Lecture o5 Geometry of Least Squares

16 September 2015

Taylor B. Arnold Yale Statistics STAT 312/612



Goals for today

- 1. Geometry of least squares
- 2. Projection matrix P and annihilator matrix M
- 3. Multivariate Galton Heights

GEOMETRY OF LEAST SQUARES

Last time, we established that the least squares solution to the model:

$$y = X\beta + \epsilon$$

Yields the solution:

$$\widehat{\beta} = (X^t X)^{-1} X^t y$$

As long as the matrix X^tX is invertable.

Define the column space of the matrix X as:

$$\mathcal{R}(X) = \{\theta : \theta = Xb, b \in \mathbb{R}^p\} \subset \mathbb{R}^n$$

This is the space spanned by the p columns of X sitting in n-dimensional space.

Define the column space of the matrix X as:

$$\mathcal{R}(X) = \{\theta : \theta = Xb, b \in \mathbb{R}^p\} \subset \mathbb{R}^n$$

This is the space spanned by the p columns of X sitting in n-dimensional space.

Notice that the least squares problem can be re-written as:

$$\widehat{\theta} = \operatorname*{arg\,min}_{a} \left\{ ||y - \theta||_{2}^{2}, \quad \text{s.t} \quad \theta \in \mathcal{R}(X) \right\}$$

Where then $\widehat{\beta} = X\widehat{\theta}$.

Theorem 3.2 (p.g. 37, Rao & Toutenburg) The minimum, $\widehat{\theta}$ is attained when $(y - \widehat{\theta}) \perp \mathcal{R}(X)$. In other words, $(y - \widehat{\theta})$ is

perpendicular to all vectors in \mathcal{R} .

Proof. Pick a $\widehat{\theta}$ in \mathcal{R} such that $(y - \widehat{\theta}) \perp \mathcal{R}(X)$.

Proof. Pick a $\widehat{\theta}$ in \mathcal{R} such that $(y - \widehat{\theta}) \perp \mathcal{R}(X)$. This implies that $X^{t}(y - \widehat{\theta}) = 0$. Then for all $\theta \in \mathcal{R}$:

$$||y - \theta||^2 = (y - \widehat{\theta} + \widehat{\theta} - \theta)^t (y - \widehat{\theta} + \widehat{\theta} - \theta)$$

$$|y - \theta||^2 = (y - \theta + \theta - \theta)^t (y - \theta + \theta - \theta)^t$$

Proof: Pick a $\widehat{\theta}$ in \mathcal{R} such that $(y - \widehat{\theta}) \perp \mathcal{R}(X)$. This implies that $X^t(y - \widehat{\theta}) = 0$. Then for all $\theta \in \mathcal{R}$:

$$X^{i}(y-\theta)=0$$
. Then for all $\theta \in \mathcal{R}$:

$$||y - \theta||^2 = (y - \widehat{\theta} + \widehat{\theta} - \theta)^t (y - \widehat{\theta} + \widehat{\theta} - \theta)$$

= $(y - \widehat{\theta})^t (y - \widehat{\theta}) + (\widehat{\theta} - \theta)^t (\widehat{\theta} - \theta) + 2(y - \widehat{\theta})^t (\widehat{\theta} - \theta)$

Proof. Pick a $\widehat{\theta}$ in \mathcal{R} such that $(y - \widehat{\theta}) \perp \mathcal{R}(X)$. This implies that $X^{t}(y - \widehat{\theta}) = 0$. Then for all $\theta \in \mathcal{R}$:

$$X^{t}(y-\theta) = 0$$
. Then for all $\theta \in \mathcal{R}$:
$$||y-\theta||^{2} = (y-\widehat{\theta}+\widehat{\theta}-\theta)^{t}(y-\widehat{\theta}+\widehat{\theta}-\theta)$$

 $= (v - \widehat{\theta})^t (v - \widehat{\theta}) + (\widehat{\theta} - \theta)^t (\widehat{\theta} - \theta)$

$$X^{t}(y - \theta) = 0$$
. Then for all $\theta \in \mathcal{R}$:
$$||y - \theta||^{2} = (y - \widehat{\theta} + \widehat{\theta} - \theta)^{t}(y - \widehat{\theta} + \widehat{\theta} - \theta)$$

 $= (\mathbf{v} - \widehat{\theta})^t (\mathbf{v} - \widehat{\theta}) + (\widehat{\theta} - \theta)^t (\widehat{\theta} - \theta) + 2(\mathbf{v} - \widehat{\theta})^t (\widehat{\theta} - \theta)$

Proof. Pick a $\widehat{\theta}$ in \mathcal{R} such that $(y - \widehat{\theta}) \perp \mathcal{R}(X)$. This implies that $X^{t}(y - \widehat{\theta}) = 0$. Then for all $\theta \in \mathcal{R}$:

$$X^{t}(y-\theta) = 0$$
. Then for all $\theta \in \mathcal{R}$:
 $||y-\theta||^{2} = (y-\widehat{\theta}+\widehat{\theta}-\theta)^{t}(y-\widehat{\theta}+\widehat{\theta}-\theta)$

 $= (v - \widehat{\theta})^t (v - \widehat{\theta}) + (\widehat{\theta} - \theta)^t (\widehat{\theta} - \theta)$

 $= ||\mathbf{v} - \widehat{\theta}||_2^2 + ||\widehat{\theta} - \theta||_2^2$

$$X^{t}(y - \theta) = 0$$
. Then for all $\theta \in \mathcal{R}$:
 $||y - \theta||^{2} = (y - \widehat{\theta} + \widehat{\theta} - \theta)^{t}(y - \widehat{\theta} + \widehat{\theta} - \theta)$

$$= (y - \widehat{\theta})^{t}(y - \widehat{\theta}) + (\widehat{\theta} - \theta)^{t}(\widehat{\theta} - \theta) + 2(y - \widehat{\theta})^{t}(\widehat{\theta} - \theta)$$

Proof. Pick a $\widehat{\theta}$ in \mathcal{R} such that $(y - \widehat{\theta}) \perp \mathcal{R}(X)$. This implies that $X^t(y - \widehat{\theta}) = 0$. Then for all $\theta \in \mathcal{R}$:

$$X^{t}(y - \widehat{\theta}) = 0$$
. Then for all $\theta \in \mathcal{R}$:
 $||y - \theta||^{2} = (y - \widehat{\theta} + \widehat{\theta} - \theta)^{t}(y - \widehat{\theta} + \widehat{\theta} - \theta)$

 $= (v - \widehat{\theta})^t (v - \widehat{\theta}) + (\widehat{\theta} - \theta)^t (\widehat{\theta} - \theta)$

 $= ||\mathbf{v} - \widehat{\theta}||_2^2 + ||\widehat{\theta} - \theta||_2^2$

 $> ||v - \widehat{\theta}||_2^2$

 $= (\mathbf{v} - \widehat{\theta})^t (\mathbf{v} - \widehat{\theta}) + (\widehat{\theta} - \theta)^t (\widehat{\theta} - \theta) + 2(\mathbf{v} - \widehat{\theta})^t (\widehat{\theta} - \theta)$

Proof: Pick a $\widehat{\theta}$ in \mathcal{R} such that $(y - \widehat{\theta}) \perp \mathcal{R}(X)$. This implies that $X^t(y - \widehat{\theta}) = 0$. Then for all $\theta \in \mathcal{R}$:

$$X^{t}(y-\theta) = 0. \text{ Then for all } \theta \in \mathcal{R}:$$

$$||y-\theta||^{2} = (y-\widehat{\theta}+\widehat{\theta}-\theta)^{t}(y-\widehat{\theta}+\widehat{\theta}-\theta)$$

$$= (y-\widehat{\theta})^{t}(y-\widehat{\theta}) + (\widehat{\theta}-\theta)^{t}(\widehat{\theta}-\theta) + 2(y-\widehat{\theta})^{t}(\widehat{\theta}-\theta)$$

 $= (v - \widehat{\theta})^t (v - \widehat{\theta}) + (\widehat{\theta} - \theta)^t (\widehat{\theta} - \theta)$

 $= ||\mathbf{v} - \widehat{\theta}||_2^2 + ||\widehat{\theta} - \theta||_2^2$

 $> || \mathbf{v} - \widehat{\theta} ||_2^2$

So, if such a $\widehat{\theta}$ exists it attains the minimum. To see that it does, write $\widehat{\theta}=X\widehat{\beta}$.

Proof: Pick a $\widehat{\theta}$ in \mathcal{R} such that $(y - \widehat{\theta}) \perp \mathcal{R}(X)$. This implies that $X^t(y - \widehat{\theta}) = 0$. Then for all $\theta \in \mathcal{R}$:

$$||y - \theta||^{2} = (y - \widehat{\theta} + \widehat{\theta} - \theta)^{t}(y - \widehat{\theta} + \widehat{\theta} - \theta)$$

$$= (y - \widehat{\theta})^{t}(y - \widehat{\theta}) + (\widehat{\theta} - \theta)^{t}(\widehat{\theta} - \theta) + 2(y - \widehat{\theta})^{t}(\widehat{\theta} - \theta)$$

$$= (y - \widehat{\theta})^{t}(y - \widehat{\theta}) + (\widehat{\theta} - \theta)^{t}(\widehat{\theta} - \theta)$$

$$= ||y - \widehat{\theta}||_{2}^{2} + ||\widehat{\theta} - \theta||_{2}^{2}$$

$$\geq ||y - \widehat{\theta}||_{2}^{2}$$

So, if such a $\widehat{\theta}$ exists it attains the minimum. To see that it does, write $\widehat{\theta}=X\widehat{\beta}$. Then:

$$X^{t}(y - \widehat{\theta}) = X^{t}(y - X\widehat{\beta})$$
$$= X^{t}y - X^{t}X\widehat{\beta}$$

To see that such a $\widehat{\theta}$ does exist, write $\widehat{\theta} = X \widehat{\beta}$.

To see that such a $\widehat{\theta}$ does exist, write $\widehat{\theta}=X\widehat{\beta}.$ Then:

$$X^{t}(y-\widehat{\theta}) = X^{t}(y-X\widehat{\beta})$$

To see that such a $\widehat{\theta}$ does exist, write $\widehat{\theta} = X\widehat{\beta}$. Then:

$$X^{t}(y - \widehat{\theta}) = X^{t}(y - X\widehat{\beta})$$
$$= X^{t}y - X^{t}X\widehat{\beta}$$

To see that such a $\widehat{\theta}$ does exist, write $\widehat{\theta} = X\widehat{\beta}$. Then:

$$X^t(y-\widehat{\theta}) = X^t(y-X\widehat{\beta})$$

 $= X^t y - X^t X \widehat{\beta}$

 $= X^t y - X^t X (X^t X)^{-1} X^t y$

To see that such a $\widehat{\theta}$ does exist, write $\widehat{\theta} = X\widehat{\beta}$. Then:

$$X^{t}(y-\widehat{ heta}) = X^{t}(y-X\widehat{eta})$$

$$X^t(y-\widehat{\theta}) = X^t(y-X\widehat{\beta})$$

 $= X^t y - X^t X (X^t X)^{-1} X^t y$

 $= X^t y - X^t y$

$$X^{t}(y - \widehat{\theta}) = X^{t}(y - X\widehat{\beta})$$
$$= X^{t}y - X^{t}X\widehat{\beta}$$

And therefore our proposed $\widehat{\theta} \in \mathcal{R}(X)$.

From this geometric interpretation of the least squares estimator, we introduce an important matrix P_X called the *projection matrix*.

$$P_X = X(X^t X)^{-1} X^t$$

I'll often drop the subscript as it should be understood that the projection is on the data matrix X.

Notice that PX = X:

$$PX = X(X^t X)^{-1} X^t X$$
$$= X$$

Notice that PX = X:

$$PX = X(X^t X)^{-1} X^t X$$
$$= X$$

And Py gives the fitted values \hat{y} :

$$Py = X(X^{t}X)^{-1}X^{t}Xy$$
$$= X\widehat{\beta}$$
$$= \widehat{\theta}$$
$$= \widehat{y}$$

Do you see why the projection matrix is called the projection matrix?

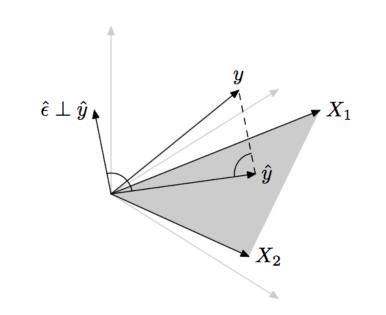
Notice that PX = X:

$$PX = X(X^t X)^{-1} X^t X$$
$$= X$$

And Py gives the fitted values \hat{y} :

$$Py = X(X^{t}X)^{-1}X^{t}Xy$$
$$= X\widehat{\beta}$$
$$= \widehat{\theta}$$
$$= \widehat{y}$$

Do you see why the projection matrix is called the projection matrix?



The projection matrix is sometimes called the <i>hat matrix</i> . Any thoughts as to why?

A closely related matrix to Die the annihilator matrix M.

A closely related matrix to P is the $annihilator\ matrix\ M$:

 $M = I_n - P$

A closely related matrix to *P* is the *annihilator matrix M*:

$$M = I_n - P$$

It gets its name because MX = 0.

$$P^2 = X(X^tX)^{-1}X^tX(X^tX)^{-1}X^t$$

$$P^{2} = X(X^{t}X)^{-1}X^{t}X(X^{t}X)^{-1}X^{t}$$

= $X(X^{t}X)^{-1}(X^{t}X)(X^{t}X)^{-1}X^{t}$

$$P^{2} = X(X^{t}X)^{-1}X^{t}X(X^{t}X)^{-1}X^{t}$$

$$= X(X^{t}X)^{-1}(X^{t}X)(X^{t}X)^{-1}X^{t}$$

$$= X(X^{t}X)^{-1}X^{t}$$

$$P^{2} = X(X^{t}X)^{-1}X^{t}X(X^{t}X)^{-1}X^{t}$$

$$= X(X^{t}X)^{-1}(X^{t}X)(X^{t}X)^{-1}X^{t}$$

$$= X(X^{t}X)^{-1}X^{t}$$

$$M^{t} = (I_{n} - P)^{t}$$
$$= (I_{n} - P^{t})$$
$$= M$$

$$M^{t} = (I_{n} - P)^{t}$$
$$= (I_{n} - P^{t})$$
$$= M$$

And idempotent:

$$M^{2} = (I_{n} - P)^{2}$$
$$= (I_{n} - P)(I_{n} - P)$$

$$M^{t} = (I_{n} - P)^{t}$$
$$= (I_{n} - P^{t})$$
$$= M$$

And idempotent:

$$M^2 = (I_n - P)^2$$

= $(I_n - P)(I_n - P)$
= $I_n - 2 * P + P^2$

$$M^{t} = (I_{n} - P)^{t}$$
$$= (I_{n} - P^{t})$$
$$= M$$

And idempotent:

$$M^2 = (I_n - P)^2$$

= $(I_n - P)(I_n - P)$
= $I_n - 2 * P + P^2$
= $I_n - 2 * P + P$

M is also symmetric

$$M^{t} = (I_{n} - P)^{t}$$
$$= (I_{n} - P^{t})$$
$$= M$$

And idempotent:

$$M^{2} = (I_{n} - P)^{2}$$

$$= (I_{n} - P)(I_{n} - P)$$

$$= I_{n} - 2 * P + P^{2}$$

$$= I_{n} - 2 * P + P$$

$$= I_{n} - P$$

$$= M$$

M is also symmetric

$$M^{t} = (I_{n} - P)^{t}$$
$$= (I_{n} - P^{t})$$
$$= M$$

And idempotent:

$$M^{2} = (I_{n} - P)^{2}$$

$$= (I_{n} - P)(I_{n} - P)$$

$$= I_{n} - 2 * P + P^{2}$$

$$= I_{n} - 2 * P + P$$

$$= I_{n} - P$$

$$= M$$

These properties both make sense given the geometric interpretation of P and M as projections; into the column space of X and the compliment of the columns space of X.

$$r = v - X\widehat{\beta}$$

$$r = y - X\widehat{\beta}$$
$$= y - Py$$

$$r = y - X\widehat{\beta}$$
$$= y - Py$$
$$= (I_n - P)y$$

$$r = y - X\widehat{\beta}$$

$$= y - Py$$

$$= (I_n - P)y$$

$$= My$$

$$r = y - X\widehat{\beta}$$

$$= y - Py$$

$$= (I_n - P)y$$

$$= My$$

$$= M(X\beta - \epsilon)$$

$$r = y - X\widehat{\beta}$$

$$= y - Py$$

$$= (I_n - P)y$$

$$= My$$

$$= M(X\beta - \epsilon)$$

$$= M\epsilon$$

$$r = y - X\widehat{\beta}$$

$$= y - Py$$

$$= (I_n - P)y$$

$$= My$$

$$= M(X\beta - \epsilon)$$

$$= M\epsilon$$

The matricies P and M not only help make the derivation easier, they also give geometric insight into what we are doing.

One particularly useful formula will be writing the squared residuals as:

$$||r||_2^2 = ||M\epsilon||_2^2$$
$$= \epsilon^t M^t M \epsilon$$
$$= \epsilon^t M \epsilon$$

One particularly useful formula will be writing the squared residuals as:

$$||r||_2^2 = ||M\epsilon||_2^2$$
$$= \epsilon^t M^t M \epsilon$$
$$= \epsilon^t M \epsilon$$

So the matrix M translates the sum of squared residuals into the sum of the square errors, which are estimated by the residuals.

APPLICATIONS