

The mbbefd Package: A Package for handling MBBEFD exposure curves in R

Giorgio Alfredo Spedicato, Ph.D C.Stat ACAS

Abstract

The package models MBBEFD distribution providing density, quantile, distribution and random generation functions. In addition it provides exposure curves for the MBBEFD distribution family.

Keywords: mbbefd, exposure curves, reinsurance, non-life insurance.

1. Introduction

The **mbbefd** package provides function to use Maxwell-Bolzano, Bose-Einstein, Fermi-Dirac probability distributions, introduced by (BERNEGGER 1997), within R statistical software (R Core Team 2014).

Such kind of distributions are widely used in the pricing of non-life reinsurance contracts and yet they are not present in any R package.

The paper is structured as follows: Section 2 discusses review the theory (mathematics and actuarial application) of MBBEFD distributions, Section 3 shows the package's features, applied examples are shown in Section 4 while the issue of fitting MBBEFD curves to empirical data is discussed in Section 5.

2. A theoretical review

Within actuarial jargon, an exposure curve is a distribution that shows the ratio between the expected limited loss at various limits and the expected unlimited loss. They are usually to rate exposures from large commercial risks and in non-proportional reinsurance treaties for non - life business. In mathematical notation, if IV is the insured value and d the ratio of loss x to IV the exposure curve $G(d)$ is defined as Equation 1 displays.

$$G(d) = \frac{E[\min(d * IV, x)]}{E[x]} = \frac{\int_0^{d*IV} (1 - F(x)) dx}{\int_0^\infty x * f(x) dx} = \frac{\int_0^{d*IV} S(x) dx}{\int_0^\infty S(x) dx} \quad (1)$$

It can be shown that $G(0) = 0$, $G(1) = 1$, $dG(d) \geq 0$ and $ddG(d) \leq 0$. Whilst losses normally lie in the interval $0, \dots, \infty$, it will be assumed x to represent a normalized loss in the interval $0, \dots, PML$, being PML the so - called maximum probable loss (in other words, the

maximum loss it is thought to can happen) all along the paper. Therefore x would represent a percentage loss with respect to a maximum, i.e., a destruction rate.

BERNEGGER (1997) and Mahler (2014) provide a discussion on the actuarial theory regarding such curve. In particular, the curves discussed by BERNEGGER (1997) are of the form expressed by Equation 2.

$$\begin{cases} G(x) = \frac{\ln(a+b^x) - \ln(a+1)}{\ln(a+b) - \ln(a+1)} \\ x \in [0, 1] \\ b > 0 \end{cases} \quad (2)$$

Using some calculus on Equation 2 it can be shown that the expected value is equal to the reciprocal of exposure curve derivative at 0, Equation 3.

$$\mu = \frac{1}{dG(0)} \quad (3)$$

The probability of a total loss, p , is expressed by Equation 4.

$$p = 1 - F(1^-) = \frac{1}{g} = \frac{dG(1)}{dG(0)} = \frac{(a+1) * b}{a+b} \quad (4)$$

It is similarly possible to write expression for the Survival Function, $S(x)$. Equation 5 expresses the Survival Function for a generic exposure curve and Equation 6 expresses the same for a MBBEFD distribution family, and the density, Equation 7.

$$S(x) = \frac{G'(x)}{G'(0)} \quad (5)$$

$$S(x) = \frac{(a+1) b^x}{a+b^x} \quad (6)$$

$$f() = -S'(x) = -\frac{(a+1) * a * \ln(b^x)}{(a+b^x)^2} \quad (7)$$

The MBBEFD can be expressed using an alternative parametrization based on parameter $g = \frac{1}{p}$, the reciprocal of the total loss $p = S(1) = \frac{(a+1)*b}{a+b}$. This equation allows to rewrite MBBEFD formulas using the b and g parameters, taking into account Equation 8. Clearly, $g \geq 1$.

$$a = \frac{(g-1) * b}{1 - g * b} \quad (8)$$

Equation 9 shows the mean as a function of b and g .

$$E[X] = \frac{\ln(g * b)}{\ln(b)} * \frac{1-b}{1-g*b} \quad (9)$$

The limits shown in Equation 10 follows using some calculus on Equation 9.

$$\left\{ \begin{array}{l} \frac{\partial \mu}{\partial g} \leq 0 \\ \frac{\partial \mu}{\partial b} \leq 0 \\ b \rightarrow 0, \mu \rightarrow 1 \\ b \rightarrow \infty, \mu \rightarrow p \\ g \rightarrow 1, \mu \rightarrow 1 \\ g \rightarrow 0, \mu \rightarrow 0 \end{array} \right. \quad (10)$$

Two special cases are defines:

1. $b * g = 1, g > 1$. The distribution is an Exponential, censored from above at 1. Its exposure curve is $G = \frac{1-b^x}{1-b}$.
2. $b = 1, g > 1$. The distribution is a censored Pareto from above at 1. Its exposure curve is $G = \frac{\ln(1+(g-1)*x)}{\ln(g)}$.

3. The MBBEFD class and its related package

```
R> library(mbbefd)
```

The `mbbefdExposure` function evaluates the exposure curve for a given destruction rate x , given either a and b , or b and g . Figure 3 displays the destruction rate by level of x , for an exposure cuve of parameters $a = 0.2$ and $b = 0.04$

4. Applied examples

The curve can be use to price property coverage and associate reinsurance treaties. Suppose a property expected loss to be 40K, MPL to be 2MLN. An XL coverage is available with a retention of 1Mln. The exposure curve that characterize the property is the usual one. Therefore the percentage of loss net and ceded is determined as it follows

```
R> net<-mbbefdExposure(x=1/2, a=0.2,b=0.04)*40000
R> ceded<-40000-net
```

and the expected loss as a percentage of total insured value is

```
R> expectedLoss<-1/dG(x=0,a=0.2,b=0.04)*40000
R> expectedLoss
```

```
[1] 24000
```

Similarly, it is possible to draw the underluying suvival curve $S(x) = \frac{G'(x)}{G'(0)}$ using Figure 2. The probability of a maximum loss for such exposure curve is obtained evaluating the survival function at 1

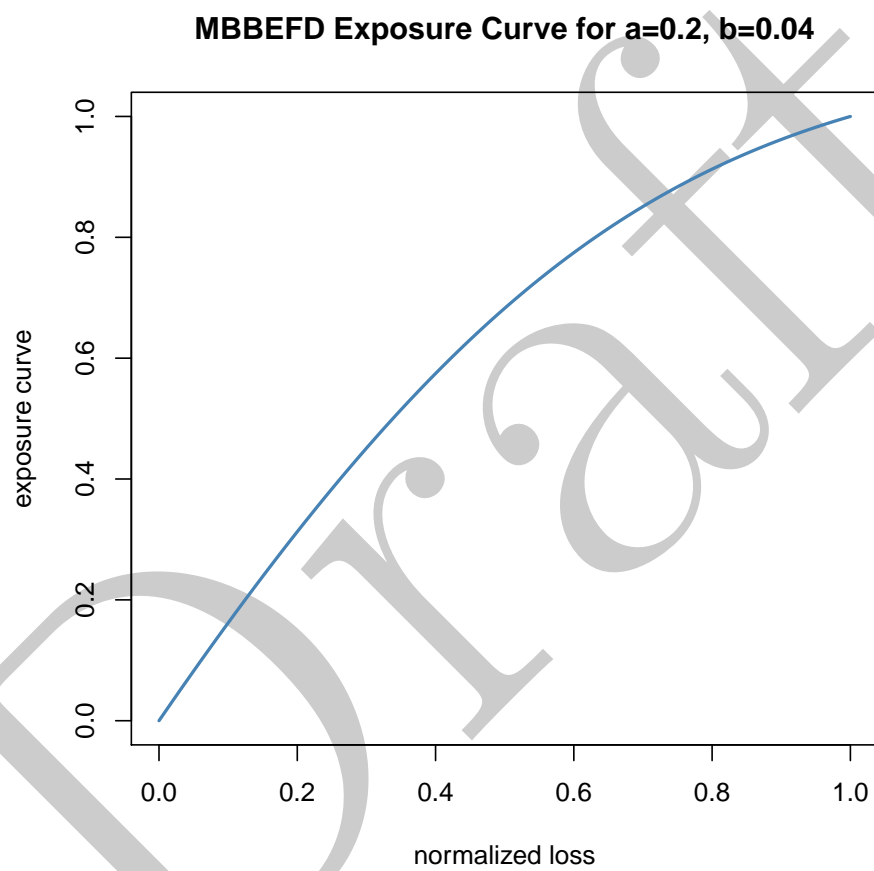


Figure 1: Exposure curve example

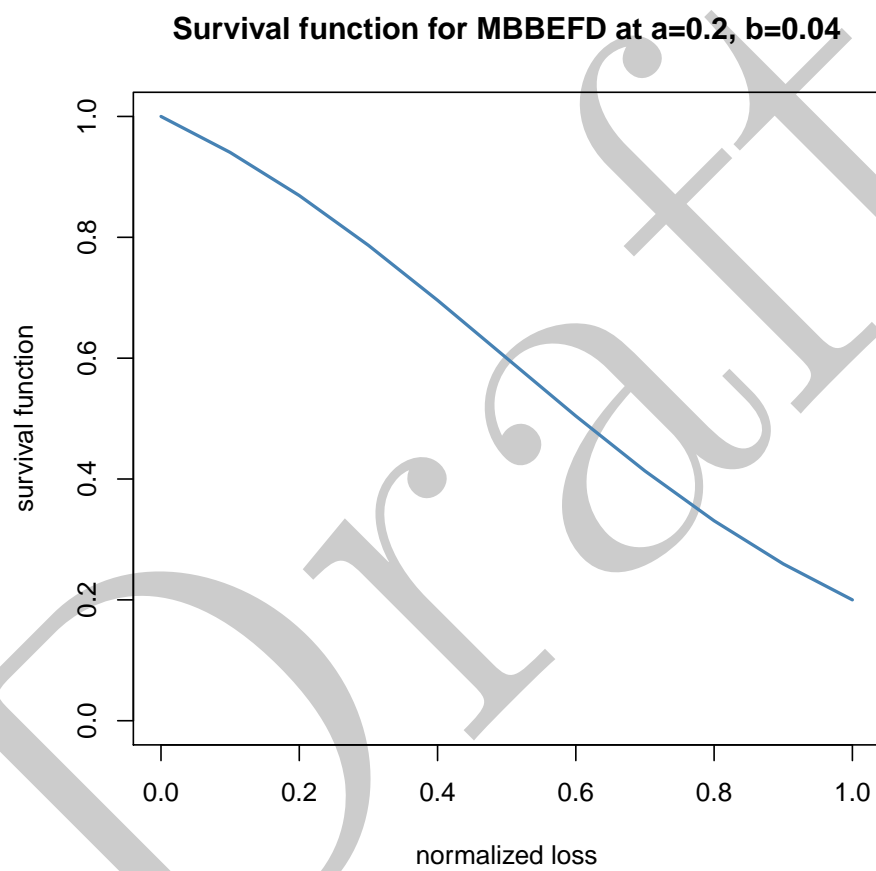


Figure 2: Underlying survival curve

```
R> pTotalLoss<-1-pmbbefd(q=1,a=0.2,b=0.04)
R> pTotalLoss
```

```
[1] 0.2
```

Similarly, it is possible to assess the mean of the distribution underlying the exposure curve. Quantile functions, distribution functions and density functions are defined as well. For example, the 60th percentile of the distribution above defined (i.e., how bad can be in 60% of cases in terms of destruction rate) is

```
R> qmbbefd(p=0.6,a=0.2,b=0.04)
```

```
[1] 0.7153383
```

whilst a loss worse than 80% of IV could happen in

```
R> 100*(1-pmbbefd(q=0.8,a=0.2,b=0.04))
```

```
[1] 33.0895
```

cases out of 100.

It would be possible to simulate variates from the MBBEFD distribution using the random generation command `rmbbefd`.

```
R> simulatedLosses<-rmbbefd(n=10000,a=0.2,b=0.04)
R> mean(simulatedLosses)
```

```
[1] 0.597828
```

```
R> sum(simulatedLosses==1)/length(simulatedLosses)
```

```
[1] 0.1949
```

Finally another way to show the probability of total loss to be greater than zero is to show that the (numerical) integral between 0 and 1 of the density function is lower than 1, that is $1 - F(1^-)$.

```
R> integrate(dmbbefd,lower=0, upper=1, a=0.2, b=0.04)
```

```
0.8 with absolute error < 2.4e-13
```

5. Fitting MBBEFD curves

? suggests an iterative process, based on the method of moments, in order to estimate the parameter of the distribution function, starting from known values of $p = \frac{1}{g}$. The algorithm outlined is:

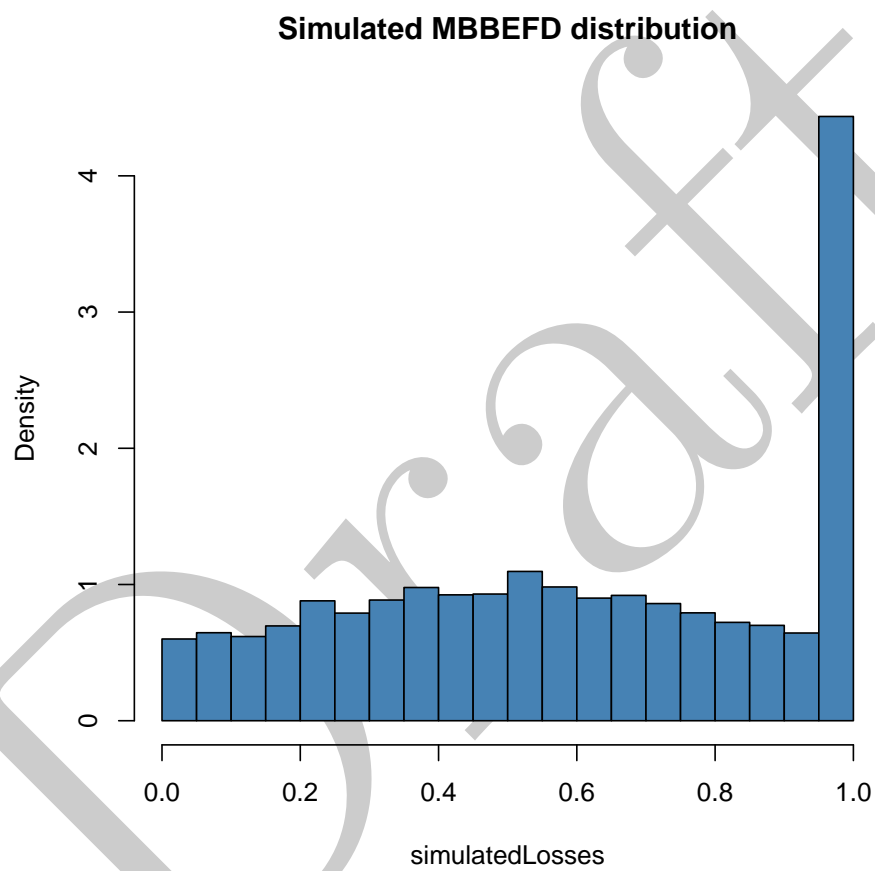


Figure 3: Exposure curve example

1. Try $p_0 = m_2$, being m_2 the second empirical moment. Obtain $g_0 = \frac{1}{p_0}$.
2. Solve for b_0 the equation $E[x] = m_0 = \frac{\ln(g_0*b_0)}{b_0} \frac{1-b_0}{1-g_0*b}$.
3. Get the second theoretical moment, $E[x^2]$ of x from estimated b_0 and g_0 .
4. Compare $E[x^2]$ to the empirical moment. Repeat the process modifying p until the theoretical second moment is close to the empirical one enough (the second moment is an increasing function of p).

Fitting a MBBEFD distribution is not easy. The result is sensible to initial values and appears to be instable. We have applied the first three steps of this process in order to obtain initial estimates of a and b to feed the Maximum Likelihood estimation process using **fitdistrplus** package, Marie Laure, Regis, Jean-Baptiste, and Christophe (2014). We show two example one using both artificial data or real one (from package **copula**, Jun Yan (2007)).

```
R> #get data
R> data1<-rmbbefd(n=1000,a = .2,b=.04)
R> data(loss, package = "copula")
R> data2<-pmin(1,pmax(0,loss$loss/loss$limit)) #capping loss data to lim
R> #functions used to initialize the parameters
R> #using one iteration of Method of Moments
R>
R> #method of moments
R>
R> giveFunction2Minimize<-function(mu,g) {
+   out = function(b) (mu - (log(g*b)*(1 - b))/( log(b)*(1 - g*b)) )^2
+   return(out)
+ }
R> giveFunction2Integrate<-function(b,g) {
+   out = function(x) x^2*dmbbefd(x,b=b,g=g)
+   return(out)
+ }
R> giveInits<-function(x) {
+   m0<-mean(x)
+   m2<-mean(x^2)
+
+   #p<=1/g
+
+   p0=m2 #m2 upper limit of p0
+   g=1/p0
+
+   #equate 1rst moment to get the mean
+   myMin<-giveFunction2Minimize(mu=m0,g=g)
+   b<-nlm(f=myMin,p=.1)$estimate
+
+   #return a
```



```

+   a=(g-1)*b/(1-g*b)
+   out<-list(a=a, b=b)
+   return(out)
+ }
R> ###fitting process
R>
R> library(fitdistrplus)
R> #using close starting points
R> est1<-fitdist(data=data1,distr = "mbbefd",method = "mle",start=list(a=.9,b=.14))
R> est1

```

Fitting of the distribution ' mbbefd ' by maximum likelihood
Parameters:

	estimate	Std. Error
a	0.16570167	0.024238705
b	0.03284038	0.005238313

```

R> #using estimated starting points
R> inits2<-giveInits(x=data2)
R> est2<-fitdist(data=data2,distr = "mbbefd",method = "mle",start=inits2)
R> est1

```

Fitting of the distribution ' mbbefd ' by maximum likelihood
Parameters:

	estimate	Std. Error
a	0.16570167	0.024238705
b	0.03284038	0.005238313

References

- BERNEGGER S (1997). "THE SWISS RE EXPOSURE CURVES AND THE MBBEFD DISTRIBUTION CLASS." *Astin Bulletin*, p. 99.
- Jun Yan (2007). "Enjoy the Joy of Copulas: With a Package copula." *Journal of Statistical Software*, **21**(4), 1–21. URL <http://www.jstatsoft.org/v21/i04/>.
- Mahler H (2014). *Mahler's Guide to Advanced Ratemaking*.
- Marie Laure DM, Regis P, Jean-Baptiste D, Christophe D (2014). *fitdistrplus: Help to Fit of a Parametric Distribution to Non-Censored or Censored Data*. R package version 1.0-3.
- R Core Team (2014). *R: A Language and Environment for Statistical Computing*. R Foundation for Statistical Computing, Vienna, Austria. URL <http://www.R-project.org/>.

Affiliation:

Giorgio Alfredo Spedicato, Ph.D C.Stat ACAS

StatisticalAdvisor

Paderno Dugnano

20037 Italy

E-mail: spedicato_giorgio@yahoo.it

URL: www.statisticaladvisor.com

Draft