# Lecture o5 Geometry of Least Squares

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#### Goals for today

- 1. Geometry of least squares
- 2. Projection matrix P and annihilator matrix M
- 3. Application to

## GEOMETRY OF LEAST SQUARES

Last time, we established that the least squares solution to the model:

$$y = X\beta + \epsilon$$

Yields the solution:

$$\widehat{\beta} = (X^t X)^{-1} X^t y$$

As long as the matrix  $X^tX$  is invertable.

Define the column space of the matrix X as:

$$\mathcal{R}(X) = \{\theta : \theta = X\beta, \beta \in \mathbb{R}^p\} \subset \mathbb{R}^n$$

This is the space spanned by the p columns on X sitting in n-dimensional space.

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Notice that the least squares problem can be re-written as:

$$\widehat{\theta} = \operatorname*{arg\,min}_{\theta} \left\{ ||y - \theta||_2^2, \quad \text{s.t} \quad \theta = X\beta \right\}$$

Where  $\widehat{\beta} = X\widehat{\theta}$ .

**Theorem 3.2 (p.g. 37, Rao & Toutenburg)** The minimum,  $\widehat{\theta}$  is attained when  $(y - \widehat{\theta}) \perp \mathcal{R}(X)$ . In other words,  $(y - \widehat{\theta})$  is

perpendicular to all vectors in  $\mathcal{R}$ .

*Proof.* Pick a  $\widehat{\theta}$  in  $\mathcal{R}$  such that  $(y - \widehat{\theta}) \perp \mathcal{R}(X)$ .

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$$||y - \theta||^2 = (y - \widehat{\theta} + \widehat{\theta} - \theta)^t (y - \widehat{\theta} + \widehat{\theta} - \theta)$$

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$$X^{i}(y-\theta)=0$$
. Then for all  $\theta \in \mathcal{R}$ :

$$||y - \theta||^2 = (y - \widehat{\theta} + \widehat{\theta} - \theta)^t (y - \widehat{\theta} + \widehat{\theta} - \theta)$$
  
=  $(y - \widehat{\theta})^t (y - \widehat{\theta}) + (\widehat{\theta} - \theta)^t (\widehat{\theta} - \theta) + 2(y - \widehat{\theta})^t (\widehat{\theta} - \theta)$ 

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$$X^{t}(y-\theta) = 0. \text{ Then for all } \theta \in \mathcal{R}:$$

$$||y-\theta||^{2} = (y-\widehat{\theta}+\widehat{\theta}-\theta)^{t}(y-\widehat{\theta}+\widehat{\theta}-\theta)$$

$$= (y-\widehat{\theta})^{t}(y-\widehat{\theta}) + (\widehat{\theta}-\theta)^{t}(\widehat{\theta}-\theta) + 2(y-\widehat{\theta})^{t}(\widehat{\theta}-\theta)$$

 $= (v - \widehat{\theta})^t (v - \widehat{\theta}) + (\widehat{\theta} - \theta)^t (\widehat{\theta} - \theta)$ 

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$$= ||y - \widehat{\theta}||_{2}^{2} + ||\widehat{\theta} - \theta||_{2}^{2}$$

$$\geq ||y - \widehat{\theta}||_{2}^{2}$$

So, if such a  $\widehat{\theta}$  exists it attains the minimum. To see that it does, write  $\widehat{\theta}=X\widehat{\beta}$ . Then:

$$X^{t}(y - \widehat{\theta}) = X^{t}(y - X\widehat{\beta})$$
$$= X^{t}y - X^{t}X\widehat{\beta}$$

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 $= X^t y - X^t y$ 

$$X^{t}(y - \widehat{\theta}) = X^{t}(y - X\widehat{\beta})$$
$$= X^{t}y - X^{t}X\widehat{\beta}$$

And therefore our proposed  $\widehat{\theta} \in \mathcal{R}(X)$ .

From this geometric interpretation of the least squares estimator, we introduce an important matrix  $P_X$  called the *projection matrix*.

$$P_X = X(X^t X)^{-1} X^t$$

I'll often drop the subscript as it should be understood that the projection is on the data matrix X.

Notice that PX = X:

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$$Py = X(X^{t}X)^{-1}X^{t}Xy$$
$$= X\widehat{\beta}$$
$$= \widehat{\theta}$$
$$= \widehat{y}$$

Do you see why the projection matrix is called the projection matrix?

Notice that PX = X:

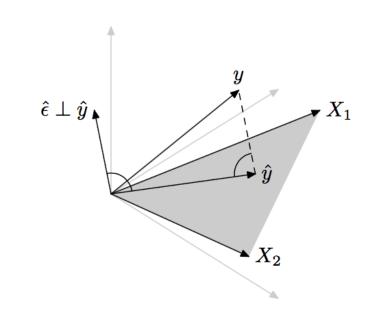
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Do you see why the projection matrix is called the projection matrix?

The projection matrix is sometimes called the <i>hat matrix</i> . Any thoughts as to why?



#### A closely related matrix to Die the annihilator matrix M.

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 $M = I_n - P$ 

A closely related matrix to *P* is the *annihilator matrix M*:

$$M = I_n - P$$

It gets its name because MX = 0.

The matrix  $P = X(X^tX)^{-1}X^t$  is clearly symettric. It is also

idempotent: 
$$P^2 = X(X^tX)^{-1}X^tX(X^tX)^{-1}X^t$$

 $= X(X^tX)^{-1}X^t$ 

= P

 $= X(X^{t}X)^{-1}(X^{t}X)(X^{t}X)^{-1}X^{t}$ 

$$M^{t} = (I_{n} - P)^{t}$$
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$$M^{t} = (I_{n} - P)^{t}$$
$$= (I_{n} - P^{t})$$
$$= M$$

$$M^2 = (I_n - P)^2$$
  
=  $(I_n - P)(I_n - P)$   
=  $I_n - 2 * P + P^2$ 

$$M^{t} = (I_{n} - P)^{t}$$
$$= (I_{n} - P^{t})$$
$$= M$$

$$M^{2} = (I_{n} - P)^{2}$$

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And idempotent:

$$M^{2} = (I_{n} - P)^{2}$$

$$= (I_{n} - P)(I_{n} - P)$$

$$= I_{n} - 2 * P + P^{2}$$

$$= I_{n} - 2 * P + P$$

$$= I_{n} - P$$

$$= M$$

These properties both make sense given the geometric interpretation of P and M as projections; into the column space of X and the compliment of the columns space of X.

$$r = v - X\widehat{\beta}$$

$$r = y - X\widehat{\beta}$$
$$= y - Py$$

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$$= y - Py$$
$$= (I_n - P)y$$

$$r = y - X\widehat{\beta}$$

$$= y - Py$$

$$= (I_n - P)y$$

$$= My$$

$$r = y - X\widehat{\beta}$$

$$= y - Py$$

$$= (I_n - P)y$$

$$= My$$

$$= M(X\beta - \epsilon)$$

$$r = y - X\widehat{\beta}$$

$$= y - Py$$

$$= (I_n - P)y$$

$$= My$$

$$= M(X\beta - \epsilon)$$

$$= M\epsilon$$

$$r = y - X\widehat{\beta}$$

$$= y - Py$$

$$= (I_n - P)y$$

$$= My$$

$$= M(X\beta - \epsilon)$$

$$= M\epsilon$$

The matricies P and M not only help make the derivation easier, they also give geometric insight into what we are doing.

One particularly useful formula will be writing the squared residuals as:

$$||r||_2^2 = ||M\epsilon||_2^2$$
$$= \epsilon^t M^t M \epsilon$$
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So the matrix M translates the sum of squared residuals into the sum of the square errors, which are estimated by the residuals.

### APPLICATIONS