机器学习 Machine learning

第二章 贝叶斯学习 Bayesian Learning

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第二章 贝叶斯学习

- 2.1 概述
- 2.2 贝叶斯决策论
- 2.3 贝叶斯分类器
- 2.4 贝叶斯学习与参数估计问题

预备知识

贝叶斯分类器:基于 Bayesian 决策的分类器

变量和参数:

类别 C: $C = \{c_1, c_2, \dots c_M\}$,

数据 D 和样本 x: $D=\{x_i\}$

贝叶斯学习

$$P(c_i \mid \boldsymbol{x}) \propto P(\boldsymbol{x} \mid c_i) P(c_i)$$

核心是估计

$$P(c_i \mid \mathbf{x}) \propto P(\mathbf{x} \mid c_i) P(c_i)$$

预备知识

贝叶斯决策

类别相似性函数:

$$g_{i}(\mathbf{x}) = p(c_{i} \mid \mathbf{x}) = \frac{p(\mathbf{x} \mid c_{i})p(c_{i})}{\sum_{j=1}^{c} p(\mathbf{x} \mid c_{j})p(c_{j})}$$

$$g_{i}(\mathbf{x}) = p(\mathbf{x} \mid c_{i})p(c_{i})$$

$$g_{i}(\mathbf{x}) = \ln p(\mathbf{x} \mid c_{i}) + \ln p(c_{i})$$

决策函数:

$$g(\mathbf{x}) = g_1(\mathbf{x}) - g_2(\mathbf{x})$$

$$g(\mathbf{x}) = p(c_1 \mid \mathbf{x}) - p(c_2 \mid \mathbf{x})$$

$$g(\mathbf{x}) = \ln \frac{p(\mathbf{x} \mid c_1)}{p(\mathbf{x} \mid c_2)} + \ln \frac{p(c_1)}{p(c_2)}$$

预备知识

贝叶斯分类器

- 朴素贝叶斯分类器: 假设P(x|c)中x特征向量的各维属性独立;
- 半朴素贝叶斯分类器: 假设P(x|c)中x的各维属性存在依赖;
- 正态分布的贝叶斯分类器: 假设 $P(x|c(\theta))$ 服从正态分布;

朴素贝叶斯分类器

采用了"属性条件独立性假设"

$$P(c \mid \mathbf{x}) = \frac{P(c)P(\mathbf{x} \mid c)}{P(\mathbf{x})} \propto P(c)P(\mathbf{x} \mid c) = P(c)\prod_{i=1}^{d} P(x_i \mid c)$$

关键问题: 由训练样本学习类别条件概率和类别先验概率

朴素贝叶斯分类器

采用了"属性条件独立性假设"

$$P(c \mid \mathbf{x}) = \frac{P(c)P(\mathbf{x} \mid c)}{P(\mathbf{x})} \quad \propto \quad P(c)P(\mathbf{x} \mid c) = P(c)\prod_{i=1}^{d} P(x_i \mid c)$$

关键问题: 由训练样本学习类别条件概率和类别先验概率

需要学习的概率分布?

k 个类别,d 个属性: p(c)和 $P(x_i|c_j)$, (i=1,...,d 个属性, j=1,...,k)

共 $1 + d^*k$ 个概率分布要统计.

朴素贝叶斯分类器

类别先验概率的估计
$$P(c) = \frac{|D_c|}{|D|}$$

类别概率密度估计

• *x_i* 离散情况:

$$P(x_i \mid c) = \frac{\left| D_{c,x_i} \right|}{\left| D_c \right|}$$

 $P(x_i | c) = \frac{|D_{c,x_i}|}{|D|}$ D_{c,x_i} 表示 D_c 中在第 i 个属性上取值为 x_i 的样本组成的集合;

 x_i 连续情况:

$$P(x_i \mid c) = \frac{1}{\sqrt{2\pi}\sigma_{c,i}} \exp\left(-\frac{\left(x_i - \mu_{c,i}\right)^2}{2\sigma_{c,i}^2}\right)$$
 (由某一概率分布估计类别概率)

朴素贝叶斯分类器

学习过程

(1) 类别先验估计
$$P(c) = \frac{|D_c|}{|D|}$$

(2) 类别条件概率估计
$$P(x_i | c) = \frac{\left| D_{c,x_i} \right|}{\left| D_c \right|}$$

朴素贝叶斯分类器

决策过程

- (1) 类别先验估计 $P(c) = \frac{|D_c|}{|D|}$
- (2) 类别条件概率估计 $P(x|c) = \prod_{i=1}^{d} P(x_i|c)$
- (3) 贝叶斯决策 $h(x) = \underset{c \in y}{\operatorname{argmax}} P(c) \prod_{i=1}^{a} P(x_i | c)$

朴素贝叶斯分类器

拉普拉斯平滑

避免因训练集样本不充分而导致概率估计值为零.

避免
$$P(c|x) \propto P(c) \prod_{i=1}^{d} P(x_i|c)$$
中, $P(c)$ 或 $P(x_i|c)$ 为 0 (即 $|D_c| = 0$ 或 $|D_{c,xi}| = 0$)

进行拉普拉斯平滑

$$\hat{P}(c) = \frac{|D_c|+1}{|D|+N}, \text{ N为类别数}$$

$$\hat{P}(x_i|c) = \frac{|D_{c,x_i}|+1}{|D_c|+N_i}, N_i \text{为} x_i \text{的可能取值个数}$$

朴素贝叶斯分类器

例子:《机器学习》教材 p.151

• 数据: pp.84

• 离散属性: 色泽、根蒂、敲声、纹理、脐部、触感;

连续属性:密度、含糖绿

类别属性: 好瓜? (是与否)

朴素贝叶斯分类器

• 训练数据

编号	色泽	根蒂	敲声	纹理	脐部	触感	密度	含糖率	好瓜
1	青绿	蜷缩	浊响	清晰	凹陷	硬滑	0.697	0.460	是
2	乌黑	蜷缩	沉闷	清晰	凹陷	硬滑	0.774	0.376	是
3	乌黑	蜷缩	浊响	清晰	凹陷	硬滑	0.634	0.264	是
4	青绿	蜷缩	沉闷	清晰	凹陷	硬滑	0.608	0.318	是
5	浅白	蜷缩	浊响	清晰	凹陷	硬滑	0.556	0.215	是
6	青绿	稍蜷	浊响	清晰	稍凹	软粘	0.403	0.237	是
7	乌黑	稍蜷	浊响	稍糊	稍凹	软粘	0.481	0.149	是
8	乌黑	稍蜷	浊响	清晰	稍凹	硬滑	0.437	0.211	是
9	乌黑	稍蜷	沉闷	稍糊	稍凹	硬滑	0.666	0.091	否
10	青绿	硬挺	清脆	清晰	平坦	软粘	0.243	0.267	否
11	浅白	硬挺	清脆	模糊	平坦	硬滑	0.245	0.057	否
12	浅白	蜷缩	浊响	模糊	平坦	软粘	0.343	0.099	否
13	青绿	稍蜷	浊响	稍糊	凹陷	硬滑	0.639	0.161	否
14	浅白	稍蜷	沉闷	稍糊	凹陷	硬滑	0.657	0.198	否
15	乌黑	稍蜷	浊响	清晰	稍凹	软粘	0.360	0.370	否
16	浅白	蜷缩	浊响	模糊	平坦	硬滑	0.593	0.042	否
17	青绿	蜷缩	沉闷	稍糊	稍凹	硬滑	0.719	0.103	否

测试数据

编号	色泽	根蒂	敲声	纹理	脐部	触感	密度	含糖率	好瓜
測1	青绿	蜷缩	浊响	清晰	凹陷	硬滑	0.697	0.460	?

朴素贝叶斯分类器

学习过程:

(1) 类别先验估计

首先估计类别先验概率 P(c) ,显然有

$$P(好瓜 = 是) = \frac{8}{17} \approx 0.471,$$

$$P(好瓜=否) = \frac{9}{17} \approx 0.529$$
.

朴素贝叶斯分类器

(2) 类别条件概率估计

对于离散值属性:
$$P(x_i \mid c) = \frac{\left| D_{c,xi} \right|}{\left| D_c \right|} \Rightarrow \begin{cases} P(x_i \mid \mathcal{G} \square = \mathbb{E}) = \frac{\left| D_{\mathcal{G} \square = \mathbb{E},xi} \right|}{\left| D_c \right|} \\ P(x_i \mid \mathcal{G} \square = \mathbb{E}) = \frac{\left| D_{\mathcal{G} \square = \mathbb{E},xi} \right|}{\left| D_c \right|} \end{cases}$$

对于连续值属性:

$$P(x_{i}|c) = \frac{1}{\sqrt{2\pi}\sigma_{c,i}} \exp\left(-\frac{(x_{i} - \mu_{c,i})^{2}}{2\sigma^{2}_{c,i}}\right) \Rightarrow \begin{cases} P(x_{i}|) \notin \mathbb{Z} = \mathbb{E} = \frac{1}{\sqrt{2\pi}\sigma_{\mathcal{H}\mathbb{Z}} = \mathbb{E},i} \exp\left(-\frac{(x_{i} - \mu_{\mathcal{H}\mathbb{Z}} = \mathbb{E},i)^{2}}{2\sigma^{2}_{\mathcal{H}\mathbb{Z}} = \mathbb{E},i}\right) \\ P(x_{i}|) \notin \mathbb{Z} = \mathbb{E} = \frac{1}{\sqrt{2\pi}\sigma_{\mathcal{H}\mathbb{Z}} = \mathbb{E},i} \exp\left(-\frac{(x_{i} - \mu_{\mathcal{H}\mathbb{Z}} = \mathbb{E},i)^{2}}{2\sigma^{2}_{\mathcal{H}\mathbb{Z}} = \mathbb{E},i}\right) \end{cases}$$

均值、方差作为参数,可用 ML 估计,

$$P_{\text{青绿|L}} = P$$
(色泽 = 青绿|好瓜 = 是) = $\frac{3}{8}$ = 0.375

$$P_{\text{青绿|A}} = P(色泽 = 青绿 | 好瓜 = 否) = \frac{3}{9} = 0.333$$

$$P_{\text{蜷缩是}} = P(根蒂 = 蜷缩 | 好瓜 = 是) = \frac{5}{8} = 0.625$$

$$P_{\text{蜷缩衙}} = P(根蒂 = 蜷缩 | 好瓜 = 否) = \frac{3}{9} = 0.333$$

$$P_{\text{浊响是}} = P($$
 敲声 = 浊响 | 好瓜 = 是) = $\frac{6}{8}$ = 0.750
$$= \frac{1}{\sqrt{2\pi} \cdot 0.129} \exp\left(-\frac{(0.697 - 0.574)^2}{2 \cdot 0.129^2}\right) \approx 1.959$$

$$P_{\text{独响否}} = P($$
敲声 = 浊响 | 好瓜 = 否) = $\frac{4}{9}$ = 0.444

$$P_{\text{清晰是}} = P(纹理 = 清晰 | 好瓜 = 是) = \frac{7}{8} = 0.875$$

$$P_{\text{清晰|否}} = P(纹理 = 清晰 | 好瓜 = 否) = \frac{2}{9} = 0.222$$

$$P_{\text{凹陷|是}} = P(脐部 = 凹陷 | 好瓜 = 是) = \frac{6}{8} = 0.750$$

$$P_{\text{凹陷否}} = P(脐部 = 凹陷 | 好瓜 = 否) = \frac{2}{9} = 0.222$$

$$P_{\text{@滑|}} = P(触感 = 硬滑 | 好瓜 = 是) = \frac{6}{8} = 0.750$$

$$P_{\text{@滑|否}} = P(触感 = 硬滑 | 好瓜 = 否) = \frac{6}{9} = 0.667$$

为每个属性估计条件概率 $P(x_i | c)$

$$P_{\text{密度:0.697}, \mathbb{E}} = p($$
密度 = 0.697 | 好瓜 = 是)

$$V2\pi \cdot 0.129$$
 (2.0.1.)
$$P_{\text{密度:0.6976}} = p(密度 = 0.697 | 好瓜 = 否)$$

$$= \frac{1}{\sqrt{2\pi} \cdot 0.195} \exp\left(-\frac{(0.697 - 0.496)^2}{2 \cdot 0.195^2}\right) \approx 1.203$$

$$P_{\text{含糖:0.460}} = p(含糖率 = 0.460 | 好瓜 = 是)$$

$$= \frac{1}{\sqrt{2\pi} \cdot 0.101} \exp\left(-\frac{(0.460 - 0.279)^2}{2 \cdot 0.101^2}\right) \approx 0.788$$

$$P_{\text{含糖:0.460}} = p$$
(含糖率 = 0.460|好瓜 = 否)

$$= \frac{1}{\sqrt{2\pi} \cdot 0.108} \exp\left(-\frac{(0.460 - 0.154)^2}{2 \cdot 0.108^2}\right) \approx 0.066$$

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朴素贝叶斯分类器

(3) 贝叶斯决策

$$P($$
好瓜 = 是 $) \times P_{\text{青绿|E}} \times P_{\text{蜷缩|E}} \times P_{\text{浊ゅ|E}} \times P_{\text{清晰|E}} \times P_{\text{凹陷|E}} \times P_{\text{夜月|E}} \times p_{\text{密度:0.697|E}} \times p_{\text{含糖:0.460|E}} \approx 0.038$, $P($ 好瓜 = 否 $) \times P_{\text{青绿|E}} \times P_{\text{蜷缩|E}} \times P_{\text{浊ゅ|E}} \times P_{\text{清晰|E}} \times P_{\text{凹陷|E}} \times P_{\text{ცар|E}} \times P_{\text{еве:0.697|E}} \times P_{\text{shape}} \times P_{\text{hape}} \times P_{\text{hape}} \times P_{\text{luple}} \times P_{\text{lup$

由于 0.038>6.80×10⁻⁵, 因此, 朴素贝叶斯分类器将测试样本"测 1"判别为"好瓜"。

Have a break!

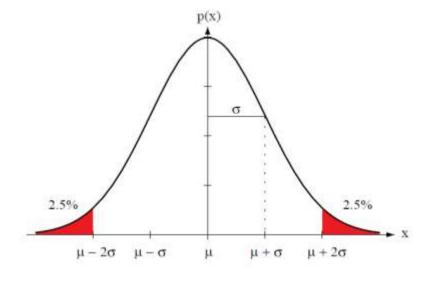
正态密度的贝叶斯分类器

类别条件概率为正态分布

$$h(x) = \underset{c \in y}{\operatorname{argmax}} P(c) P(x|c)$$
 正态分布

• 正态分布的概率密度 $N(\mu, \sigma^2)$:

$$p(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$



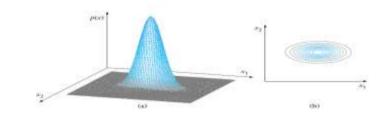
正态密度的贝叶斯分类器

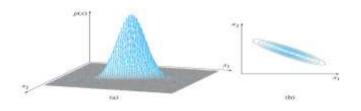
• 多维正态分布的概率密度 $N(\mu, \Sigma)$:

$$p(\mathbf{x}) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp\left[-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu})\right]$$

$$\mu \equiv \varepsilon[x] = \int x p(x) dx$$

$$\Sigma \equiv \varepsilon[(\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^T] = \int (\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^T p(\mathbf{x}) d\mathbf{x}$$





每个维度上都是正态分布 $\mu_i = \varepsilon[x_i]$; $\sigma_{ij} = \varepsilon[(x_i - \mu_i)(x_j - \mu_j)]$

正态密度的贝叶斯分类器

贝叶斯分类:

• 贝叶斯学习(结果取对数):

$$g_{i}(\mathbf{x}) = \ln\left(p\left(\mathbf{x} \mid \omega_{i}\right) p\left(\omega_{i}\right)\right) = \ln p\left(\mathbf{x} \mid \omega_{i}\right) + \ln p\left(\omega_{i}\right)$$

$$p\left(\mathbf{x} \mid \omega_{i}\right) = \frac{1}{(2\pi)^{d/2} \left|\sum_{i}\right|^{1/2}} \exp\left[-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_{i})^{T} \sum_{i}^{-1} (\mathbf{x} - \boldsymbol{\mu}_{i})\right]$$

$$g_{i}(\mathbf{x}) = -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_{i})^{T} \sum_{i}^{-1} (\mathbf{x} - \boldsymbol{\mu}_{i}) - \frac{d}{2} \ln 2\pi - \frac{1}{2} \ln\left|\sum_{i}\right| + \ln p(\omega_{i})$$
(1)

正态密度的贝叶斯分类器

• 决策函数:

$$g_{ij}(x) \equiv g_i(x) - g_j(x)$$

 $g_{ij}(x)=0$ 为决策界 如果 $g_{ij}(x)\geq 0$,则归为i类 如果 $g_{ij}(x)<0$,则归为j类

正态密度的贝叶斯分类器

不同高斯参数情况讨论

Case 1:
$$\sum_{i} = \sigma^{2} I$$

$$\begin{split} g_{i}(x) = & -\frac{1}{2}(x - \mu_{i})^{T} \sum_{i}^{-1}(x - \mu_{i}) + \ln p(\omega_{i}) + c_{i} \\ g_{i}(x) = & -\frac{\left\|x - \mu_{i}\right\|^{2}}{2\sigma^{2}} + \ln p(\omega_{i}) \\ g_{i}(x) = & -\frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i}) + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i}) + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i}) + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i}) + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i}) + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i}) + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i}) + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i}) + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i}) + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i}) + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i}) + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i})}_{1} + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i})}_{1} + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i})}_{1} + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i})}_{1} + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i})}_{1} + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i})}_{1} + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i})}_{1} + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \ln p(\omega_{i})}_{1} + \frac{1}{2\sigma^{2}} \underbrace{\left[x^{T}x\right] - 2\mu_{i}^{T}x + \mu_{i}^{T}\mu_{i}}_{1} + \frac$$

正态密度的贝叶斯分类器

不同高斯参数情况讨论

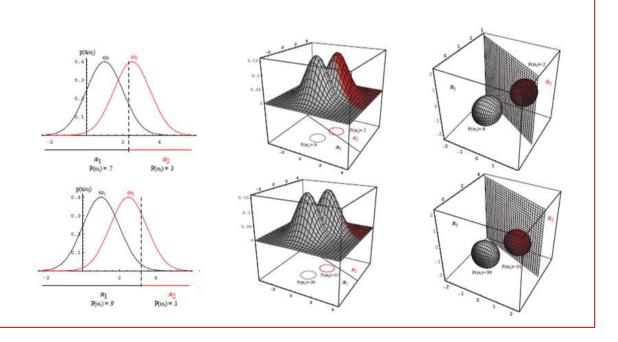
Case 1:
$$\sum_{i} = \sigma^{2} I$$

决策界: $g_i(x) - g_j(x) = 0$

$$w^T(x-x_0)=0$$

$$w = \mu_i - \mu_j$$

$$x_0 = \frac{1}{2} \left(\mu_i + \mu_j \right) - \frac{\sigma^2}{\|\mu_i - \mu_j\|^2} \ln \frac{p(\omega_i)}{p(\omega_j)} (\mu_i - \mu_j)$$



正态密度的贝叶斯分类器

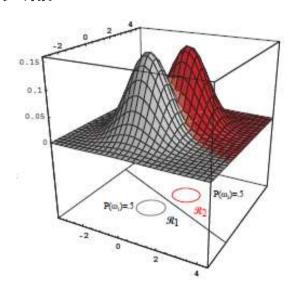
不同高斯参数情况讨论

Case 1:
$$\sum_{i} = \sigma^{2} I$$

特殊情况,当各个类别先验相等时,退化为最小距离分类器。

$$w = \mu_{i} - \mu_{j}$$

$$x_{0} = \frac{1}{2} (\mu_{i} + \mu_{j})$$



正态密度的贝叶斯分类器

不同高斯参数情况讨论

Case
$$2: \sum_{i} = \sum_{i}$$

$$g_i(x) = -\frac{1}{2}(x - \mu_i)^T \sum_{i=1}^{-1} (x - \mu_i) + \ln p(\omega_i)$$

该式分解后: $x^T \sum_i^{-1} x$ 各类都相等,可以忽略

$$g_i(x) = w_i^T x + w_{i0}$$
, $w_i = \Sigma^{-1} \mu_i$, $w_{i0} = \frac{-1}{2} \mu_i^T \Sigma^{-1} \mu_i + \ln p(\omega_i)$

正态密度的贝叶斯分类器

不同高斯参数情况讨论

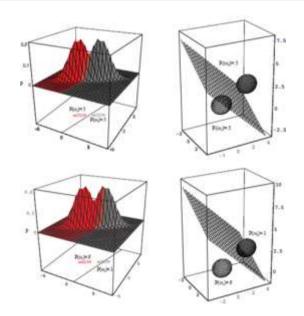
Case
$$2: \sum_{i} = \sum_{i}$$

决策界: $g_i(x) - g_j(x) = 0$

$$w^T(x-x_0)=0$$

$$w = \Sigma^{-1} (\mu_i - \mu_j)$$

$$x_0 = \frac{1}{2} \left(\mu_i + \mu_j \right) - \frac{\ln \left[p(\omega_i) / p(\omega_j) \right]}{(\mu_i - \mu_j)^T \sum^{-1} (\mu_i - \mu_j)} (\mu_i - \mu_j)$$



正态密度的贝叶斯分类器

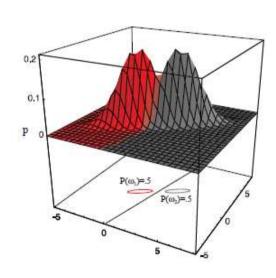
不同高斯参数情况讨论

Case
$$2: \sum_{i} = \sum_{i}$$

当各个类别先验相等时,

$$w = \sum^{-1} (\mu_i - \mu_j)$$

$$x_0 = \frac{1}{2}(\mu_i + \mu_j)$$



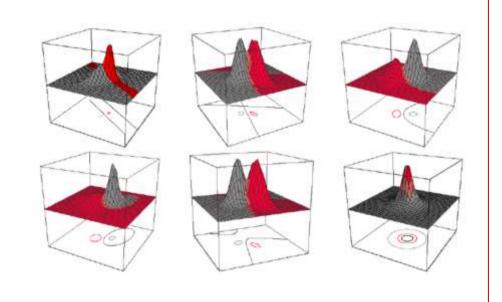
正态密度的贝叶斯分类器

不同高斯参数情况讨论

Case 3: $\sum_{i} = arbitrary$

$$\begin{split} g_{i}(x) &= x^{T} W_{i} x + w_{i}^{T} x + w_{i0} \\ W_{i} &= -\frac{1}{2} \Sigma_{i}^{-1} \\ w_{i} &= \Sigma_{i}^{-1} \mu_{i} \\ w_{i0} &= \frac{-1}{2} \mu_{i}^{T} \Sigma^{-1} \mu_{i} - \frac{1}{2} \ln |\Sigma_{i}| + \ln p(\omega_{i}) \end{split}$$

决策界: $g_i(x)-g_i(x)=0$, 情况比较复杂, 可能非线性。



正态密度的贝叶斯分类器

例子:

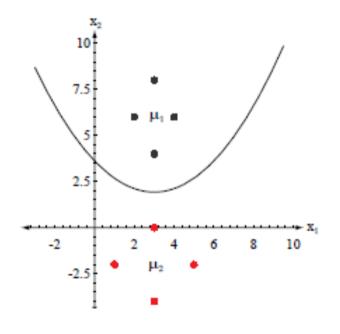
$$\mu_{1} = \begin{bmatrix} 3 \\ 6 \end{bmatrix}; \quad \Sigma_{1} = \begin{pmatrix} 1/2 & 0 \\ 0 & 2 \end{pmatrix} \text{ and } \mu_{2} = \begin{bmatrix} 3 \\ -2 \end{bmatrix}; \quad \Sigma_{2} = \begin{pmatrix} 2 & 0 \\ 0 & 2 \end{pmatrix}$$

$$\Sigma_{1}^{-1} = \begin{pmatrix} 2 & 0 \\ 0 & 1/2 \end{pmatrix} \quad \text{and} \quad \Sigma_{2}^{-1} = \begin{pmatrix} 1/2 & 0 \\ 0 & 1/2 \end{pmatrix}$$

$$p(\omega_{1}) = p(\omega_{2}) = 0.5$$

决策界:
$$g_1(x) \equiv g_2(x)$$

 $x_2 = 3.514 - 1.125x_1 + 0.1875x_1^2$



Have a break!

第二章 贝叶斯学习

- 2.1 概述
- 2.2 贝叶斯决策论
- 2.3 贝叶斯分类器
- 2.4 贝叶斯学习与参数估计问题

问题描述

 \mathcal{D} — data set

 \mathcal{M} — models (or parameters)

The probability of a model M given data set D is:

$$P(\mathcal{M}|\mathcal{D}) = \frac{P(\mathcal{D}|\mathcal{M})P(\mathcal{M})}{P(\mathcal{D})}$$

 $P(\mathcal{D}|\mathcal{M})$ is the **evidence** (or **likelihood**)

 $P(\mathcal{M})$ is the *prior* probability of \mathcal{M}

 $P(\mathcal{M}|\mathcal{D})$ is the posterior probability of \mathcal{M}

 $P(\mathcal{D}) = \int P(\mathcal{D}|\mathcal{M})P(\mathcal{M}) d\mathcal{M}$

三个基本问题: Bayes, MAP and ML

Bayesian Learning:

Assumes a prior over the model parameters. Computes the posterior distribution of the parameters: $P(\theta|\mathcal{D})$.

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 $\times \bigwedge_{i=1}^{n} = \bigwedge_{i=1}^{n}$

Maximum a Posteriori

(MAP) Learning:

Assumes a prior over the model parameters $P(\theta)$. Finds a parameter setting that maximises the posterior: $P(\theta|\mathcal{D}) \propto P(\theta) \ P(\mathcal{D}|\theta)$.

Maximum Likelihood

(ML) Learning:

Does not assume a prior over the model parameters. Finds a parameter setting that maximises the likelihood of the data: $P(D|\theta)$.

贝叶斯学习

通过观测数据 likelihood 修正模型的先验,得到后验概率分布:

$$p(\theta|\mathcal{D}, \alpha) \propto p(\mathcal{D}|\theta)p(\theta|\alpha)$$

其中, α是超参数, 不是估计的参数。

例子 1: Beta 先验分布

贝叶斯学习

例子 1: Beta 先验分布

• 观察数据

```
Coin example: we have a coin that can be biased HHTTHHTHTHTHTHTHHHHHHTHHHHTT 1 1 0 0 1 1 0 1 0 1 0 1 0 1 1 1 1 0 1 1 1 1 0
```

Data: D a sequence of outcomes x_i such that

- head $x_i = 1$
- tail $x_i = 0$

Model: probability of a head θ probability of a tail $(1-\theta)$

贝叶斯学习

例子 1: Beta 先验分布

Binary Variables

$$x \in \{0, 1\}$$
$$p(x = 1|\theta) = \theta$$
$$p(x = 0|\theta) = 1 - \theta$$

・ 贝努力分布(Bernoulli):

$$Bern(x|\theta) = \theta^x (1-\theta)^{1-x}$$

贝叶斯学习

例子 1: Beta 先验分布

Binary Variables

$$x \in \{0, 1\}$$

$$p(x = 1 | \theta) = \theta$$

$$p(x = 0 | \theta) = 1 - \theta$$

• 贝努力分布(Bernoulli):

$$Bern(x|\theta) = \theta^x (1-\theta)^{1-x}$$

• Likelihood (观察似然):

HHTTHHTHTTTTHTHHHHHHHHHH 1 1 0 0 1 1 0 1 0 1 0 0 0 1 0 1 1 1 1 0 1 1 1 1 0

$$P(D \mid \theta) = \prod_{i=1}^{n} \theta^{x_i} (1 - \theta)^{(1 - x_i)}$$

$$P(D \mid \theta) = \theta^{N_1} (1 - \theta)^{N_2} \qquad (N_1 + N_2 = N)$$

贝叶斯学习

例子 1: Beta 先验分布

Binary Variables

$$x \in \{0, 1\}$$
$$p(x = 1|\theta) = \theta$$
$$p(x = 0|\theta) = 1 - \theta$$

• 贝努力分布(Bernoulli):

$$Bern(x|\theta) = \theta^x (1-\theta)^{1-x}$$

• Likelihood (观察似然):

HHTTHHTHTTTTHTHHHHHHHHHH 1 1 0 0 1 1 0 1 0 1 0 0 0 1 0 1 1 1 1 0 1 1 1 1 0

$$P(D \mid \theta) = \prod_{i=1}^{n} \theta^{x_i} (1 - \theta)^{(1 - x_i)}$$

$$P(D \mid \theta) = \theta^{N_1} (1 - \theta)^{N_2} \qquad (N_1 + N_2 = N)$$

(对比)二项式分布:

$$Bin(m|N,\theta) = \binom{N}{m} \theta^m (1-\theta)^{N-m}$$

贝叶斯学习

例子 1: Beta 先验分布

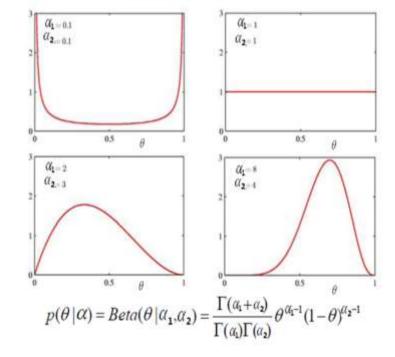
Prior

Choice of prior: Beta distribution

$$p(\theta \mid \alpha) = Beta(\theta \mid \alpha_1, \alpha_2) = \frac{\Gamma(\alpha_1 + \alpha_2)}{\Gamma(\alpha_1)\Gamma(\alpha_2)} \theta^{\alpha_1 - 1} (1 - \theta)^{\alpha_2 - 1}$$

 $\Gamma(x)$ - a Gamma function $\Gamma(x) = (x-1)\Gamma(x-1)$ For integer values of x $\Gamma(n) = (n-1)!$

Beta distribution



贝叶斯学习

例子 1: Beta 先验分布

Prior

Choice of prior: Beta distribution

$$p(\theta \mid \alpha) = Beta(\theta \mid \alpha_1, \alpha_2) = \frac{\Gamma(\alpha_1 + \alpha_2)}{\Gamma(\alpha_1)\Gamma(\alpha_2)} \theta^{\alpha_1 - 1} (1 - \theta)^{\alpha_2 - 1}$$

 $\Gamma(x)$ - a Gamma function $\Gamma(x) = (x-1)\Gamma(x-1)$ For integer values of x $\Gamma(n) = (n-1)!$

Why to use Beta distribution?

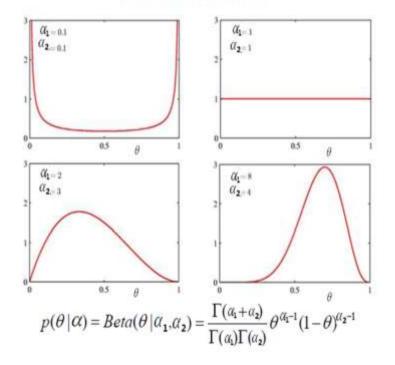
Beta distribution "fits" Bernoulli trials - conjugate choices

$$P(D \mid \theta) = \theta^{N_1} (1 - \theta)^{N_2}$$

Posterior distribution is again a Beta distribution

$$p(\theta \mid D, \alpha) = \frac{P(D \mid \theta) Beta(\theta \mid \alpha_1, \alpha_2)}{P(D \mid \alpha)} = Beta(\theta \mid \alpha_1 + N_1, \alpha_2 + N_2)$$

Beta distribution



贝叶斯学习

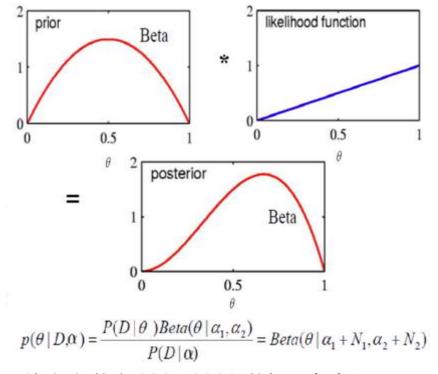
例子 1: Beta 先验分布

Posterior:

$$p(\theta \mid D, \alpha) = \frac{P(D \mid \theta) Beta(\theta \mid \alpha_1, \alpha_2)}{P(D \mid \alpha)} = Beta(\theta \mid \alpha_1 + N_1, \alpha_2 + N_2)$$

$$= \frac{\Gamma(\alpha_1 + \alpha_2 + N_1 + N_2)}{\Gamma(\alpha_1 + N_1)\Gamma(\alpha_2 + N_2)} \theta^{N_1 + \alpha_1 - 1} (1 - \theta)^{N_2 + \alpha_2 - 1}$$
Notice that parameters of the prior act like counts of heads and tails
(sometimes they are also referred to as **prior counts**)

Posterior distribution



极大似然估计

问题描述

• 最大化观察数据的概率

$$p(\theta|\mathcal{D}, \alpha) \propto p(\mathcal{D}|\theta) p(\theta|\alpha)$$
 最大化

似然函数 likelihood:

$$p(\mathcal{D}|\boldsymbol{\theta}) = p(\mathbf{x}_1, \dots, \mathbf{x}_n | \boldsymbol{\theta}) = \prod_{i=1}^n p(\mathbf{x}_i | \boldsymbol{\theta})$$

Maximum Likelihood

$$\hat{\boldsymbol{\theta}} = \arg \max_{\boldsymbol{\theta}} p(\mathcal{D}|\boldsymbol{\theta})$$

转化为求 log-likelihood 极大的问题

$$\hat{\boldsymbol{\theta}} = \arg \max_{\boldsymbol{\theta}} \sum_{i=1}^{n} \log p(\mathbf{x}_i | \boldsymbol{\theta})$$

求解过程

$$\sum_{i=1}^{n} \nabla_{\boldsymbol{\theta}} \log p(\mathbf{x}_i | \boldsymbol{\theta}) = 0$$

极大似然估计

例子 1: 二项式分布的 ML

likelihood

$$p(\mathcal{D}|\theta) = \prod_{n=1}^{N} p(x_n|\theta) = \prod_{n=1}^{N} \theta^{x_n} (1-\theta)^{1-x_n}.$$

Log-likelihood

$$\ln p(\mathcal{D}|\theta) = \sum_{n=1}^{N} \ln p(x_n|\theta) = \sum_{n=1}^{N} \{x_n \ln \theta + (1 - x_n) \ln(1 - \theta)\}$$

• 最优的参数

$$\theta_{\rm ML} = \frac{1}{N} \sum_{n=1}^{N} x_n$$
 $\theta_{\rm ML} = \frac{m}{N}$

极大似然估计

例子 1: 二项式分布的 ML

- 实例:
- Assume the unknown and possibly biased coin
- Probability of the head is θ
- · Data:

HHTTHHTHTTTTHTHHHHTHHHHT

- **Heads:** 15
- Tails: 10

What is the ML estimate of the probability of head and tail?

Likelihood:
$$P(D \mid \theta) = \theta^{N_1} (1 - \theta)^{N_2}$$

最优的参数:

Head:
$$\theta_{ML} = \frac{N_1}{N} = \frac{N_1}{N_1 + N_2} = \frac{15}{25} = 0.6$$
Tail: $(1 - \theta_{ML}) = \frac{N_2}{N} = \frac{N_2}{N_1 + N_2} = \frac{10}{25} = 0.4$

极大似然估计

例子 2: 高斯分布的 ML-估计 u

Let x_1, x_2, \ldots, x_N be vectors stemmed from a normal distribution with known covariance matrix and unknown mean, that is,

$$p(\mathbf{x}_k; \boldsymbol{\mu}) = \frac{1}{(2\pi)^{l/2} |\Sigma|^{1/2}} \exp\left(-\frac{1}{2} (\mathbf{x}_k - \boldsymbol{\mu})^T \Sigma^{-1} (\mathbf{x}_k - \boldsymbol{\mu})\right)$$

· Log-likelihood:

$$L(\boldsymbol{\mu}) \equiv \ln \prod_{k=1}^{N} p(x_k; \boldsymbol{\mu}) = -\frac{N}{2} \ln((2\pi)^l |\Sigma|) - \frac{1}{2} \sum_{k=1}^{N} (x_k - \boldsymbol{\mu})^T \Sigma^{-1} (x_k - \boldsymbol{\mu})$$

极大似然估计

例子 2: 高斯分布的 ML-估计 u

Taking the gradient with respect to μ , we obtain

$$\frac{\partial L(\boldsymbol{\mu})}{\partial \boldsymbol{\mu}} = \begin{bmatrix} \frac{\partial L}{\partial \mu_1} \\ \frac{\partial L}{\partial \mu_2} \\ \vdots \\ \frac{\partial L}{\partial \mu_l} \end{bmatrix} = \sum_{k=1}^{N} \Sigma^{-1} (\boldsymbol{x}_k - \boldsymbol{\mu}) = 0$$

or

$$\hat{\boldsymbol{\mu}}_{ML} = \frac{1}{N} \sum_{k=1}^{N} x_k$$

极大似然估计

例子 3: 高斯分布的 ML-估计方差

Assume that N data points, x_1, x_2, \ldots, x_N , have been generated by a one-dimensional Gaussian pdf of known mean, μ , but of unknown variance. Derive the ML estimate of the variance.

The log-likelihood function for this case is given by

$$L(\sigma^2) = \ln \prod_{k=1}^{N} p(x_k; \sigma^2) = \ln \prod_{k=1}^{N} \frac{1}{\sqrt{2\pi}\sqrt{\sigma^2}} \exp\left(-\frac{(x_k - \mu)^2}{2\sigma^2}\right)$$
$$= -\frac{N}{2} \ln(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{k=1}^{N} (x_k - \mu)^2$$

Taking the derivative of the above with respect to σ^2 and equating to zero, we obtain

$$-\frac{N}{2\sigma^2} + \frac{1}{2\sigma^4} \sum_{k=1}^{N} (x_k - \mu)^2 = 0 \qquad \qquad \hat{\sigma}_{ML}^2 = \frac{1}{N} \sum_{k=1}^{N} (x_k - \mu)^2$$

Chapter 2 Bayesian Learning

-47- 中国科学院大学网络安全学院 2021-2022 学年研究生课程

最大后验估计

问题描述

求使后验概率最大的模型或参数(θ)。

$$p(\theta|\mathcal{D},\alpha) \propto p(\mathcal{D}|\theta)p(\theta|\alpha)$$
 贝叶斯公式中 最大化

$$p(\theta \mid D, \alpha) = \frac{P(D \mid \theta) \ p(\theta \mid \alpha)}{P(D \mid \alpha)}$$

$$\hat{\boldsymbol{\theta}}_{MAP}$$
: $\frac{\partial}{\partial \boldsymbol{\theta}} p(\theta \mid D, \alpha) = 0$ or $\frac{\partial}{\partial \boldsymbol{\theta}} P(D \mid \theta) p(\theta \mid \alpha) = 0$

最大后验估计

例子 1: Beta 先验分布的 MAP

Maximum a posteriori estimate

- Selects the mode of the **posterior distribution**

$$p(\theta \mid D, \alpha) = \frac{P(D \mid \theta)Beta(\theta \mid \alpha_1, \alpha_2)}{P(D \mid \alpha)} = Beta(\theta \mid \alpha_1 + N_1, \alpha_2 + N_2)$$

$$= \frac{\Gamma(\alpha_1 + \alpha_2 + N_1 + N_2)}{\Gamma(\alpha_1 + N_1)\Gamma(\alpha_2 + N_2)} \theta^{N_1 + \alpha_1 - 1} (1 - \theta)^{N_2 + \alpha_2 - 1}$$
Notice that parameters of the prior

(sometimes they are also referred to as prior counts)

MAP Solution:
$$\theta_{MAP} = \frac{\alpha_1 + N_1 - 1}{\alpha_1 + \alpha_2 + N_1 + N_2 - 2}$$

act like counts of heads and tails

最大后验估计

例子 1: Beta 先验分布的 MAP

- 实例:
- Assume the unknown and possibly biased coin
 - Probability of the head is θ
 - · Data:

HHTTHHTHTTTTHTHHHHTHHHHT

- **Heads:** 15
- Tails: 10
- Assume $p(\theta \mid \alpha) = Beta(\theta \mid 5,5)$

What is the MAP estimate?

$$\theta_{MAP} = \frac{N_1 + \alpha_1 - 1}{N - 2} = \frac{N_1 + \alpha_1 - 1}{N_1 + N_2 + \alpha_1 + \alpha_2 - 2} = \frac{19}{33}$$

与ML比较: $\theta_{ML} = 15/25 = 0.6$, $\theta_{MAP} = 19/33 = 0.5758$

最大后验估计

例子 2: 高斯分布的 MAP-估计 u

Let x_1, x_2, \ldots, x_N be vectors stemmed from a normal distribution with known covariance matrix and unknown mean, that is,

$$p(x_k; \boldsymbol{\mu}) = \frac{1}{(2\pi)^{l/2} |\Sigma|^{1/2}} \exp\left(-\frac{1}{2}(x_k - \boldsymbol{\mu})^T \Sigma^{-1} (x_k - \boldsymbol{\mu})\right)$$

$$p(\mu) = \frac{1}{(2\pi)^{l/2} \sigma_{\mu}^{l}} \exp\left(-\frac{1}{2} \frac{\|\mu - \mu_{0}\|^{2}}{\sigma_{\mu}^{2}}\right)$$

The MAP estimate is given by the solution of

$$\frac{\partial}{\partial \boldsymbol{\mu}} \ln \left(\prod_{k=1}^{N} p(\boldsymbol{x}_{k} | \boldsymbol{\mu}) p(\boldsymbol{\mu}) \right) = 0$$

or, for
$$\Sigma = \sigma^2 I$$
,
$$\sum_{k=1}^N \frac{1}{\sigma^2} (x_k - \hat{\boldsymbol{\mu}}) - \frac{1}{\sigma_{\mu}^2} (\hat{\boldsymbol{\mu}} - \boldsymbol{\mu}_0)$$

$$\hat{\mu}_{MAP} = \frac{\mu_0 + \frac{\sigma_{\mu}^2}{\sigma^2} \sum_{k=1}^{N} x_k}{1 + \frac{\sigma_{\mu}^2}{\sigma^2} N}$$

小 结

- 1. beyas 决策准则
- 2. 几种贝叶斯分类器
- 3. 贝叶斯学习与参数估计问题
 - --Beyas Learning
 - --M L 参数估计
 - --MAP参数估计

-52-

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