

Mathematical Statistics - Assignment 6

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Problem 4.8

Suppose that Y has a density function

$$f(y) = \begin{cases} ky(1-y), & 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

(a)

Find some $k \in \mathbb{R}$ such that $f(y)$ is a probability density function. Notice that $\forall y \in [0, 1], \quad y(1-y) > 0$. This indicates that $k \geq 0$. The integral $\int_{-\infty}^0 f(y)dy + \int_1^{\infty} f(y)dy = 0$ since we know that the function is constantly zero on these open intervals. Then $P(-\infty, \infty) = P(0 \leq Y \leq 1) = \int_0^1 f(y)dy$. Thus we want to find some k such that $\int_0^1 ky(y-1)dy = 1$. Let us use rules of algebra and integration to derive the following

$$\begin{aligned} 1 &= \int_0^1 ky(1-y)dy \\ &= k \int_0^1 y - y^2 dy \\ &= k \left(\frac{y^2}{2} - \frac{y^3}{3} \right) \Big|_0^1 \\ &= \frac{k}{6} \end{aligned}$$

Since we must have $\frac{k}{6} = 1$, of course our answer will be $k = 6$.

(b)-(e)

Compute various probabilities using this value of k and the normalized density function $f(y)$. First we must find $P(.4 \leq Y \leq 1)$.

$$\begin{aligned} P(.4 \leq Y \leq 1) &= \int_{.4}^1 6y(1-y)dy \\ &= 6 \int_{.4}^1 y - y^2 dy \\ &= 6 \left[\frac{y^2}{2} - \frac{y^3}{3} \right]_{.4}^1 \\ &= 6 \left[\frac{1}{6} - \left(\frac{2}{25} - \frac{2^3}{(3)5^3} \right) \right] \\ &= .648 \end{aligned}$$

The next probability to compute is $P(.4 \leq Y < 1)$. However, this probability will be the same as the one we just computed because $P(.4 \leq Y \leq 1) = P(.4 \leq Y < 1) + P(Y = 1)$. Since $P(Y = 1) = \int_1^1 f(y)dy = 0$, both probabilities must be equal.

To compute $P(Y \leq .4 | Y \leq .8)$ by definition the probability equals

$$\frac{P(Y \leq .4 \text{ and } Y \leq .8)}{P(Y \leq .8)}$$

The probability in the numerator will simply be $P(Y \leq .4)$ since this event is a subset of $Y \leq .8$. For the denominator, the integral must be computed.

$$6 \int_0^{.8} f(y) dy = 6 \left[\frac{y^2}{2} - \frac{y^3}{3} \right]_0^{.8} = 6 \left[\frac{4^2}{(2)5^2} - \frac{4^3}{(3)5^3} \right] = .896$$

Knowing both the numerator and denominator we say $P(Y \leq .4 | Y \leq .8) = \frac{.352}{.896} = .393$.

Problem 4.14

We have the following probability density function

$$f(y) = \begin{cases} y, & 0 < y < 1 \\ 2 - y, & 1 \leq y < 2 \\ 0, & \text{otherwise} \end{cases}$$

(a)

We sketch the function $f(y)$.

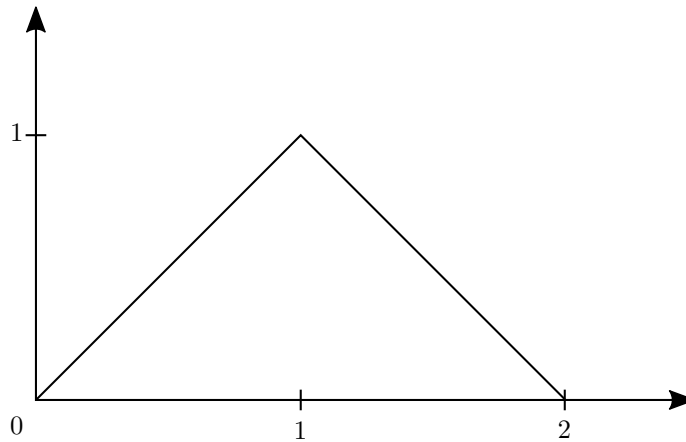


Figure 1: Probability Density Function $f(y)$

(b)

$$F(y) = \begin{cases} \int_{-\infty}^y 0 dt = 0, & y \leq 0 \\ \int_{-\infty}^0 0 dt + \int_0^y t dt = \frac{t^2}{2} \Big|_0^y = \frac{y^2}{2}, & 0 < y < 1 \\ \int_{-\infty}^0 0 dt + \int_0^y t dt + \int_1^y 2 - t dt = \frac{1}{2} + \left[2t - \frac{t^2}{2} \right]_1^y = \frac{1}{2} + \left[2y - \frac{y^2}{2} - \left(2 - \frac{1}{2} \right) \right] = -\frac{y^2}{2} + 2y - 1, & 1 \leq y \leq 2 \\ 1, & y > 2 \end{cases}$$

(c)

To find the probability that the station pumps between 8000 and 12000 gallons of gas, we say $P(.8 < Y < 1.2) = F(1.2) - F(.8) = -\frac{1.2^2}{2} + 2(1.2) - 1 - \left[-\frac{.8^2}{2} \right] = .36$

(d)

$P(Y > 1.5 | Y > 1) = \frac{P(Y > 1.5)}{P(Y > 1)}$. We must compute each of these in order to find the conditional probability.

$$P(Y > 1.5) = (.5) \int_{1.5}^{\infty} f(y) dy = (.5) \frac{(.5)^2}{2} = (.5) \frac{.25}{2} = (.5).125 = .0625$$

$$P(Y > 1) = (.5) \int_1^{\infty} f(y) dy = (.5) \frac{1}{2} = .25$$

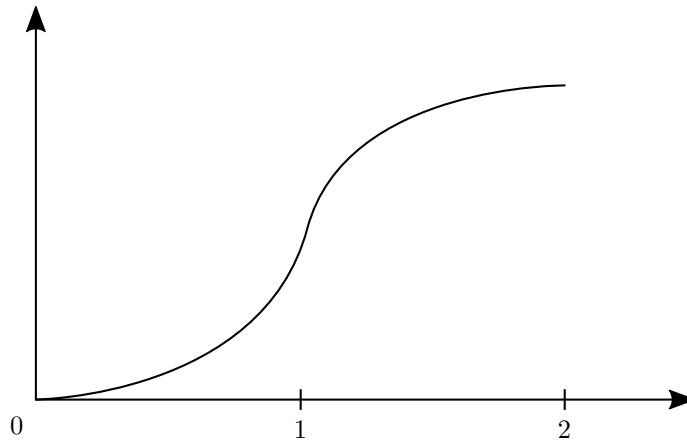


Figure 2: The Anti-Derivative $F(y)$

This is clear by the geometric argument where we are simply taking the area of square right triangles. Finally our conditional probability is going to be $\frac{.064}{.25} = .256$.

Problem 4.18

For this problem we have the density function

$$f(y) = \begin{cases} .2, & -1 < y \leq 0 \\ .2 + cy, & 0 < y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

(a)

Find the value of c such that f is a proper probability density function. We know that the integral up to 0 will be .2 (area of a rectangle), which means the integral from 0 onward must be .8. We say that $.2 + \frac{c}{2} = .8 \Leftrightarrow c = 1.2$

(b)

$$F(y) = \begin{cases} 0, & y \leq -1 \\ .2y|_{-1}^y = .2y + .2, & -1 < y \leq 0 \\ .2(0) + .2 + \left[.2y + \frac{1.2y^2}{2} \right]_0^y = .2 + .2y + \frac{1.2y^2}{2}, & 0 < y \leq 1 \\ 1, & y > 1 \end{cases}$$

(c)

Graph $f(y)$, $F(y)$. See Figure 3.

(d)

$$F(-1) = 0, F(0) = .2, F(1) = 1$$

(e)

$$P(0 \leq Y \leq .5) = F(.5) - F(0) = .2 + .2(.5) + \frac{1.2(.5)^2}{2} - [.2] = .25$$

(f)

$$P(Y > .5 | Y > .1) = \frac{P(Y > .5)}{P(Y > .1)} = \frac{1 - P(Y \leq .5)}{1 - P(Y \leq .1)} = \frac{1 - .45}{1 - .2 + .2(.1) + \frac{1.2(.1)^2}{2}} = .67.$$

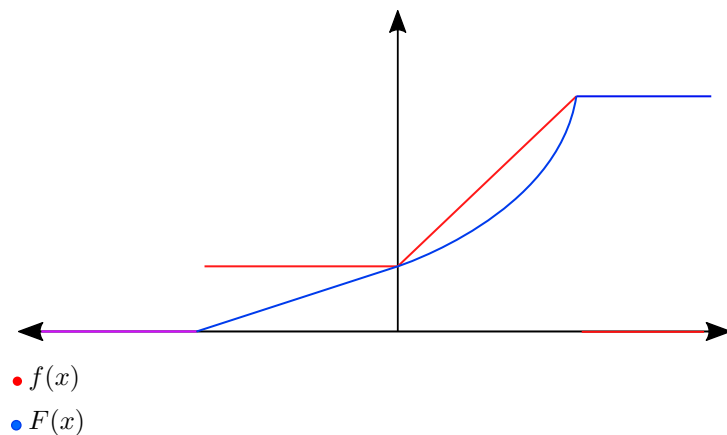


Figure 3: Graphs of $f(x)$ and $F(x)$

Problem 4.28

$$f(y) = \begin{cases} cy^2(1-y)^4, & 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

(a)

Find c such that $f(y)$ is a probability density function.

$$\int_0^1 cy^2(1-y)^4 dy = \frac{c}{105} \implies c = 105$$

(b)

Find $E(Y)$.

$$E(Y) = \int_{-\infty}^{\infty} yf(y)dy = 105 \int_0^1 y^3(1-y)^4 dy = \frac{3}{8} = .375$$

Problem 4.32

$$f(y) = \begin{cases} \frac{3}{64}y^2(4-y), & 0 \leq y \leq 4 \\ 0, & \text{otherwise} \end{cases}$$

(a)

Find $E(Y)$ and $V(Y)$.

$$\begin{aligned} E(Y) &= \int_{-\infty}^{\infty} yf(y)dy \\ &= \int_0^4 yf(y)dy \\ &= \int_0^4 y \frac{3y^2(4-y)}{64} dy \\ &= \int_0^4 \frac{3y^3(4-y)}{64} dy \\ &= \int_0^4 \frac{12y^3 - 3y^4}{64} dy \\ &= \left. \frac{12y^4}{4 \cdot 64} - \frac{3y^5}{5 \cdot 64} \right|_0^4 \\ &= \frac{12}{5} = 2.4 \end{aligned}$$

$$\begin{aligned} V(Y) &= E(Y^2) - E(Y)^2 \\ &= \int_0^4 \frac{12y^4 - 3y^5}{64} dy - 2.4^2 \\ &= .64 \end{aligned}$$

(b)

If Y is our weekly CPU time in hours, and the cost is \$200 dollars per hour, we can find the expected value and variance of CPU cost by multiplying $V(Y)$ and $E(Y)$ by 200. The expected weekly cost is $200 \cdot 2.4 = 480$ dollars. The variance will be 128 dollars.

(c)

We expect this to occur somewhat often. For this to occur we must use at least 3 hours of CPU. $P(Y > 3) = \int_3^4 \frac{3}{64} y^2(4-y) dy = \frac{3}{64} \int_3^4 (4y^2 - y^3) dy = \frac{67}{256} = .263$. This means our probability of this occurring will be more than $\frac{1}{4}$, which depending on how often is defined is a fairly reasonable chance.

Problem 4.40

Here we have a continuous uniform distribution, which must have a constant function between the endpoints A, B and must integrate to 1 so we say, $P(x_i \in (A, B) \text{ for some } i \in \{1, 2, 3\}) = 1 - P(x_i \notin (A, B) \forall i = 1, 2, 3) = 1 - \left(\frac{1}{2}\right)^3 = 1 - \frac{1}{8} = \frac{7}{8} = .875$

Problem 4.48

(a)

Here we have another continuous uniform distribution on the interval $[0, 500]$. It must be the case that $f(y)$ is constant and that $\int_0^{500} f(y) dy = 1$ so $f(y) = \frac{1}{500} \cdot \frac{1}{500} \int_{475}^{500} 1 dy = \frac{1}{500} [500 - 475] = \frac{25}{500} = .05$

(b)

We have the same probability except we are integrating over $(0, 25)$ but since the length is still 25, we get the same value. .05

(c)

$$P(Y > 250) = \int_{250}^{500} \frac{1}{500} dy = \frac{250}{500} = .5.$$

Problem 4.50

We have success from 12:00-1:00, failure from 1:00-3:00, success from 3:00-4:00, and failure from 4:00-5:00. Our total probability is 1 distributed evenly over the 5 hour period. Since $\frac{2}{5}$ of the possible time lies within the windows of success, we say the probability of success is $\frac{2}{5} = .4$.