ECONOMETRICS 1 - HOMEWORK

PILI Pierre

December 8, 2023

1 The Basics

1.1 Create a table of descriptive statistics of the variables in the dataset.

var	median	mean	min	max	sd	NAs
agro_emp	18.6	25.1	0.1	86.3	22.3	30
bribery	11.7	17.0	0.0	67.1	14.7	87
gfce	16.5	17.7	5.1	62.9	8.4	36
literacy	93.0	83.6	24.2	100.0	19.3	61
\log_{gdp}	9.4	9.3	6.6	11.6	1.2	22
pop_total	$6.2\mathrm{e}{+06}$	$3.4\mathrm{e}{+07}$	$1.1\mathrm{e}{+04}$	$1.4\mathrm{e}{+09}$	$1.4\mathrm{e}{+08}$	2
$\operatorname{self} \operatorname{\underline{-emp}}$	35.0	40.9	0.4	94.8	27.0	30
stocks	6.4	28.8	0.0	538.7	66.8	131
$sample_size$	715.1	$3.6\mathrm{e}{+03}$	120.1	$1.4\mathrm{e}{+05}$	$1.3\mathrm{e}{+04}$	2

Table 1: Descriptive statistics

The data is made of 217 countries and certain variables contains a significant amount of missing values (NAs) (See 1). Bribery is the worst variable with up to 87 NAs. Using this variable in a regression would yield a low precision as the size of the sample is quite small. This table demonstrates how different countries are, indeed the min-max interval is rather close to one for rate variables and the standart deviations (sd) are very significant compared to the mean value of each variable.

1.2.a Is it the case that self-employment is correlated with how rich a country is (in-terms of log GDP per-capita)?

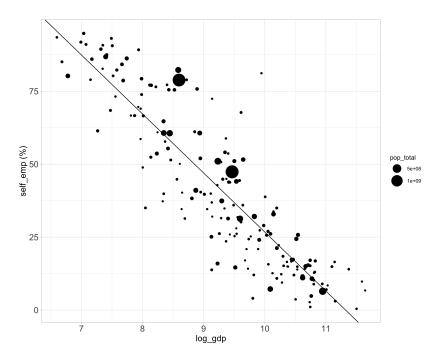


Figure 1: Self employment rate with respect to GDP

Figure 1 shows a clear negative relationship between the share of self employment and gdp. The empirical correlation coefficient is equal to -0.89 which is very close to -1. The anticorrelation is very strong.

1.2.b Is it the case that countries with higher share of employment in agriculture also have higher self-employment rates?

As in the previous question, Figure 2 shows a clear positive relationship between the share of self employment and share of employment in the agricultural sector. The empirical correlation coefficient is equal to 0.91 which is very close to 1. The correlation is very strong.

1.2.c Present a bar graph comparing the mean self-employment rates in each of these 3 literacy-based categories of countries.

Figure 3 seems to demonstrate a negative relationship between the literacy category of the population and the self employment rate.

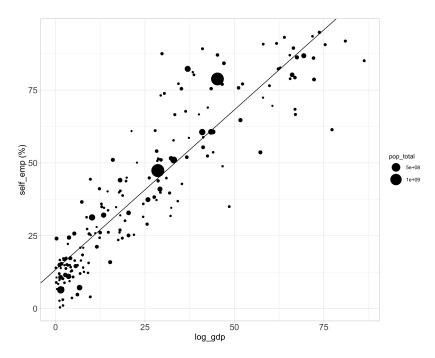


Figure 2: Self employment rate with respect to employment share in the agricultural sector

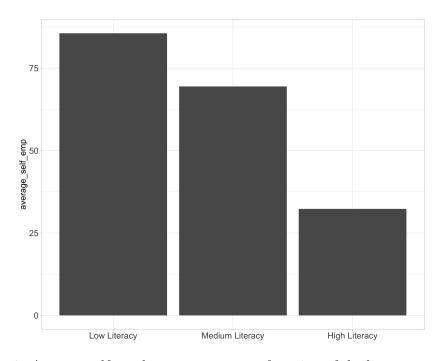


Figure 3: Average self employement rate as a function of the literacy category

1.3 Estimate the model parameters described using OLS, report your results and summarize them

Denoting this linear model as (1) (we shall refer to it as the "simple estimation" in the rest of this section) and estimating it yields the results displayed in Table 2 (first column). This OLS estimation was based on only 143 observations, which is far from 217. This raises a question, are NAs equally distributed across the sample? That being aknowledged, this regression seems to be very significant. The overall significance is very high as the F-stat demonstrates. Every of the three coefficients are significantly different from zero at the 1% level. The signs of the variables are in line with the previous discussion in question 1.1. According to this estimation, a 1% increase in GDP corresponds to a 6.5% decrease in the share of self employement. A 1% increase in the literacy rate is associated with a 0.3% decrease in the dependent variable, while a 0.59% increase in self employment shares can be explained by a 1% increase in the share of the agricultural sector.

Table 2: Simple and Extended Models - Exercise 1 $\,$

	Dependent	variable:		
	self_emp			
	(1)	(2)		
log_gdp	-6.506***	-5.520		
3_3	(1.755)	(4.042)		
literacy	-0.313***	-0.358**		
v	(0.070)	(0.175)		
agro emp	0.592***	0.628***		
<u> </u>	(0.080)	(0.176)		
gfce		-0.922**		
		(0.380)		
stocks		0.110*		
		(0.061)		
bribery		-0.111		
V		(0.156)		
Constant	113.219***	121.953***		
	(16.361)	(41.265)		
Observations	143	49		
\mathbb{R}^2	0.845	0.828		
Adjusted R^2	0.841	0.804		
Residual Std. Error	$10.574 \; (\mathrm{df} = 139)$	9.262 (df = 42)		
F Statistic	$252.152^{***} (df = 3; 139)$,		
Note:	*p<0.1: **p<0.05: ***p<0.0			

Note:

*p<0.1; **p<0.05; ***p<0.01

1.4 Describe how you could estimate β_3 using a 3-step procedure based on the Frisch-Waugh Theorem.

The Frisch-Waugh theorem can be used for estimating specific coefficients in a multiple regression model while controlling for other variables. Here are the simplified three steps:

Step 1:

Identify control variables, here \log_{gdp} and literacy, and regress it on the dependent variable to get the residuals \hat{r} . The residuals of this regression is the part of self_emp that is not explained by the \log_{gdp} and literacy variables.

self emp =
$$\alpha_0 + \alpha_1 \log \operatorname{gdp} + \alpha_2 \operatorname{literacy} + r$$

Step 2:

Regress the variable of interest on the control variables and get the residuals \hat{u} . The residuals of this regression is the part of agro_emp that is not explained by the log_gdp and literacy variables.

$$agro_emp = \gamma_0 + \gamma_1 log_gdp + \gamma_2 literacy + u$$

Step 3:

Regress \hat{r} on \hat{u} and get the coefficient of interest β_3 .

$$\hat{r} = \beta_0 + \beta_3 \hat{u} + \epsilon$$

The coefficient β_3 is the estimate of the parameter for agro_emp after controlling for log_gdp and literacy.

1.5 Implement this 3-step procedure and compare your estimate and standard error of β_3 .

Table 3 and Table 4 display the results of the 3-step procedure previously described, where \hat{r} , \hat{u} are respectively denoted as r_hat and u_hat. The first regression yields coefficients that are very different from the first estimation (first column of Table 2). This is an expected result, indeed as agro_emp is assumed to be part of the data generating process, the first step estimation is inconsistent because of endogeneity issues. The higher the correlation between agro_emp and control variable, the larger is the discrepency between coefficients from the first step of the Frisch-Waugh procedure and the simple estimation. The third regression yields an estimation for β_3 and its standard error which are very close to what we got in the simple estimation. Those results were expected and a direct implication of the Frisch-Waugh theorem.

Table 3: First and Second Regressions in 3-step Procedure - Exercise 1

self_emp (1)	agro_emp
(1)	
\ /	(2)
-15.784***	-15.670***
(1.453)	(1.308)
-0.377^{***}	-0.108
(0.081)	(0.073)
219.877***	180.127***
(9.247)	(8.326)
143	143
0.783	0.743
0.780	0.739
12.454	11.215
252.747***	201.952***
	(1.453) -0.377*** (0.081) 219.877*** (9.247) 143 0.783 0.780 12.454

Note:

*p<0.1; **p<0.05; ***p<0.01

Table 4: Third Regression in 3-step Procedure - Exercise 1

	Dependent variable:
	r_hat
u hat	0.592***
_	(0.079)
Constant	0.000
	(0.878)
Observations	143
R^2	0.284
Adjusted R ²	0.279
Residual Std. Error	$10.499 \; (\mathrm{df} = 141)$
F Statistic	$56.008^{***} (df = 1; 141)$
Note:	*p<0.1; **p<0.05; ***p<0.

1.6.a Estimate a linear model that includes the expanded set of covariates using OLS. Report your results and interpret them

Denoting this model as (2) (we shall refer to it as the "extended model"), the results are displayed in Table 2 along with the simple model estimation. The common coefficients among both model are in line with each other regarding values and errors. The bribery variable is not significantly different from zero at the 10%. The stocks variable is significantly different from zero at the 10% level which is not a very strong result. The gfce variable is significantly different from zero at the 5% level and means that a 1% increase in the value of stocks traded as a share of GDP is associated with a roughly 1% increase in the share of self employement. In other words, the larger the financial sector, the larger the share of self employement in the economy.

1.6.b How many observations was this model estimated on? Is it identical or different from the number of observations in question 3 and why?

The first thing to notice by comparing the simple and extended models is that the sample size is much smaller for the extended model as the individuals with at least one NA among the 6 variables were removed. The

immediate consequence is that the simple model is roughly twice as precise as the extended one.

1.6.c We are told that many of these variables, including the variable on self-employment rates, are measured using sample surveys in each country. This implies that the variance of each observation is directly proportional to the sample size used in each country. What does this imply for the OLS estimator?

The statement that the variance of each observation is directly proportional to the sample size implies heteroscedasticity in the data. Heteroscedasticity violates one of the assumptions of OLS regression, which assumes that the variance of the error term is constant across all levels of the independent variables. Indeed, testing for heteroscedasticity in the simple model using a White test leads to reject the homoskedascity assumption at the 10% level with a p-value of 0.058. Running the White test on the extended model yields no proof of heteroscedasticity mostly because the sample is not large enough. Assuming that the variance of each observation is directly proportional to the sample size, it is enough to intensify the variables and divide them all by the sample size. What we obtain is thus a sphericalized model which should not bear any heteroscedasticity issues. Table 5 displays the results of the sphericalized estimation. The coefficient β_3 is significantly different from zero at the 1% level and positive but its value is very different from the extended OLS. This was an expected result, in case of heteroscedasticity, the OLS estimated variance has no reason to be true as the variance-covariance matrix of the noise used to compute it is wrongly specified.

Table 5: Weighted Least Squares Extended Estimation - Exercise 1

	Dependent variable:
	self_emp
log_gdp	-0.585
	(1.698)
literacy	0.239
	(0.186)
agro_emp	1.043***
· _ ·	(0.107)
gfce	-0.441^{**}
	(0.193)
stocks	0.383
	(0.288)
bribery	0.404***
	(0.130)
Constant	-0.001
	(0.002)
Observations	49
\mathbb{R}^2	0.929
Adjusted \mathbb{R}^2	0.918
Residual Std. Error	$0.007~({ m df}=42)$
F Statistic	$91.073^{***} (df = 6; 42)$
Note:	*p<0.1; **p<0.05; ***p<0.01

var	median	mean	\min	max	sd	NAs
business_crea	0.1	0.1	5.6e-02	0.3	3.1e-02	0
nb_crimes	73.0	149.8	9.0	$1.0\mathrm{e}{+04}$	451.3	0
$nb_households$	$9.8\mathrm{e}{+03}$	$1.9\mathrm{e}{+04}$	$3.7\mathrm{e}{+03}$	$1.4\mathrm{e}{+06}$	$5.3\mathrm{e}{+04}$	0
pop	$2.0\mathrm{e}{+04}$	$3.5\mathrm{e}{+04}$	$1.0\mathrm{e}{+04}$	$2.2\mathrm{e}{+06}$	$8.8\mathrm{e}{+04}$	0
income	2.0	2.9	1.0	6.0	1.1	0
$\operatorname{crime}_{\operatorname{rate}}$	7.2e-03	7.9e-03	1.0e-03	3.0e-02	4.1e-03	0

Table 6: Descriptive statistics

1.7 What can you conclude about the factors driving differences in self-employment rates at the macro-level based on your above answers? Can these results be interpreted as causal?

2 Heteroskedasticity & Monte Carlo Simulations

3 Instrumental Variables

3.1

Table 6 shows the descriptive statistics of the table made of 900 hundred municipalities. First, the data is remarkably clean as their is no NA values in it.

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Table 7: Results

	$Dependent\ variable:$
	crime_rate
business crea	0.018***
_	(0.005)
$\log(\text{pop})$	0.001***
	(0.0002)
income	0.00001
	(0.0001)
com_typeC	-0.003***
_ , ,	(0.0003)
com typeI	-0.002***
_ v -	(0.001)
Constant	0.001
	(0.002)
Observations ————————————————————————————————————	899
R^2	0.150
Adjusted R^2	0.145
Residual Std. Error	0.004 (df = 893)
F Statistic	$31.568^{***} (df = 5; 893)$
Note:	*p<0.1; **p<0.05; ***p<0.01