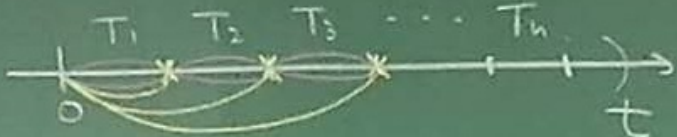


$N(t)$ $P(N(t)=k) = \frac{(\lambda t)^k}{k!} \exp(-\lambda t)$. Waiting.

Jump Processes  $\{T_k\}$ continuous r.v.

$$\frac{1}{\lambda} F_{T_1}(t) = P(T_1 \leq t) = 1 - P(T_1 > t) = 1 - P(N(t) = 0)$$

$$\frac{1}{\lambda^2} = 1 - \exp(-\lambda t), \quad f_{T_1}(t) = \frac{d}{dt} F_{T_1}(t) = \lambda \exp(-\lambda t), \quad t \geq 0$$

T_1 : Exponential Distribution. T_k i.i.d!

$$E(N(t)) = \lambda t \Rightarrow \lambda = \frac{E(N(t))}{t} \text{ Intensity.}$$

$$S_1 = T_1, \quad S_2 = T_1 + T_2, \quad S_3 = T_1 + T_2 + T_3, \quad \dots$$

$$\begin{aligned} \Phi_{S_k}(\omega) &= E(\exp(j\omega S_k)) = E(\exp(j\omega \sum_{i=1}^k T_i)) = \prod_{i=1}^k E(\exp(j\omega T_i)) \\ &= \prod_{i=1}^k \Phi_{T_i}(\omega) = (\Phi_{T_1}(\omega))^k = \left(\frac{\lambda}{\lambda - j\omega}\right)^k \Rightarrow \dots \end{aligned}$$

$$\Phi_{T_1}(\omega) = \int_0^\infty \lambda \exp(-\lambda x) \exp(j\omega x) dx = \lambda \int_0^\infty \exp(-(\lambda - j\omega)x) dx = \frac{\lambda}{\lambda - j\omega}$$

$$F_{S_k}(t) = P(S_k \leq t) = \sum_{i=k}^{\infty} \frac{(\lambda t)^i}{i!} \exp(-\lambda t).$$

$$\begin{aligned} f_{S_k}(t) &= \frac{d}{dt} F_{S_k}(t) = \sum_{i=k}^{\infty} \frac{\lambda (\lambda t)^{i-1}}{(i-1)!} \exp(-\lambda t) - \lambda \frac{(\lambda t)^i}{i!} \exp(-\lambda t) \\ &= \lambda \exp(-\lambda t) \sum_{i=k}^{\infty} \left(\frac{(\lambda t)^{i-1}}{(i-1)!} - \frac{(\lambda t)^i}{i!} \right) = \lim_{m \rightarrow \infty} \left(\frac{(\lambda t)^{k-1}}{(k-1)!} - \frac{(\lambda t)^m}{m!} \right) = \frac{(\lambda t)^{k-1}}{(k-1)!} \\ f_{S_k}(t) &= \frac{\lambda (\lambda t)^{k-1}}{(k-1)!} \exp(-\lambda t), (t \geq 0). \end{aligned}$$

Gamma Distribution

$$\Gamma(n+1) = n \Gamma(n) \Rightarrow \Gamma(n) = (n-1)! \quad \Gamma(1) = 1$$


Gamma Function. $\Gamma(p) = \int_0^{\infty} x^{p-1} \exp(-x) dx.$

$$\begin{aligned} \Gamma(p+1) &= \int_0^{\infty} x^p \exp(-x) dx = - \int_0^{\infty} x^p d \exp(-x) = -x^p \exp(-x) \Big|_0^{\infty} \\ &\quad + \int_0^{\infty} \exp(-x) dx^p = p \int_0^{\infty} x^{p-1} \exp(-x) dx = p \Gamma(p) \end{aligned}$$

$$f_{S_k}(t) = \frac{\lambda (\lambda t)^{k-1}}{\Gamma(k)} \exp(-\lambda t). \quad S_k \sim T(k, \lambda) \frac{\lambda^p}{\Gamma(p)} \int_0^{\infty} t^{p-1} \exp(-\alpha t) dt = 1$$

$$f_{\bar{x}}(t) = \frac{\alpha^p t^{p-1}}{\Gamma(p)} \exp(-\alpha t), t \geq 0, \quad \bar{x} \sim T(p, \alpha)$$

$$F_{Y^2}(x) = P(Y^2 \leq x) = P(-\sqrt{x} \leq Y \leq \sqrt{x}) = 2 \int_0^{\sqrt{x}} f_Y(y) dy.$$



$$= 2 \int_0^{\sqrt{x}} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{y^2}{2}\right) dy. \quad f_{Y^2}(x) = \frac{d}{dx} F_{Y^2}(x).$$

$$= \frac{1}{\sqrt{2\pi}} \cdot x^{-\frac{1}{2}} \exp\left(-\frac{x}{2}\right), \quad x \geq 0. \quad \text{Chi-Square Distribution}$$

$$N(t). \quad R_N(t, s) = E(N(t)N(s)) = E((N(t) - N(s)) + N(s))N(s)$$

$$= E(N(t) - N(s))E(N(s)) + E(N^2(s)) = \lambda(t-s) \cdot \lambda s + (\lambda s)^2 + \lambda s.$$

$\lambda(t-s) \quad \lambda s \quad \lambda^2 s^2 + \lambda s$

$$\Gamma\left(\frac{1}{2}\right) = \int_0^\infty x^{-\frac{1}{2}} \exp(-x) dx = 2 \int_0^\infty \exp(-x) dx^{\frac{1}{2}}$$

$$\stackrel{y=x^{\frac{1}{2}}}{=} 2 \int_0^\infty \exp(-y^2) dy = 2 \int_0^\infty \exp\left(-\frac{y^2}{2 \cdot \frac{1}{2}}\right) dy.$$

$$= 2 \cdot \sqrt{2\pi} \cdot \sqrt{\frac{1}{2}} \cdot \frac{1}{2} = \sqrt{\pi}$$

$$Z \sim \Gamma\left(\frac{1}{2}, \frac{1}{2}\right) \quad f_Z(x) = \frac{\left(\frac{1}{2}\right)^{\frac{1}{2}}}{\Gamma\left(\frac{1}{2}\right)} \cdot x^{-\frac{1}{2}} \exp\left(-\frac{x}{2}\right), \quad x \geq 0.$$

$$Y \sim N(0, 1) \Rightarrow Z = Y^2 = \frac{1}{\sqrt{2\pi}} x^{-\frac{1}{2}} \exp\left(-\frac{x}{2}\right), \quad x \geq 0$$

$$= \lambda^2 ts + \lambda \min(t, s). \quad N(t) \xrightarrow{\quad} \boxed{H} \longrightarrow Y(t).$$

$$Y(t) = N(t) - N(t-1). \quad R_Y(t, s) = E(Y(t)Y(s)).$$

$$= E(N(t)N(s)) - E(N(t-1)N(s)) - E(N(t)N(s-1)) + E(N(t-1)N(s-1))$$

$$= \lambda^2 ts + \lambda \min(t, s) - \lambda^2 (t-1)s - \lambda \min(t-1, s) - \lambda^2 t(s-1) - \lambda \min(t, s-1) \\ + \lambda^2 (t-1)(s-1) + \lambda \min(t-1, s-1)$$

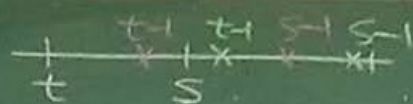
$$= \min(t, s) - \min(t-1, s) - \min(t, s-1) + \lambda \min(t-1, s-1)$$

$$\Gamma(r) = \int_0^\infty x^r \exp(-x) dx = \int_0^\infty x^r$$

$$+ \int_0^\infty \exp(-x) dx = P \int_0^\infty x^r$$

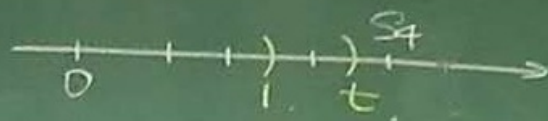
$$f_{S_r}(t) = \frac{\lambda (P)^r}{\Gamma(r)} \exp(-\lambda t) \quad S_r \sim \text{Gamma}(r, \lambda) \quad (P)$$

$$f_{\Sigma}(t) = \frac{P^r t^{r-1}}{\Gamma(r)} \exp(-\alpha t), \quad t \geq 0, \quad \Sigma \sim \text{Gamma}(r, \alpha)$$



$$S = S - (S-1)$$

$$S - (t-1) = (S-1)$$



$$S_k = T_1 + \dots + T_k. \quad E(S_k) = E(T_1 + \dots + T_k) = \sum_{i=1}^k E(T_i) = \frac{k}{\lambda}$$

$$E(S_4) = \frac{4}{\lambda}. \quad E(S_4 | N(1)=2) \quad F_{S_4}(t | N(1)=2) = P(S_4 \leq t | N(1)=2)$$

$$= P(N(t) \geq 4 | N(1)=2) = 1 - P(N(t) \leq 3 | N(1)=2)$$

$$= 1 - P(N(t)=2 | N(1)=2) - P(N(t)=3 | N(1)=2)$$

$$P(N(t)=2 | N(1)=2) = \frac{P(N(1)=2, N(t)-N(1)=0)}{P(N(1)=2)} = \frac{P(N(1)=2) P(N(t)-N(1)=0)}{P(N(1)=2)}$$

$$= P(N(t)-N(1)=0) = \exp(-\lambda(t-1))$$

$$P(N(t)=3 | N(1)=2) = P(N(t)-N(1)=1) = \lambda(t-1) \exp(-\lambda(t-1))$$

$$F_{S_4}(t | N(1)=2) = 1 - \exp(-\lambda(t-1)) - \lambda(t-1) \exp(-\lambda(t-1)), \quad t \geq 1$$

$$f_{S_4}(t | N(1)=2) = \frac{d}{dt} F_{S_4}(t | N(1)=2) = \lambda \exp(-\lambda(t-1)) - \lambda \exp(-\lambda(t-1)) + \lambda^2(t-1) \exp(-\lambda(t-1))$$

$$f_{S_4}(t|N(1)=2) = \lambda^2(t-1)\exp(-\lambda(t-1)), \quad t \geq 1.$$

$$E(\bar{x}) = \frac{1}{\lambda} \int_1^{\infty} \lambda^2(t-1)\exp(-\lambda(t-1))dt = \lambda^2 \int_0^{\infty} t \exp(-\lambda t) dt = 1.$$

$$E(\bar{x}^2) = \frac{2}{\lambda^2} \lambda^2 \int_1^{\infty} t(t-1)\exp(-\lambda(t-1))dt = \lambda^2 \int_0^{\infty} (t+1)t \exp(-\lambda t) dt$$

$$= \lambda^2 \int_0^{\infty} t^2 \exp(-\lambda t) dt + \lambda^2 \int_0^{\infty} t \exp(-\lambda t) dt$$

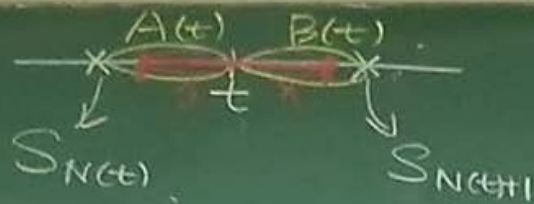
$$E(S_4|N(1)=2) = \lambda E(\bar{x}^2) + \lambda E(\bar{x}) = 1 + \frac{2}{\lambda} \quad \frac{4}{\lambda} \quad 0 \quad 1$$

$$\bar{x} \sim \text{Exp}(\lambda). \quad F_{\bar{x}}(x) = P(\bar{x} \leq x) = 1 - \exp(-\lambda x), \quad x \geq 0.$$

$$F_{\bar{x}}(x|\bar{x} > a) = P(\bar{x} \leq x | \bar{x} > a) = \frac{P(a < \bar{x} \leq x)}{P(\bar{x} > a)}$$

$$= \frac{\exp(-\lambda a) - \exp(-\lambda x)}{\exp(-\lambda a)} = 1 - \exp(-\lambda(x-a)), \quad (x \geq a)$$

$$E(\bar{x}) > E(Y) \quad \boxed{\text{Inspection Paradoxes}}$$



$$A(t) = t - S_{N(t)}$$

$$P(B(t) > x)$$

$$B(t) = S_{N(t)+1} - t = P(A(t) > 0, B(t) > x)$$

$$P(A(t) > y, B(t) > x)$$

$$P(B(t) \leq x) = F_{B(t)}(x)$$

$$= P(N(x+t) = 0) = \exp(-\lambda(x+t))$$

$$P(B(t) > x) = \exp(-\lambda x) = 1 - F_{B(t)}(x)$$

$$E(S_{N(t)+1} - S_{N(t)}) \neq E(S_{k+1} - S_k), \forall k$$

$$E(S_{N(t)+1} - S_{N(t)}) = E_N(E_S(S_{k+1} - S_k \mid N(t) = k))$$

$$= E\left(\frac{1}{\lambda}\right) = \frac{1}{\lambda}$$