Show that the Gaussian pdf integrates to one. Consider the square of the integral of the pdf:

$$\left[\frac{1}{\sqrt{2\pi}}\int_{-\infty}^{\infty} e^{-x^{2}/2} dx\right]^{2} = \frac{1}{2\pi}\int_{-\infty}^{\infty} e^{-x^{2}/2} dx \int_{-\infty}^{\infty} e^{-x^{2}/2} dy$$
$$= \frac{1}{2\pi}\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-(x^{2}+y^{2})/2} dx dy.$$

Let $x = r \cos \theta$ and $y = r \sin \theta$ and carry out the change from Cartesian to

polar coordinates, then we obtain:

$$\frac{1}{2\pi} \int_{0}^{\infty} \int_{0}^{2\pi} e^{-\frac{\pi}{2}} r dr d\theta = \int_{0}^{\infty} r e^{-\frac{\pi}{2}} dr$$

$$= \left[-e^{-\frac{\pi}{2}} \right]_{0}^{\infty}$$

$$= 1.$$

EXAMPLE .

A communication system accepts a positive voltage V as input and outputs a voltage $Y = \alpha V + N$, where $\alpha = 10^{-2}$ and N is a Gaussian random variable with parameters m = 0 and $\sigma = 2$. Find the value of V that gives P[Y < $0] = 10^{-6}$.

The probability P[Y < 0] is written in terms of N as follows: $P[Y < 0] = P[\alpha V + N < 0]$ $P[Y < 0] = P[\alpha V + N < 0]$

The probability
$$P[Y < 0]$$
 is written in the second of $P[Y < 0] = P[\alpha V + N < 0]$

$$= P[N < -\alpha V] = \Phi\left(\frac{-\alpha V}{\sigma}\right) = Q\left(\frac{\alpha V}{\sigma}\right) = 10^{-6}.$$

N = N(0,2)

From Table 3.4 we see that the argument of the Q-function should be $\alpha V/\sigma = 4.753$. Thus $V = (4.753)\sigma/\alpha = 950.6$.

EXAMPLE 413

Show that the pdf of a gamma random variable integrates to one.

The integral of the pdf is
$$\int_{0}^{\infty} f_{X}(x) dx = \int_{0}^{\infty} \frac{\lambda(\lambda x)^{\alpha - 1} e^{-\lambda x}}{\Gamma(\alpha)} dx$$

$$= \frac{\lambda^{\alpha}}{\Gamma(\alpha)} \int_{0}^{\infty} x^{\alpha - 1} e^{-\lambda x} dx.$$

Let y = ix, then $dx = dy/\lambda$ and the integral becomes

where we used the fact that the integral equals
$$\Gamma(\alpha)$$
.

Let the function $h(x) = (x)^+$ be defined as follows:

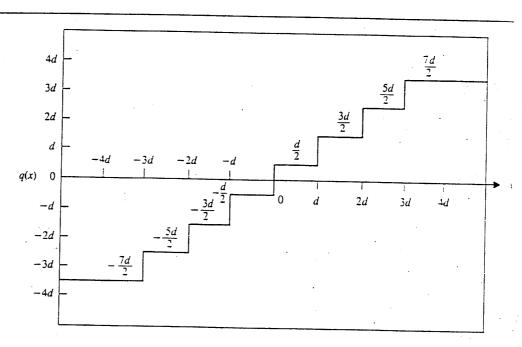
$$(x)^{+} = \begin{cases} 0 & \text{if } x < 0 \\ x & \text{if } x \ge 0. \end{cases}$$

For example, let X be the number of active speakers in a group of N speakers, and let Y be the number of active speakers in excess of M, then $Y = (X - M)^+$. In another example, let X be a voltage input to a half-wave rectifier, then $Y = (X)^+$ is the output.

EXAMPLE ATT

Let the function q(x) be defined as shown in Fig. 418 Where The School points on the real-line are mapped into the nearest representation point from the set $S_0 = \{-3.5d, -2.5d, \dots -(0.5...)\}$ 2.513.513. Thus for example, all the points in the interval (0,d) are mapped into the point d/2. The function q(x) represents an eight-level uniform quantizer.

FIGURE 4.9 A uniform quantizer maps the input x into the cissest point from the set $\{\pm d/2, \pm 3d/2, \pm 5d/2, \pm 7d/2\}$



4.30

EXAMPLE TO

Let X be a sample voltage of a speech waveform, and suppose that X has a uniform distribution in the interval [-4d, 4d]. Let Y = q(X), where the quantizer input-output characteristic is as shown in Fig. \blacksquare Find the pmf for Y.

The event $\{Y = q\}$ for q in S_Y is equivalent to the event $\{X \text{ in } I_q\}$, where I_q is an interval of points mapped into the representation point q. The pmf of Y is therefore found by evaluating q(A)

 $P[Y=q] = \int_{Q} f_{x}(t)dt = \int_{Q} dt = \frac{1}{8}$ This easy to see that the representation point has an interval of length d

His cost to see that the representation point has an interval of length d mapped into it. Thus the eight possible outputs are equiprobable, that is, P[Y=q]=1/8 for q in S_Y .

Let the random variable Y be defined by

$$Y = aX + b$$

where a is a nonzero constant. Suppose that X has cdf $F_X(x)$, then find $F_Y(y)$. The event $\{Y \le y\}$ occurs when $A = \{aX + b \le y\}$ occurs. If a > 0,

then $A = \{X \le (y - b)/a\}$ (see Fig. 3.16), and thus

$$F_Y(y) = P\left[X \leq \frac{y-b}{a}\right] = F_X\!\!\left(\!\frac{y-b}{a}\right) \qquad a>0.$$

On the other hand, if a < 0, then $A = \{X \ge (y - b)/a\}$, and

$$F_{\nu}(y) = P\left[X \geq \frac{y-b}{a}\right] = 1 - F_X\left(\frac{y-b}{a}\right) \qquad a < 0.$$

We can obtain the pdf of Y by differentiating with respect to v. To do this we need to use the chain rule for derivatives:

$$\frac{dF}{dy} = \frac{dF \, du}{du \, dy},$$

where u is the argument of F. In this case, u = (y - b)/a, and we then obtain

$$f_Y(y) = \frac{1}{a} f_X\left(\frac{y-b}{a}\right) \qquad a > 0$$

and

$$f_Y(y) = \frac{1}{-a} f_X\left(\frac{y-b}{a}\right) \qquad a < 0.$$

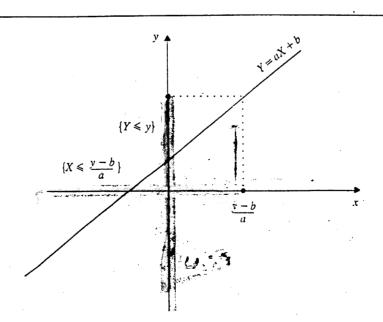
The above two results can be written compactly as

$$f_Y(y) = \frac{1}{|a|} f_X\left(\frac{y-b}{a}\right).$$

447 (343)

FIGURE VI I

The equivalent event for $\{Y \le y\}$ is the event $\{X \le (y - b)/a\}$, if a > 0.





Let X be a random variable with a Gaussian pdf with mean m and standard deviation σ :

$$f_X(x) = \frac{1}{\sqrt{2\pi} \sigma} e^{-(x-m)^2/2\sigma^2} \qquad -\infty < x < \infty. \tag{3.50}$$

Let Y = aX + b, then find the pdf of Y.

Substitution of Eq. (3.50) into Eq. (3.49) yields

$$f_Y(y) = \frac{1}{\sqrt{2\pi} |a\sigma|} e^{-(y-b-am)^2/2(a\sigma)^2}.$$

Note that Y also has a Gaussian distribution with mean b + am and standard deviation $|a| \sigma$. Therefore a linear function of a Gaussian random variable is also a Gaussian random variable.

EXAMPLE 3

Let the random variable Y be defined by

Where X is a Continuous Y.V. Find cdf, pdf off

The event $\{Y \le y\}$ occurs when $\{X^2 \le y\}$ or equivalently when $\{X^2 \le y\}$ for y nonnegative; see Fig. 3.17. The event is null when y is negative. Thus

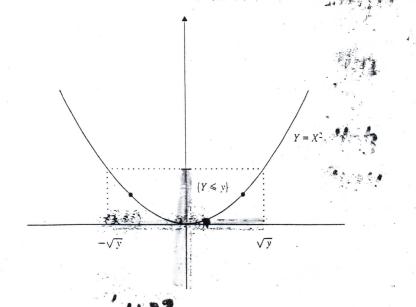
$$F_Y(y) = \begin{cases} 0 & y < 0 \\ F_X(\sqrt{y}) - F = \sqrt{y}) & y > 0 \end{cases}$$

and differentiating with respect to y,

$$f_{Y}(y) = \frac{f_{X}(\sqrt{y})}{2\sqrt{y}} - \frac{f_{X}(-\sqrt{y})}{-2\sqrt{y}} \qquad y > 0$$

$$= \frac{f_{X}(\sqrt{y})}{2\sqrt{y}} + \frac{f_{X}(-\sqrt{y})}{2\sqrt{y}}.$$
(3.51)

FIGURE 3.17 The equivalent event for $\{Y \le y\}$ is the event $\{-\sqrt{y} \le X \le \sqrt{y}\}$, if $y \ge 0$.



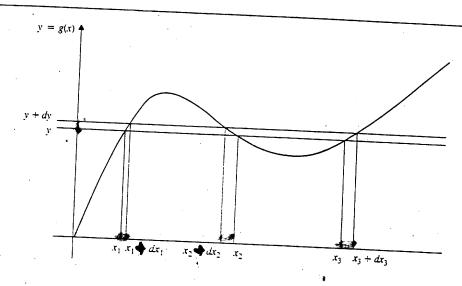
The result in Example x_0 suggests that if the equation $y_0 = g(x)$ has n solutions, x_0, x_1, \ldots, x_n , then $f_{Y_0}(y_0)$ will be equal to n terms of the type on the right-hand side of Eq. (3.51). We now show that this is generally true by using a method for directly obtaining the pdf of Y in terms of the pdf of X.

Consider a nonlinear function Y = g(X) such as the one shown in Fig. 3.18. Consider the event $C_y = \{y < Y < y + dy\}$ and let B_y be its equivalent event. For y indicated in the figure, the equation g(x) = y has three solutions x_1, x_2 , and x_3 , and the equivalent event B_y has a segment corresponding to each

Consider a nonlinear function Y = g(X) such as the one shown in Fig. 3 Consider the event $C_y = \{y < Y < y + dy\}$ and let B_y be its equivalent event. For y indicated in the figure, the equation g(x) = y has three solutions x_1, x_2 , and x_3 , and the equivalent event B_y has a segment corresponding to each



The equivalent event of $\{y < 1\}$ Y < y + dy is $\{x_1 < X < x_1 + x_2 < x_3 + x_4 < x_4 < x_4 + x_4 < x_4 < x_4 + x_4 < x$ dx_1 $\cup \{x_2 + dx_2 < X < x_2\} \cup$ $\{x_3 < X < x_3 + dx_3\}$



the equivalent event By

solution:

$$\Rightarrow B_y = \{x_1 < X < x_1 + dx_1\} \cup \{x_2 + dx_2 < X < x_2^*\}$$

 $\cup \{x_3 < X < x_3 + dx_3\}.$

The probability of the event C_y is approximately

$$P[C_y] = f_Y(y) |dy|,$$

where |dy| is the length of the interval $y < Y \le y + dy$. Similarly, the probability of the event B_y is approximately

$$P[B_y] = f_X(x_1) |dx_1| + f_X(x_2) |dx_2| + f_X(x_3) |dx_3|.$$

Since C_0 and B_0 are equivalent events, their probabilities must be equal. By equating Eqs. (3.53) and (3.54) we obtain

$$f_Y(y) = \sum_{k} \frac{f_X(x)}{|dy/dx|} \Big|_{x = \mathbf{x}_K}$$
$$= \sum_{k} f_X(x) \left| \frac{dx}{dy} \right|_{x = \mathbf{x}_K}$$

It is clear that if the equation g(x) = y has n solutions, the expression for the pdf of Y at that point is given by Eqs. and and contains n terms.