

Probabilistic model of random co-localization in a cell

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1 Aim

For biological studies purpose we image a cell in 3D using fluorescent microscopy and we are interested in quantifying the interactions between different types of single molecules. From those images we have no means of quantifying directly the interaction between single molecules, instead we want to see how frequently they localize together in space. Our goal is to build a general model that will modelise the likeliness that a single molecule from distribution i localize together or **co-localize** with a single molecule from distribution j in a cell of volume V **assuming single molecules take random postions within the cells independently of one another**.

Our images are produced through a sequential fish microscope in fixed cells that divides, according to its resolution cells in a set of v voxels localized with their set of coordinates (z,y,x) in the volume V .

2 Co-localization events

2.1 Self co-localization event

We note $C(i,i)$ a **self co-localization** event between two single molecules of distribution i occupy the same voxel.

2.2 Co-localization event

We note $C(i,j)$ a **co-localization** event between two single molecules of distribution i and (j) when a single molecule from distribution i occupy the same voxel as a single molecule from distribution j .

3 Probabilistic model

3.1 Presentation

The process of assigning a position to single molecules amongst the v possible positions can be modeled as a probabilistic game where positions are uniquely numbered balls placed in an urn. Assigning coordinates to a single molecule is drawing a ball in the urn, before each drawn balls are replaced in the urn.

3.2 Unique distribution and self co-localization

To begin, let us consider a system containing a unique distribution \mathbf{I} of single molecules randomly placed amongst the v positions of the cell. The probability that a specific single molecule i is found at position m is the probability to draw the ball numbered m :

$$p_i(X = m) = \frac{1}{v} \quad \forall i \in I; \forall m \in V$$

Then the probability to not draw the specific location m is :

$$p_i(\overline{m}) = \frac{1}{v} \quad \forall i \in I; \forall m \in V$$

After k draws, the probability to never draw a the specific location m follows a binomial law and is :

$$p_k(X_m = 0) = \left(1 - \frac{1}{v}\right)^k \quad k \in \mathbb{N}$$

On the contrary the probability that the location m was drawn at least once is :

$$p_k(X_m \geq 1) = 1 - \left(1 - \frac{1}{v}\right)^k \quad (\text{i})$$

To study self co-localization we are interested to know how many different positions have been drawn in k -trials. To do so let us define the observation variable ϵ .

$$\forall m \in V, I(m) = \begin{cases} 0 & \text{if } X_m = 0 \\ 1 & \text{if } X_m \geq 1 \end{cases}$$

The expected number of **different** positions N_{pos}^{unique} drawn is then the expectancy of ϵ after k draws.

$$N_{uniquepos} = \sum_{m \in V} p(X_m = 0).0 + p(X_m \geq 1).1$$

$$N_{uniquepos} = v\left(1 - \left(1 - \frac{1}{v}\right)^k\right) \quad (\text{ii})$$

The expected number of different positions drawn is, in other words, the expected number of draws **that discovered a new position**. To address self co-localization probability we are interested to know the number of draws that **didn't** discover a new position, in other words, duplicates draws or **self co-localization events**. To do so, we remove from the total number of picks (k) the number of picks that discover a new position.

$$N_{selfcolocalization} = k - v(1 - (1 - \frac{1}{v})^k)$$

$$p_{selfcolocalization} = 1 - \frac{v}{k}(1 - (1 - \frac{1}{v})^k)$$

Notes:

- This is true for any distribution I of single molecule i of abundance k_i (i.e. number of single molecule).

3.3 2 distributions and co-localization probability

We now consider positions assignement of a second distribution I of abundance k_i after the assignement of a first distribution J of abundance k_j with $I \neq J$. The number of unique positions J occupy can be estimated with (ii).

$$N_{uniquepos}^j = v(1 - (1 - \frac{1}{v})^{k_j})$$

A colocalisation event is drawing during assignement of I a position that was already drawn when assigning positions to J . To understand this in our probabilistic game, let us consider again a pool of v uniquely numbered balls where all balls that were drawn while assigning postions to J distribution have been colored in **red**. Again all balls are replaced in the pool after each draw. Then the co-localization probability $p(C(i, j))$ is the probability to draw a red ball.

3.3.1 Probability of co-localization

Drawing ball in the urn is a sequence of independent and uniformly random events. The probability of picking one of the $N_{uniquepos}^j$ red ball amongst the v balls is :

$$p(X \in \mathbb{J}) = \frac{N_{uniquepos}^j}{v}$$

where \mathbb{J} is the set of positions drawn for the J distribution. Again co-localization events follow a bionomial process of sucess probability $p = p(X \in \mathbb{J})$. This co-localization probability can also be interpreted as the co-localization rate of molecules i with molecules j .

We can deduce the expectancy and standard deviation of co-localization events :

$$E(C(i, i)) = k_i \frac{N_{uniquepos}^j}{v} \tag{iii}$$

$$std(C(i, j)) = k_i \frac{N_{uniquepos}^j}{v} (1 - \frac{N_{uniquepos}^j}{v})$$

Notes:

- $E(C(i, j)) \neq C(j, i)$

3.3.2 Number of unique positions occupied by a pair (i,j)

In section 3.1 we deduced the probability that a specific location m to be drawn at least once in k draws as :

$$p_k(X_m \geq 1) = 1 - (1 - \frac{1}{v})^k \tag{i}$$

To know how many red balls have been drawn at least once (**i.e. number of unique pair (i,j)**) let us use again the observation variable I .

$$N_{uniquepos}^{(i,j)} = \sum_{m \in \mathbb{J}} p(X_m \geq 1) \cdot 1 + 0 \cdot \dots$$

$$N_{uniquepos}^{(i,j)} = N_{uniquepos}^j \left(1 - \left(1 - \frac{1}{v}\right)^{k_i}\right) \quad (\text{iv})$$

or :

$$N_{uniquepos}^{(i,j)} = v \left(1 - \left(1 - \frac{1}{v}\right)^{k_j}\right) \left(1 - \left(1 - \frac{1}{v}\right)^{k_i}\right)$$

Notes:

- Though $E(C(i,j)) \neq E(C(j,i))$, explicit writing in last equation shows that $N_{uniquepos}^{(i,j)} = N_{uniquepos}^{(j,i)}$.
- Last equation also seems to show it will be easy to generalise the count of unique position occupied by any combination of molecules c of dimension n as $N_{uniquepos}^c = v \prod_{i \in c} \left(1 - \left(1 - \frac{1}{v}\right)^{k_i}\right)$.

3.4 2 distributions and standard deviation

We have determined expectancies above but in order to perform relevant statistics we need to find standard deviation of our random variables.

3.4.1 Co-localization probability standard deviation

In above section we argue that co-localization events $C(i, j)$ follows a binomial distribution of success probability $p = p(X \in \mathbb{J})$. This reasoning is exact when considering expectancies but inexact when considering standard deviation. Indeed it is exact only if the distribution of singles j is fixed but we would like to extend it for i and j following, individually, random distributions, as described in section 3.2. To do so the calculus of variance for $C(i, j)$ should take into consideration the variance of $N_{uniquepos}^j$ and use the law of total variance, which states :

$$Var(C(i, j)) = E[Var(C|N_{uniquepos}^j)] + Var(E[C|N_{uniquepos}^j])$$

$Var(C|N_{uniquepos}^j)$ is the case where distribution j is fixed and standard deviation can be deduced from binomial law (iii) :

$$Var(C|N_{uniquepos}^j) = np(1-p) = k_i \frac{N_{uniquepos}^j}{v} (1 - \frac{N_{uniquepos}^j}{v})$$

$$E[Var(C|N_{uniquepos}^j)] = k_i (1 - (1 - \frac{1}{v})^{k_i}) (1 - \frac{1}{v})^{k_i}$$

3.4.2 Dependency of any distribution within the volume

Although, distributions i_i are independent, X_m^c and X_l^c are **not independent** which means we can't consider the sum of $I_c(m)$ as a binomial distribution. To understand this let us consider the probability that a single molecule i is found at voxel m and at voxel $l \neq m$. We use well known inclusion-exclusion principle.

$$P(X_m \geq 1 \cap X_l \geq 1) = P(X_m \geq 1) + P(X_l \geq 1) - P(X_m \geq 1 \cup X_l \geq 1)$$

$$P(X_m \geq 1 \cap X_l \geq 1) = 2p(i) - P(X_m \geq 1 \cup X_l \geq 1)$$

The probability that a single molecule localize at l or m is the chance to draw one of 2 voxels out of the volume V .

$$P(m \cup l) = \frac{2}{v}$$

Thus from:

$$P(X_m \geq 1 \cup X_l \geq 1) = 1 - (1 - \frac{2}{v})^{k_i}$$

We can already conclude X_m and X_l are not independent since their intersection is non null, it also highlights that the dependency is made through the volume of

voxels.

Conclusion : All distributions **are** independent from each other, but the occupancy of a voxel **is not** independent of the occupancy of other voxels. This result is important for the computation of variance for the occupancy of voxel, for single molecules or any combination single molecules since it means we must introduce a covariance term.

3.5 Generalisation to n distributions

Let us consider n distributions : i_1, i_2, \dots, i_n . What is the probability that any combination of single molecule c to be found in a unique position m ? Under the assumption of randomness, any distribution i is independent from other distributions, thus the probability to find i_1 and i_2 at any given location is the product of the probabilities of distributions to localize at this location.

$$p(c) = \prod_{i \in c} p(i)$$

with

$$p(i) = 1 - \left(1 - \frac{1}{v}\right)^{k_i} \quad (\text{i})$$

Again to know how many different positions contains a combination of single molecule c , we define the observation variable I .

$$\forall m \in V, I_c(m) = \begin{cases} 0 & \text{if } X_m^c = 0 \\ 1 & \text{if } X_m^c \geq 1 \end{cases}$$