FIT2086 Lecture 2 Probability and Probability Distributions

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Outline

- Random Variables and Probability Distributions
 - Random Variables
 - Expectations of Random Variables
- Statistical Models as Probability Distributions
 - Parametric Probability Distributions
 - Two Probability Results

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Some important notation – refresher

- We will use several bits of set notation in this lecture
 - We use $\{a, b, c\}$ to denote a set with elements a, b and c
 - We use $x \in \mathcal{X}$ to denote that x is an element of the set \mathcal{X}
 - Example: $3 \in \{1, 2, 3, 4, 5\}$
 - We use $A \subseteq \mathcal{X}$ to denote that A is a subset of the set \mathcal{X}
 - Example: $\{2,3,4\} \subseteq \{1,2,3,4,5\}$
- Some important sets:
 - ullet $\mathbb Z$ is the set of all integers;
 - \mathbb{Z}_+ is the set of non-negative integers;
 - \bullet \mathbb{R} is the set of all real numbers:
 - \mathbb{R}_+ is the set of non-negative numbers.

Random Variables (1)

- A random variable (RV) is a variable that takes on a value from a set of possible values with specified probabilities
 - ullet We can let ${\mathcal X}$ denote the possible set of values
 - ullet For now, let's just consider cases where ${\mathcal X}$ is discrete
- We often use capital letters to denote a random variable
- Example: let X be a random variable over $\mathcal{X} = \{1, 2, 3\}$ with:

$$X = \begin{cases} 1 & \text{with probability } 1/2 \\ 2 & \text{with probability } 1/4 \\ 3 & \text{with probability } 1/4 \end{cases}$$

Random Variables (2)

- ullet A *realisation* of a random variable is a particular value from ${\mathcal X}$ drawn at random
- ullet Consider our example distribution over $\mathcal{X}=\{1,2,3\}$ with:

$$X = \left\{ \begin{array}{ll} 1 & \text{with probability } 1/2 \\ 2 & \text{with probability } 1/4 \\ 3 & \text{with probability } 1/4 \end{array} \right.,$$

Twenty-two sample realisations are:

$$3, 3, 1, 3, 2, 1, 1, 1, 2, 3, 3, 2, 1, 3, 3, 2, 1, 2, 1, 2, 1, 1$$

- There are nine 1s, six 2s and seven 3s
 - We would expect 1s to appear more frequently the more realisations we take

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Probability Distributions (1)

- We use the language of probability distributions to describe random variables
- The notation

$$\mathbb{P}(X=x), \ x \in \mathcal{X}$$

describes the probability that the RV X takes on the value x from \mathcal{X} .

ullet We can use this notation to describe the example random variable X from the previous slides

$$\mathbb{P}(X=1) = 1/2, \ \mathbb{P}(X=2) = 1/4, \ \mathbb{P}(X=3) = 1/4$$

Probability Distributions (2)

- Review of facts regarding probability distributions
- Fact 1: A probability distribution satisfies:

$$\mathbb{P}(X=x) \in [0,1]$$
 for all $x \in \mathcal{X}$

and

$$\sum_{x \in \mathcal{X}} \mathbb{P}(X = x) = 1$$

Probability Distributions (3)

• Fact 2: The probability of $(X \in A_1 \text{ OR } X \in A_2)$, with $A_1,A_2 \subset \mathcal{X}$ and $A_1 \cap A_2 = \emptyset$ is

$$\mathbb{P}(X \in A_1 \cup A_2) = \mathbb{P}(X \in A_1) + \mathbb{P}(X \in A_2),$$

with " \cap " set intersection, " \cup " set union and \emptyset the empty set.

• Example: If X follows the probability distribution

$$\mathbb{P}(X=1) = 1/2, \ \mathbb{P}(X=2) = 1/4, \ \mathbb{P}(X=3) = 1/4$$

then $\mathbb{P}(X \geq 2)$ is

$$\mathbb{P}(X \in \{2\} \cup \{3\}) = \mathbb{P}(X = 2) + \mathbb{P}(X = 3)$$

$$= 1/4 + 1/4$$

$$= 1/2$$

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Probability Distributions of Two RVs (1)

- Now let us consider the case of two RVs $X \in \mathcal{X}$ and $Y \in \mathcal{Y}$
 - ullet ${\mathcal X}$ and ${\mathcal Y}$ are the sets of values X and Y can take, respectively
 - ullet $\mathcal{X} imes \mathcal{Y}$ is the set of values the pair can assume
- Example: If $\mathcal{X}=\{1,2,3\}$ and $\mathcal{Y}=\{1,2\}$, then $\mathcal{X}\times\mathcal{Y}=\{\{1,1\},\{2,1\},\{3,1\},\{1,2\},\{2,2\},\{3,2\}\}$

• Example: An example distribution over $\mathcal{X} \times \mathcal{Y}$:

	X = 1	X = 2	X = 3
Y = 1	0.05	0.15	0.1
Y = 2	0.25	0.15	0.3

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$$\mathcal{X} \times \mathcal{Y} = \{\{1,1\},\{2,1\},\{3,1\},\{1,2\},\{2,2\},\{3,2\}\}$$

• Example: An example distribution over $\mathcal{X} \times \mathcal{Y}$:

$$X = 1$$
 $X = 2$ $X = 3$
 $Y = 1$ 0.05 0.15 0.1
 $Y = 2$ 0.25 0.15 0.3

Probability Distributions of Two RVs (2)

• We can define a probability distribution over (X,Y) as before:

$$\mathbb{P}(X=x,Y=y) \in [0,1]$$
 for all $x \in \mathcal{X}, y \in \mathcal{Y}$

which satisfies

$$\sum_{x \in \mathcal{X}, y \in \mathcal{Y}} \mathbb{P}(X = x, Y = y) = 1$$

- $\mathbb{P}(X=x,Y=y)$ is the *joint* probability of X=x and Y=y• That is, the probability of X=x AND Y=y
- Example: The example distribution from previous slide

$$\mathbb{P}(X = 1, Y = 1) = 0.05$$

 $\mathbb{P}(X = 1, Y = 2) = 0.25$

$$\mathbb{P}(X=2, Y=1) = 0.15$$

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and so on.



The Sum Rule (1)

The Sum Rule

The sum rule is given by:

$$\mathbb{P}(X = x) = \sum_{y \in \mathcal{Y}} P(X = x, Y = y)$$

The probability $\mathbb{P}(X=x)$ is called the *marginal* probability.

 \bullet The marginal probability $\mathbb{P}(X=x)$ is the probability of seeing X=x irrespective of what value Y takes on

The Sum Rule (2)

Example:

$$X = 1$$
 $X = 2$ $X = 3$
 $Y = 1$ 0.05 0.15 0.1
 $Y = 2$ 0.25 0.15 0.3

Then

$$\mathbb{P}(Y=1) = 0.05 + 0.15 + 0.1 = 0.3$$

 $\mathbb{P}(Y=2) = 0.25 + 0.15 + 0.3 = 0.7$

so that the probability of seeing a Y=2 is significantly higher than the probability of seeing a Y=1, irrespective of the value of X.

Conditional Probability (1)

Conditional Probability

$$\mathbb{P}(X = x \mid Y = y) = \frac{\mathbb{P}(X = x, Y = y)}{\mathbb{P}(Y = y)}$$

The probability $\mathbb{P}(X=x\,|\,Y=y)$ is called the probability of X=x, conditional on Y=y.

• The conditional probability $\mathbb{P}(X=x\,|\,Y=y)$ is the (joint) probability of seeing X=x and Y=y, divided by the (marginal) probability that we have observed Y=y.

Conditional Probability (2)

• Example:

$$X = 1$$
 $X = 2$ $X = 3$
 $Y = 1$ 0.05 0.15 0.1
 $Y = 2$ 0.25 0.15 0.3

Then

$$\mathbb{P}(X = 1 | Y = 1) = \mathbb{P}(X = 1, Y = 1) / \mathbb{P}(Y = 1)$$
$$= 0.05 / 0.3 \approx 0.1667$$

and

$$\mathbb{P}(X=1 \,|\, Y=2) = \mathbb{P}(X=1,Y=2)/\mathbb{P}(Y=2) \\ = 0.25/0.7 \approx 0.3571$$

so that seeing X=1 is twice as likely when Y=2 as compared to the case that Y=1.



Independent Random Variables (4)

- Independent random variables are very important
- X and Y are considered independent if

$$\mathbb{P}(X = x, Y = y) = \mathbb{P}(X = x)\mathbb{P}(Y = y)$$

for all $x \in \mathcal{X}$, $y \in \mathcal{Y}$.

This implies that

$$\mathbb{P}(X = x \mid Y = y) = \mathbb{P}(X = x).$$

- \Rightarrow Knowing about Y tells us nothing new about X
- An even more special class are independent and identically distributed (i.i.d.) random variables
 - $X_1 \in \mathcal{X}, X_2 \in \mathcal{X}$ are i.i.d. if they are independent and



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- An even more special class are independent and identically distributed (i.i.d.) random variables
 - $X_1 \in \mathcal{X}$, $X_2 \in \mathcal{X}$ are i.i.d. if they are independent and $\mathbb{P}(X_1 = x) = \mathbb{P}(X_2 = x)$ for all $x \in \mathcal{X}$

Continuous Random Variables (1)

- So far we have considered only discrete random variables
- The ideas extend to the case that the values X can take on form a continuum, that is, $\mathcal{X} \subseteq \mathbb{R}$
- X now follows a probability density function (pdf) p(x).
- A pdf satisfies:

$$p(x) \ge 0$$
 for all $x \in \mathcal{X}$

and

$$\int_{\mathcal{X}} p(x)dx = 1$$

Continuous Random Variables (2)

• The probability that X lies in an interval (a, b) is

$$\mathbb{P}(a < X < b) = \int_{a}^{b} p(x)dx.$$

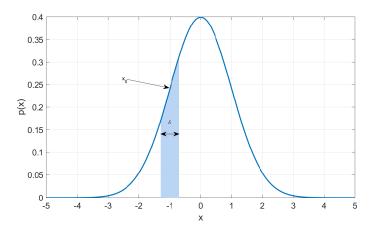
• More generally, the probability $X \in A$, where $A \subset \mathcal{X}$ is

$$\mathbb{P}(X \in A) = \int_A p(x)dx.$$

- This implies that $\mathbb{P}(X=x)=0$ ⇒ One of the most confusing aspects of continous RVs

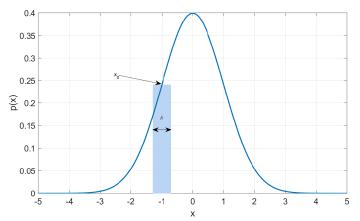
Continuous Random Variables (3)

• Example: Probability of $(x_0 - \delta/2 < X < x_0 + \delta/2)$



Continuous Random Variables (4)

• If δ is small enough then $\int_{x_0-\delta/2}^{x_0+\delta/2} p(x) dx \approx p(x_0) \delta$ \Rightarrow Take $\delta \to 0$ and it is clear why $\mathbb{P}(X=x)=0$.



Cumulative Distribution Functions (1)

 The cumulative distribution function (cdf) of a continuous RV is:

$$\mathbb{P}(X \le x) = \int_{-\infty}^{x} p(x')dx'$$

that is, the probability that X is less than some value x

• Let's introduce some shorthand notation for discrete RVs:

$$\mathbb{P}(X=x) \equiv p(x)$$

ullet Then, if X is a discrete RV over the integers (or a subset)

$$\mathbb{P}(X \le x) = \sum_{x' \le x} p(x')$$

It follows that

$$\mathbb{P}(X > x) = 1 - \mathbb{P}(X \le x)$$

Cumulative Distribution Functions (2)

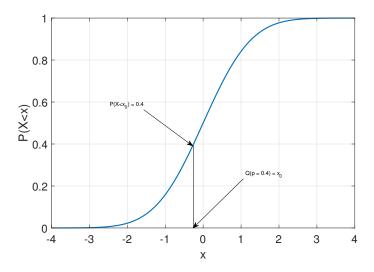
The inverse cdf is

$$Q(p) = \{ x \in \mathcal{X} : \mathbb{P}(X \le x) = p \}$$

which is sometimes called the quantile function.

- In words, the quantile function says: find the the value x such that the probability that $X \le x$ is p
- For example:
 - Q(p=1/2) is the median;
 - Q(p = 1/4) is the first quartile; and
 - Q(p = 3/4) is the third quartile.

Cumulative Distribution Functions (3)



Expected Values (1)

 Given a distribution, we can define the expected value of the RV:

$$\mathbb{E}\left[X\right] = \sum_{x \in \mathcal{X}} x \, p(x)$$

recalling that $p(x) \equiv \mathbb{P}(X = x)$.

- The expected value is the average value over \mathcal{X} , weighted by the probability of each particular $x \in \mathcal{X}$ appearing.
- For continuous RVs, replace the sum with an integral:

$$\mathbb{E}\left[X\right] = \int x \, p(x) dx$$

• Example: $\mathbb{P}(X=1) = 0.5$, $\mathbb{P}(X=2) = 0.4$, $\mathbb{P}(X=3) = 0.1$:

$$\mathbb{E}[X] = 1 \cdot 0.5 + 2 \cdot 0.4 + 3 \cdot 0.1 = 1.6$$

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Expected Values (2)

More generally:

$$\mathbb{E}\left[f(X)\right] = \sum_{x \in \mathcal{X}} f(x)p(x)$$

where f(x) is any function of x.

• Example: $\mathbb{P}(X=1) = 0.5$, $\mathbb{P}(X=2) = 0.4$, $\mathbb{P}(X=3) = 0.1$:

$$\mathbb{E} \left[\log X \right] = \log 1 \cdot 0.5 + \log 2 \cdot 0.4 + \log 3 \cdot 0.1 = 0.3871$$

where $\log x$ is the natural logarithm (sometimes called \ln).

Variance (1)

This lets us define important properties such as the variance:

$$\mathbb{V}[X] = \mathbb{E}\left[(X - \mathbb{E}[X])^2\right]$$
$$= \sum_{x \in \mathcal{X}} (x - \mathbb{E}[X])^2 p(x)$$

- \Rightarrow The expected squared deviation around the mean
- ullet The larger $\mathbb{V}\left[X
 ight]$ the more variation around the mean
- The standard deviation is equal to $\sqrt{\mathbb{V}[X]}$.
- Example: $\mathbb{P}(X=1) = 0.5$, $\mathbb{P}(X=2) = 0.4$, $\mathbb{P}(X=3) = 0.1$; recall that in this case, $\mathbb{E}[X] = 1.6$, so:

$$V[X] = (1 - 1.6)^2 \cdot 0.5 + (2 - 1.6)^2 \cdot 0.4 + (3 - 1.6)^2 \cdot 0.1 = 0.44$$

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Variance (2)

A useful alternative expression for variance is:

$$V[X] = \mathbb{E}\left[(X - \mathbb{E}[X])^2\right]$$

$$= \mathbb{E}\left[X^2 - 2X\mathbb{E}[X] + \mathbb{E}[X]^2\right]$$

$$= \mathbb{E}\left[X^2\right] - 2\mathbb{E}[X]\mathbb{E}[X] + \mathbb{E}[X]^2$$

$$= \mathbb{E}\left[X^2\right] - \mathbb{E}[X]^2$$

where the third step follows from properties of sums/integrals

- \bullet Variance is sum of expected squared value of X, minus square of expected value of X
 - ⇒ Use this to find variance for our example on previous slide

Covariance/Correlation (1)

• For two variables X and Y we can define the covariance:

$$cov (X, Y) = \mathbb{E} [(X - \mathbb{E} [X])(Y - \mathbb{E} [Y])]$$
$$= \mathbb{E} [XY] - \mathbb{E} [X] \mathbb{E} [Y]$$

and from this, we can define the correlation:

$$\operatorname{corr}(X, Y) = \frac{\operatorname{cov}(X, Y)}{\sqrt{\mathbb{V}[X] \mathbb{V}[Y]}}$$

 \Rightarrow Compare to the sample correlation formula in Lecture 1.

Covariance/Correlation (2)

- Positive covariance/correlation:
 - \Rightarrow if X greater than $\mathbb{E}[X]$ then likely Y is greater than $\mathbb{E}[Y]$
- Negative covariance/correlation:
 - \Rightarrow if X greater than $\mathbb{E}[X]$ then likely Y is *less* than $\mathbb{E}[Y]$
- Covariance between $(-\infty, \infty)$,
 - ullet Depends on scale (unit of measurement) of variables X and Y
- Correlation between [-1, 1],
 - Independent of scale of variables
- If X, Y independent, cov(X,Y) = corr(X,Y) = 0
 - ⇒ Converse is not true

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- If X, Y independent, cov(X,Y) = corr(X,Y) = 0 \Rightarrow Converse is **not** true!

Expectations and Independent RVs

In general, expectation of a function of two RVs is

$$\mathbb{E}\left[f(X,Y)\right] = \sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} f(x,y)p(x,y)$$

Fact 1: Due to linearity of expectation, we have

$$\mathbb{E}\left[f(X) + g(Y)\right] = \mathbb{E}\left[f(X)\right] + \mathbb{E}\left[g(Y)\right]$$

for all RVs X and Y, and

Fact 2: For independent RVs, we have

$$\mathbb{E}\left[f(X)f(Y)\right] = \mathbb{E}\left[f(X)\right]\mathbb{E}\left[g(Y)\right]$$

implying that

$$\mathbb{V}\left[X+Y\right] = \mathbb{V}\left[X\right] + \mathbb{V}\left[Y\right]$$

for X and Y independent.



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Existence of Expected Values

- Expected values do not always exist
- ullet If ${\mathcal X}$ is *finite*, then ${\mathbb E}\left[X
 ight]$ always exists
- ullet However, in general, ${\mathcal X}$ will not be finite
- $\mathcal X$ is usually the set of integers $\mathbb Z$ or real numbers $\mathbb R$ \Rightarrow In this case, expectations are not guaranteed to exist
- In contrast, the quantiles (such as median) always exist

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- 2 Statistical Models as Probability Distributions
 - Parametric Probability Distributions
 - Two Probability Results

Probability Distributions as Models

- We can use probability distributions as models of reality
- ullet For example, imagine we knew the distribution of heights of people in Australia, say p(h)
- We could then make predictions/statements about heights
 - ullet For example, the proportion of people taller than 1.7m

$$\mathbb{P}(H > 1.7) = \int_{1.7}^{\infty} p(h)dh$$

 Then a company clothing could use this information to determine how much product of different sizes to stock

Parametric Probability Distributions (1)

- So far we have built probability distributions by directly specifying the probabilities for each element $x \in \mathcal{X}$
- This is fine if X is a small finite set
- But if \mathcal{X} is large, or infinite (for example, all the integers), this approach no longer works
- Instead it is usual to use parametric probability distributions
- We will look at several important distributions:
 - The Gaussian distribution;
 - The Bernoulli distribution;
 - The binomial distribution;
 - The Poisson distribution.

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Parametric Probability Distributions (2)

We specify the probability density function by

$$p(x \mid \boldsymbol{\theta}), \ x \in \mathcal{X}, \ \boldsymbol{\theta} \in \Theta$$

or, for discrete RVs we use the shorthand notation:

$$\mathbb{P}(X = x \mid \boldsymbol{\theta}) \equiv p(x \mid \boldsymbol{\theta}), \ x \in \mathcal{X}, \ \boldsymbol{\theta} \in \Theta$$

where

- $\theta = (\theta_1, \dots, \theta_k)$ are the parameters that control distribution of the probabilities;
- \bullet Θ is the set of valid parameters for the model.
- \Rightarrow by changing θ we can change the distribution.
- ullet Usually, the number of parameters $k \ll |\mathcal{X}|$

Parametric Probability Distributions (3)

- The properties of the RV are determined by $p(x | \theta)$
- For example, the mean is

$$\mathbb{E}\left[X\right] = f(\boldsymbol{\theta}),$$

where $f(\cdot)$ is a function that depends on θ and $p(x \mid \theta)$.

- The same applies to the variance, cdf, quantiles, etc.
- The parameterisation will not be unique
 ⇒ there are often several common parameterisations for the same distribution

Gaussian Distribution (1)

- ullet Let's begin with the case that $\mathcal{X}=\mathbb{R}$
 - ⇒ that is, we want a distribution over all the real numbers
- Probably the most important distribution for real numbers is the Gaussian (normal) distribution
 - ⇒ named after Carl Friedrich Gauss (1777-1855)
- The pdf for a Gaussian distribution is given by

$$p(x \mid \mu, \sigma^2) = \left(\frac{1}{2\pi\sigma^2}\right)^{\frac{1}{2}} \exp\left(-\frac{(x-\mu)^2}{\sigma^2}\right)$$

where

- \bullet μ is the mean of the distribution;
- σ^2 is the variance of the distribution;

so that $\theta = (\mu, \sigma^2)$ for the Gaussian distribution.

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 ⇒ named after Carl Friedrich Gauss (1777-1855)
- The pdf for a Gaussian distribution is given by

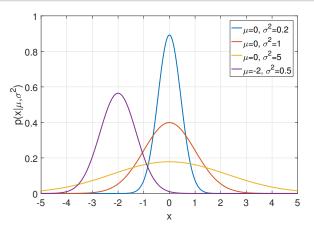
$$p(x \mid \mu, \sigma^2) = \left(\frac{1}{2\pi\sigma^2}\right)^{\frac{1}{2}} \exp\left(-\frac{(x-\mu)^2}{\sigma^2}\right)$$

where

- \bullet μ is the mean of the distribution;
- σ^2 is the variance of the distribution;

so that $\theta = (\mu, \sigma^2)$ for the Gaussian distribution.

Gaussian Distribution (2)



Probability density functions for several normal (Gaussian) distributions. The orange curve is the *standard normal distribution*. Note that the normal distribution is symmetric and tails off to zero as $|x| \to \infty$.

Gaussian Distribution (3)

If X follows a Gaussian distribution, we write that

$$X \sim N(\mu, \sigma^2)$$

where " \sim " is read as "is distributed per a"

- An important property of Gaussian RVs is self-similarity
- \bullet Every Gaussian distribution is a translated and scaled version of the standard normal distribution N(0,1)
- If $Z \sim N(0,1)$, then

$$X = \sigma Z + \mu$$

is distributed as per $N(\mu, \sigma^2)$

Gaussian Distribution (4)

• If $X \sim N(\mu, \sigma^2)$, then

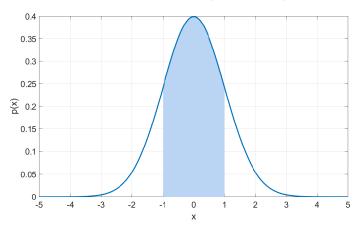
$$\mathbb{E}[X] = \mu,$$

$$\mathbb{V}[X] = \sigma^2.$$

- The Gaussian distribution is symmetric around μ , so that:
 - its mode is μ ;
 - \bullet its median is μ .
- The cdf for the Gaussian has no closed form
 - Most packages have algorithms to evaluate it numerically
 - Some well-known rules regarding the cdf are ...

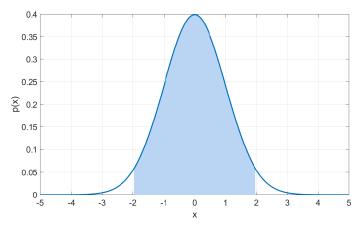
Gaussian Distribution (5)

- For any $N(\mu, \sigma^2)$:
 - 68.27% of probability falls within $(\mu \sigma, \mu + \sigma)$



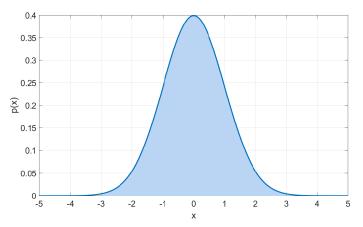
Gaussian Distribution (6)

- For any $N(\mu, \sigma^2)$:
 - 95.45% of probability falls within $(\mu-2\sigma,\mu+2\sigma)$



Gaussian Distribution (7)

- For any $N(\mu, \sigma^2)$:
 - 99.73% of probability falls within $(\mu 3\sigma, \mu + 3\sigma)$



Bernoulli Distribution (1)

- ullet Let's consider the case of discrete, binary RVs, i.e., $\mathcal{X}=\{0,1\}$
- The Bernoulli distribution models these variables

$$\mathbb{P}(X=1\,|\,\theta)=\theta,\;\theta\in[0,1]$$

so that the parametric probability distribution follows:

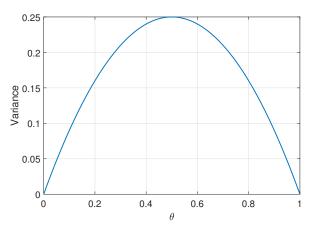
$$p(x \mid \theta) = \theta^x (1 - \theta)^{(1-x)}$$

- ullet The parameter heta is the probability of observing a "success"
- If X follows a Bernoulli distribution, we write $X \sim \text{Be}(\theta)$
- It is easy to see that

$$\mathbb{E}[X] = \theta$$

$$\mathbb{V}[X] = \theta(1 - \theta)$$

Bernoulli Distribution (2)



Variance of a Bernoulli random variable as a function of θ . The variance is maximum when $\theta=1/2$ and smallest for $\theta=0$ and $\theta=1$.

Binomial Distribution (1)

- Now consider n binary RVs $\mathbf{X} = (X_1, \dots, X_n)$.
 - Example realisation: $\mathbf{x} = (0, 1, 1, 1, 0, 1, 0, 0, 1, 1)$
- The sum

$$m(\mathbf{x}) \equiv m = \sum_{j=1}^{n} x_j$$

counts the number of "successes"

- \Rightarrow in our example, m=6
- Given n, the count is a RV, say M, over the sample space $\{0,1,2,\ldots,n\}$

Binomial Distribution (2)

ullet The binomial distribution describes the probability that M takes a particular value m

$$p(m \mid \theta) = \binom{n}{m} \prod_{i=1}^{n} p(x_i \mid \theta) = \binom{n}{m} \theta^m (1 - \theta)^{(n-m)}$$

where

$$\binom{n}{m} = \frac{n!}{(n-m)!m!}$$

is the number of ways of choosing m objects out of n identical objects (the binomial coefficient)

- $\Rightarrow m! = 1 \times 2 \times 3 \times \ldots \times m$ is the factorial function
- This captures the fact that, for $1 \le m \le (n-1)$ there are multiple sequences with m successes out of n trials

Binomial Distribution (3)

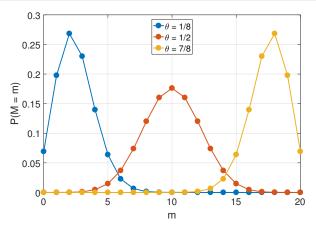
• Example: The following six sequences have m=2 successes out of n=4 trials:

$$1, 1, 0, 0$$
 $1, 0, 1, 0$
 $1, 0, 0, 1$
 $0, 1, 1, 0$
 $0, 1, 0, 1$
 $0, 0, 1, 1$

so that

$$p(m = 2 \mid \theta) = \binom{n = 4}{m = 2} \theta^2 (1 - \theta)^{(4-2)}$$

Binomial Distribution (4)



Binomial distribution for n=20 and $\theta=1/8$, $\theta=1/12$, $\theta=7/8$. The distribution is defined only on the integers – the connecting lines are only guides for the eye. Note that $\theta=7/8$ is a mirror of $\theta=1/8$.

Binomial Distribution (5)

• If M follows a binomial distribution, we write

$$M \sim \text{Bin}(\theta, n)$$

ullet As m is a sum of independent Bernoulli RVs we have

$$\mathbb{E}[M] = n\theta$$

$$\mathbb{V}[M] = n\theta(1-\theta)$$

Poisson Distribution (1)

- What if our data is non-negative integers; for example:
 - number of telephone calls made in an hour
 - number of people kicked to death by horses in a year
 - ullet Sample space is then $\mathbb{Z}_+ = \{0,1,2,\ldots\}$
- One suitable distribution is the Poisson distribution
 - Named after Simeon Poisson (1781–1840)
- Has the form

$$p(k \mid \lambda) = \frac{\lambda^k \exp(-\lambda)}{k!}$$

where λ is often called the *rate*.

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Poisson Distribution (2)

ullet If X is distributed per a Poisson distribution we write

$$X \sim \text{Pois}(\lambda)$$

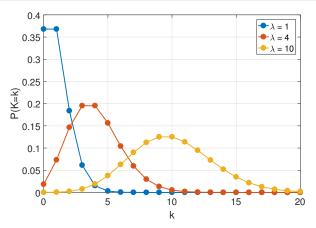
The Poisson distribution has

$$\mathbb{E}[X] = \lambda$$

$$\mathbb{V}[X] = \lambda$$

 The Poisson distribution is an example of a distribution in which the variance grows with the mean

Poisson Distribution (3)



Poisson distribution for $\lambda=1$, $\lambda=4$ and $\lambda=10$. The distribution is defined only on the integers – the connecting lines are only guides for the eye.

Poisson Distribution (4)

- The Poisson distribution models the number of events in an interval of time
- When is the Poisson appropriate?
 - The occurrence of one event does not affect the probability that a second event will occur. That is, events occur independently.
 - The rate at which events occur is constant. The rate cannot be higher in some intervals and lower in other intervals.
 - Two events cannot occur at exactly the same instant.
 - The probability of an event in a small interval is proportional to the length of the interval.

(taken from Wikipedia)

Chebyshev's Inequality (1)

- Named after P. Chebyshev (1821-1894)
- If X is a RV with mean μ and variance σ^2 , then for any k>0

$$\mathbb{P}(|X - \mu| \ge k) \le \frac{\sigma^2}{k^2}$$

 This inequality allows us to compute (bounds on) probabilities even when only the mean and variance are known

Chebyshev's Inequality (2)

- Chebyshev's bound if only $\mathbb{E}[X] = 0$, $\mathbb{V}[X] = 1$ is known:
 - $\mathbb{P}(|X| \ge 1) \le 1$;
 - $\mathbb{P}(|X| \ge 2) \le 0.25$;
 - $\mathbb{P}(|X| \ge 3) \le 0.1112;$
- Compare to the situation that we know $X \sim N(0,1)$:
 - $\mathbb{P}(|X| \ge 1) = 0.3173;$
 - $\mathbb{P}(|X| \ge 2) = 0.0455;$
 - $\mathbb{P}(|X| \ge 3) = 0.0027$.
 - ⇒ Chebyshev's bounds very general but not always accurate.

Weak Law of Large Numbers (1)

- An important application of Chebyshev's inequality is to prove the weak law of large numbers.
- Let X_1,\ldots,X_n be RVs with $\mathbb{E}\left[X_i\right]=\mu$; then for any $\varepsilon>0$

$$\mathbb{P}\left\{\left|\frac{X_1+\cdots+X_n}{n}-\mu\right|>\varepsilon\right\}\to 0 \text{ as } n\to\infty.$$

 Informally, you can think of this result as saying that the (sample) mean of a realisation of random variables converges to the expected value as the number of realisations grows larger and larger.

Reading/Terms to Revise

- Reading for this week: Chapters 4 and 5 of Ross.
- Terms you should know:
 - Random variable;
 - Conditional Probability;
 - Probability density function;
 - Expectations;
 - Variance and co-variance;
 - Normal, Bernoulli, binomial and Poisson distributions
 - Law of large numbers