &\propto \int p(\mu, \sigma^2 | y) d\mu \\ &\propto \int \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \left[(n-1)s^2 + n(\bar{y} - \mu)^2\right]\right) d\mu \\ &\propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} (n-1)s^2\right) \\ &\quad \int \exp\left(-\frac{n}{2\sigma^2} (\bar{y} - \mu)^2\right) d\mu \\ &\quad \int \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2} (y - \theta)^2\right) dy = 1 \\ &\propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} (n-1)s^2\right) \sqrt{2\pi\sigma^2/n} \\ &\propto (\sigma^2)^{-(n+1)/2} \exp\left(-\frac{(n-1)s^2}{2\sigma^2}\right) \\ p(\sigma^2 | y) &= \text{Inv-}\chi^2(\sigma^2 | n-1, s^2) \end{aligned}$$

Gaussian - non-informative prior

Known mean

$$\sigma^2 \mid y \sim \text{Inv-}\chi^2(n, \nu)$$

$$\text{where } \nu = \frac{1}{n} \sum_{i=1}^n (y_i - \theta)^2$$

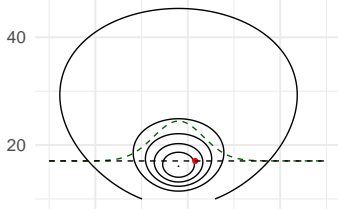
Unknown mean

$$\sigma^2 \mid y \sim \text{Inv-}\chi^2(n-1, s^2)$$

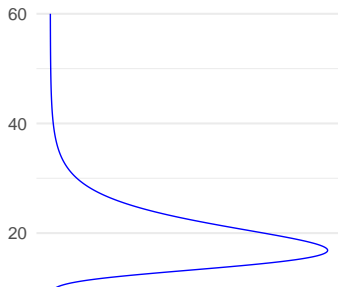
$$\text{where } s^2 = \frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2$$

Joint posterior

60
- Exact contour plot — Cond. distribution of μ
Sample from joint post. — Sample from the marg.



Marginal of sigma

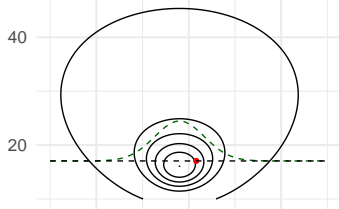


Factorization

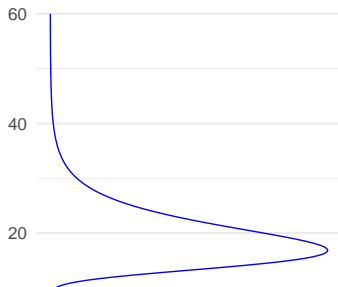
$$p(\mu, \sigma^2 | y) = p(\mu | \sigma^2, y) p(\sigma^2 | y)$$

Joint posterior

60
 - Exact contour plot — Cond. distribution of μ
 Sample from joint post. — Sample from the marg.



Marginal of sigma



Factorization

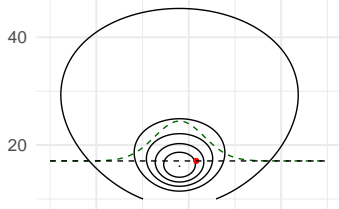
$$p(\mu, \sigma^2 | y) = p(\mu | \sigma^2, y) p(\sigma^2 | y)$$

$$p(\sigma^2 | y) = \text{Inv-}\chi^2(\sigma^2 | n - 1, s^2)$$

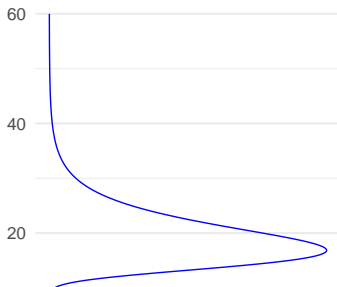
$$(\sigma^2)^{(s)} \sim p(\sigma^2 | y)$$

Joint posterior

60
 - Exact contour plot — Cond. distribution of μ
 Sample from joint post. — Sample from the marg.



Marginal of sigma



Factorization

$$p(\mu, \sigma^2 | y) = p(\mu | \sigma^2, y) p(\sigma^2 | y)$$

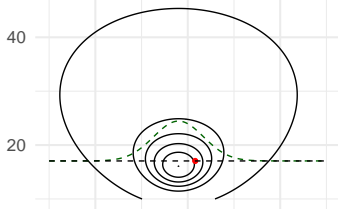
$$p(\sigma^2 | y) = \text{Inv-}\chi^2(\sigma^2 | n - 1, s^2)$$

$$(\sigma^2)^{(s)} \sim p(\sigma^2 | y)$$

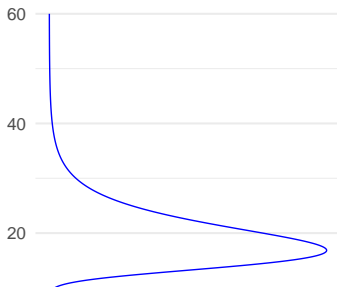
$$p(\mu | \sigma^2, y) = \text{N}(\mu | \bar{y}, \sigma^2/n)$$

Joint posterior

60
 - Exact contour plot - Cond. distribution of μ
 Sample from joint post. — Sample from the marg.



Marginal of sigma



Factorization

$$p(\mu, \sigma^2 | y) = p(\mu | \sigma^2, y) p(\sigma^2 | y)$$

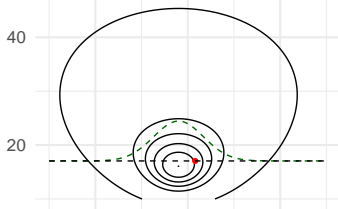
$$p(\sigma^2 | y) = \text{Inv-}\chi^2(\sigma^2 | n-1, s^2)$$

$$(\sigma^2)^{(s)} \sim p(\sigma^2 | y)$$

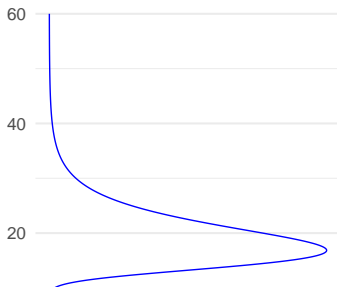
$$p(\mu | \sigma^2, y) = \text{N}(\mu | \bar{y}, \sigma^2/n) \propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} [(n-1)s^2 + n(\bar{y} - \mu)^2]\right)$$

Joint posterior

60
 - Exact contour plot - Cond. distribution of μ
 Sample from joint post. - Sample from the marg.



Marginal of sigma



Factorization

$$p(\mu, \sigma^2 | y) = p(\mu | \sigma^2, y) p(\sigma^2 | y)$$

$$p(\sigma^2 | y) = \text{Inv-}\chi^2(\sigma^2 | n - 1, s^2)$$

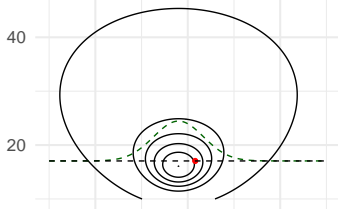
$$(\sigma^2)^{(s)} \sim p(\sigma^2 | y)$$

$$p(\mu | \sigma^2, y) = \text{N}(\mu | \bar{y}, \sigma^2/n)$$

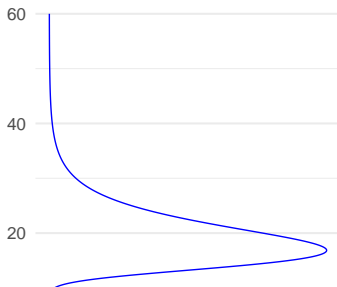
$$\mu^{(s)} \sim p(\mu | \sigma^2, y)$$

Joint posterior

60
 - Exact contour plot - Cond. distribution of μ
 Sample from joint post. - Sample from the marg.



Marginal of sigma



Factorization

$$p(\mu, \sigma^2 | y) = p(\mu | \sigma^2, y) p(\sigma^2 | y)$$

$$p(\sigma^2 | y) = \text{Inv-}\chi^2(\sigma^2 | n - 1, s^2)$$

$$(\sigma^2)^{(s)} \sim p(\sigma^2 | y)$$

$$p(\mu | \sigma^2, y) = \text{N}(\mu | \bar{y}, \sigma^2/n)$$

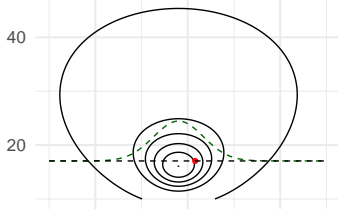
$$\mu^{(s)} \sim p(\mu | \sigma^2, y)$$

$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma | y)$$

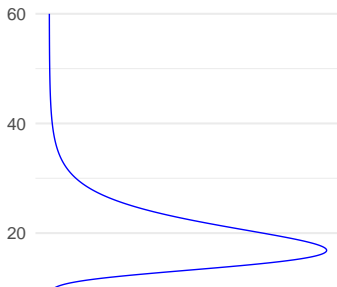
Joint posterior

60
40
20

- Exact contour plot — Cond. distribution of μ
Sample from joint post. — Sample from the marg.



Marginal of sigma

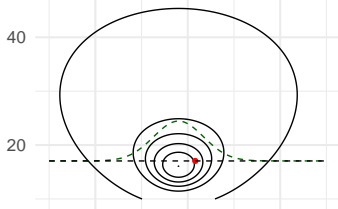


Factorization

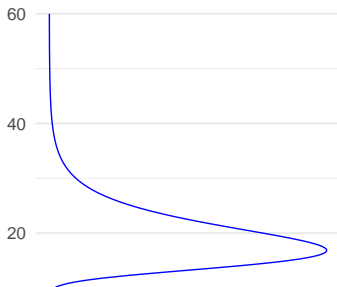
$$p(\mu, \sigma^2 | y) = p(\mu | \sigma^2, y) p(\sigma^2 | y)$$

Joint posterior

60
 - Exact contour plot — Cond. distribution of μ
 Sample from joint post. — Sample from the marg.



Marginal of sigma



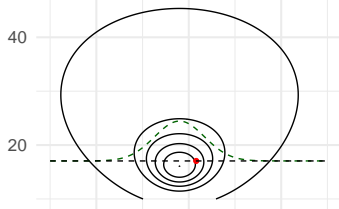
Factorization

$$p(\mu, \sigma^2 | y) = p(\mu | \sigma^2, y) p(\sigma^2 | y)$$

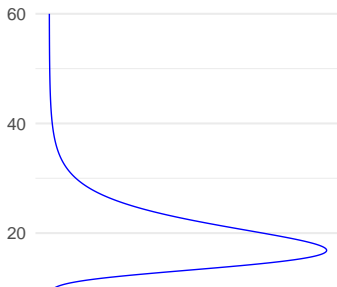
$$(\sigma^2)^{(s)} \sim p(\sigma^2 | y)$$

Joint posterior

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 - Exact contour plot — Cond. distribution of μ
 Sample from joint post. — Sample from the marg.



Marginal of sigma



Factorization

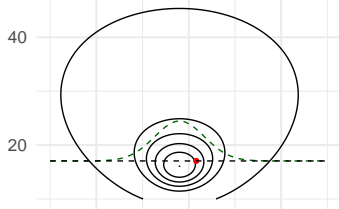
$$p(\mu, \sigma^2 | y) = p(\mu | \sigma^2, y) p(\sigma^2 | y)$$

$$(\sigma^2)^{(s)} \sim p(\sigma^2 | y)$$

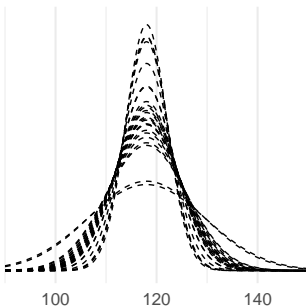
$$p(\mu | (\sigma^2)^{(s)}, y) = \text{N}(\mu | \bar{y}, (\sigma^2)^{(s)}/n)$$

Joint posterior

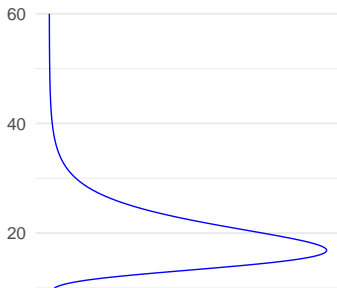
60
 - Exact contour plot - Cond. distribution of μ
 Sample from joint post. — Sample from the marg.



Cond distr of mu for 25 draws



Marginal of sigma



Factorization

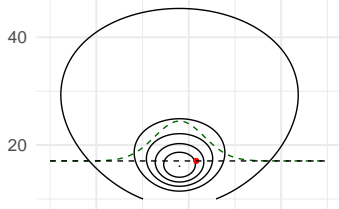
$$p(\mu, \sigma^2 | y) = p(\mu | \sigma^2, y) p(\sigma^2 | y)$$

$$(\sigma^2)^{(s)} \sim p(\sigma^2 | y)$$

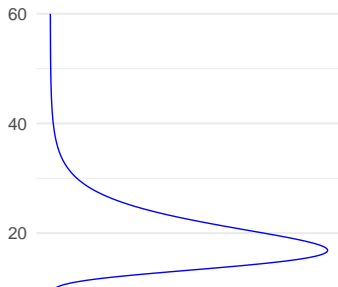
$$p(\mu | (\sigma^2)^{(s)}, y) = \text{N}(\mu | \bar{y}, (\sigma^2)^{(s)}/n)$$

Joint posterior

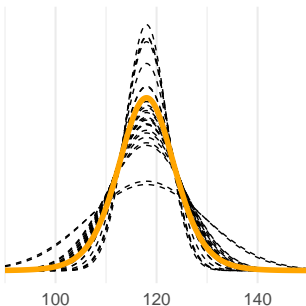
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 - Exact contour plot — Cond. distribution of μ
 Sample from joint post. — Sample from the marg.



Marginal of sigma



Cond distr of mu for 25 draws



Factorization

$$p(\mu, \sigma^2 | y) = p(\mu | \sigma^2, y) p(\sigma^2 | y)$$

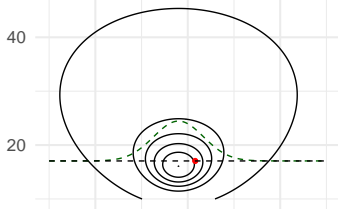
$$(\sigma^2)^{(s)} \sim p(\sigma^2 | y)$$

$$p(\mu | (\sigma^2)^{(s)}, y) = \text{N}(\mu | \bar{y}, (\sigma^2)^{(s)}/n)$$

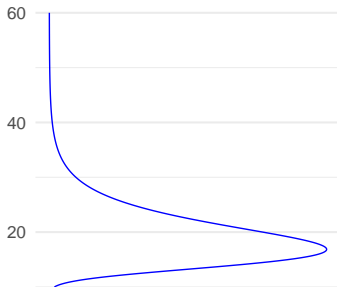
$$p(\mu | y) \approx \frac{1}{S} \sum_{s=1}^S \text{N}(\mu | \bar{y}, (\sigma^2)^{(s)}/n)$$

Joint posterior

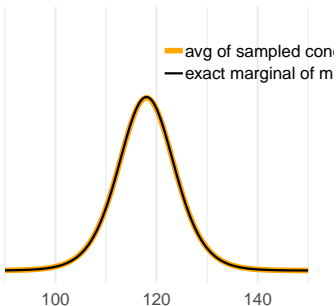
60
 - Exact contour plot - Cond. distribution of μ
 Sample from joint post. — Sample from the marg.



Marginal of sigma



Cond. distr of μ



Factorization

$$p(\mu, \sigma^2 | y) = p(\mu | \sigma^2, y) p(\sigma^2 | y)$$

$$(\sigma^2)^{(s)} \sim p(\sigma^2 | y)$$

$$p(\mu | (\sigma^2)^{(s)}, y) = N(\mu | \bar{y}, (\sigma^2)^{(s)}/n)$$

$$p(\mu | y) \approx \frac{1}{S} \sum_{s=1}^S N(\mu | \bar{y}, (\sigma^2)^{(s)}/n)$$

Marginal posterior $p(\mu \mid y)$

$$p(\mu \mid y) = \int_0^\infty p(\mu, \sigma^2 \mid y) d\sigma^2$$

Marginal posterior $p(\mu \mid y)$

$$\begin{aligned} p(\mu \mid y) &= \int_0^\infty p(\mu, \sigma^2 \mid y) d\sigma^2 \\ &\propto \int_0^\infty \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \left[(n-1)s^2 + n(\bar{y} - \mu)^2\right]\right) d\sigma^2 \end{aligned}$$

Marginal posterior $p(\mu | y)$

$$\begin{aligned} p(\mu | y) &= \int_0^\infty p(\mu, \sigma^2 | y) d\sigma^2 \\ &\propto \int_0^\infty \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \left[(n-1)s^2 + n(\bar{y} - \mu)^2\right]\right) d\sigma^2 \end{aligned}$$

Transformation

$$A = (n-1)s^2 + n(\mu - \bar{y})^2$$

Marginal posterior $p(\mu | y)$

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Transformation

$$A = (n-1)s^2 + n(\mu - \bar{y})^2 \quad \text{and} \quad z = \frac{A}{2\sigma^2}$$

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$$p(\mu | y) \propto A^{-n/2} \int_0^\infty z^{(n-2)/2} \exp(-z) dz$$

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Recognize gamma integral $\Gamma(u) = \int_0^\infty x^{u-1} \exp(-x) dx$

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Recognize gamma integral $\Gamma(u) = \int_0^\infty x^{u-1} \exp(-x) dx$

$$\begin{aligned} &\propto [(n-1)s^2 + n(\mu - \bar{y})^2]^{-n/2} \\ &\propto \left[1 + \frac{n(\mu - \bar{y})^2}{(n-1)s^2}\right] \end{aligned}$$

Marginal posterior $p(\mu | y)$

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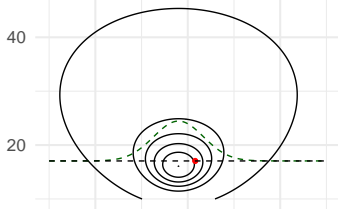
$$\propto [(n-1)s^2 + n(\mu - \bar{y})^2]^{-n/2}$$

$$\propto \left[1 + \frac{n(\mu - \bar{y})^2}{(n-1)s^2}\right]$$

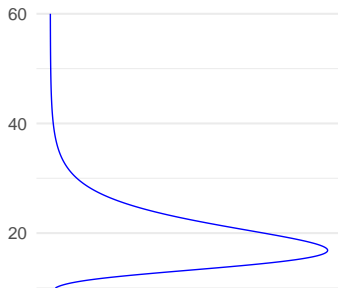
$$p(\mu | y) = t_{n-1}(\mu | \bar{y}, s^2/n) \quad \text{Student's } t$$

Joint posterior

60
 - Exact contour plot — Cond. distribution of μ
 Sample from joint post. — Sample from the marg.



Marginal of sigma

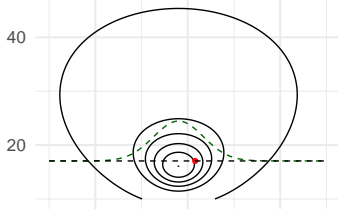


Predictive distribution for new \tilde{y}

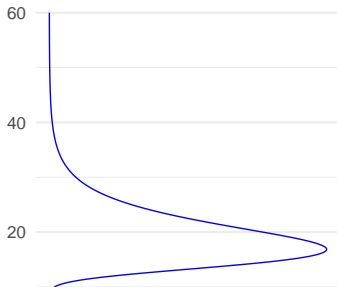
$$p(\tilde{y} | y) = \int p(\tilde{y} | \mu, \sigma) p(\mu, \sigma | y) d\mu d\sigma$$

Joint posterior

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 - Exact contour plot - Cond. distribution of μ
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Marginal of sigma



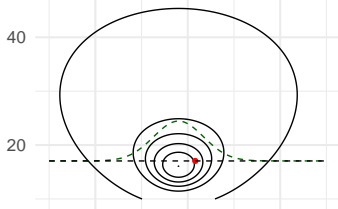
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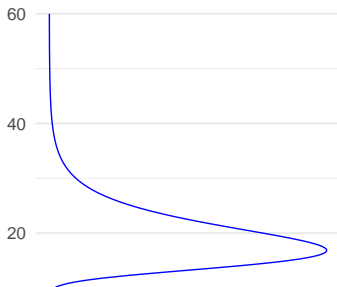
$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma | y)$$

Joint posterior

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Marginal of sigma



Predictive distribution for new \tilde{y}

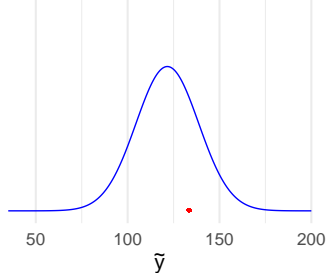
$$p(\tilde{y} | y) = \int p(\tilde{y} | \mu, \sigma) p(\mu, \sigma | y) d\mu d\sigma$$

$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma | y)$$

$$\tilde{y}^{(s)} \sim p(\tilde{y} | \mu^{(s)}, \sigma^{(s)})$$

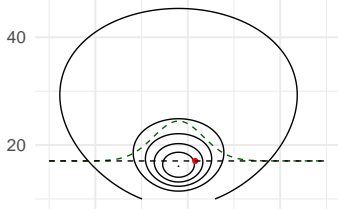
Posterior predictive distribution

Sample from the predictive distribution
 Predictive distribution given the posterior sample

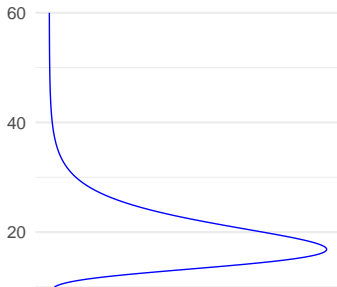


Joint posterior

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 - Exact contour plot - Cond. distribution of μ
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Marginal of sigma



Predictive distribution for new \tilde{y}

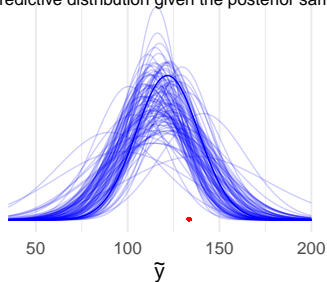
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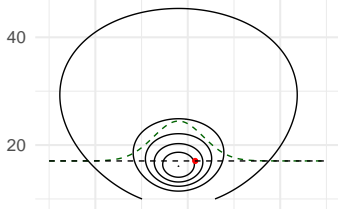
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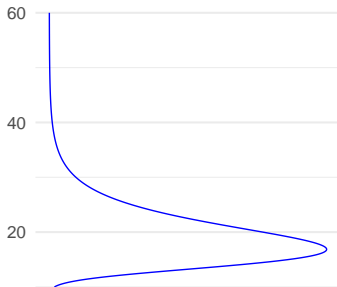


Joint posterior

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Marginal of sigma



Predictive distribution for new \tilde{y}

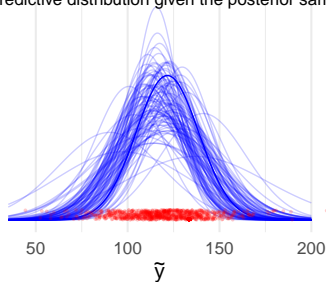
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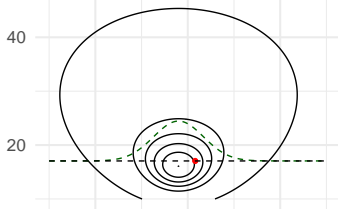
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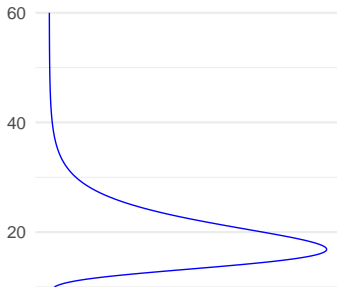


Joint posterior

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Marginal of sigma



Predictive distribution for new \tilde{y}

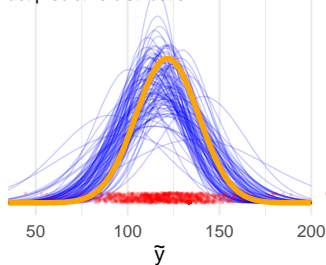
$$p(\tilde{y} | y) = \int p(\tilde{y} | \mu, \sigma) p(\mu, \sigma | y) d\mu d\sigma$$

$$\mu^{(s)}, \sigma^{(s)} \sim p(\mu, \sigma | y)$$

$$\tilde{y}^{(s)} \sim p(\tilde{y} | \mu^{(s)}, \sigma^{(s)})$$

Posterior predictive distribution

• Sample from the predictive distribution
 - Predictive distribution given the posterior sample
 — Exact predictive distribution



Gaussian - posterior predictive distribution

Posterior predictive distribution given known variance

$$p(\tilde{y} \mid \sigma^2, y) = \int p(\tilde{y} \mid \mu, \sigma^2) p(\mu \mid \sigma^2, y) d\mu$$

Gaussian - posterior predictive distribution

Posterior predictive distribution given known variance

$$\begin{aligned} p(\tilde{y} \mid \sigma^2, y) &= \int p(\tilde{y} \mid \mu, \sigma^2) p(\mu \mid \sigma^2, y) d\mu \\ &= \int \mathcal{N}(\tilde{y} \mid \mu, \sigma^2) \mathcal{N}(\mu \mid \bar{y}, \sigma^2/n) d\mu \end{aligned}$$

Gaussian - posterior predictive distribution

Posterior predictive distribution given known variance

$$\begin{aligned} p(\tilde{y} \mid \sigma^2, y) &= \int p(\tilde{y} \mid \mu, \sigma^2) p(\mu \mid \sigma^2, y) d\mu \\ &= \int \mathcal{N}(\tilde{y} \mid \mu, \sigma^2) \mathcal{N}(\mu \mid \bar{y}, \sigma^2/n) d\mu \\ &= \mathcal{N}(\tilde{y} \mid \bar{y}, (1 + \frac{1}{n})\sigma^2) \end{aligned}$$

Gaussian - posterior predictive distribution

Posterior predictive distribution given known variance

$$\begin{aligned} p(\tilde{y} \mid \sigma^2, y) &= \int p(\tilde{y} \mid \mu, \sigma^2) p(\mu \mid \sigma^2, y) d\mu \\ &= \int \mathcal{N}(\tilde{y} \mid \mu, \sigma^2) \mathcal{N}(\mu \mid \bar{y}, \sigma^2/n) d\mu \\ &= \mathcal{N}(\tilde{y} \mid \bar{y}, (1 + \frac{1}{n})\sigma^2) \end{aligned}$$

this is up to scaling factor same as $p(\mu \mid \sigma^2, y)$

Gaussian - posterior predictive distribution

Posterior predictive distribution given known variance

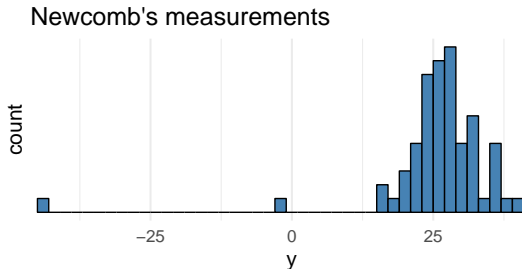
$$\begin{aligned}p(\tilde{y} \mid \sigma^2, y) &= \int p(\tilde{y} \mid \mu, \sigma^2) p(\mu \mid \sigma^2, y) d\mu \\&= \int \mathcal{N}(\tilde{y} \mid \mu, \sigma^2) \mathcal{N}(\mu \mid \bar{y}, \sigma^2/n) d\mu \\&= \mathcal{N}(\tilde{y} \mid \bar{y}, (1 + \frac{1}{n})\sigma^2)\end{aligned}$$

this is up to scaling factor same as $p(\mu \mid \sigma^2, y)$

$$p(\tilde{y} \mid y) = t_{n-1}(\tilde{y} \mid \bar{y}, (1 + \frac{1}{n})s^2)$$

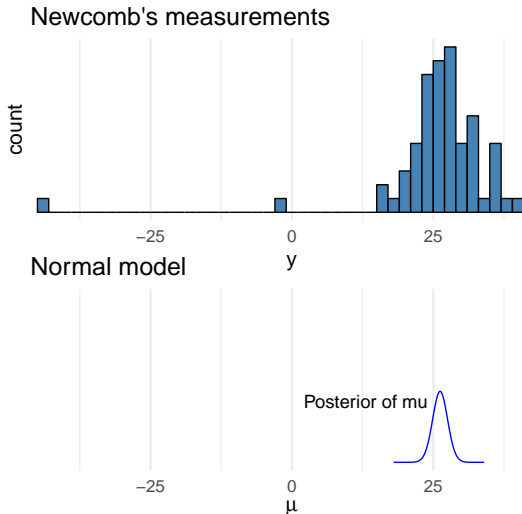
Simon Newcomb's light of speed experiment in 1882

Newcomb measured ($n = 66$) the time required for light to travel from his laboratory on the Potomac River to a mirror at the base of the Washington Monument and back, a total distance of 7422 meters.



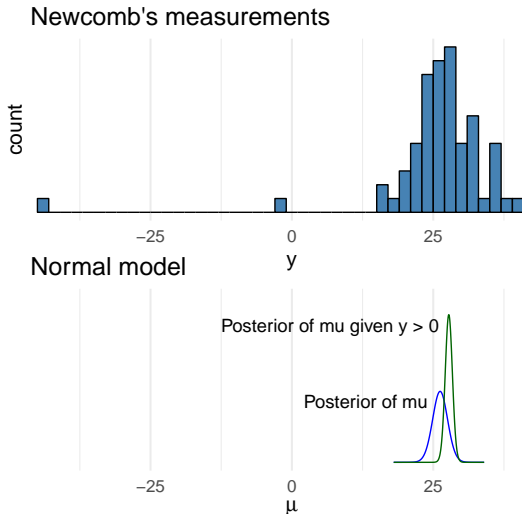
Simon Newcomb's light of speed experiment in 1882

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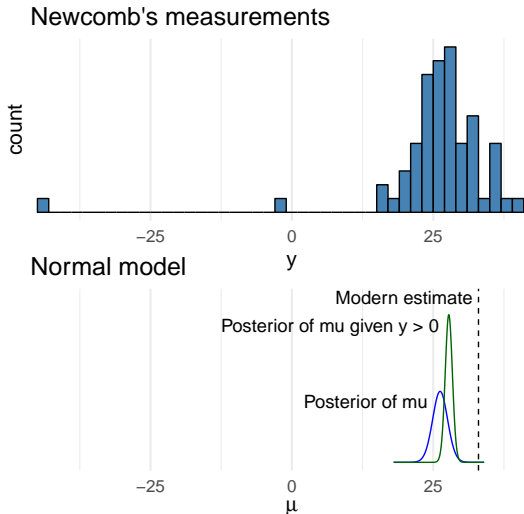
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Simon Newcomb's light of speed experiment in 1882

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Gaussian - conjugate prior

- Conjugate prior has to have a form $p(\sigma^2)p(\mu | \sigma^2)$
(see the chapter notes)

Gaussian - conjugate prior

- Conjugate prior has to have a form $p(\sigma^2)p(\mu \mid \sigma^2)$ (see the chapter notes)
- Handy parametrization

$$\begin{aligned}\mu \mid \sigma^2 &\sim \text{N}(\mu_0, \sigma^2/\kappa_0) \\ \sigma^2 &\sim \text{Inv-}\chi^2(\nu_0, \sigma_0^2)\end{aligned}$$

which can be written as

$$p(\mu, \sigma^2) = \text{N-Inv-}\chi^2(\mu_0, \sigma_0^2/\kappa_0; \nu_0, \sigma_0^2)$$

Gaussian - conjugate prior

- Conjugate prior has to have a form $p(\sigma^2)p(\mu | \sigma^2)$ (see the chapter notes)
- Handy parametrization

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which can be written as

$$p(\mu, \sigma^2) = \text{N-Inv-}\chi^2(\mu_0, \sigma_0^2/\kappa_0; \nu_0, \sigma_0^2)$$

- μ and σ^2 are a priori dependent
 - if σ^2 is large, then μ has wide prior

Gaussian - conjugate prior

Joint posterior (exercise 3.9)

$$p(\mu, \sigma^2 \mid y) = \text{N-Inv-}\chi^2(\mu_n, \sigma_n^2/\kappa_n; \nu_n, \sigma_n^2)$$

where

$$\mu_n = \frac{\kappa_0}{\kappa_0 + n} \mu_0 + \frac{n}{\kappa_0 + n} \bar{y}$$

$$\kappa_n = \kappa_0 + n$$

$$\nu_n = \nu_0 + n$$

$$\nu_n \sigma_n^2 = \nu_0 \sigma_0^2 + (n-1)s^2 + \frac{\kappa_0 n}{\kappa_0 + n} (\bar{y} - \mu_0)^2$$

Multinomial model for categorical data

- Extension of binomial
- Observation model

$$p(y \mid \theta) \propto \prod_{j=1}^k \theta_j^{y_j},$$

- BDA3 p. 69-

Multivariate Gaussian

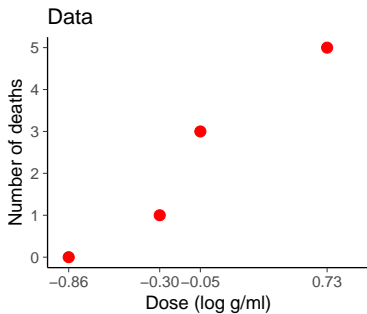
- Observation model

$$p(y \mid \mu, \Sigma) \propto |\Sigma|^{-1/2} \exp \left(-\frac{1}{2} (y - \mu)^T \Sigma^{-1} (y - \mu) \right),$$

- BDA3 p. 72-
- New recommended LKJ-prior mentioned in Appendix A, see more in Stan manual

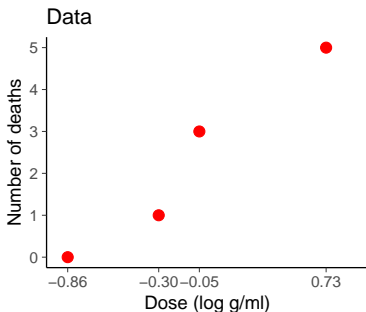
Bioassay

Dose, x_i (log g/ml)	Number of animals, n_i	Number of deaths, y_i
-0.86	5	0
-0.30	5	1
-0.05	5	3
0.73	5	5



Bioassay

Dose, x_i (log g/ml)	Number of animals, n_i	Number of deaths, y_i
-0.86	5	0
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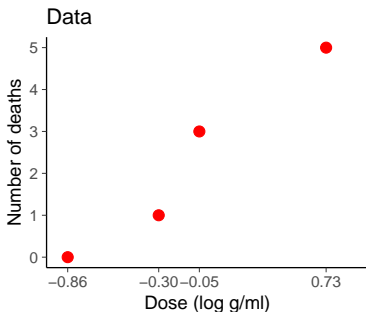


Find out lethal dose 50% (LD50)

- used to classify how hazardous chemical is
- 1984 EEC directive has 4 levels (see the chapter notes)

Bioassay

Dose, x_i (log g/ml)	Number of animals, n_i	Number of deaths, y_i
-0.86	5	0
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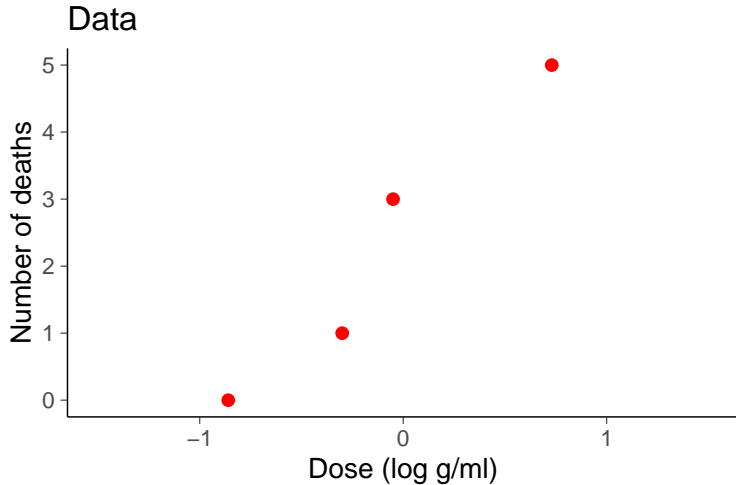
Find out lethal dose 50% (LD50)

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- 1984 EEC directive has 4 levels (see the chapter notes)

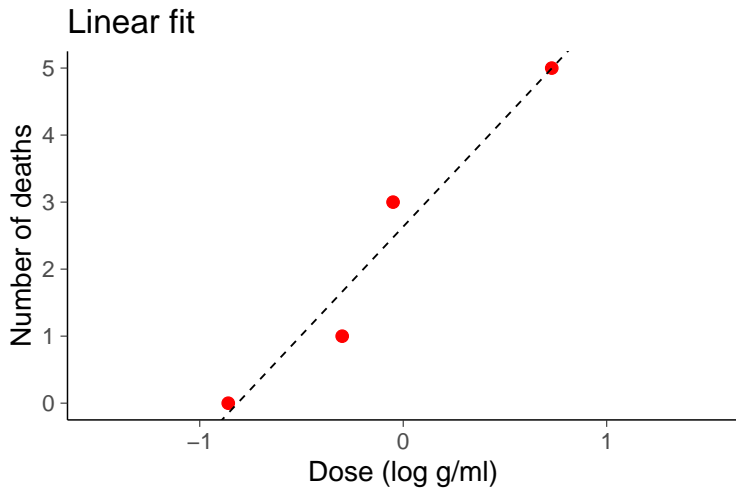
Bayesian methods help to

- reduce the number of animals needed
- easy to make sequential experiment and stop as soon as desired accuracy is obtained

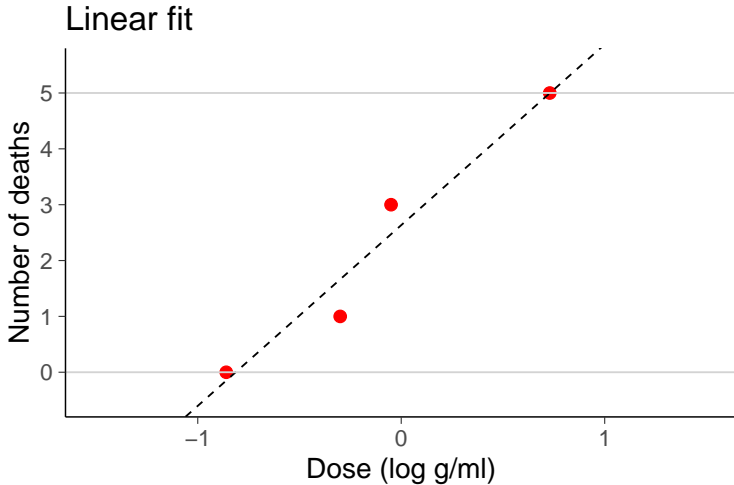
Bioassay



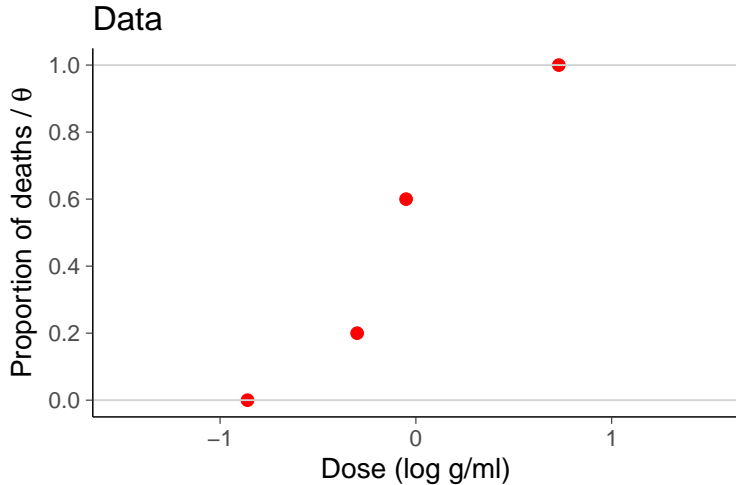
Bioassay



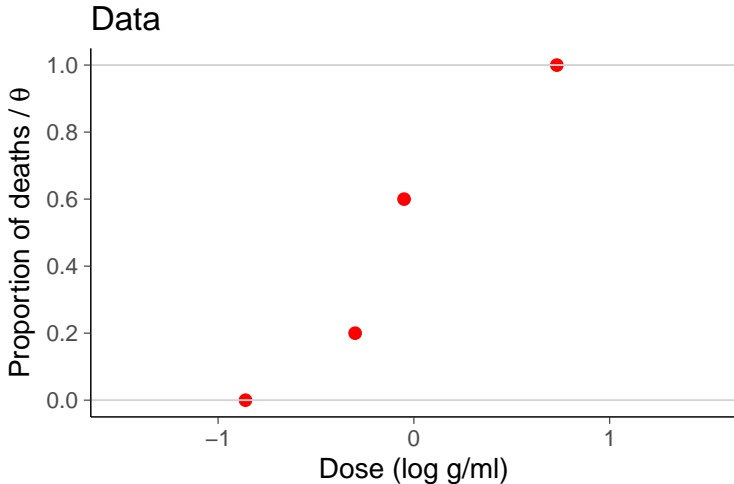
Bioassay



Bioassay



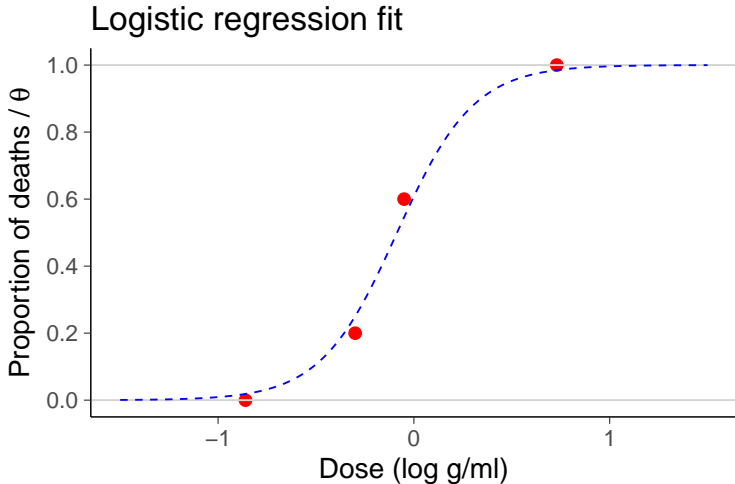
Bioassay



Binomial model

$$y_i \mid \theta_i \sim \text{Bin}(\theta_i, n_i)$$

Bioassay



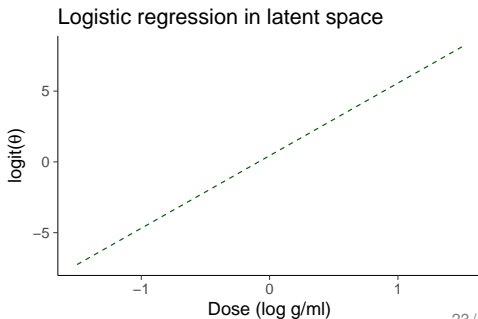
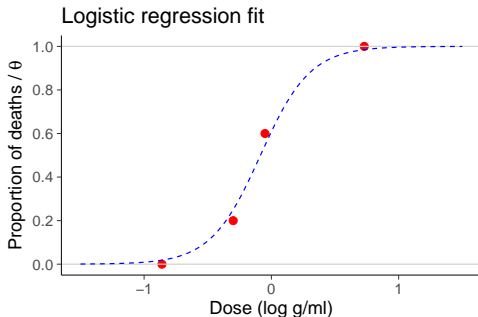
Binomial model

$$y_i \mid \theta_i \sim \text{Bin}(\theta_i, n_i), \quad \text{logit}(\theta_i) = \log\left(\frac{\theta_i}{1 - \theta_i}\right) = \alpha + \beta x_i$$

Bioassay

$$y_i \mid \theta_i \sim \text{Bin}(\theta_i, n_i)$$

$$\begin{aligned}\text{logit}(\theta_i) &= \log\left(\frac{\theta_i}{1 - \theta_i}\right) \\ &= \alpha + \beta \mathbf{x}_i\end{aligned}$$



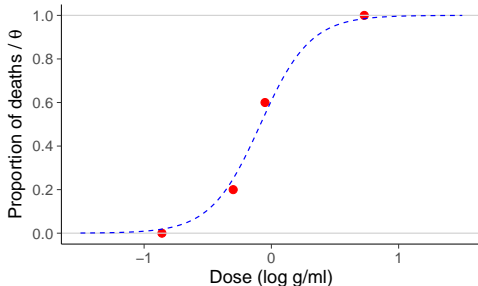
Bioassay

$$y_i \mid \theta_i \sim \text{Bin}(\theta_i, n_i)$$

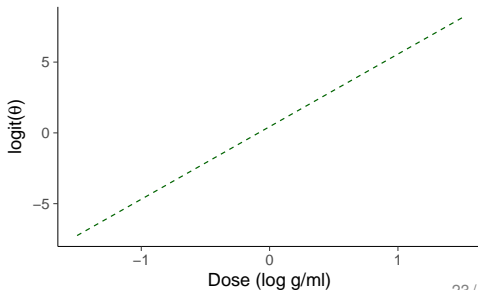
$$\begin{aligned}\text{logit}(\theta_i) &= \log\left(\frac{\theta_i}{1 - \theta_i}\right) \\ &= \alpha + \beta x_i\end{aligned}$$

$$\theta = \frac{1}{1 + \exp(\alpha + \beta x)}$$

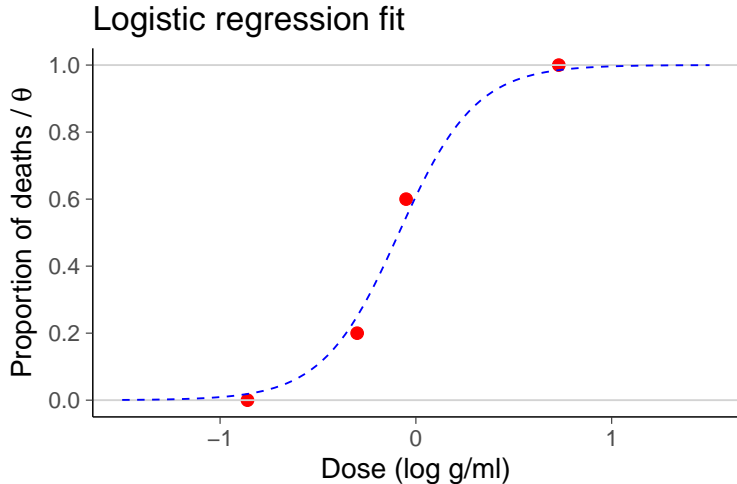
Logistic regression fit



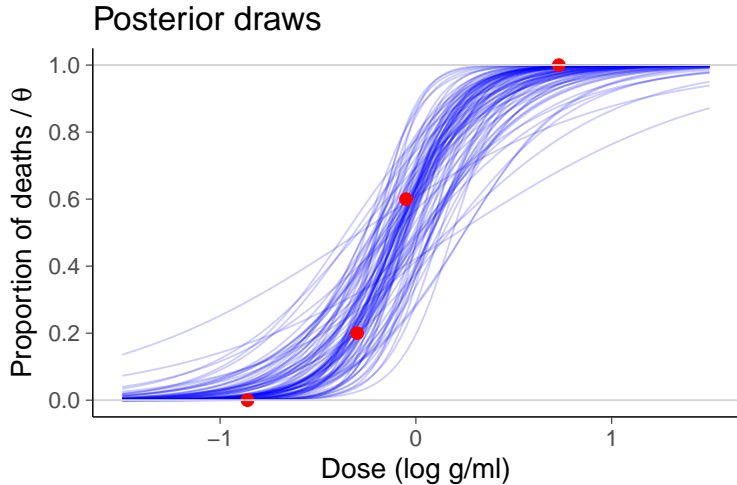
Logistic regression in latent space



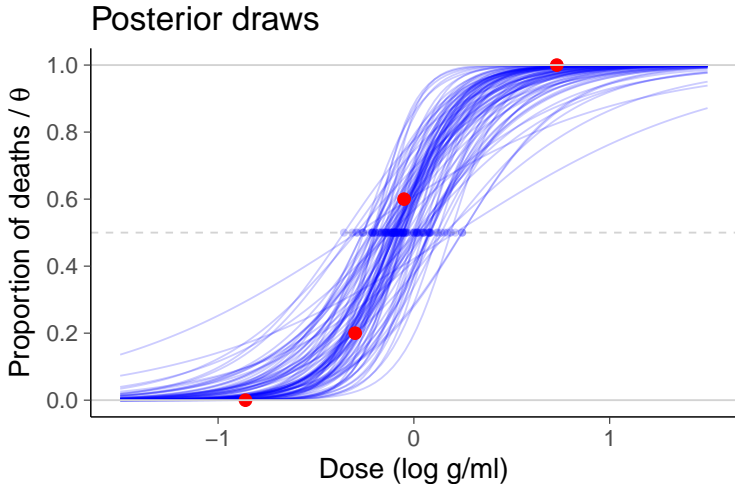
Bioassay



Bioassay

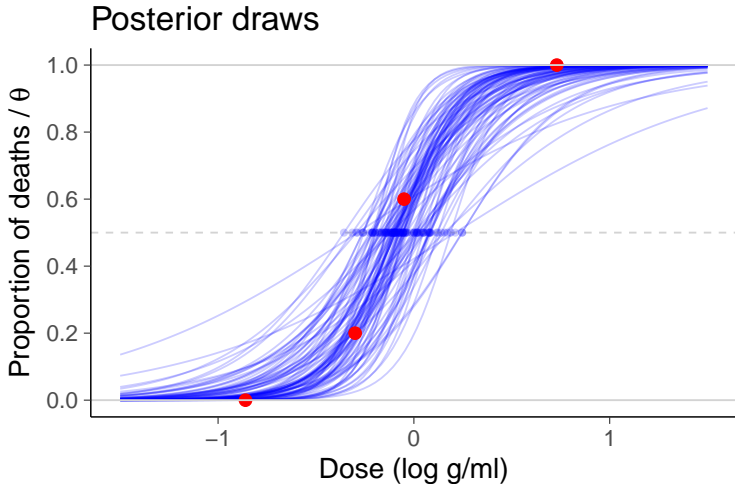


Bioassay



$$\text{LD50: } E\left(\frac{y}{n}\right) = \text{logit}^{-1}(\alpha + \beta x) = 0.5$$

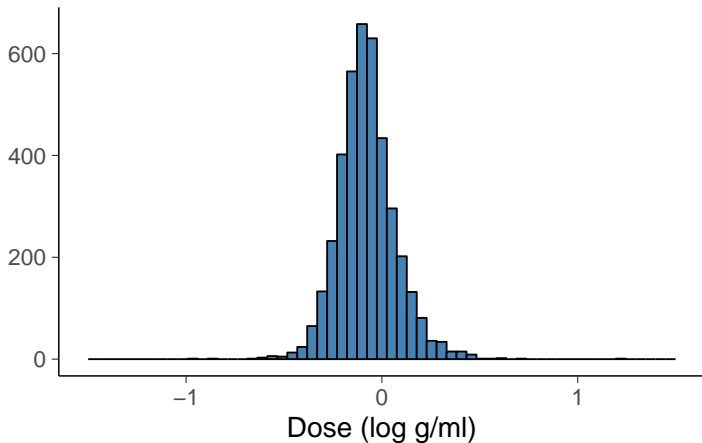
Bioassay



$$\text{LD50: } E\left(\frac{y}{n}\right) = \text{logit}^{-1}(\alpha + \beta x) = 0.5 \Rightarrow x_{\text{LD50}} = -\alpha/\beta$$

Bioassay

Bioassay LD50



$$\text{LD50: } E\left(\frac{y}{n}\right) = \text{logit}^{-1}(\alpha + \beta x) = 0.5 \quad \Rightarrow \quad x_{\text{LD50}} = -\alpha/\beta$$

Bioassay posterior

Binomial model

$$y_i \mid \theta_i \sim \text{Bin}(\theta_i, n_i)$$

Link function

$$\text{logit}(\theta_i) = \alpha + \beta x_i$$

Bioassay posterior

Binomial model

$$y_i \mid \theta_i \sim \text{Bin}(\theta_i, n_i)$$

Link function

$$\text{logit}(\theta_i) = \alpha + \beta x_i$$

Likelihood

$$p(y_i \mid \alpha, \beta, n_i, x_i) \propto \theta_i^{y_i} [1 - \theta_i]^{n_i - y_i}$$

Bioassay posterior

Binomial model

$$y_i \mid \theta_i \sim \text{Bin}(\theta_i, n_i)$$

Link function

$$\text{logit}(\theta_i) = \alpha + \beta x_i$$

Likelihood

$$p(y_i \mid \alpha, \beta, n_i, x_i) \propto \theta_i^{y_i} [1 - \theta_i]^{n_i - y_i}$$

$$p(y_i \mid \alpha, \beta, n_i, x_i) \propto [\text{logit}^{-1}(\alpha + \beta x_i)]^{y_i} [1 - \text{logit}^{-1}(\alpha + \beta x_i)]^{n_i - y_i}$$

Bioassay posterior

Binomial model

$$y_i \mid \theta_i \sim \text{Bin}(\theta_i, n_i)$$

Link function

$$\text{logit}(\theta_i) = \alpha + \beta x_i$$

Likelihood

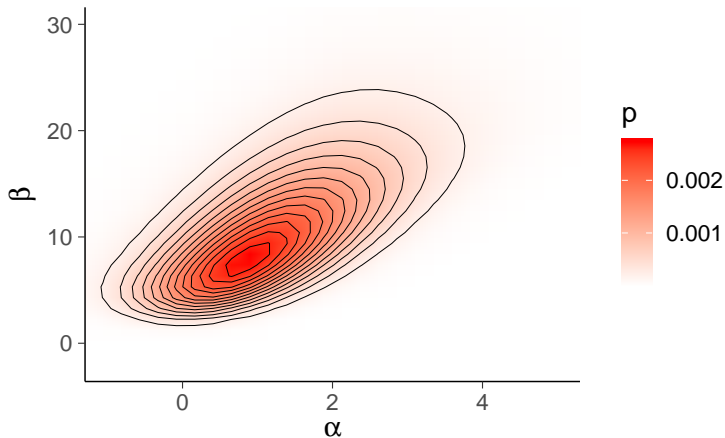
$$p(y_i \mid \alpha, \beta, n_i, x_i) \propto \theta_i^{y_i} [1 - \theta_i]^{n_i - y_i}$$

$$p(y_i \mid \alpha, \beta, n_i, x_i) \propto [\text{logit}^{-1}(\alpha + \beta x_i)]^{y_i} [1 - \text{logit}^{-1}(\alpha + \beta x_i)]^{n_i - y_i}$$

Posterior (with uniform prior on α, β)

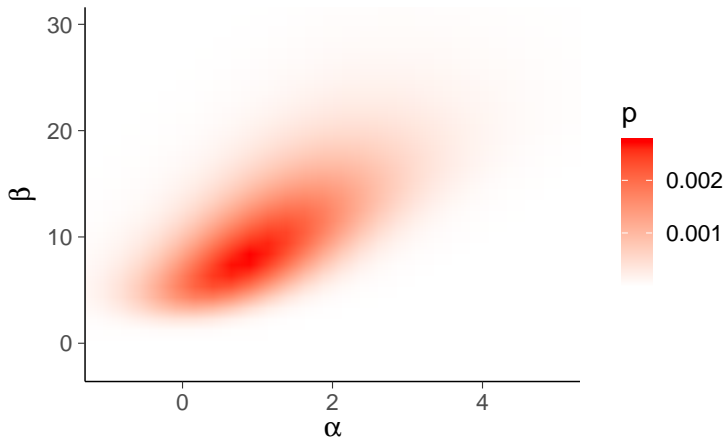
$$p(\alpha, \beta \mid y, n, x) \propto p(\alpha, \beta) \prod_{i=1}^n p(y_i \mid \alpha, \beta, n_i, x_i)$$

Posterior density evaluated in a grid



Bioassay

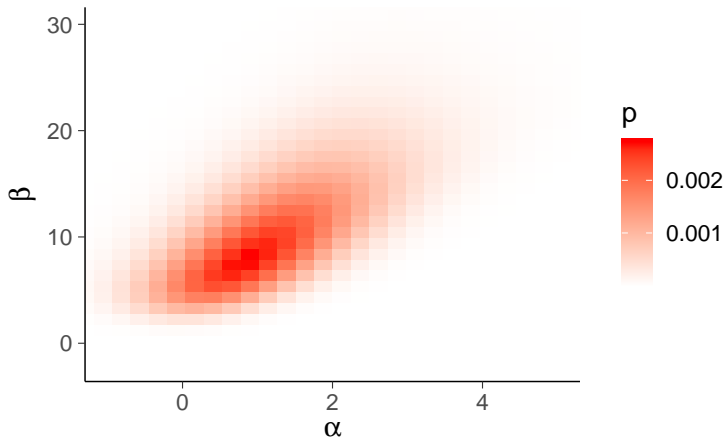
Posterior density evaluated in a grid



Density evaluated in grid, but plotted using interpolation

Bioassay

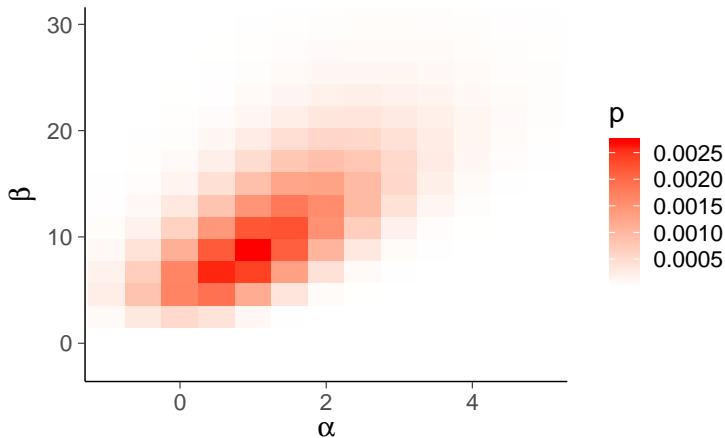
Posterior density evaluated in a grid



Density evaluated in grid, and plotted without interpolation

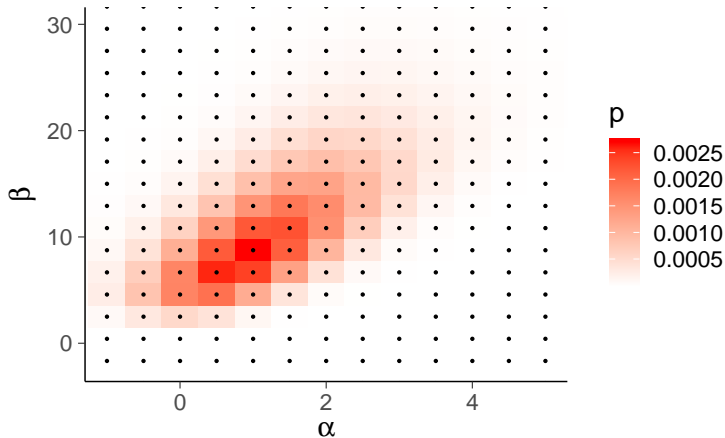
Bioassay

Posterior density evaluated in a grid



Density evaluated in a coarser grid

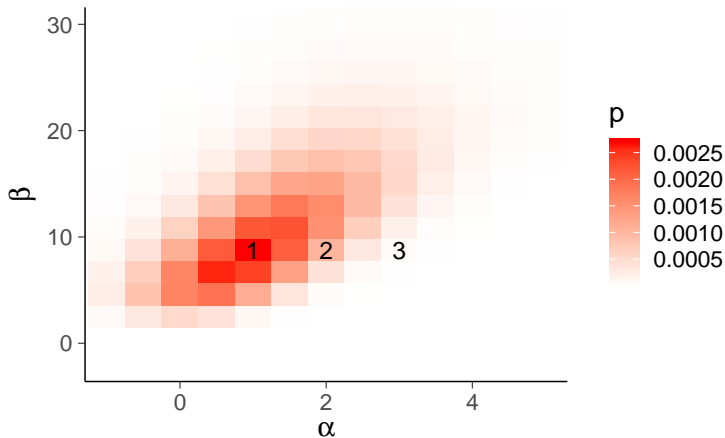
Posterior density evaluated in a grid



- Approximate the density as piecewise constant function
- Evaluate density in a grid over some finite region
- Density times cell area gives probability mass in each cell

Bioassay

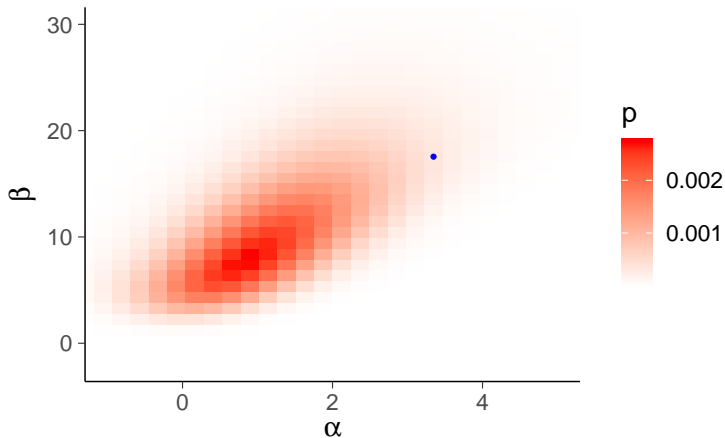
Posterior density evaluated in a grid



- Densities at 1, 2, and 3: 0.0027 0.0010 0.0001
- Probabilities of cells 1, 2, and 3: 0.0431 0.0166 0.0010
- Probabilities of cells sum to 1

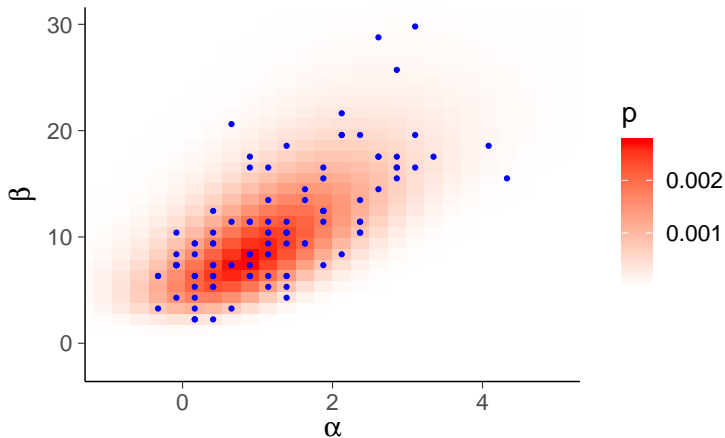
Bioassay

Posterior density and draws in a grid



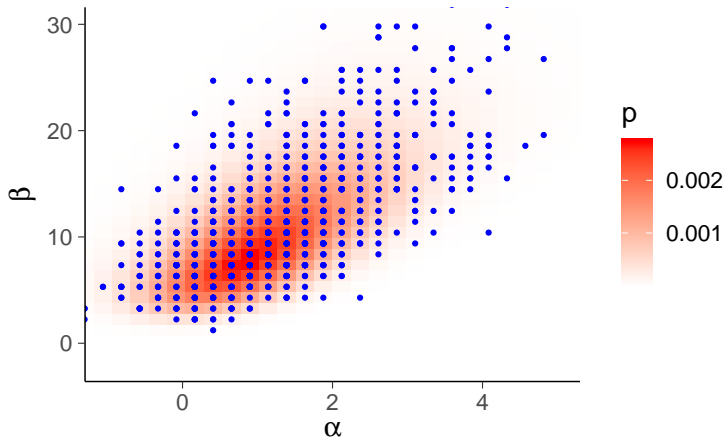
- Sample according to grid cell probabilities

Posterior density and draws in a grid



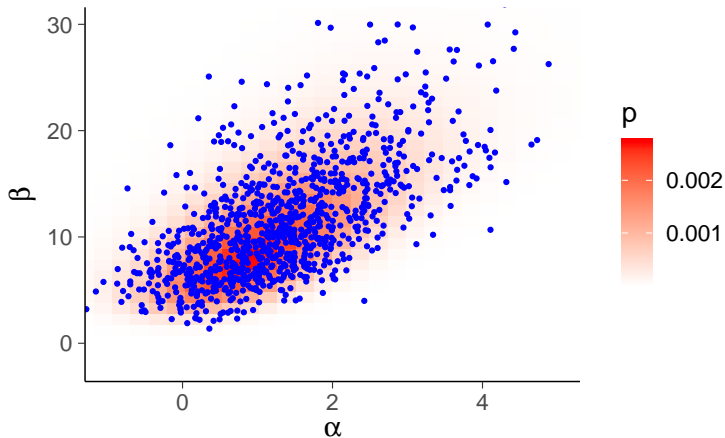
- Sample according to grid cell probabilities

Posterior density and draws in a grid



- Sample according to grid cell probabilities
- Several draws can be from the same grid cell

Posterior density in a grid and jittered draws



- Jitter can be added to improve visualization

Grid sampling

- Draws can be used to estimate expectations, for example

$$E[x_{\text{LD50}}] = E[-\alpha/\beta] \approx \frac{1}{S} \sum_{s=1}^S \frac{\alpha^{(s)}}{\beta^{(s)}}$$

Grid sampling

- Draws can be used to estimate expectations, for example

$$E[x_{LD50}] = E[-\alpha/\beta] \approx \frac{1}{S} \sum_{s=1}^S \frac{\alpha^{(s)}}{\beta^{(s)}}$$

- Instead of sampling, grid could be used to evaluate functions directly, for example

$$E[-\alpha/\beta] \approx \sum_{t=1}^T w_{\text{cell}}^{(t)} \frac{\alpha^{(t)}}{\beta^{(t)}},$$

where $w_{\text{cell}}^{(t)}$ is the normalized probability of a grid cell t , and $\alpha^{(t)}$ and $\beta^{(t)}$ are center locations of grid cells

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- Grid sampling gets computationally too expensive in high dimensions