

Advanced Section #2

Model Selection & Information Criteria

Akaike Information Criterion

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CS109A Introduction to Data Science
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Outline

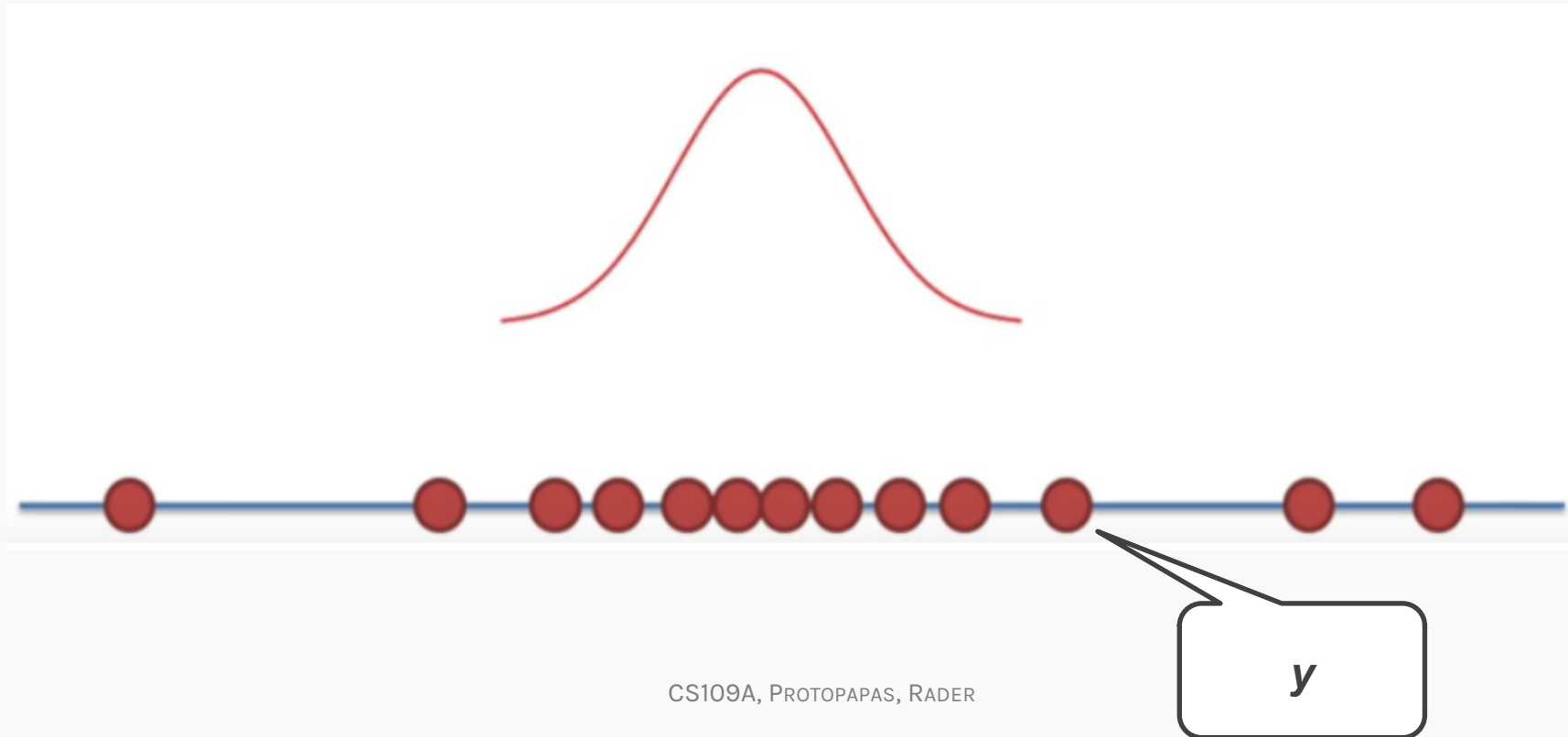
- Maximum Likelihood Estimation (MLE). Fit a distribution
 - Exponential distribution
 - Normal (Linear Regression Model)
- Model Selection & Information Criteria
 - KL divergence
 - MLE justification through KL divergence
 - Model Comparison
 - Akaike Information Criterion (AIC)

Maximum Likelihood Estimation (MLE) & Parametric Models

Maximum Likelihood Estimation (MLE)

Fit your data with a parametric distribution $q(\mathbf{y}|\boldsymbol{\theta})$.

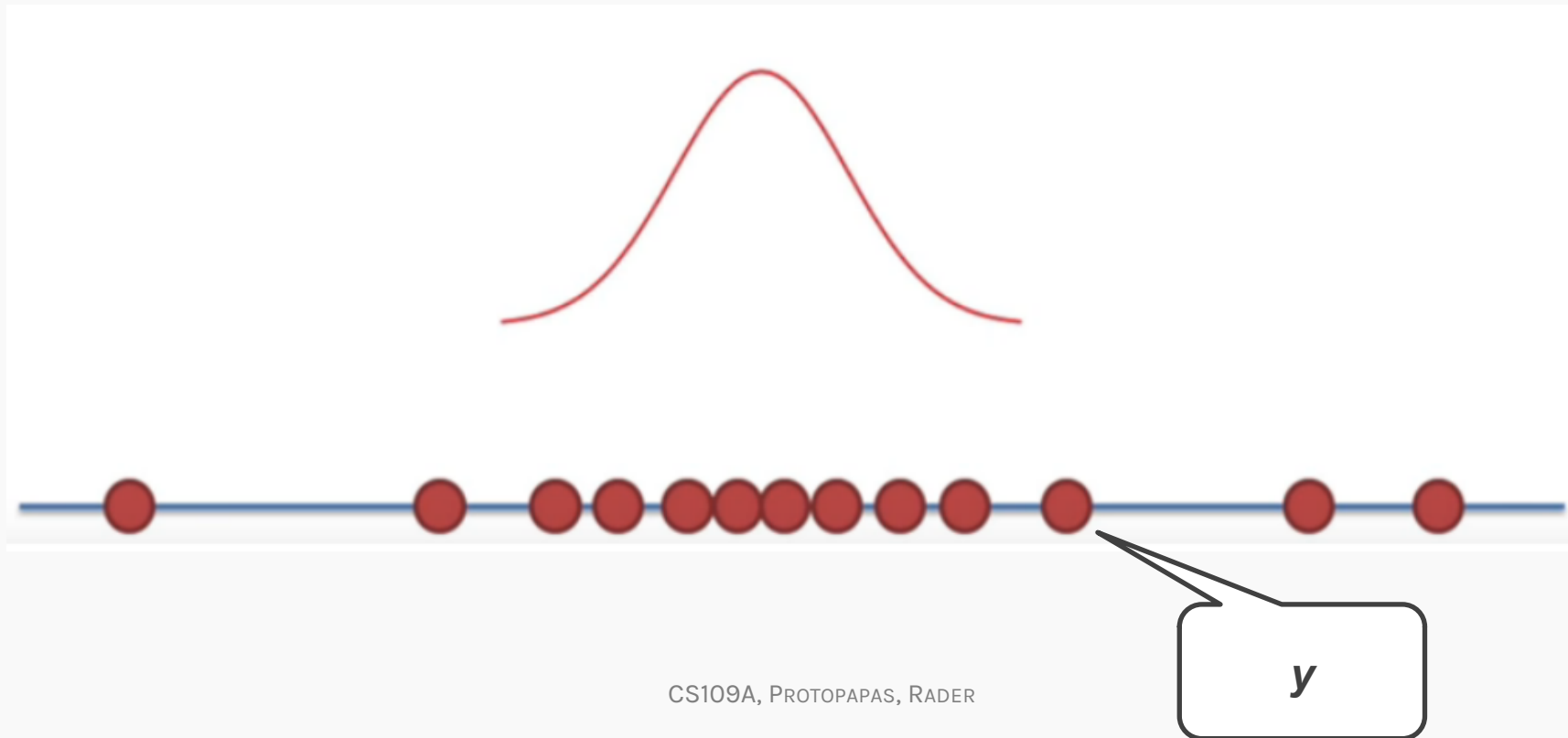
$\boldsymbol{\theta}=(\theta_1, \dots, \theta_k)$ is a parameter set to be estimated.



Maximum Likelihood Estimation (MLE)

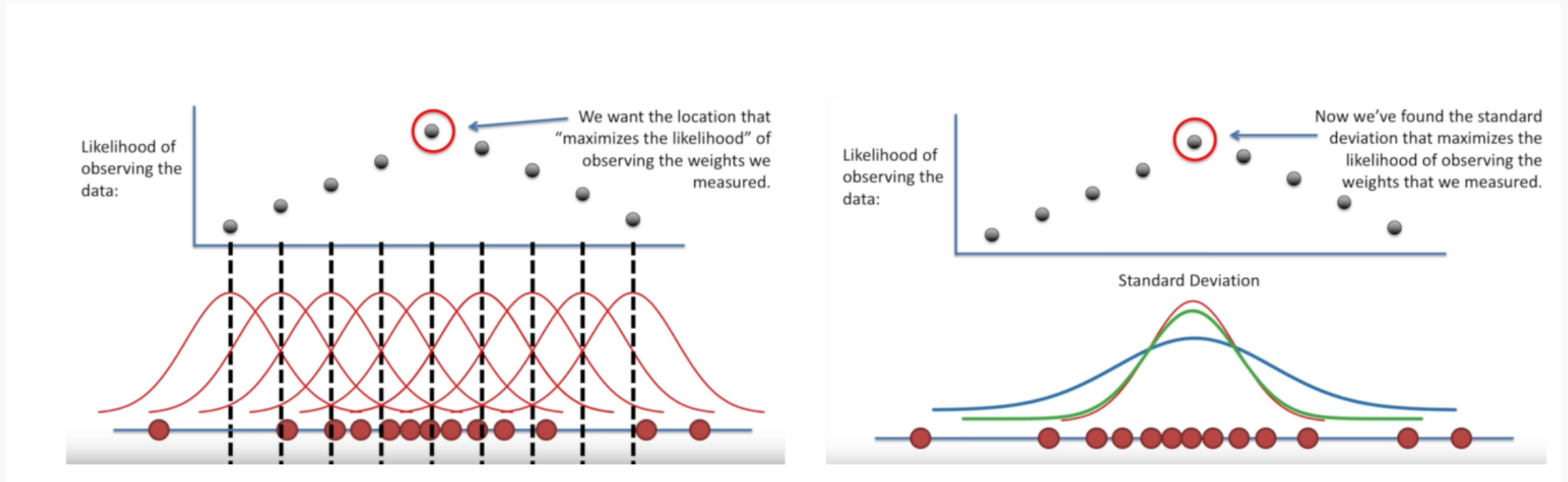
Fit your data with a parametric distribution $q(\mathbf{y}|\boldsymbol{\theta})$.

$\boldsymbol{\theta}=(\theta_1, \dots, \theta_k)$ is a parameter set to be estimated.



Maximize the Likelihood L

Scanning over all the parameters until find the maximum L



...but this is a too time-consuming approach.

Maximum Likelihood Estimation (MLE)

A formal and efficient method is given by MLE

Observations: $\mathbf{y}=(y_1, \dots, y_n)$

$$L(\boldsymbol{\theta}) = \prod_{i=1}^N q(y_i|\boldsymbol{\theta}),$$

Easier and numerically more stable to work with log-likelihood

$$\ell(\boldsymbol{\theta}) = \sum_{i=1}^n \log (q(y_i|\boldsymbol{\theta}))$$

Maximum Likelihood Estimation (MLE)

Easier and numerically more stable to work with log-likelihood

$$\frac{\partial \ell}{\partial \boldsymbol{\theta}} = \frac{\partial}{\partial \boldsymbol{\theta}} \log L = \frac{1}{L} \frac{\partial L}{\partial \boldsymbol{\theta}} \quad \Rightarrow \quad \left. \frac{\partial}{\partial \boldsymbol{\theta}} L(\boldsymbol{\theta}) \right|_{\boldsymbol{\theta}=\boldsymbol{\theta}_{\text{MLE}}} = \left. \frac{\partial}{\partial \boldsymbol{\theta}} \ell(\boldsymbol{\theta}) \right|_{\boldsymbol{\theta}=\boldsymbol{\theta}_{\text{MLE}}} = 0$$

Exponential distribution: A simple and useful example

A one parameter distribution: *rate parameter* λ

$$f(y_i|\lambda) = \begin{cases} \lambda e^{-\lambda y_i} & y_i \geq 0 \\ 0 & y_i < 0 \end{cases}$$

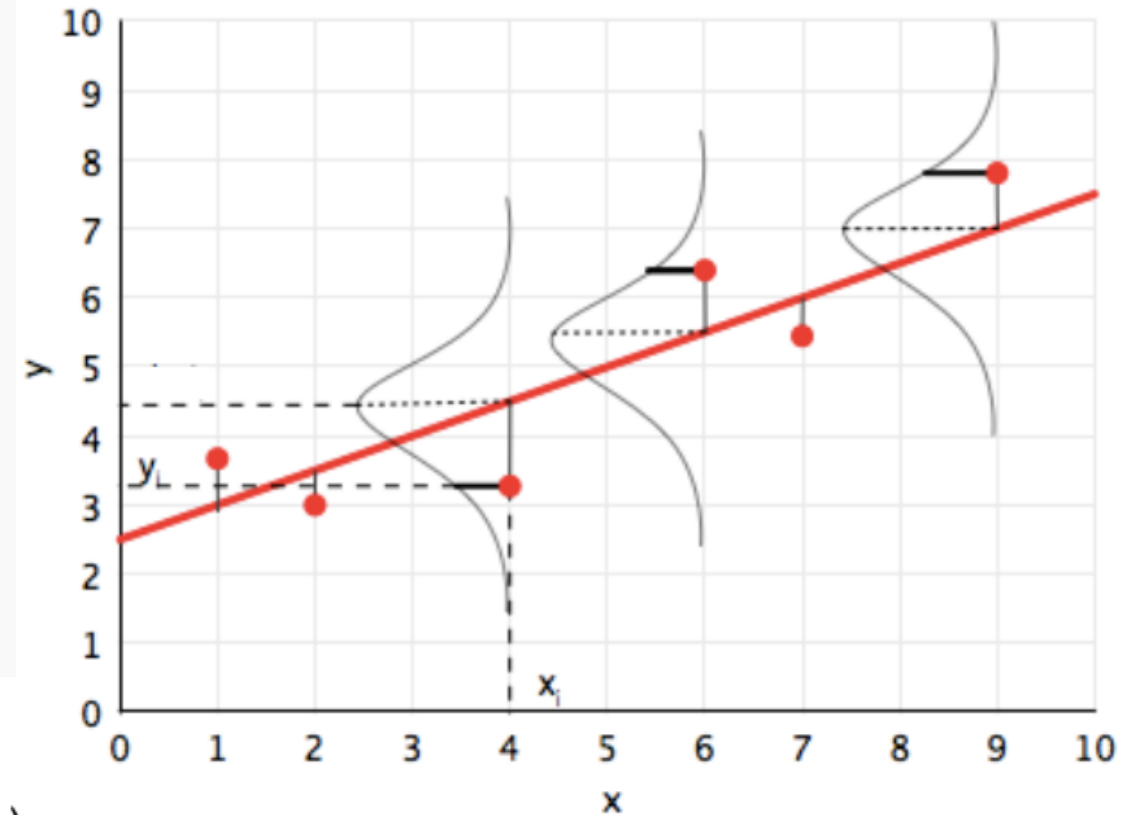
$$\ell(\lambda) = \sum_{i=1}^n \log(\lambda e^{-\lambda y_i}) = \sum_{i=1}^n (\log(\lambda) - \lambda y_i)$$

$$\lambda_{\text{MLE}} = \left(\frac{1}{n} \sum_{i=1}^n y_i \right)^{-1}$$

Linear Regression Model with gaussian error

$$\begin{aligned}y_i &= \sum_{j=0}^k x_{ij}\beta_j + \epsilon_i \\&= \mathbf{x}_i \cdot \boldsymbol{\beta} + \epsilon_i \\&= \mathbf{x}_i^T \boldsymbol{\beta} + \epsilon_i\end{aligned}$$

$$\begin{aligned}y_i &= q(y_i|\mu_i, \sigma^2) = \mathcal{N}(\mu_i, \sigma^2) = \mathcal{N}(\mathbf{x}_i^T \boldsymbol{\beta}, \sigma^2) \\&= \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y_i - \mathbf{x}_i^T \boldsymbol{\beta})^2}{2\sigma^2}\right)\end{aligned}$$



Linear Regression Model through MLE

$$L(\beta, \sigma^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y_i - \mathbf{x}_i^T \beta)^2}{2\sigma^2}\right)$$

$$\begin{aligned}\ell(\beta, \sigma^2) &= \sum_{i=1}^n \log\left(\frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y_i - \mathbf{x}_i^T \beta)^2}{2\sigma^2}\right)\right) \\ &= -\sum_{i=1}^n \left(\frac{1}{2} \log(2\pi) + \frac{1}{2} \log(\sigma^2) + \frac{(y_i - \mathbf{x}_i^T \beta)^2}{2\sigma^2}\right) \\ &= -\frac{n}{2} \log(2\pi) - \frac{n}{2} \log(\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mathbf{x}_i^T \beta)^2\end{aligned}$$

Loss
Function

Linear Regression Model: Standard Formulas

Minimize the loss essentially maximize the likelihood,
and we get

$$\beta_{\text{MLE}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

$$\sigma_{\text{MLE}}^2 = \frac{1}{n} \sum_{i=1}^n \left(y_i - \mathbf{x}_i^T \beta_{\text{MLE}} \right)^2$$

\mathbf{X} is called *the design matrix*

$$\mathbf{X} = \begin{pmatrix} \mathbf{x}_1^T \\ \mathbf{x}_2^T \\ \vdots \\ \mathbf{x}_n^T \end{pmatrix} = \begin{pmatrix} 1 & \mathbf{x}_{11} & \cdots & \mathbf{x}_{1v} \\ 1 & \mathbf{x}_{21} & \cdots & \mathbf{x}_{2v} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & \mathbf{x}_{n1} & \cdots & \mathbf{x}_{nv} \end{pmatrix}$$

Model Selection & Information Theory: Akaike Information Criterion

Kullback-Leibler (KL) divergence (or relative entropy)

How good do we fit the data?

What additional uncertainty have we introduced?

p is the *real* distribution
 q is the model distribution

$$\begin{aligned}\mathcal{D}_{\text{KL}}(p \parallel q) &= \sum_{i=1}^n p(y_i) \log \left(\frac{p(y_i)}{q(y_i | \boldsymbol{\theta})} \right) \\ &= \int_{-\infty}^{\infty} p(\mathbf{y}) \log \left(\frac{p(\mathbf{y})}{q(\mathbf{y} | \boldsymbol{\theta})} \right) d\mathbf{y}\end{aligned}$$

$$\begin{aligned}\mathcal{D}_{\text{KL}}(p \parallel q) &= \mathbb{E}_p \left[\log \left(\frac{p(\mathbf{y})}{q(\mathbf{y} | \boldsymbol{\theta})} \right) \right] \\ &= \mathbb{E}_p [\log(p(\mathbf{y})) - \log(q(\mathbf{y} | \boldsymbol{\theta}))]\end{aligned}$$

KL divergence

The KL divergence shows the *distance* between two distributions, hence it is a non-negative quantity.

With Jensen's inequality for convex functions $f(\mathbf{y})$, $\mathbb{E}[f(\mathbf{y})] \geq f(\mathbb{E}[\mathbf{y}])$:

$$\begin{aligned}\mathcal{D}_{\text{KL}}(p \parallel q) &= \mathbb{E}_p \left[\log \left(\frac{p(\mathbf{y})}{q(\mathbf{y}|\boldsymbol{\theta})} \right) \right] \\ &= \mathbb{E}_p \left[-\log \left(\frac{q(\mathbf{y}|\boldsymbol{\theta})}{p(\mathbf{y})} \right) \right] \geq -\log \left(\mathbb{E}_p \left[\frac{q(\mathbf{y}|\boldsymbol{\theta})}{p(\mathbf{y})} \right] \right) = 0\end{aligned}$$

KL divergence is a non-symmetric quantity $\mathcal{D}_{\text{KL}}(p \parallel q) \neq \mathcal{D}_{\text{KL}}(q \parallel p)$

MLE justification through KL divergence

Empirical distribution

$$p(\mathbf{y}) \simeq \frac{1}{n} \sum_{i=1}^n \delta(\mathbf{y} - y_i),$$

Minimize KL divergence is the same with maximize likelihood (empirical distribution)

$$\begin{aligned} \mathcal{D}_{\text{KL}}(p \parallel q) &\simeq \int_{-\infty}^{\infty} p(\mathbf{y}) \log \left(\frac{p(\mathbf{y})}{q(\mathbf{y}|\boldsymbol{\theta})} \right) d\mathbf{y} \\ &= \frac{1}{n} \sum_{i=1}^n \int_{-\infty}^{\infty} \delta(\mathbf{y} - y_i) \log \left(\frac{p(\mathbf{y})}{q(\mathbf{y}|\boldsymbol{\theta})} \right) d\mathbf{y} = \frac{1}{n} \sum_{i=1}^n \log \left(\frac{p(y_i)}{q(y_i|\boldsymbol{\theta})} \right) \\ &= \frac{1}{n} \sum_{i=1}^n (\log p(y_i) - \log q(y_i|\boldsymbol{\theta})), \end{aligned} \quad \text{log-likelihood}$$

Model Comparison

Consider to model distributions $q(\mathbf{y}|\boldsymbol{\theta})$ and $r(\mathbf{y}|\boldsymbol{\theta})$

$$\begin{aligned}\mathcal{D}_{\text{KL}}(p \parallel q) - \mathcal{D}_{\text{KL}}(p \parallel r) &= \mathbb{E}_p [\log(p(\mathbf{y})) - \log(q(\mathbf{y}|\boldsymbol{\theta}))] - \mathbb{E}_p [\log(p(\mathbf{y})) - \log(r(\mathbf{y}|\boldsymbol{\theta}))] \\ &= \mathbb{E}_p [\log(r(\mathbf{y}|\boldsymbol{\theta})) - \log(q(\mathbf{y}|\boldsymbol{\theta}))] = \mathbb{E}_p \left[\log \left(\frac{r(\mathbf{y}|\boldsymbol{\theta})}{q(\mathbf{y}|\boldsymbol{\theta})} \right) \right]\end{aligned}$$

By using the empirical distribution:

$$\mathcal{D}_{\text{KL}}(p \parallel q) - \mathcal{D}_{\text{KL}}(p \parallel r) = \frac{1}{n} \log \left(\frac{L_r(\mathbf{y}|\boldsymbol{\theta})}{L_q(\mathbf{y}|\boldsymbol{\theta})} \right)$$

p is eliminated.

Akaike Information Criterion (AIC)

AIC is a trade off between the number of parameters k and the error that is introduced (overfitting).

AIC is an asymptotic approximation of the KL-divergence $\mathcal{D}_{\text{KL}}(p \parallel q)$

The data are being used twice: first for MLE and second for the KL-divergence estimation.

AIC estimates which is the optimal number of parameters k

Polynomial Regression Model Example

Suppose a polynomial regression model

$$y_i = \beta_0 + \sum_{j=1}^k \beta_j x_{ij},$$

Which is the optimal k ?

For k smaller than the optimal: Underfitting

For k larger than the optimal: Overfitting

Minimizing real and empirical KL-divergence

Suppose many models indicated by index j

Work with the j -th model which has k_j parameters

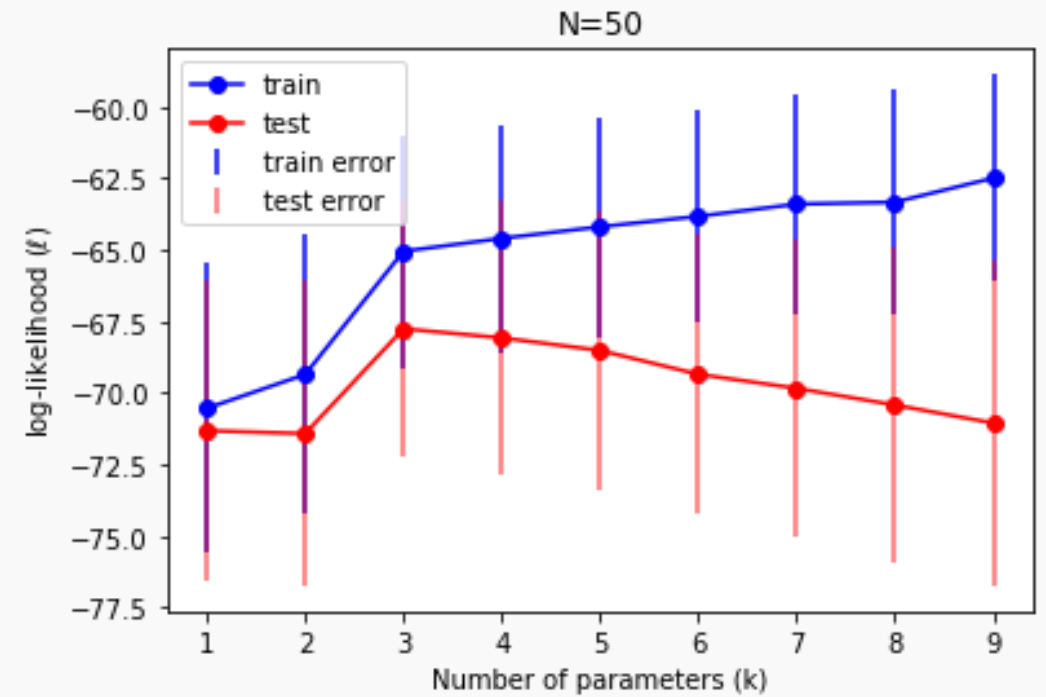
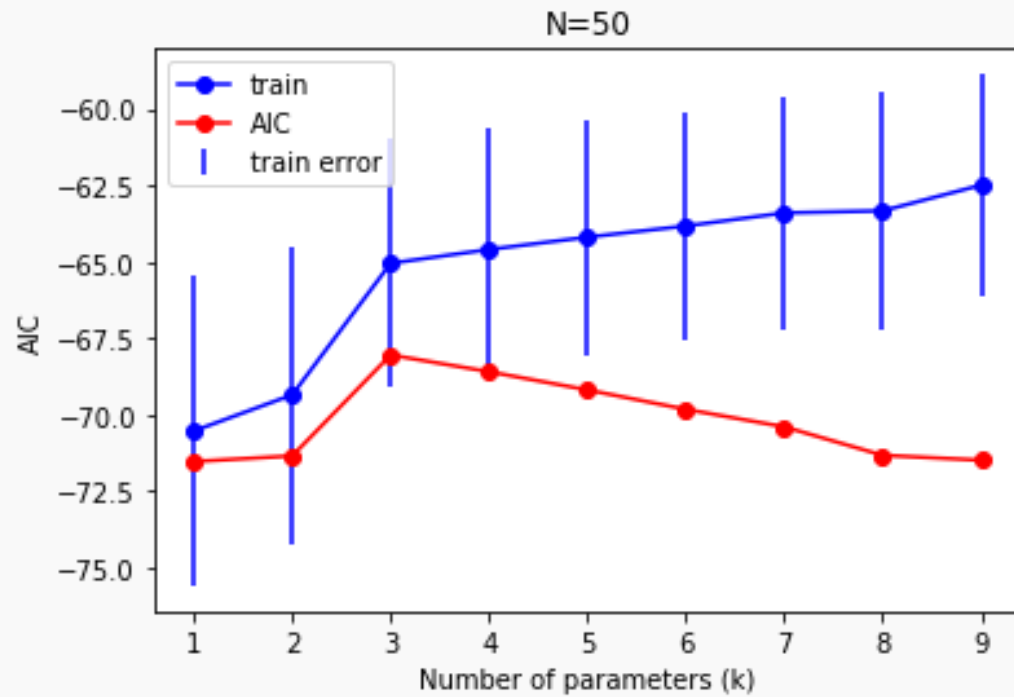
$$K_j = \int p(\mathbf{y}) \log q_j(\mathbf{y} | \boldsymbol{\theta}_{\text{MLE}}^{(j)}) d\mathbf{y}.$$

$$\bar{K}_j = \frac{1}{n} \sum_{i=1}^n \log q_j(y_i | \boldsymbol{\theta}_{\text{MLE}}^{(j)}) = \frac{\ell_j(\boldsymbol{\theta}_{\text{MLE}}^{(j)})}{n}$$

$$\begin{aligned} K_j &= \bar{K}_j - \frac{k_j}{n} \\ &= \frac{\ell_j(\boldsymbol{\theta}_{\text{MLE}}^{(j)})}{n} - \frac{k_j}{n}. \end{aligned}$$

$$\begin{aligned} \text{AIC}(j) &= 2nK_j \\ &= 2\ell_j(\boldsymbol{\theta}_{\text{MLE}}^{(j)}) - 2k_j. \end{aligned}$$

Numerical verification of AIC



Akaike Information Criterion (AIC): Proof

Asymptotic Expansion around true ideal MLE $\boldsymbol{\theta}_0$

$$\begin{aligned} K_j &\simeq \int p(\mathbf{y}) \left(\log q(\mathbf{y}|\boldsymbol{\theta}_0) + (\boldsymbol{\theta}_{\text{MLE}} - \boldsymbol{\theta}_0)^T s(\mathbf{y}|\boldsymbol{\theta}_0) + \frac{1}{2}(\boldsymbol{\theta}_{\text{MLE}} - \boldsymbol{\theta}_0)^T H(\mathbf{y}|\boldsymbol{\theta}_0)(\boldsymbol{\theta}_{\text{MLE}} - \boldsymbol{\theta}_0) \right) d\mathbf{y} \\ &= K_0 + \frac{1}{2n} Z^T J(\mathbf{y}|\boldsymbol{\theta}_0) Z, \end{aligned}$$

$$\begin{aligned} \bar{K}_j &\simeq \frac{1}{n} \sum_{i=1}^n \left(\log q(y_i|\boldsymbol{\theta}_0) + (\boldsymbol{\theta}_{\text{MLE}} - \boldsymbol{\theta}_0)^T s(y_i|\boldsymbol{\theta}_0) + \frac{1}{2}(\boldsymbol{\theta}_{\text{MLE}} - \boldsymbol{\theta}_0)^T H(y_i|\boldsymbol{\theta}_0)(\boldsymbol{\theta}_{\text{MLE}} - \boldsymbol{\theta}_0) \right) \\ &= K_0 + A_n + \frac{Z^T S_n}{\sqrt{n}} - \frac{1}{2n} Z^T J_n Z^T, \end{aligned}$$

Akaike Information Criterion (AIC): Proof

$$J(y|\boldsymbol{\theta}) = -\mathbb{E}_p [H(y|\boldsymbol{\theta})]$$

$$Z = \sqrt{n} (\boldsymbol{\theta}_{\text{MLE}} - \boldsymbol{\theta}_0) \quad (\text{with } Z_i \text{ given by } \mathcal{N}(0, V_Z)),$$

$$S_n = \frac{1}{n} \sum_{i=1}^n s(y_i|\boldsymbol{\theta}_0)$$

$$A_n = \frac{1}{n} \sum_{i=1}^n (\log q(y_i|\boldsymbol{\theta}_0) - K_0)$$

$$\begin{aligned} \bar{K} - K &\simeq A_n + \frac{\sqrt{n} Z^T S_n}{n} \\ &= A_n + \frac{Z^T J Z}{n}, \end{aligned}$$

$$\mathbb{E}_p [\bar{K} - K] = \mathbb{E}_p [A_n] + \mathbb{E}_p \left[\frac{Z^T J Z}{n} \right]$$

Akaike Information Criterion (AIC): Proof

$$\mathbb{E}_p [\bar{K} - K] = 0 + \text{trace} \left(\frac{J J^{-1} V J^{-1}}{n} \right) = \frac{1}{n} \text{trace} (J^{-1} V).$$

$$K \simeq \bar{K} - \frac{1}{n} \text{trace} (J^{-1} V).$$

In the limit of a correct model: $\boldsymbol{\theta}_{\text{MLE}} = \boldsymbol{\theta}_0$, and thus, $J^{-1} = V$.

$$K \simeq \bar{K} - \frac{k}{n}$$

Review

- Maximum Likelihood Estimation (MLE)
 1. A powerful method to estimate the ideal fitting parameters of a model.
 2. Exponential distribution, a simple but useful example.
 3. Linear Regression Model as a special paradigm of MLE implementation.
- Model Selection & Information Criteria
 1. KL-divergence quantifies the “distance” between the fitting model and the “real” distribution.
 2. KL-divergence justifies the MLE and is used for model comparison.
 3. AIC: Estimates the number of model parameters and protects from overfitting.

Thank you

Office hours are:

Monday 6-7:30 (Marios)

Tuesday 6:30-8 (Trevor)