

# Bag of Results

**Lemma .1.**

$$q_t = \mathcal{Q} \left( w_t + \frac{X_t^T u_{t-1}}{\|X_t\|_2^2} \right).$$

*Proof.* This follows simply by completing a square. Provided  $X_t \neq 0$ , we have by the definition of  $q_t$

$$\begin{aligned} q_t &= \arg \min_{p \in \{-1, 0, 1\}} \|u_{t-1} + (w_t - p)X_t\|_2^2 = \arg \min_{p \in \{-1, 0, 1\}} (w_t - p)^2 + 2(w_t - p) \frac{X_t^T u_{t-1}}{\|X_{t-1}\|_2^2} \\ &= \arg \min_{p \in \{-1, 0, 1\}} \left( (w_t - p) + \frac{X_t^T u_{t-1}}{\|X_{t-1}\|_2^2} \right)^2 - \left( \frac{X_t^T u_{t-1}}{\|X_{t-1}\|_2^2} \right)^2. \end{aligned}$$

Because the former term is always non-negative, it must be the case that the minimizer is  $\mathcal{Q} \left( w_t + \frac{X_t^T u_{t-1}}{\|X_t\|_2^2} \right)$ .  $\square$

**Lemma .2.** Suppose  $0 < w_t < 1$ , and  $X_t$  a random vector in  $\mathbb{R}^d$ . Then

$$\mathbb{P}_{X_t} \left( (w_t - q_t)^2 + 2(w_t - q_t)X_t^T u_{t-1} > \alpha \mid u_{t-1} \right) = \begin{cases} \mu_y \left( \frac{\alpha - (w_t + 1)^2}{2(w_t + 1)}, \frac{\alpha - (w_t - 1)^2}{2(w_t - 1)} \right) & \alpha < -w_t - w_t^2 \\ \mu_y \left( \frac{\alpha - w_t^2}{2w_t}, \frac{\alpha - (w_t - 1)^2}{2(w_t - 1)} \right) & -w_t - w_t^2 \leq \alpha \leq w_t - w_t^2 \\ 0 & \alpha > w_t - w_t^2 \end{cases}$$

where  $\mu_y$  is the probability measure over  $\mathbb{R}$  induced by the random variable  $y = X_t^T u_{t-1}$ .

*Proof.* Let  $A_b$  denote the event that  $q_t = b$  for  $b \in \{-1, 0, 1\}$ . Then by the law of total probability

$$\mathbb{P} \left( (w_t - q_t)^2 + 2(w_t - q_t)y > \alpha \mid u_{t-1} \right) = \sum_{b \in \{-1, 0, 1\}} \mathbb{P} \left( (w_t - b)^2 + 2(w_t - b)y > \alpha \mid u_{t-1}, A_b \right).$$

- **b = 0**  $q_t = 0$  precisely when  $-1/2 - w_t \leq y \leq 1/2 - w_t$ . So we have

$$\begin{aligned} \mathbb{P} \left( w_t^2 + 2w_t y > \alpha \mid u_{t-1}, A_0 \right) &= \mathbb{P} \left( y > \frac{\alpha - w_t^2}{2w_t} \mid u_{t-1}, -1/2 - w_t \leq y \leq 1/2 - w_t \right) \\ &= \begin{cases} \mu_y(-1/2 - w_t, 1/2 - w_t) & \alpha < -w_t - w_t^2 \\ \mu_y \left( \frac{\alpha - w_t^2}{2w_t}, 1/2 - w_t \right) & -w_t - w_t^2 \leq \alpha \leq w_t - w_t^2 \\ 0 & \alpha > w_t - w_t^2 \end{cases}. \end{aligned}$$

- **b = 1**  $q_t = 1$  precisely when  $y > 1/2 - w_t$ . Noting that  $w_t - 1 < 0$ , we have

$$\begin{aligned} \mathbb{P} \left( (w_t - 1)^2 + 2(w_t - 1)y > \alpha \mid u_{t-1}, A_1 \right) &= \mathbb{P} \left( y < \frac{\alpha - (w_t - 1)^2}{2(w_t - 1)} \mid u_{t-1}, y > 1/2 - w_t \right) \\ &= \begin{cases} \mu_y \left( 1/2 - w_t, \frac{\alpha - (w_t - 1)^2}{2(w_t - 1)} \right) & \alpha \leq w_t - w_t^2 \\ 0 & \alpha > w_t - w_t^2 \end{cases}. \end{aligned}$$

- **b = -1**  $q_t = -1$  precisely when  $y < -1/2 - w_t$ . So we have

$$\begin{aligned} \mathbb{P} \left( (w_t + 1)^2 + 2(w_t + 1)y > \alpha \mid u_{t-1}, A_{-1} \right) &= \mathbb{P} \left( y > \frac{\alpha - (w_t + 1)^2}{2(w_t + 1)} \mid u_{t-1}, y < -1/2 - w_t \right) \\ &= \begin{cases} \mu_y \left( \frac{\alpha - (w_t + 1)^2}{2(w_t + 1)}, -1/2 - w_t \right) & \alpha \leq -w_t - w_t^2 \\ 0 & \alpha > -w_t - w_t^2 \end{cases}. \end{aligned}$$

Summing these three functions yields the result.  $\square$

**Corollary .3.** When  $-1 < w_t < 0$ , we have

$$\mathbb{P}_{X_t} \left( (w_t - q_t)^2 + 2(w_t - q_t) X_t^T u_{t-1} > \alpha \mid u_{t-1} \right) = \begin{cases} \mu_y \left( \frac{\alpha - (w_t + 1)^2}{2(w_t + 1)}, \frac{\alpha - (w_t - 1)^2}{2(w_t - 1)} \right) & \alpha < w_t - w_t^2 \\ \mu_y \left( \frac{\alpha - (w_t + 1)^2}{2(w_t + 1)}, \frac{\alpha - w_t^2}{2w_t} \right) & w_t - w_t^2 \leq \alpha \leq -w_t - w_t^2 \\ 0 & \alpha > w_t - w_t^2 \end{cases}.$$

**Lemma .4.** Let  $X_t \sim \text{Unif}(S^{d-1})$ . Then

$$\mathbb{E}[q_t X_t \mid u_{t-1}] = \frac{\text{Area}(S^{d-2})}{(d-1) \cdot \text{Area}(S^{d-1})} \left( \left( 1 - \min \left\{ \left( \frac{1/2 - w_t}{\|u_{t-1}\|} \right)^2, 1 \right\} \right)^{\frac{d-1}{2}} + \left( 1 - \min \left\{ \left( \frac{1/2 + w_t}{\|u_{t-1}\|} \right)^2, 1 \right\} \right)^{\frac{d-1}{2}} \right) \frac{u_{t-1}}{\|u_{t-1}\|_2},$$

where  $\mu_y$  is the probability measure over  $\mathbb{R}$  induced by the random variable  $y = X_t^T u_{t-1}$ .

*Proof.* By unitary invariance of the  $X_t$ , we may assume that  $u_{t-1} = \|u_{t-1}\|_2 e_1$ . Under this assumption, we have

$$\mathbb{E}[q_t X_t \mid u_{t-1}] = \mathbb{E}[\mathcal{Q}(w_t + \|u_{t-1}\|_2 X_{t,1}) X_t \mid u_{t-1}].$$

Breaking this apart into the events  $A_{-1} := \{X_{t,1} \leq \frac{-1/2 - w_t}{\|u_{t-1}\|}\}$  and  $A_1 := \{X_{t,1} \geq \frac{1/2 - w_t}{\|u_{t-1}\|}\}$ , we note that these regions on  $S^{d-1}$  are symmetric about the  $u_{t-1} = e_1$  axis. So,

$$\mathbb{E}[q_t X_t \mid u_{t-1}] = \mathbb{E}[X_t \mid u_{t-1}, A_1] - \mathbb{E}[X_t \mid u_{t-1}, A_{-1}].$$

Using the hyperspherical coordinates

$$f \left( \begin{bmatrix} \varphi_1 \\ \varphi_2 \\ \vdots \\ \varphi_{d-1} \end{bmatrix} \right) = \begin{bmatrix} \cos(\varphi_1) \\ \sin(\varphi_1) \cos(\varphi_2) \\ \sin(\varphi_1) \sin(\varphi_2) \cos(\varphi_3) \\ \vdots \\ \sin(\varphi_1) \dots \sin(\varphi_{d-2}) \cos(\varphi_{d-1}) \\ \sin(\varphi_1) \dots \sin(\varphi_{d-2}) \sin(\varphi_{d-1}) \end{bmatrix}, \quad (1)$$

we have  $\mathbb{E}[X_{t,1} \mid u_{t-1}, A_1] = 0$  if  $\frac{1/2 - w_t}{\|u_{t-1}\|_2} > 1$ . Otherwise,

$$\begin{aligned} \mathbb{E}[X_{t,1} \mid u_{t-1}, A_1] &= \frac{1}{\text{Area}(S^{d-1})} \int_0^{\arccos(\frac{1/2 - w_t}{\|u_{t-1}\|_2})} \cos(\varphi_1) dA \\ &= \frac{\text{Area}(S^{d-2})}{\text{Area}(S^{d-1})} \int_0^{\arccos(\frac{1/2 - w_t}{\|u_{t-1}\|_2})} \cos(\varphi_1) \sin^{d-2}(\varphi_1) d\varphi_1 = \frac{\text{Area}(S^{d-2})}{(d-1) \text{Area}(S^{d-1})} \left( 1 - \left( \frac{1/2 - w_t}{\|u_{t-1}\|_2} \right)^2 \right)^{\frac{d-1}{2}}. \end{aligned}$$

Similarly,  $\mathbb{E}[X_{t,1} \mid u_{t-1}, A_{-1}] = 0$  if  $\frac{1/2 - w_t}{\|u_{t-1}\|_2} \leq -1$ . Otherwise,

$$\mathbb{E}[X_{t,1} \mid u_{t-1}, A_{-1}] = \frac{\text{Area}(S^{d-2})}{(d-1) \text{Area}(S^{d-1})} \left( 1 - \left( \frac{1/2 + w_t}{\|u_{t-1}\|_2} \right)^2 \right)^{\frac{d-1}{2}}.$$

□

**Lemma .5.** Let  $d \geq 2$ . Then

$$\sqrt{\frac{d}{4e^3}} \leq \frac{\text{Area}(S^{d-2})}{\text{Area}(S^{d-1})} \leq \frac{e}{\pi} \sqrt{d}$$

*Proof.* Recall that the Stirling bounds on the Gamma function give

$$\sqrt{2\pi n}^{n+1/2} e^{-n} \leq \Gamma(n) \leq e n^{n+1/2} e^{-n}$$

for all  $n \in \mathbb{N}$ . Using these bounds along with the fact that  $\text{Area}(S^{d-1}) = \frac{d\pi^{d/2}}{\Gamma(\frac{d}{2}+1)}$ , we have for the lower bound **[\*\*EL Note: Of course, it can never be the case that both  $\frac{d}{2}$  and  $\frac{d-1}{2}$  are both integers. This seems like a small, annoying detail though.]**

$$\begin{aligned} \frac{\text{Area}(S^{d-2})}{\text{Area}(S^{d-1})} &= \frac{d-1}{d} \frac{1}{\sqrt{\pi}} \frac{\Gamma(\frac{d}{2}+1)}{\Gamma(\frac{d-1}{2}+1)} \geq \frac{1}{2} \frac{\sqrt{2\pi}}{e\sqrt{\pi}} \frac{(\frac{d+2}{2})^{\frac{d+3}{2}} e^{-\frac{d+2}{2}}}{(\frac{d+1}{2})^{\frac{d+2}{2}} e^{-\frac{d+1}{2}}} = \frac{1}{\sqrt{2e^3}} \frac{(\frac{d+2}{2})^{\frac{d+3}{2}}}{(\frac{d+1}{2})^{\frac{d+2}{2}}} \\ &= \sqrt{\frac{d+2}{4e^3}} \left(\frac{d+2}{d+1}\right)^{\frac{d+2}{2}} \geq \sqrt{\frac{d}{4e^3}}. \end{aligned}$$

Now, for the upper bound,

$$\begin{aligned} \frac{\text{Area}(S^{d-2})}{\text{Area}(S^{d-1})} &= \frac{d-1}{d} \frac{1}{\sqrt{\pi}} \frac{\Gamma(\frac{d}{2}+1)}{\Gamma(\frac{d-1}{2}+1)} \leq \frac{1}{\sqrt{\pi}} \frac{e^{\left(\frac{d+2}{2}\right)^{\frac{d+3}{2}} e^{-\frac{d+2}{2}}}}{\sqrt{2\pi} \left(\frac{d+1}{2}\right)^{\frac{d+2}{2}} e^{-\frac{d+1}{2}}} = \frac{\sqrt{e}}{\pi\sqrt{2}} \left(\frac{d+2}{2}\right)^{\frac{1}{2}} \left(\frac{d+2}{d+1}\right)^{\frac{d+2}{2}} \\ &\leq \frac{\sqrt{e}}{\pi\sqrt{2}} \sqrt{d} \left(1 + \frac{1}{d+1}\right)^{\frac{d+1}{2}} \cdot \left(1 + \frac{1}{d+1}\right)^{\frac{1}{2}} \leq \frac{e}{\pi} \sqrt{d}. \end{aligned}$$

□

**Lemma .6.** For any  $\varepsilon > 0$  and  $C_\varepsilon = \frac{(4+\varepsilon)e^{21/8}}{2}$ , we have

$$\mathbb{E} \left[ \left| \|u_t\|_2^2 - \|u_{t-1}\|_2^2 \right| \middle| \|u_{t-1}\|_2^2 > C^2(d-1) \right] \leq -\varepsilon.$$

*Proof.* Let  $\varepsilon > 0$ . For any fixed  $u_{t-1}$  with  $\|u_{t-1}\|_2 > C_\varepsilon \sqrt{d-1}$ , we have by Lemma .4 and Lemma .5

$$\begin{aligned} \mathbb{E} \left[ \left| \|u_t\|_2^2 - \|u_{t-1}\|_2^2 \right| \middle| u_{t-1} \right] &= \mathbb{E} \left[ (w_t - q_t)^2 + 2(w_t - q_t) X_t^T u_{t-1} \middle| u_{t-1} \right] \leq 4 - 2\mathbb{E} \left[ q_t X_t^T u_{t-1} \middle| u_{t-1} \right] \\ &= 4 - 2\|u_{t-1}\|_2 \frac{\text{Area}(S^{d-2})}{(d-1) \cdot \text{Area}(S^{d-1})} \left( \left(1 - \left(\frac{1/2 - w_t}{\|u_{t-1}\|}\right)^2\right)^{(d-1)/2} + \left(1 - \left(\frac{1/2 + w_t}{\|u_{t-1}\|}\right)^2\right)^{(d-1)/2} \right) \\ &\leq 4 - 2C_\varepsilon \frac{\text{Area}(S^{d-2})}{\sqrt{d}\text{Area}(S^{d-1})} \left( \left(1 - \frac{(1/2 - w_t)^2}{d-1}\right)^{(d-1)/2} + \left(1 - \frac{(1/2 + w_t)^2}{d}\right)^{(d-1)/2} \right) \\ &\leq 4 - 2C_\varepsilon \frac{\text{Area}(S^{d-2})}{\sqrt{d}\text{Area}(S^{d-1})} \left( \exp\left(-\frac{(1/2 - w_t)^2}{2}\right) + \exp\left(-\frac{(1/2 + w_t)^2}{2}\right) \right) \\ &\leq 4 - 2C_\varepsilon \sqrt{\frac{1}{4e^3}} \left( \exp\left(-\frac{(1/2 - w_t)^2}{2}\right) + \exp\left(-\frac{(1/2 + w_t)^2}{2}\right) \right) \\ &\leq 4 - 4C_\varepsilon e^{-\frac{9}{8}} \sqrt{\frac{1}{4e^3}} = 4 - \frac{4C_\varepsilon}{2e^{21/8}} = -\varepsilon < 0. \end{aligned}$$

□

**Lemma .7.** Let  $w_t$  be fixed. If  $X_t \sim \text{Unif}(S^{d-1})$ , then for any  $\lambda \in (0, 1)$  and  $\beta > 2$

$$\mathbb{E} \left[ e^{\lambda \Delta \|u_t\|_2^2} \middle| \|u_{t-1}\|_2 \geq \beta \right] \leq \frac{e\sqrt{d}}{\beta\pi} \cdot \frac{\max\{e^{w_t^2}, e^{(1-w_t)^2}, e^{(1+w_t)^2}\}^\lambda}{2\lambda \min\{w_t, 1-w_t, 1+w_t\}} \left( e^{-2\lambda(1-w_t)(1/2-w_t)} + e^{-2\lambda(w_t+1)(1/2+w_t)} + e^{2\lambda w_t(1/2-w_t)} \right).$$

*Proof.* As we usually have done, we'll split the conditional expectation up into the three events which determine the values of  $q$ . Note that whenever  $\|u_{t-1}\| \geq 2$ , each event has non-zero probability of

occurring regardless of  $w_t$  (this is why we assume  $\beta > 2$ ). By rotational invariance, we may assume that  $u_{t-1} = \|u_{t-1}\|_2 e_1$ . Using the hyperspherical coordinates as in (1), we have

$$\begin{aligned} \mathbb{E} \left[ e^{\lambda \Delta \|u_t\|_2^2} \middle| \|u_{t-1}\|_2 \geq \beta \right] &= \frac{\text{Area}(S^{d-2})}{\text{Area}(S^{d-1})} \left( e^{\lambda(w_t-1)^2} \int_{\arccos(1)}^{\arccos\left(\frac{1/2-w_t}{\|u_{t-1}\|_2}\right)} e^{2\lambda(w_t-1)\|u_{t-1}\|_2 \cos(\varphi)} \sin^{d-2}(\varphi) d\varphi \right. \\ &\quad + e^{\lambda w_t^2} \int_{\arccos\left(\frac{1/2-w_t}{\|u_{t-1}\|_2}\right)}^{\arccos\left(\frac{-1/2-w_t}{\|u_{t-1}\|_2}\right)} e^{2\lambda w_t \|u_{t-1}\|_2 \cos(\varphi)} \sin^{d-2}(\varphi) d\varphi \\ &\quad \left. + e^{\lambda(w_t+1)^2} \int_{\arccos\left(\frac{-1/2-w_t}{\|u_{t-1}\|_2}\right)}^{\arccos(-1)} e^{2\lambda(w_t+1)\|u_{t-1}\|_2 \cos(\varphi)} \sin^{d-2}(\varphi) d\varphi \right). \end{aligned}$$

We'll control each integral individually. Note for the first in the sum of three, we may decompose it as

$$\begin{aligned} &\int_{\arccos(1)}^{\arccos\left(\frac{1/2-w_t}{\|u_{t-1}\|_2}\right)} e^{2\lambda(w_t-1)\|u_{t-1}\|_2 \cos(\varphi)} \sin^{d-2}(\varphi) d\varphi \\ &= \int_0^{\frac{\pi}{2}} e^{2\lambda(w_t-1)\|u_{t-1}\|_2 \cos(\varphi)} \sin^{d-2}(\varphi) d\varphi + \int_{\frac{\pi}{2}}^{\arccos\left(\frac{1/2-w_t}{\|u_{t-1}\|_2}\right)} e^{2\lambda(w_t-1)\|u_{t-1}\|_2 \cos(\varphi)} \sin^{d-2}(\varphi) d\varphi \\ &\leq \int_0^{\frac{\pi}{2}} \sin^{d-2}(\varphi) d\varphi + \left| \arccos\left(\frac{1/2-w_t}{\|u_{t-1}\|_2}\right) - \arccos(0) \right| \max_{\varphi \in \left[\pi/2, \arccos\left(\frac{1/2-w_t}{\|u_{t-1}\|_2}\right)\right]} e^{2\lambda(w_t-1)\|u_{t-1}\|_2 \cos(\varphi)} \\ &\leq \int_0^{\frac{\pi}{2}} \sin^{d-2}(\varphi) d\varphi + \max_{|x| \leq \frac{1/2-w_t}{\|u_{t-1}\|_2}} \left| \frac{1}{\sqrt{1-x^2}} \right| \frac{1}{\|u_{t-1}\|_2} \max \left\{ e^{2\lambda(w_t-1)(1/2-w_t)}, 1 \right\} \\ &\leq \frac{\text{Area}(S^{d-1})}{2\text{Area}(S^{d-2})} + \frac{16}{7\|u_{t-1}\|_2} \max \left\{ e^{2\lambda(w_t-1)(1/2-w_t)}, 1 \right\} \leq \frac{\text{Area}(S^{d-1})}{2\text{Area}(S^{d-2})} + \frac{16}{7\beta} \max \left\{ e^{2\lambda(w_t-1)(1/2-w_t)}, 1 \right\}. \end{aligned}$$

The third integral follows a similar argument, namely

$$\begin{aligned} &\int_{\arccos\left(\frac{-1/2-w_t}{\|u_{t-1}\|_2}\right)}^{\arccos(-1)} e^{2\lambda(w_t+1)\|u_{t-1}\|_2 \cos(\varphi)} \sin^{d-2}(\varphi) d\varphi \\ &\leq \int_{\arccos\left(\frac{-1/2-w_t}{\|u_{t-1}\|_2}\right)}^{\frac{\pi}{2}} e^{2\lambda(w_t+1)\|u_{t-1}\|_2 \cos(\varphi)} \sin^{d-2}(\varphi) d\varphi + \int_{\frac{\pi}{2}}^{\pi} \sin^{d-2}(\varphi) d\varphi \\ &\leq \frac{16}{7\beta} \max \left\{ e^{2\lambda(w_t+1)(-1/2-w_t)}, 1 \right\} + \frac{\text{Area}(S^{d-1})}{2\text{Area}(S^{d-2})}. \end{aligned}$$

Finally, for the second integral we use the same upper bound + mean value theorem trick as we've used twice above

$$\begin{aligned} &\int_{\arccos\left(\frac{1/2-w_t}{\|u_{t-1}\|_2}\right)}^{\arccos\left(\frac{-1/2-w_t}{\|u_{t-1}\|_2}\right)} e^{2\lambda w_t \|u_{t-1}\|_2 \cos(\varphi)} \sin^{d-2}(\varphi) d\varphi \\ &= \int_{\arccos\left(\frac{1/2-w_t}{\|u_{t-1}\|_2}\right)}^{\frac{\pi}{2}} e^{2\lambda w_t \|u_{t-1}\|_2 \cos(\varphi)} \sin^{d-2}(\varphi) d\varphi + \int_{\frac{\pi}{2}}^{\arccos\left(\frac{-1/2-w_t}{\|u_{t-1}\|_2}\right)} e^{2\lambda w_t \|u_{t-1}\|_2 \cos(\varphi)} \sin^{d-2}(\varphi) d\varphi \\ &\leq \frac{32}{7\beta} \max \left\{ e^{2\lambda w_t(1/2-w_t)}, e^{2\lambda w_t(-1/2-w_t)}, 1 \right\}. \end{aligned}$$

Combining these three upper bounds give

$$\mathbb{E} \left[ e^{\lambda \Delta \|u_t\|_2^2} \middle| \|u_{t-1}\|_2 \geq \beta \right] \leq \max \left\{ e^{w_t^2}, e^{(1+w_t)^2} \right\} \frac{\text{Area}(S^{d-2})}{\text{Area}(S^{d-1})} \left( \frac{\text{Area}(S^{d-1})}{\text{Area}(S^{d-2})} \right)$$

**[\*\*EL Note:** This new proof also seems to be very sensitive to how close  $w_t$  is to  $\pm 1$ . If you can't modify this version to get it to work, uncomment out the block below and use it instead. On second thought, this is entirely unsurprising.  $\Delta \|u_{t-1}\|_2^2 = 0$  if  $w_t = 0, \pm 1$ , so the m.g.f can't be bounded away from the origin in these cases.]

□

**Theorem .8.** For any  $t \in \mathbb{N}$ , we have

$$\mathbb{P}(\|u_t\|_2^2 > \alpha) \leq \dots$$

*Proof.* The proof essentially follows the argument in the main result of Hajek's work [?], [?], [?] using Lemma .7 to, in the notation of Hajek, secure condition D1 and the fact that  $\mathbb{E}\left[e^{\lambda\Delta\|u_t\|_2^2}\left|\|u_{t-1}\|_2^2\right.\right] \leq e^{\frac{\lambda}{4}}$  in place of condition D2. For the sake of completeness, we include that argument here. Using the familiar Laplace transform trick with Chebyshev's inequality, we have for any  $\lambda > 0$

$$\mathbb{P}(\|u_t\|_2^2 \geq \alpha) \leq e^{-\lambda\alpha} \mathbb{E}\left[e^{\lambda\|u_t\|_2^2}\right] \leq e^{-\lambda\alpha} \mathbb{E}\left[e^{\lambda\|u_{t-1}\|_2^2} e^{\lambda\Delta\|u_t\|_2^2}\right] = e^{-\lambda\alpha} \mathbb{E}\left[\mathbb{E}\left[e^{\lambda\|u_{t-1}\|_2^2} e^{\lambda\Delta\|u_t\|_2^2} \middle| \mathcal{F}_{t-1}\right]\right].$$

We can now slip this conditional expectation of the moment generating function of the increments into two pieces based on the norm of the previous residual as

$$\mathbb{E}\left[e^{\lambda\|u_{t-1}\|_2^2} e^{\lambda\Delta\|u_t\|_2^2} \middle| \mathcal{F}_{t-1}\right] = \mathbb{E}\left[e^{\lambda\|u_{t-1}\|_2^2} e^{\lambda\Delta\|u_t\|_2^2} \middle| \mathcal{F}_{t-1}, \|u_{t-1}\|_2 \geq \beta\right] + \mathbb{E}\left[e^{\lambda\|u_{t-1}\|_2^2} e^{\lambda\Delta\|u_t\|_2^2} \middle| \mathcal{F}_{t-1}, \|u_{t-1}\|_2 < \beta\right]$$

Specifically, we note by Lemma .7 when **\*\*EL Note:  $\beta \gtrsim \sqrt{d}$**  we have

$$\mathbb{E}\left[e^{\lambda\Delta\|u_t\|_2^2} \middle| \mathcal{F}_{t-1}, \|u_{t-1}\|_2^2 \geq \beta\right] = \textbf{**EL Note: some quantity...} := \rho < 1. \quad (2)$$

Moreover, by Lemma .2 and Corollary .3

$$\mathbb{E}\left[e^{\lambda\|u_{t-1}\|_2^2} e^{\lambda\Delta\|u_t\|_2^2} \middle| \mathcal{F}_{t-1}, \|u_{t-1}\|_2^2 < \beta\right] \leq e^{\frac{\lambda}{4}} e^{\lambda\beta}. \quad (3)$$

Combining (2), (3) then gives

$$\mathbb{P}(\|u_t\|_2^2 \geq \alpha) \leq e^{-\lambda\alpha} \left(\rho \mathbb{E}\left[e^{\lambda\|u_{t-1}\|_2^2}\right] + e^{\frac{\lambda}{4}}\right).$$

Proceeding inductively thus yields

$$\mathbb{P}(\|u_t\|_2^2 \geq \alpha) \leq \rho^t e^{\lambda(\|u_0\|_2^2 - \alpha)} + \frac{1 - \rho^t}{1 - \rho} e^{\lambda(\beta + \frac{1}{4} - \alpha)}.$$

**\*\*EL Note: Go through and choose  $\beta, \lambda$  appropriately.**

□