

Midterms Revision Guide

30.003 Probability and Statistics, Term 4 2019

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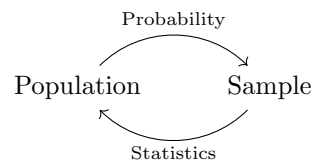
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1 W1: Probability and Statistics

1.1 Definitions

- Population: well defined collection of objects
- Sample: subset of population selected in certain manner
- Variable: any characteristic whose value may change from one object to another in population
- Probability: properties of populations known, question regarding sample taken from population are investigated (**deductive reasoning**)
- Statistics: characteristics of sample known from experiments, conclusions regarding population are made (**inductive reasoning**)



1.2 Frequency

- Frequency: number of times value occurs in data set
- Relative frequency: fraction or proportion of times the value occurs

1.3 Range and mean

- Range: difference between largest and smallest sample values
- Mean: average of all values
- Population mean is denoted by μ
- Sample mean is denoted by \bar{x} , where

$$\bar{x} = \frac{\sum x_i}{n}, \text{ and } n \text{ denoting the number of data points}$$

1.4 Variance and standard deviation

- Variance: measures variability of data set
- Population variance is denoted by σ^2 , where

$$\sigma^2 = \frac{1}{N} \sum_{i=1}^N (x_i - \mu)^2, \text{ and } N \text{ denoting the size of the population}$$

- Sample variance is denoted by s^2 , where

$$s^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2, \text{ and } n \text{ denoting the size of the sample}$$

- Standard deviation is denoted as σ for population variance and s for sample variance, and is calculated either by:

$$\sigma = \sqrt{\sigma^2}, \text{ or } s = \sqrt{s^2}$$

where σ^2 is the population variance and s^2 is the sample variance

- Shortcut to calculate population variance:

$$\sigma^2 = \frac{1}{N} \sum_{i=1}^N x_i^2 - \mu^2$$

1.5 Median

- Median: the middle value in a data set

1.6 Percentile

- value below which a given percentage of observations falls
- data set is ordered as $x'_1 \leq x'_2 \leq \dots \leq x'_n$,
where x'_1 and x'_n are the smallest and largest data values respectively
- x'_i corresponds to the $\frac{100(i-0.5)}{n}$ th percentile

1.7 Sample space and events

- Sample space: the set of all possible outcomes of an experiment
 1. Collectively exhaustive
 2. Mutually exclusive
- Event: collection of outcomes contained in sample space Ω
 1. Simple event: exactly one outcome
 2. Compound event: > 1 outcome

1.8 Sample Space vs Population

- Sample space: contains mutually exclusive events
- Population: events can repeat many times

1.9 Set Theory

- Null event, \emptyset : event with no outcome
- Events A and B are mutually exclusive/disjoint if $A \cap B = \emptyset$

1.10 De Morgan's Laws

$$(A \cup B)^c = A^c \cap B^c$$

$$(A \cap B)^c = A^c \cup B^c$$

$$A \cup B = A + B - A \cap B$$

1.11 Axiom of Probability

1. For any event A, $P(A) \geq 0$.
2. $P(\Omega) = 1$
3. Any infinite collection of mutually exclusive/disjoint events $A_1, A_2, A_3, \dots, A_n$ satisfies

$$P(A_1 \cup A_2 \cup A_3 \cup \dots \cup A_n) = \sum_{i=1}^{\infty} P(A_i)$$

1.12 Properties of Probability

- For any event A, $P(A) + P(A^c) = 1$.
- $P(\Omega) = P(A \cup A^c) = P(A) + P(A^c)$
 $\because A$ and A^c are disjoint
- For any event A, $P(A) \leq 1$.
- For a null event \emptyset , $P(\emptyset) = 0$
 - Does **NOT** suggest $A = \emptyset$
- Similarly, $P(A) = 1$ does **NOT** suggest $A = \Omega$

1.13 Equally likely outcomes

$$P(\text{equally likely event}) = \frac{1}{n}, \text{ where } n \text{ is the number of equally likely events}$$

2 W1: Counting Technique

2.1 Finding probability

- Computing probability \rightarrow counting

$$P(A) = \frac{N(A)}{N}$$

- where $N(A)$ is the number of outcomes for event A ,
and N is the number of outcomes in the sample space

2.2 Tuple

- Group of k elements: k -tuple
- The 1st element is selected in n_1 ways; the 2nd element is selected in n_2 ways; the k^{th} element is selected in n_k ways; such that *the elements are selected independently*.

2.3 Permutation

- Ordered subset
- Number of permutations of size k formed from n objects:

$$P_{k,n} = \frac{n!}{(n-k)!}$$

2.4 Combination

- Unordered subset of a group
- Number of combinations of size k formed from n objects:

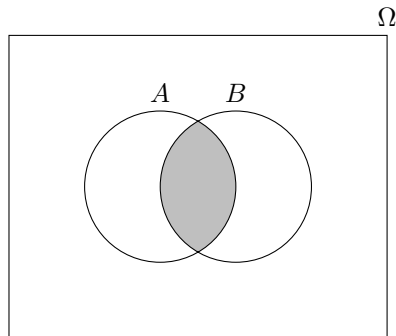
$$\binom{n}{k} \text{ or } C_{k,n} = \frac{P_{k,n}}{k!} = \frac{n!}{k!(n-k)!}$$

- Disregards the different outcomes due to order

3 W2: Conditional Probability

- Probability of event A given that event B has occurred: $P(A|B)$

$$P(A | B) = \frac{P(A \cap B)}{P(B)}$$

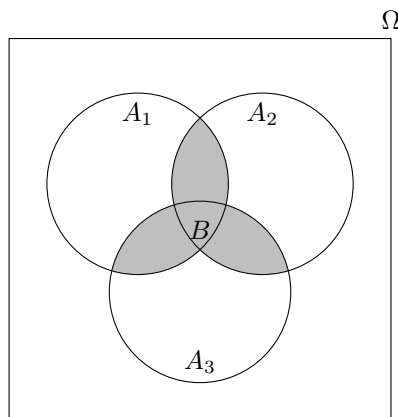


3.1 Law of Total Probability

- Let A_1, A_2, \dots, A_k be mutually exclusive and exhaustive events.

For any other event B,

$$P(B) = \sum_{i=1}^k P(B | A_i)P(A_i)$$



3.2 Bayes' Theorem

$$\begin{aligned} P(A_j | B) &= \frac{P(A_j \cap B)}{P(B)} \\ &= \frac{P(B \cap A_j)}{P(B)} \\ &= \frac{P(B | A_j)P(A_j)}{\sum_{i=1}^k P(B | A_i)P(A_i)} \end{aligned}$$

3.3 Independence of Random Variables

- Independence: occurrence/non-occurrence of one event has no bearing on the chance that the other will occur
 - $P(A | B) = P(A)$: A and B are independent
 - $P(A | B) \neq P(A)$: A and B are not independent
- Independence of A and B also implies $P(B | A) = P(B)$ if $P(A) > 0$

3.3.1 Multiplication Rule

- A and B are independent iff. $P(A \cap B) = P(A)P(B)$

3.3.2 Independence of several events

$$P(A_{i1} \cap A_{i2} \cap \dots \cap A_{ik}) = P(A_{i1})P(A_{i2}) \dots P(A_{ik})$$

4 W2: Discrete Random Variable

4.1 Probability Mass Function (PMF) for Discrete RV

- For any pmf, $p(x) \geq 0$ and $\sum_{\text{all possible } x} p(x) = 1$

4.2 Bernoulli RV

- pmf of any Bernoulli RV:

$$p(x; \alpha) = \begin{cases} 1 - \alpha, & \text{if } x = 0 \\ \alpha, & \text{if } x = 1 \\ 0, & \text{otherwise} \end{cases}$$

- α is a parameter, where $0 < \alpha < 1$

4.3 Bernoulli process

- A process with repeated independent trials
- 2 outcomes: 1 (success), 0 (failure)
- Success rate of trials is the same

4.4 Binomial distribution

- pmf of binomial RV:

$$p(x; n, p) = \begin{cases} C_{x,n} p^x (1-p)^{n-x}, & x = 0, 1, \dots, n \\ 0, & \text{otherwise} \end{cases}$$

- where n is the number of trials, and p is the success rate of each trial

4.5 Geometric distribution

- Probability distribution of number of Bernoulli trials X needed to get 1 success
- If $X = x$, $x - 1$ failures followed by success
- pmf of geometric RV:

$$p(x) = \begin{cases} p(1-p)^{x-1}, & x = 1, 2, \dots \\ 0, & \text{otherwise} \end{cases}$$

- where p is the success rate of each trial

4.6 Poisson distribution

- Used to model the number of occurrences of events in a time interval, where the average occurrence is λ

- pmf of Poisson RV:

$$p(x; \lambda) = \begin{cases} \frac{\lambda^x e^{-\lambda}}{x!}, & x = 0, 1, \dots \\ 0, & \text{otherwise} \end{cases}$$

- where λ is the parameter of Poisson distribution

4.7 Cumulative Distribution Function (CDF)

- CDF $F(x)$ of discrete RV X with pmf $p(x)$:

$$F(x) = P(X \leq x) = \sum_{y: y \leq x} p(y)$$

- $F(x)$ is the probability that the observed value is at most x
- Graph of $F(x)$ for discrete RV X is the linear combination of step functions, such that

$$\lim_{x \rightarrow -\infty} F(x) = 0 \text{ and } \lim_{x \rightarrow \infty} F(x) = 1$$

5 W3: Expectation

5.1 Expected Value

- Expected value $E(X)$

$$E(X) = \mu_x = \sum_{x \in D} x \cdot p(x), \text{ provided that } \sum_{x \in D} |x| \cdot p(x) < \infty$$

- Expected value of a function $E[h(X)]$

$$E[h(X)] = \mu_{h(x)} = \sum_{x \in D} h(x) \cdot p(x)$$

- Expected value of a linear function $aX + b$

$$E(aX + b) = aE(X) + b$$

5.2 Variance

- Variance $V(X)$

$$V(X) = \sum_{x \in D} (x - \mu)^2 p(x) = E[(X - \mu)^2], \text{ provided that the expectation exists}$$

OR

$$\text{Population variance, } \sigma^2 = V(X) = E(X^2) - [E(X)]^2$$

- Variance of a function $V[h(X)]$

$$V[h(X)] = \sum_{x \in D} \{h(x) - [E(X)]\}^2 \cdot p(x)$$

- Variance of a linear function $aX + b$

$$V(aX + b) = a^2 V(X)$$

$$\sigma_{aX+b} = |a| \sigma_x$$

5.3 Expected Value and Variance of Discrete PMFs

5.3.1 Bernoulli RV

- Expected value $E(X) = p$
- Variance $V(X) = p(1 - p)$

5.3.2 Binomial RV

- Expected value $E(X) = np$
- Variance $V(X) = np(1 - p)$

5.3.3 Geometric RV

- Expected value $E(X) = \frac{1}{p}$
- Variance $V(X) = \frac{1-p}{p^2}$

5.3.4 Poisson RV

- Expected value $E(X) = \lambda$
- Variance $V(X) = \lambda$

6 W3: Continuous Random Variable

6.1 Probability Density Function (PDF) for Continuous RV

- Probability described by the probability density function (pdf), measured between an interval

$$P(a \leq X \leq b) = \int_a^b f(x)dx$$

6.2 Uniform Distribution

$$\text{pdf } f(x; a, b) = \begin{cases} \frac{1}{b-a}, & a \leq x \leq b \\ 0, & \text{otherwise} \end{cases}$$

6.3 Exponential Distribution

$$\text{pdf } f(x; \lambda) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0, & \text{otherwise} \end{cases}$$

6.4 Normal/Gaussian Distribution

$$\text{pdf } f(x; \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

6.5 Cumulative Distribution Function (CDF)

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(u)du$$

- Capital F means CDF, while small f means PDF
- For any a : $P(x > a) = 1 - F(a)$
- Between a and b : $P(a \leq X \leq b) = F(b) - F(a)$

6.5.1 Obtaining PDF from CDF

$$f(x) = F'(x)$$

- The PDF is the derivative of the CDF.

6.6 Expected Value

- Expected value $E(X)$

$$E(X) = \mu_x = \sum_{x \in D} x \cdot p(x), \text{ provided that } \int_{-\infty}^{\infty} |x| \cdot p(x) < \infty$$

- Expected value of a function $E[h(X)]$

$$E[h(X)] = \mu_{h(x)} = \int_{-\infty}^{\infty} h(x)f(x)dx$$

- Expected value of a linear function $aX + b$

$$E(aX + b) = aE(X) + b$$

6.7 Variance

- Variance $V(X)$

$$\begin{aligned} V(X) &= \mu_X^2 = E[(X - \mu)^2] \\ &= E(X^2) - [E(X)]^2 \end{aligned}$$

- Variance of a linear function $aX + b$

$$V(aX + b) = a^2V(X)$$

$$\sigma_{aX+b} = |a|\sigma_x$$

6.8 Expected Value and Variance of Continuous PDFs

6.8.1 Uniform RV

- Expected value $E(X) = \frac{1}{2}(a + b)$
- Variance $V(X) = \frac{1}{12}(b - a)^2$

6.8.2 Exponential RV

- Expected value $E(X) = \frac{1}{\lambda_E}$
- Variance $V(X) = \frac{1}{\lambda^2}$

7 W4: Useful Distributions

7.1 Poisson Approximation of Binomial Distributions

For any binomial distribution where n is large and p is small, such that $np > 0$,

$$b(x; n, p) \approx p(x; \lambda), \text{ where } \lambda = np$$

- Approximation can be safely applied if $n > 50$ and $np < 5$

7.2 Relationship between Poisson and Exponential Distributions

- Poisson distribution: Often used to model the number of occurrence of events in a time interval
- Exponential distribution: Often used to model the elapsed time between two successive events

Let X_1, X_2, \dots be the time when the 1st, 2nd, ... event occur.

The probability of waiting not more than t for the first event is $P(X_1 \leq t)$.

Deriving via Poisson Distribution

$$\begin{aligned} P(X_1 \leq t) &= 1 - P(X_1 > t) \\ &= 1 - P(\text{no event in } [0, t]) \\ &= 1 - \frac{\lambda^0 e^{-\lambda}}{0!} \\ &= 1 - e^{-\lambda} \\ &= 1 - e^{-\alpha t}, \text{ where } \lambda = \alpha t \end{aligned}$$

Deriving via Exponential Distribution

$$\begin{aligned} P(X_1 \leq t) &= 1 - P(X_1 > t) \\ &= 1 - \int_t^\infty \alpha e^{-\alpha x} dx \\ &= 1 - \left[\frac{\alpha}{-\alpha} e^{-\alpha x} \right]_t^\infty \\ &= 1 - e^{-\alpha t} \end{aligned}$$

The rate of occurrence α in the Poisson distribution is the parameter of the exponential distribution.

7.3 Memoryless Property of Exponential Distribution

- Distribution of waiting time until a certain event does not depend on how much time has elapsed
- e.g. $P(\text{bulb can last for 600 h}) = P(\text{bulb can last for 900 h} \mid \text{bulb can last for 300 h})$

7.4 Standard Normal Distribution

- Parameters: mean $\mu = 0$, variance $\sigma^2 = 1$
- Abbreviated $Z \sim N(0, 1)$
- pdf of Z:

$$f(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}, \quad -\infty < z < \infty$$

- cdf of Z:

$$\Phi(z) = P(Z \leq z) = \int_{-\infty}^z f(u) \, du$$

– Result can be found using standard normal table

7.4.1 z_α Notation

- Denotes value on the z axis for which α of the area under the z curve lies to the **RIGHT** of z_α
- $100(1 - \alpha)$ th percentile of the standard normal distribution

7.5 Standardizing A Normal Distribution

- Normal RV: $X \sim N(\mu, \sigma^2)$
- Standard Normal RV: $Z = \frac{X - \mu}{\sigma}$
- Similarly,

$$\begin{aligned} P(a \leq X \leq b) &= P\left(\frac{a - \mu}{\sigma} \leq \frac{X - \mu}{\sigma} \leq \frac{b - \mu}{\sigma}\right) \\ &= \Phi\left(\frac{b - \mu}{\sigma}\right) - \Phi\left(\frac{a - \mu}{\sigma}\right) \end{aligned}$$

8 W4: Joint Probability Distribution

8.1 Joint Probability Mass Function

The joint probability mass function $p(x, y)$ is defined for each pair of numbers (x, y) by

$$p(x, y) = P(X = x \text{ and } Y = y)$$

It must satisfy the following conditions:

1. $p(x, y) \geq 0$
2. $\sum_x \sum_y p(x, y) = 1$

The probability $P[(X, Y) \in A]$ is obtained by summing the joint pmf over pairs in A :

$$P[(X, Y) \in A] = \sum_{(x, y) \in A} p(x, y)$$

8.2 Marginal Probability Mass Function

The marginal probability mass function of x , $p_X(x)$ is given by

$$p_X(x) = \sum_{y: p(x, y) > 0} p(x, y) \text{ for each possible value of } x.$$

Similarly, the marginal probability mass function of y , $p_Y(y)$ is given by

$$p_Y(y) = \sum_{x: p(x, y) > 0} p(x, y) \text{ for each possible value of } y.$$

- The word "marginal" indicates that the pmf is obtained from the joint probability distribution.
- We can obtain the marginal pmf from the joint pmf, however the reverse is not always true.

8.3 Joint Probability Density Function

The joint probability density function $f(x, y)$ for two different RV is satisfies two conditions:

1. $f(x, y) \geq 0$
2. $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = 1$

For any two dimensional set A , where $a \leq x \leq b$, $c \leq y \leq d$,

$$\begin{aligned} P[(X, Y) \in A] &= \iint_A f(x, y) dx dy \\ &= \int_a^b \int_c^d f(x, y) dx dy \end{aligned}$$

- $P[(X, Y) \in A]$ is the volume beneath the surface above the region A

8.4 Marginal Probability Density Function

The marginal probability density function of X and Y, denoted by $f_X(x)$ and $f_Y(y)$ respectively, are

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy, \quad -\infty < x < \infty$$
$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx, \quad -\infty < y < \infty$$

- Marginal pdf of X is the pdf of X
- The word "marginal" indicates that the pdf is obtained from the joint probability distribution.
- We can obtain the marginal pdf from the joint pdf, however the reverse is not always true.

8.5 Multiple Random Variables

If X_1, X_2, \dots, X_n are all discrete RVs, the joint pmf of the variables is

$$p(x_1, x_2, \dots, x_n) = P(X_1 = x_1, X_2 = x_2, \dots, X_n = x_n)$$

If X_1, X_2, \dots, X_n are all continuous RVs, the joint pdf of the variables with intervals $[a_1, b_1], \dots, [a_n, b_n]$ is

$$P(a_1 \leq X_1 \leq b_1, a_2 \leq X_2 \leq b_2, \dots, a_n \leq X_n \leq b_n)$$
$$= \int_{a_1}^{b_1} \int_{a_2}^{b_2} \dots \int_{a_n}^{b_n} f(x_1, x_2, \dots, x_n) dx_n \dots dx_2 dx_1$$

8.6 Independence of Random Variables

Two RVs X and Y are said to be independent if for **every pair** of x and y values:

$$p(x, y) = p_X(x) \cdot p_Y(y) \quad \text{for discrete RV}$$

$$f(x, y) = f_X(x) \cdot f_Y(y) \quad \text{for continuous RV}$$

If the above is not satisfied for all (x, y), then X and Y are dependent.

9 W5: Conditional Distribution

9.1 Conditional Probability Mass Function

Let X and Y be two discrete RVs with pmf $p(x, y)$.

For any value x for which $p(x) > 0$, the conditional probability mass function of Y given that $X = x$ is

$$p_{Y|X}(y | x) = \frac{p(x, y)}{p_X(x)}$$

where $p_X(x)$ is the marginal pmf of X .

9.2 Conditional Probability Density Function

Let X and Y be two continuous RVs with pdf $f(x, y)$. For any value x for which $f(x) > 0$, the conditional probability density function of Y given that $X = x$ is

$$f_{Y|X}(y | x) = \frac{f(x, y)}{f_X(x)}$$

where $f_X(x)$ is the marginal pdf of X .

9.3 Conditional Distribution

- The summation of the conditional pmf or pdf over the entire sample space is 1.

$$\begin{aligned} \sum_y p_{Y|X}(y | x) &= 1 \quad \text{for discrete RVs } X \text{ and } Y \\ \int_{-\infty}^{\infty} f_{Y|X}(y | x) dy &= 1 \quad \text{for continuous RVs } X \text{ and } Y \end{aligned}$$

9.4 Conditional Expectation

Let X and Y be jointly distributed RVs with pmf $p(x, y)$ or pdf $f(x, y)$. The expected value of a function $h(X, Y)$, denoted by $E[h(X, Y)]$ or $\mu_{h(X, Y)}$ is given by

$$E[h(X, Y)] = \begin{cases} \sum_x \sum_y h(x, y) p(x, y) & \text{for discrete RVs } X \text{ and } Y \\ \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(x, y) f(x, y) & \text{for continuous RVs } X \text{ and } Y \end{cases}$$

9.5 Conditional Mean

Let X and Y be jointly distributed RVs with pmf $p(x, y)$ or pdf $f(x, y)$. The conditional mean of Y , given that $X = x$, denoted by $\mu_{Y|x}$ is given by

$$\mu_{Y|x} = E(Y | x) = \begin{cases} \sum_y y p(y | x) & \text{for discrete RVs } X \text{ and } Y \\ \sum_y h(y) f(y | x) dy & \text{for continuous RVs } X \text{ and } Y \end{cases}$$

9.6 Conditional Variance

Let X and Y be jointly distributed RVs with pmf $p(x, y)$ or pdf $f(x, y)$. The conditional mean of Y , given that $X = x$, denoted by $\sigma_{Y|x}^2$ is given by

$$\begin{aligned}\sigma_{Y|x}^2 &= E\{[Y - E(Y | x)]^2\} \\ &= E(Y^2 | x) - [E(Y | x)]^2\end{aligned}$$

9.7 Law of Total Expectation

If X is a RV, and Y is a RV in the same probability space, then

$$E[E(X | Y)] = E(X)$$

i.e. expected value of the conditional expected value of X given Y is the = expected value of X

9.8 Covariance

The covariance between two variables X and Y , denoted by $\sigma_{X,Y}$ is given by

$$\begin{aligned}\sigma_{X,Y} &= K(X, Y) = E[(X - \mu_x)(Y - \mu_y)] \\ &= \begin{cases} \sum_x \sum_y (x - \mu_x)(y - \mu_y) p(x, y) & \text{for discrete RVs } X \text{ and } Y \\ \int_x \int_y (x - \mu_x)(y - \mu_y) f(x, y) dx dy & \text{for continuous RVs } X \text{ and } Y \end{cases}\end{aligned}$$

- Shortcut formula: $K(X, Y) = E(XY) - E(X)E(Y)$
- Value of covariance:
 - Positive $\sigma_{X,Y}$: positive linear relationship between X and Y
 - Near-zero $\sigma_{X,Y}$: no linear relationship between X and Y
 - Negative $\sigma_{X,Y}$: negative linear relationship between X and Y

9.9 Correlation

- Correlation coefficient $\rho_{X,Y}$: measure of degree of linear relationship between two RVs X and Y

$$\rho_{X,Y} = \tilde{K}(X,Y) = \frac{K(X,Y)}{\sigma_X \sigma_Y}$$

- It is always true that $-1 \leq \rho_{X,Y} \leq 1$
- If X and Y are independent, then $\rho_{X,Y} = 0$
 - **BUT** $\rho_{X,Y}$ does not imply independence between X and Y
- Measure of linear relationship:
 - $|\rho| = 1$: Strong linear relationship between X and Y
 - $|\rho| \neq 1$: Not completely linear relationship between X and Y ; could be strong non-linear relationship
 - $\rho = 0$: X and Y are uncorrelated

10 W5: Central Limit Theorem

10.1 Linear Combination of One RV

For a linear combination of one RV X , denoted by $aX + b$, the mean and variance are as follows:

- Mean, $E(aX + b) = aE(X) + b$
- Variance, $V(aX + b) = a^2V(X)$

10.2 Linear Combination of Two RVs

For a linear combination of two RVs X and Y , where $W = aX + bY$, the mean and variance are as follows:

	X, Y independent	X, Y dependent
Mean, $E(W)$	$aE(X) + bE(Y)$	
Variance, $V(W)$	$a^2V(X) + b^2V(Y)$	$a^2V(X) + b^2V(Y) + 2abK(X, Y)$

10.3 Linear Combination of Multiple RVs

For a linear combination of multiple RVs X_1, X_2, \dots, X_n , where $W = \sum_{i=1}^n a_i x_i$, the mean and variance are as follows:

	RVs independent	RVs dependent
Mean, $E(W)$	$\sum_{i=1}^n a_i E(X_i)$	
Variance, $V(W)$	$\sum_{i=1}^n a_i^2 V(X_i)$	$\sum_{i=1}^n a_i^2 V(X_i) + 2 \sum_{i=1}^n \sum_{j=i+1}^n a_i a_j K(X_i, X_j)$

10.4 Linear Combination of Independent and Identically Distributed RVs

For a linear combination of independent and identically distributed (iid) RVs X_1, X_2, \dots, X_n where $W = \sum_{i=1}^n X_i$ with mean μ and variance σ^2 , the mean and variance are as follows:

- Mean, $E(W) = \sum_{i=1}^n E(X_i) = \sum_{i=1}^n \mu = n\mu$
- Variance, $V(W) = \sum_{i=1}^n V(X_i) = \sum_{i=1}^n \sigma^2 = n\sigma^2$

10.5 Linear Combination of Normal RVs

For two normal RVs X and Y , where $X \sim N(\mu_X, \sigma_X^2)$ and $Y \sim N(\mu_Y, \sigma_Y^2)$, the linear combination $W = X + Y$ is also a normal RV with mean $\mu_X + \mu_Y$ and variance $\sigma_X^2 + \sigma_Y^2$, i.e.

$$W \sim N(\mu_X + \mu_Y, \sigma_X^2 + \sigma_Y^2)$$

10.6 Sample Mean

Let X_1, X_2, \dots, X_n be iid RVs with mean μ and variance σ^2 .

The sample mean \bar{X} can be calculated using the formula $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$.

The mean and variance of \bar{X} is as follows:

- Mean, $E(\bar{X}) = \mu$
- Variance, $V(\bar{X}) = \frac{\sigma^2}{n}$

10.7 Central Limit Theorem

Let X_1, X_2, \dots, X_n be a random sample from a distribution with mean μ and variance σ^2 . The sample mean \bar{X} can be calculated using the formula $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$.

For a sufficiently large n , i.e. $n \leq 30$, \bar{X} has approximately a normal distribution with mean $E(\bar{X})$ and variance $V(\bar{X})$ as follows:

- Mean, $E(\bar{X}) = \mu$
- Variance, $V(\bar{X}) = \frac{\sigma^2}{n}$

If the distribution is close to a normal pdf, a small n yields a good approximation to a normal distribution.