

# ADA2: Class 14, Ch 07b, Analysis of Covariance

[Advanced Data Analysis 2](https://StatAcumen.com/teach/ada12, Stat 428/528, Spring 2023, Prof. Erik Erhardt, UNM)

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This is a challenging dataset, in part because it's real and messy. I will guide you through a simplified sensible analysis, but other models are possible.

*Note that I needed to set `cache=FALSE` to assure all output was updated.*

## ANCOVA model: Albuquerque NM 87108, House and Apartment listing prices

Prof Erhardt constructed a dataset of listing prices for dwellings (homes and apartments) for sale from [Zillow.com](https://www.zillow.com) on Feb 26, 2016 at 1 PM for Albuquerque NM 87108. In this assignment we'll develop a model to help understand which qualities that contribute to a **typical dwelling's listing price**. We will then also predict the listing prices of new listings posted on the following day, Feb 27, 2016 by 2 PM.

Because we want to model a *typical dwelling*, it is completely reasonable to remove "unusual" dwellings from the dataset. Dwellings have a distribution with a [long tail](#)!

## Unusual assignment, not top-down, but up-down-up-down

This is an unusual assignment because the workflow of this assignment isn't top-down; instead, you'll be scrolling up and down as you make decisions about the data and model you're fitting. Yes, I have much of the code worked out for you. However, there are data decisions to make early in the code (such as excluding observations, transforming variables, etc.) that depend on the analysis (model checking) later. Think of it as a "choose your own adventure" that I've written for you.

### Keep a record of your decisions

It is always desirable to make your work reproducible, either by someone else or by your future self. For each step you take, keep a diary of (a) what the next minor goal is, (b) what evidence/information you have, (c) what decision you make, and (d) what the outcome was.

For example, here's the first couple steps of your diary:

1. Include only "typical dwellings". Based on scatterplot, remove extreme observations. Keep only HOUSE and APARTMENT.
2. Exclude a few variables to reduce multicollinearity between predictor variables. Exclude `Baths` and `LotSize`.
3. etc.

## (2 p) (Step 1) Restrict data to "typical" dwellings

**Step 1:** After looking at the scatterplot below, identify what you consider to be a “typical dwelling” and exclude observations far from that range. For example, there are only a couple `TypeSale` that are common enough to model; remember to run `factor()` again to remove factor levels that no longer appear.

```
library(erikmisc)
```

— Attaching packages ————— erikmisc 0.1.18 —

✓ tibble 3.1.7      ✓ dplyr 1.0.10

— Conflicts ————— erikmisc\_conflicts() —

✗ dplyr::filter() masks stats::filter()

✗ dplyr::lag() masks stats::lag()

erikmisc, solving common complex data analysis workflows  
by Dr. Erik Barry Erhardt <erik@StatAcumen.com>

```
library(tidyverse)
```

— Attaching packages —————

tidyverse 1.3.2 —

✓ ggplot2 3.4.0      ✓ purrr 1.0.1

✓ tidyr 1.2.1      ✓ stringr 1.5.0

✓ readr 2.1.2      ✓ forcats 0.5.2

— Conflicts ————— tidyverse\_conflicts() —

✗ dplyr::filter() masks stats::filter()

✗ dplyr::lag() masks stats::lag()

```
library(ggplot2)
library(dplyr)
# First, download the data to your computer,
# save in the same folder as this Rmd file.

# read the data, skip the first two comment lines of the data file
dat_abq <-
  read_csv("ADA2_CL_14_HomePricesZillow_Abq87108.csv", skip=2) %>%
  mutate(
    id = 1:n()
  , TypeSale = factor(TypeSale)
    # To help scale the intercept to a more reasonable value
    # Scaling the x-variables are sometimes done to the mean of each x.
    # center year at 1900 (negative values are older, -10 is built in 1890)
  , YearBuilt_1900 = YearBuilt - 1900
  ) %>%
  select(
    id, everything()
  , -Address, -YearBuilt
  )
```

Rows: 143 Columns: 9

— Column specification —

Delimiter: ","

chr (2): Address, TypeSale

dbl (7): PriceList, Beds, Baths, Size\_sqft, LotSize, YearBuilt, DaysListed

i Use `spec()` to retrieve the full column specification for this data.

i Specify the column types or set `show\_col\_types = FALSE` to quiet this message.

```
head(dat_abq)
```

# A tibble: 6 × 9

	id	TypeSale	PriceList	Beds	Baths	Size_sqft	LotSize	DaysListed
	<int>	<fct>	<dbl>	<dbl>	<dbl>	<dbl>	<dbl>	<dbl>
1	1	HOUSE	186900	3	2	1305	6969	0
2	2	APARTMENT	305000	1	1	2523	6098	0
3	3	APARTMENT	244000	1	1	2816	6098	0
4	4	CONDO	108000	3	2	1137	NA	0
5	5	CONDO	64900	2	1	1000	NA	1
6	6	HOUSE	275000	3	3	2022	6098	1

# ... with 1 more variable: YearBuilt\_1900 <dbl>

## RETURN HERE TO SUBSET THE DATA

```
dat_abq <-
  dat_abq %>%
  filter(
    TypeSale %in% c("APARTMENT", "HOUSE")
    #, !id %in% c(120, 130, 50) # (X <= z) # keep observations where variable X <= value z
  ) %>%
  mutate(
    TypeSale = factor(TypeSale)
  )
# note, if you remove a level from a categorical variable, then run factor() again
```

# SOLUTION

# these deletions are based only on the scatter plot in order to have  
# "typical" dwellings

```
summary(dat_abq)
```

id	TypeSale	PriceList	Beds
Min. : 1.00	APARTMENT:41	Min. : 65000	Min. :1.00
1st Qu.: 39.75	HOUSE :91	1st Qu.: 139250	1st Qu.:1.00
Median : 72.50		Median : 169250	Median :3.00
Mean : 73.32		Mean : 226403	Mean :2.28
3rd Qu.:108.25		3rd Qu.: 249250	3rd Qu.:3.00
Max. :143.00		Max. :3110000	Max. :5.00

Baths	Size_sqft	LotSize	DaysListed
Min. :1.000	Min. : 783	Min. : 3049	Min. : 0.0
1st Qu.:1.000	1st Qu.: 1310	1st Qu.: 6534	1st Qu.: 33.5
Median :1.000	Median : 1748	Median : 6969	Median : 88.0
Mean :1.542	Mean : 2272	Mean : 13571	Mean : 122.4
3rd Qu.:2.000	3rd Qu.: 2559	3rd Qu.: 8712	3rd Qu.: 174.0
Max. :5.000	Max. :33000	Max. :609840	Max. :1867.0
NA's :1		NA's :18	

YearBuilt_1900
Min. : 30.00
1st Qu.: 50.00
Median : 52.00
Mean : 56.72
3rd Qu.: 60.00
Max. :106.00
NA's :3

```
#filter(dat_abq, id == 21)
```

```
library(ggplot2)
library(GGally)
```

Registered S3 method overwritten by 'GGally':

```
method from
+.gg ggplot2
```

```
#ggpairs(dat_abq[,c("TypeSale", "PriceList", "Beds", "Baths", "Size_sqft")])
ggpairs(dat_abq %>% dplyr::select(everything(), -id, -Baths), mapping = ggplot2::aes(color=TypeSale,
```

`stat\_bin()` using `bins = 30`. Pick better value with `binwidth`.

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`stat\_bin()` using `bins = 30`. Pick better value with `binwidth`.



```
dat_abq_reduced = dat_abq %>% dplyr::select(everything(), -Baths, -LotSize) %>% na.omit()

ggpairs(dat_abq_reduced %>% dplyr::select(everything(), -id), mapping = ggplot2::aes(color=TypeSale,
```

```
`stat_bin()` using `bins = 30`. Pick better value with `binwidth`.
`stat_bin()` using `bins = 30`. Pick better value with `binwidth`.
`stat_bin()` using `bins = 30`. Pick better value with `binwidth`.
`stat_bin()` using `bins = 30`. Pick better value with `binwidth`.
`stat_bin()` using `bins = 30`. Pick better value with `binwidth`.
```



```
cor(dat_abq_reduced %>% dplyr::select(everything(), -id, -TypeSale))
```

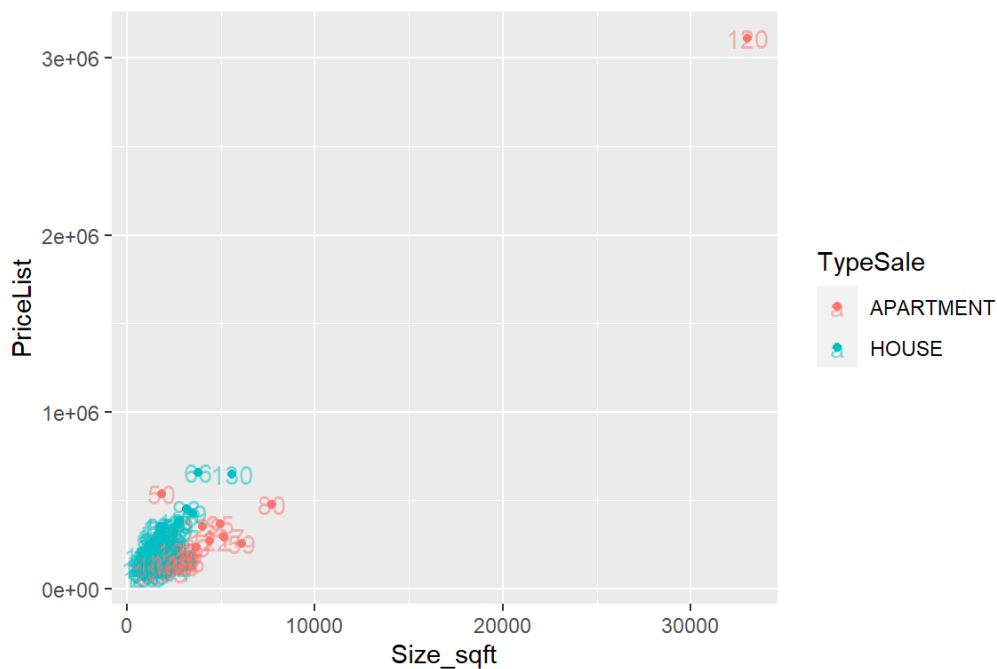
	PriceList	Beds	Size_sqft	DaysListed	YearBuilt_1900
PriceList	1.00000000	0.01050194	0.93729170	0.01589238	0.04649946
Beds	0.01050194	1.00000000	-0.17577972	-0.12561607	-0.21475246
Size_sqft	0.93729170	-0.17577972	1.00000000	0.04687242	0.16943291
DaysListed	0.01589238	-0.12561607	0.04687242	1.00000000	0.07070996
YearBuilt_1900	0.04649946	-0.21475246	0.16943291	0.07070996	1.00000000

```
ggplot(data=dat_abq,
       aes(x=Size_sqft,
           y=PriceList,
```

```

    color=TypeSale,
    label=id))+
  geom_point() +
  geom_text(alpha = .5,
            nudge_x = 0.3)

```



## (2 p) (Step 3) Transform response, if necessary.

**Step 3:** Does the response variable require a transformation? If so, what transformation is recommended from the model diagnostic plots (Box-Cox)?

### Solution

Yes we need transformation based of COX-BOX plot (it contain zero) we do log transformation. [answer]

```
library(car)
```

Loading required package: carData

Attaching package: 'car'

The following object is masked from 'package:purrr':

some

The following object is masked from 'package:dplyr':

recode

```
full.model.lm = lm(
  PriceList ~ (TypeSale + Beds + Size_sqft + DaysListed + YearBuilt_1900)^2,
  data = dat_abq_reduced)
summary(full.model.lm)
```

Call:

```
lm(formula = PriceList ~ (TypeSale + Beds + Size_sqft + DaysListed +
  YearBuilt_1900)^2, data = dat_abq_reduced)
```

Residuals:

Min	1Q	Median	3Q	Max
-187238	-39119	521	33801	392702

Coefficients: (1 not defined because of singularities)

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	1.932e+04	1.473e+05	0.131	0.8959
TypeSaleHOUSE	3.098e+05	1.970e+05	1.573	0.1185
Beds	-1.404e+05	8.320e+04	-1.688	0.0942 .
Size_sqft	1.538e+02	7.229e+01	2.127	0.0356 *
DaysListed	-3.851e+02	4.083e+02	-0.943	0.3476
YearBuilt_1900	1.936e+02	2.276e+03	0.085	0.9324
TypeSaleHOUSE:Beds	NA	NA	NA	NA
TypeSaleHOUSE:Size_sqft	1.714e+01	3.703e+01	0.463	0.6443
TypeSaleHOUSE:DaysListed	3.850e+02	2.529e+02	1.522	0.1307
TypeSaleHOUSE:YearBuilt_1900	-5.723e+03	3.110e+03	-1.841	0.0683 .
Beds:Size_sqft	2.946e+00	1.036e+01	0.284	0.7766
Beds:DaysListed	-1.298e+02	1.056e+02	-1.230	0.2214
Beds:YearBuilt_1900	2.817e+03	1.352e+03	2.084	0.0394 *
Size_sqft:DaysListed	2.118e-01	1.063e-01	1.992	0.0487 *
Size_sqft:YearBuilt_1900	-1.748e+00	9.800e-01	-1.784	0.0771 .
DaysListed:YearBuilt_1900	-5.985e-02	5.324e+00	-0.011	0.9911

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 73770 on 114 degrees of freedom

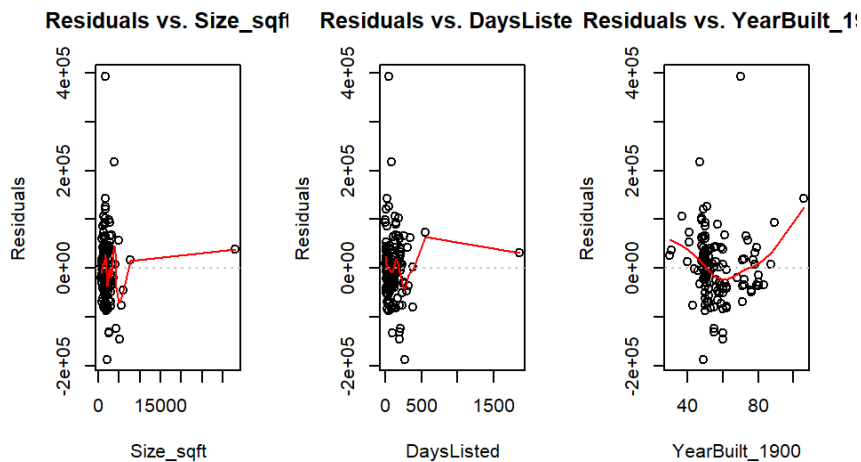
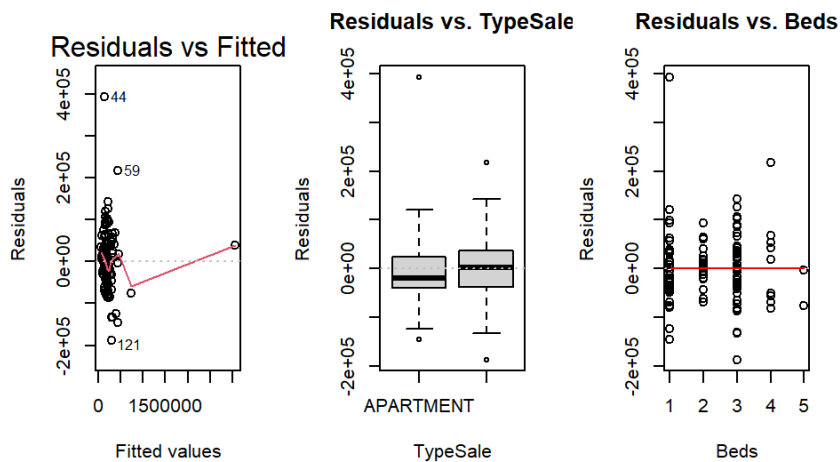
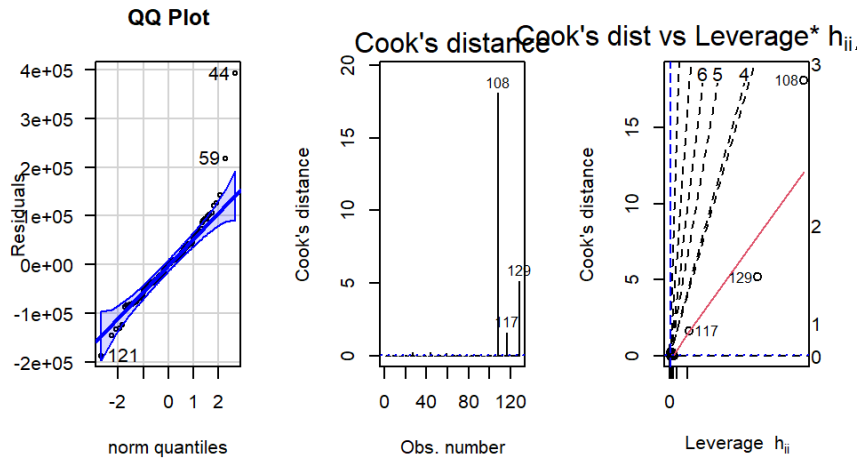
Multiple R-squared: 0.9367, Adjusted R-squared: 0.9289

F-statistic: 120.5 on 14 and 114 DF, p-value: < 2.2e-16

```
#car::Anova(full.model.lm, type=3)
```

```
e_plot_lm_diagnostics(full.model.lm)
```





Non-constant Variance Score Test

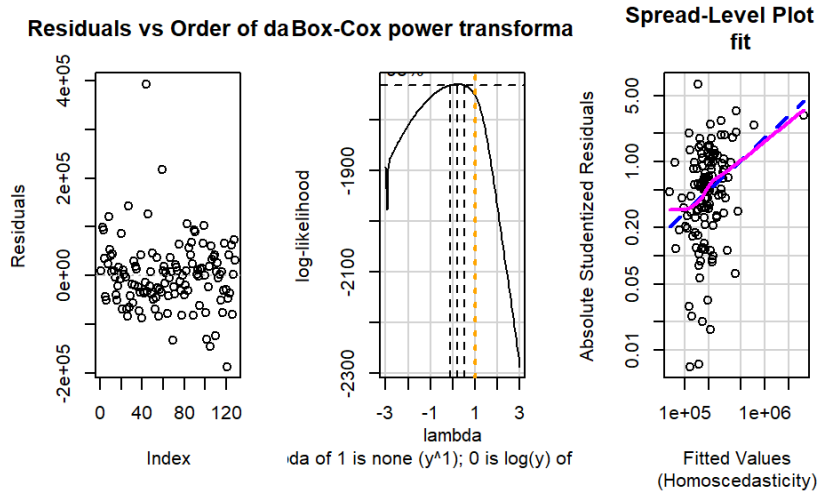
Variance formula:  $\sim$  fitted.values

Chisquare = 0.1788738, Df = 1, p = 0.67234

there are higher-order terms (interactions) in this model

consider setting type = 'predictor'; see ?vif

Error in vif.default(fit): there are aliased coefficients in the model



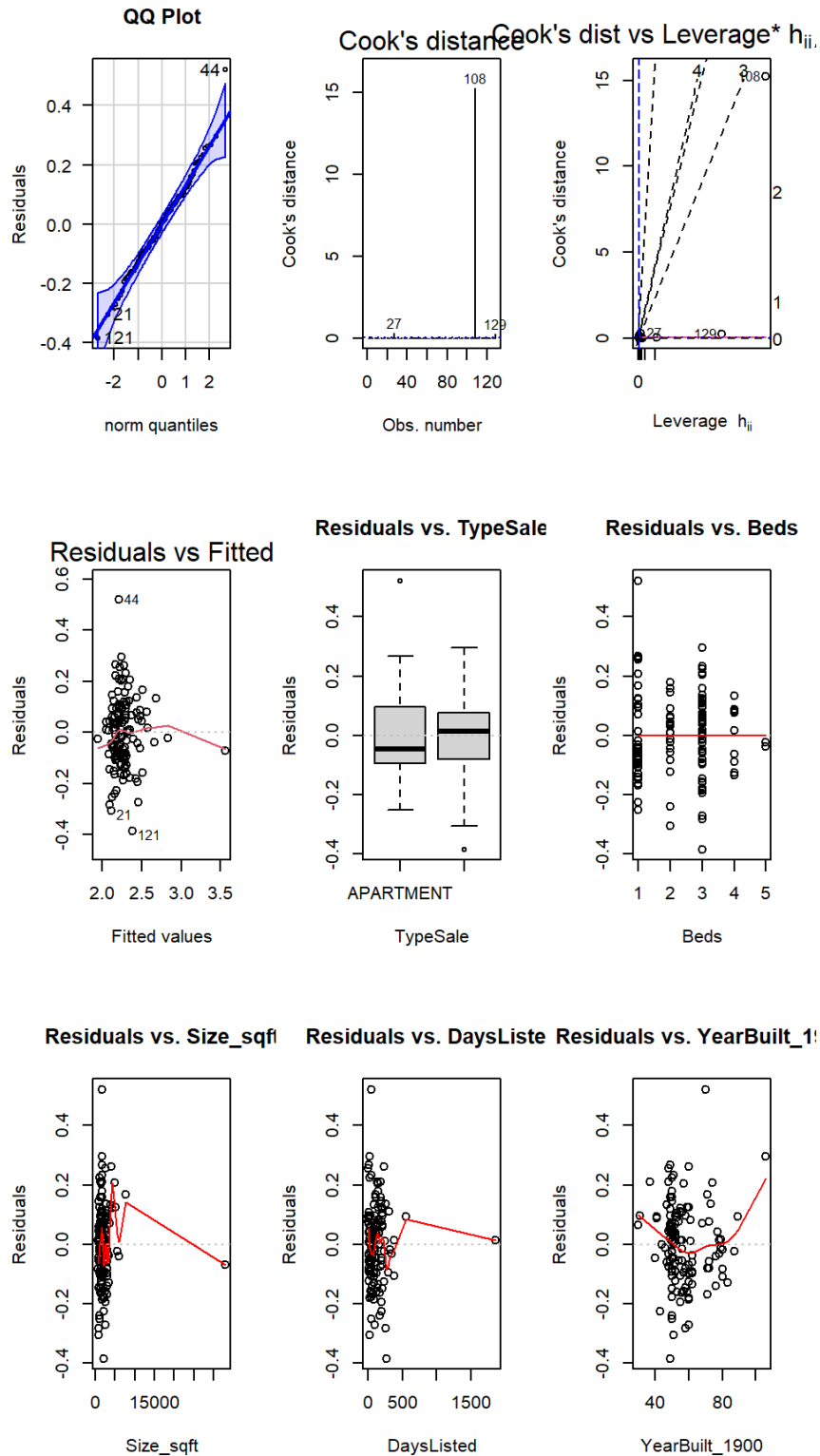
```
dat_abq_trans <-
  dat_abq_reduced %>%
  mutate(
    # Price in units of $1000
    PriceListK = (PriceList / 1000)

    # SOLUTION
  ) %>%
  select(
    -PriceList
  )
str(dat_abq_trans)
```

```
tibble [129 × 7] (S3: tbl_df/tbl/data.frame)
 $ id          : int [1:129] 1 2 3 6 7 9 10 12 13 14 ...
 $ TypeSale    : Factor w/ 2 levels "APARTMENT","HOUSE": 2 1 1 2 2 2 2 1 2 2 ...
 $ Beds        : num [1:129] 3 1 1 3 2 3 3 1 4 2 ...
 $ Size_sqft    : num [1:129] 1305 2523 2816 2022 1440 ...
 $ DaysListed   : num [1:129] 0 0 0 1 1 1 2 2 6 6 ...
 $ YearBuilt_1900: num [1:129] 54 48 89 52 52 58 52 49 41 53 ...
 $ PriceListK   : num [1:129] 187 305 244 275 133 ...
 - attr(*, "na.action")= 'omit' Named int [1:3] 57 83 98
 ... attr(*, "names")= chr [1:3] "57" "83" "98"
```

```
full.model.log = lm(
  log10(PriceListK) ~ (TypeSale + Beds + Size_sqft + DaysListed + YearBuilt_1900)^2,
  data = dat_abq_trans)
```

```
e_plot_lm_diagnostics(full.model.log)
```



Non-constant Variance Score Test

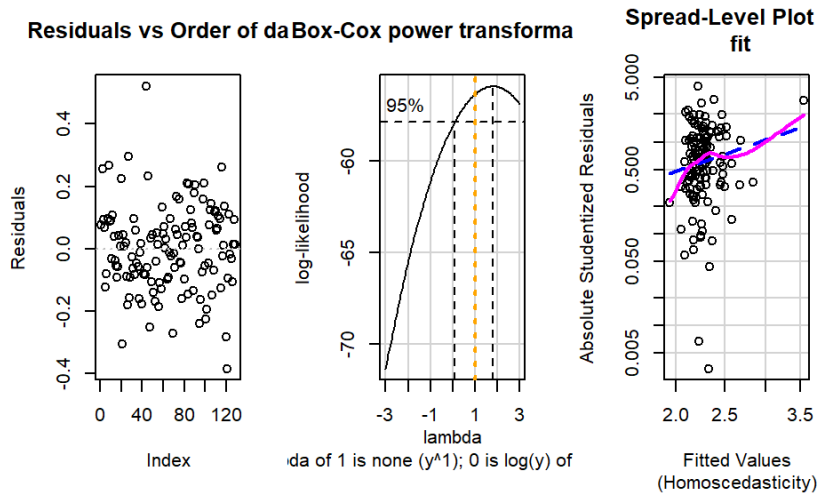
Variance formula:  $\sim$  fitted.values

Chisquare = 0.3182669, Df = 1, p = 0.57265

there are higher-order terms (interactions) in this model

consider setting type = 'predictor'; see ?vif

Error in vif.default(fit): there are aliased coefficients in the model



## (2 p) (Step 4) Remove extremely influential observations.

**Step 4:** The goal is to develop a model that will work well for the typical dwellings. If an observation is highly influential, then it's unusual.

based on Cooks dist vs Leverage plot we noticed observations with id 132 are highly influential and also we remove  $\text{PriceListK} < 500$ ,  $\text{Size\_sqft} < 4000$ ,  $\text{Beds} < 5$ , and  $\text{DaysListed} < 300$  and  $\text{YearBuilt\_1900} < 100$  to develop a model that will work well for the typical dwellings.

After transformation and removing influential observation it seems all assumptions are met

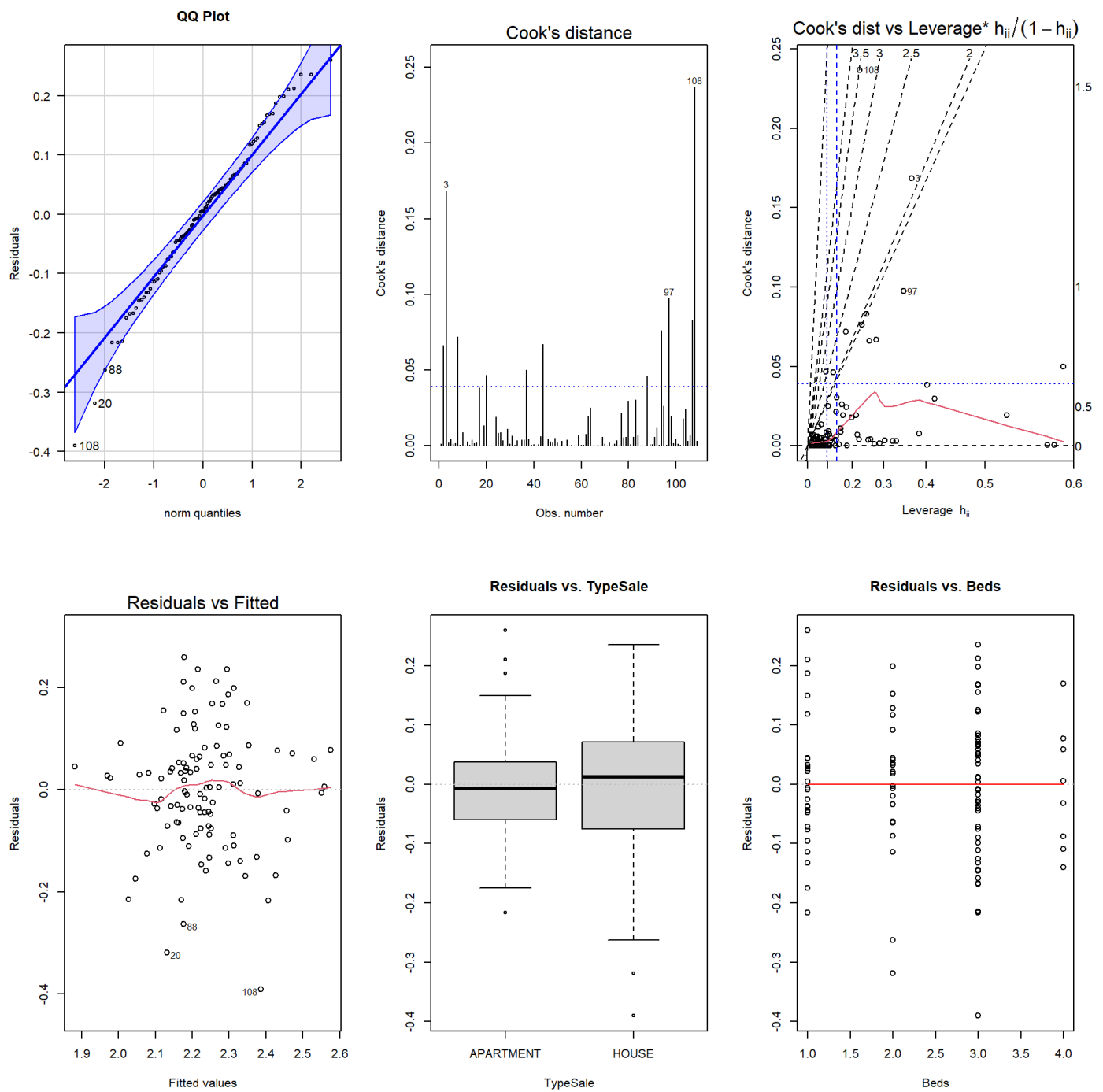
```
names(dat_abq_trans)
```

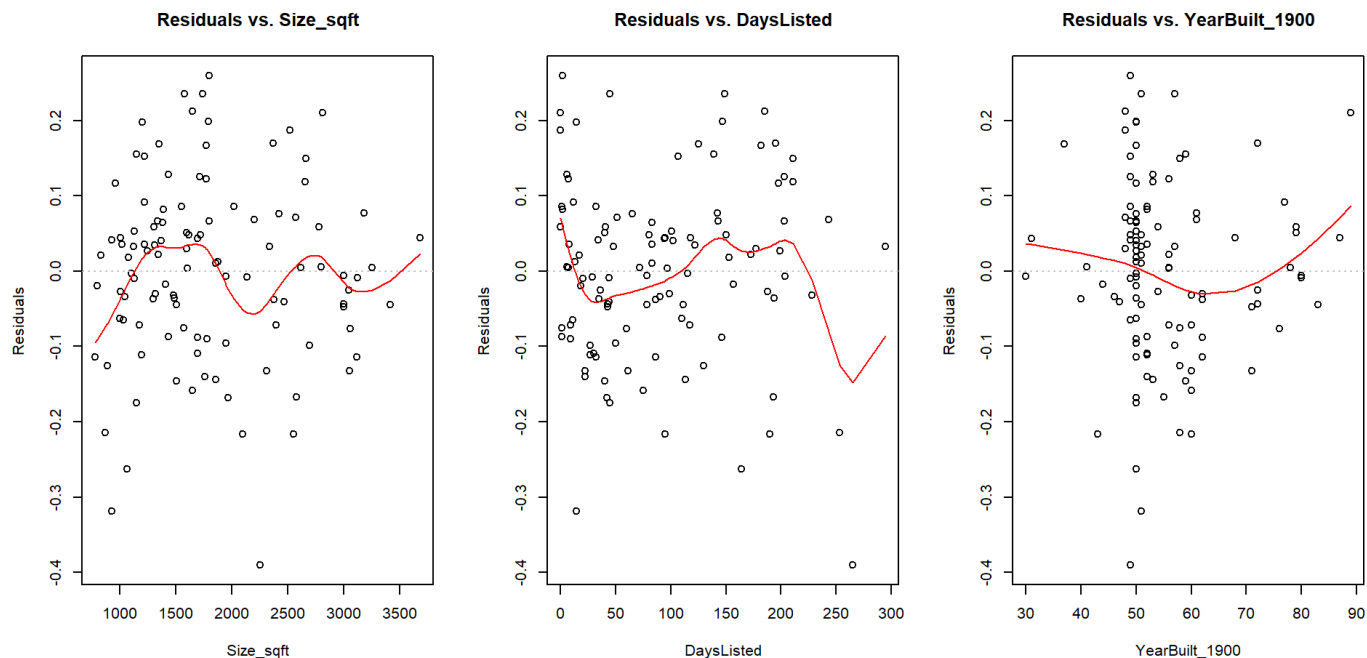
```
[1] "id"           "TypeSale"      "Beds"          "Size_sqft"
[5] "DaysListed"   "YearBuilt_1900" "PriceListK"
```

```
## Remove influential observation
#dat_abq_rem_influen[107,]
dat_abq_rem_influen <-
  dat_abq_trans %>%
  dplyr::filter(
    # !(id %in% c(120, 143, 130, 32, 50, 134, 140)),
    !(id %in% c(132)),
    PriceListK < 500,
    Size_sqft < 4000,
    Beds < 5,
    DaysListed < 300,
    YearBuilt_1900 < 100
  )

# SOLUTION
full.model.log = lm(
  log10(PriceListK) ~ (TypeSale + Beds + Size_sqft + DaysListed + YearBuilt_1900)^2,
  data = dat_abq_rem_influen)
```

```
e_plot_lm_diagnostics(full.model.log)
```





Non-constant Variance Score Test

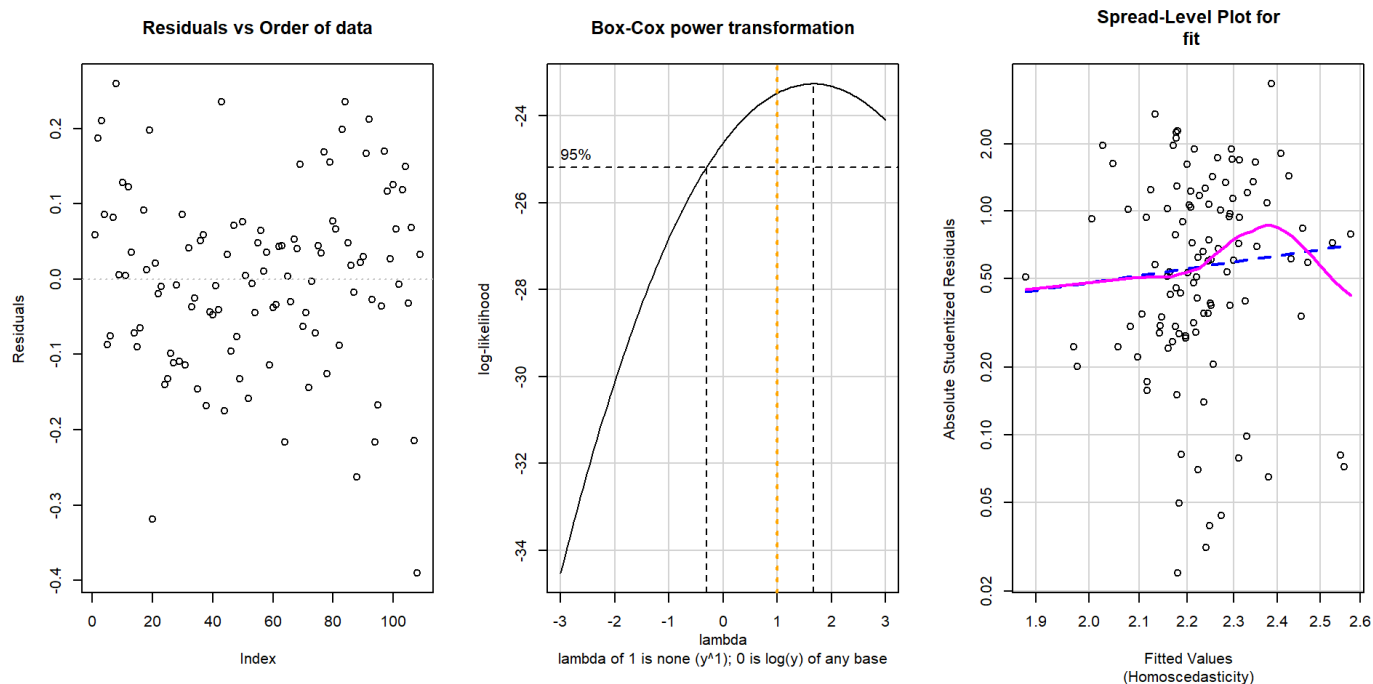
Variance formula:  $\sim \text{fitted.values}$

Chisquare = 0.4715461, Df = 1, p = 0.49228

there are higher-order terms (interactions) in this model

consider setting `type = 'predictor'`; see `?vif`

Error in `vif.default(fit)`: there are aliased coefficients in the model



## Subset data for model building and prediction

Create a subset of the data for building the model, and another subset for prediction later on.

```

# remove observations with NAs
dat_abq_rem_influen <-
  dat_abq_rem_influen %>%
  na.omit()

# the data subset we will use to build our model
dat_sub <-
  dat_abq_rem_influen %>%
  filter(
    DaysListed > 0
  )

# the data subset we will predict from our model
dat_pred <-
  dat_abq_rem_influen %>%
  filter(
    DaysListed == 0
  ) %>%
  mutate(
    # the prices we hope to predict closely from our model
    PricelistK_true = PricelistK
    # set them to NA to predict them later
    , PricelistK = NA
  )

```

Scatterplot of the model-building subset.

```

# NOTE, this plot takes a long time if you're repeatedly recompiling the document.
# comment the "print(p)" line so save some time when you're not evaluating this plot.
library(GGally)
library(ggplot2)
p <-
  ggpairs(
    dat_sub
    , mapping = ggplot2::aes(colour = TypeSale, alpha = 0.5)
    , lower = list(continuous = "points")
    , upper = list(continuous = "cor")
    , progress = FALSE
  )
print(p)

```

Warning in cor(x, y): the standard deviation is zero

`stat\_bin()` using `bins = 30`. Pick better value with `binwidth`.

`stat\_bin()` using `bins = 30`. Pick better value with `binwidth`.

Warning in cor(x, y): the standard deviation is zero

Warning in cor(x, y): the standard deviation is zero

Warning in cor(x, y): the standard deviation is zero

Warning in cor(x, y): the standard deviation is zero

```
`stat_bin()` using `bins = 30`. Pick better value with `binwidth`.
`stat_bin()` using `bins = 30`. Pick better value with `binwidth`.
`stat_bin()` using `bins = 30`. Pick better value with `binwidth`.
`stat_bin()` using `bins = 30`. Pick better value with `binwidth`.
```



There are clearly some unusual observations. Go back to the first code chunk and remove some observations that don't represent a "typical" dwelling.

For example, remove these dwellings (in code above):

- Super expensive dwelling
- Dwellings with huge lots



- Dwellings that were listed for years
- Because most dwellings were APARTMENTS and HOUSEs, remove the others (there are only 1 or so of each).

Discuss the observed correlations or other outstanding features in the data.

## Solution

[answer]

Features of data: 1. "TypeSale"

2. "Beds"

3. "Size\_sqft"

4. "DaysListed"

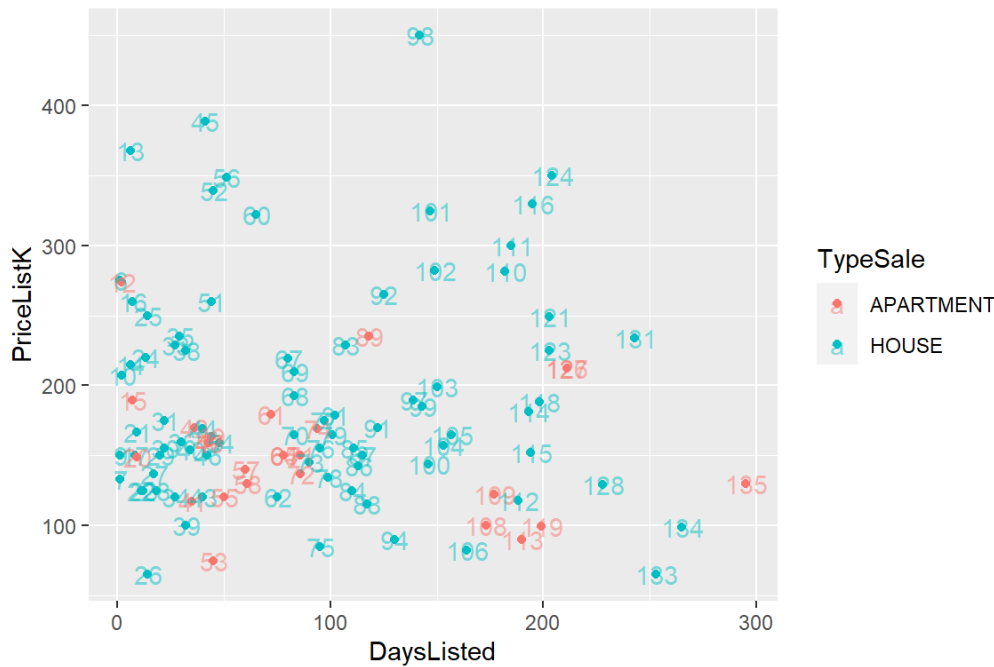
5. "YearBuilt\_1900"

There was high correlation between beds and baths and also ypeSale:Beds which cause coliniarity problem in or model. in addition we did transformation and removed influential observations.

```
names(dat_sub)
```

```
[1] "id"          "TypeSale"    "Beds"        "Size_sqft"
[5] "DaysListed"  "YearBuilt_1900" "PriceListK"
```

```
ggplot(data=dat_sub,
       aes(x=DaysListed,
           y=PriceListK,
           color=TypeSale,
           label=id))+
geom_point() +
geom_text(alpha = .5,
          nudge_x = 0.3)
```



## (2 p) (Step 2) Fit full two-way interaction model.

*You'll revisit this section after each modification of the data above.*

**Step 2:** Let's fit the full two-way interaction model and assess the assumptions. However, some of the predictor variables are highly correlated. Recall that the interpretation of a beta coefficient is "the expected increase in the response for a 1-unit increase in  $x$  with all other predictors held constant". It's hard to hold one variable constant if it's correlated with another variable you're increasing. Therefore, we'll make a decision to retain some variables but not others depending on their correlation values. (In the PCA chapter, we'll see another strategy.)

Somewhat arbitrarily, let's exclude `Baths` (since highly correlated with `Beds` and `Size_sqft`). Let's also exclude `LotSize` (since highly correlated with `Size_sqft`). Modify the code below. Notice that because APARTMENTS don't have more than 1 Beds or Baths, those interaction terms need to be excluded from the model; I show you how to do this manually using the `update()` function.

Note that the formula below  $y \sim (x_1 + x_2 + x_3)^2$  expands into all main effects and two-way interactions.

```
## SOLUTION
lm_full <-
  lm(
    log10(PriceListK) ~ (TypeSale + Beds + Size_sqft + DaysListed + YearBuilt_1900)^2
    , data = dat_sub
  )
#lm_full <-
# lm(
#   PriceListK ~ (Beds + Baths + Size_sqft + LotSize + DaysListed + YearBuilt_1900)^2
#   , data = dat_sub
# )
summary(lm_full)
```

Call:

```
lm(formula = log10(PriceListK) ~ (TypeSale + Beds + Size_sqft +
  DaysListed + YearBuilt_1900)^2, data = dat_sub)
```

Residuals:

Min	1Q	Median	3Q	Max
-0.38166	-0.06724	0.01320	0.06718	0.29322

Coefficients: (1 not defined because of singularities)

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	2.347e+00	3.609e-01	6.504	4.15e-09 ***
TypeSaleHOUSE	3.714e-01	4.415e-01	0.841	0.402
Beds	-1.697e-01	1.697e-01	-1.000	0.320
Size_sqft	1.350e-04	1.777e-04	0.760	0.449
DaysListed	-2.822e-04	1.418e-03	-0.199	0.843
YearBuilt_1900	-1.017e-02	7.562e-03	-1.345	0.182
TypeSaleHOUSE:Beds	NA	NA	NA	NA
TypeSaleHOUSE:Size_sqft	4.803e-06	1.190e-04	0.040	0.968
TypeSaleHOUSE:DaysListed	9.868e-04	9.567e-04	1.032	0.305
TypeSaleHOUSE:YearBuilt_1900	-5.155e-03	7.955e-03	-0.648	0.519
Beds:Size_sqft	1.177e-05	4.637e-05	0.254	0.800
Beds:DaysListed	-4.293e-04	4.047e-04	-1.061	0.292
Beds:YearBuilt_1900	3.556e-03	2.890e-03	1.230	0.222
Size_sqft:DaysListed	1.390e-07	3.598e-07	0.386	0.700
Size_sqft:YearBuilt_1900	3.126e-07	2.926e-06	0.107	0.915
DaysListed:YearBuilt_1900	2.876e-06	2.527e-05	0.114	0.910

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.1253 on 91 degrees of freedom

Multiple R-squared: 0.5245, Adjusted R-squared: 0.4513

F-statistic: 7.168 on 14 and 91 DF, p-value: 9.238e-10

```
try(Anova(lm_full, type=3))
```

Error in Anova.III.lm(mod, error, singular.ok = singular.ok, ...) :  
there are aliased coefficients in the model

```
## Note that this doesn't work because APARTMENTS only have 1 bed and 1 bath.
## There isn't a second level of bed or bath to estimate the interaction.
## Therefore, remove those two terms
lm_full <-
  update(
    lm_full
    , . ~ . - TypeSale:Beds
  )
try(Anova(lm_full, type=3))
```

Anova Table (Type III tests)

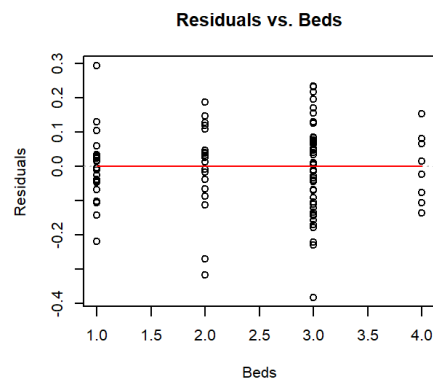
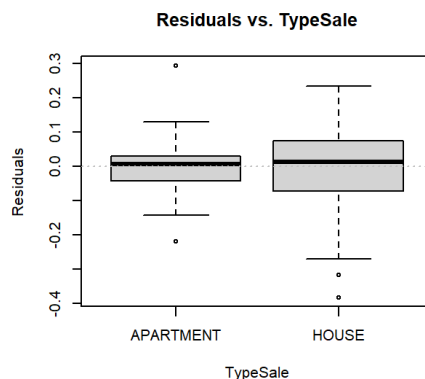
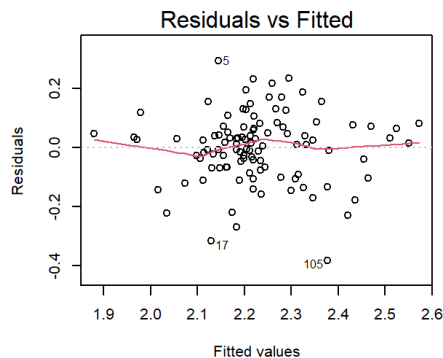
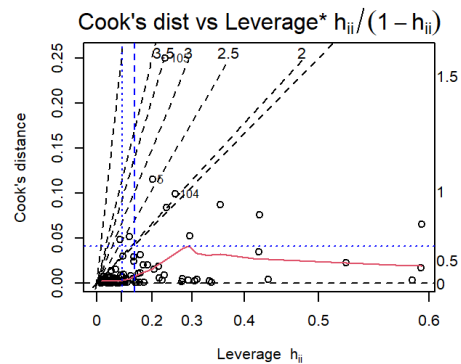
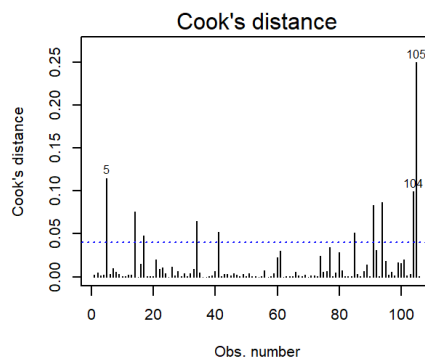
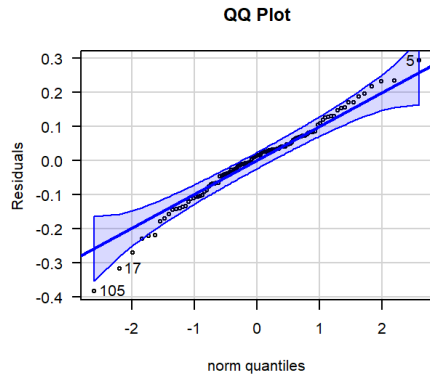
Response: log10(PriceListK)

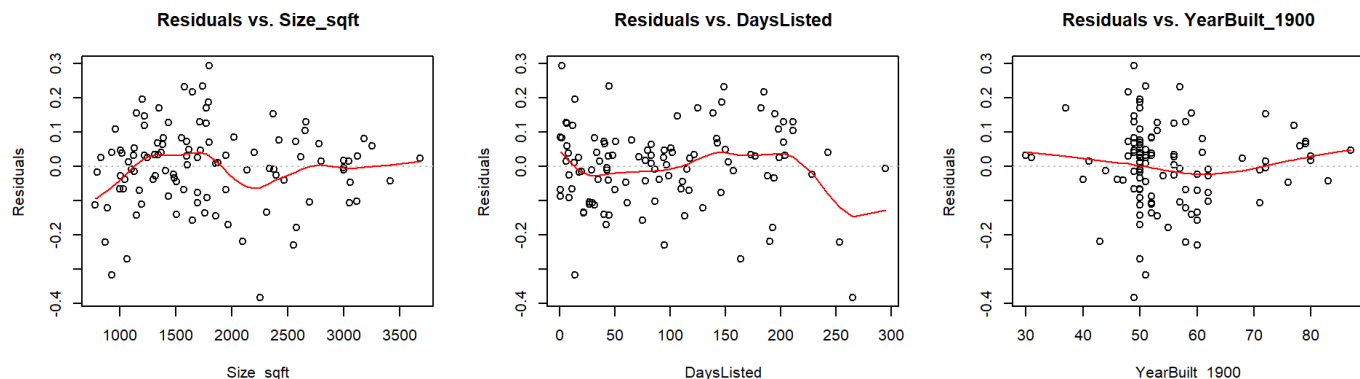
	Sum Sq	Df	F value	Pr(>F)
(Intercept)	0.66456	1	42.2963	4.151e-09 ***
TypeSale	0.01112	1	0.7078	0.4024
Beds	0.01571	1	0.9997	0.3200
Size_sqft	0.00907	1	0.5775	0.4492
DaysListed	0.00062	1	0.0396	0.8428
YearBuilt_1900	0.02842	1	1.8085	0.1820
TypeSale:Size_sqft	0.00003	1	0.0016	0.9679
TypeSale:DaysListed	0.01672	1	1.0640	0.3050
TypeSale:YearBuilt_1900	0.00660	1	0.4200	0.5186
Beds:Size_sqft	0.00101	1	0.0644	0.8003
Beds:DaysListed	0.01768	1	1.1255	0.2915
Beds:YearBuilt_1900	0.02379	1	1.5141	0.2217
Size_sqft:DaysListed	0.00234	1	0.1492	0.7002
Size_sqft:YearBuilt_1900	0.00018	1	0.0114	0.9152
DaysListed:YearBuilt_1900	0.00020	1	0.0130	0.9096
Residuals	1.42980	91		

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

```
## Uncomment this line when you're ready to assess the model assumptions
# plot diagnostics
e_plot_lm_diagnostics(lm_full)
```





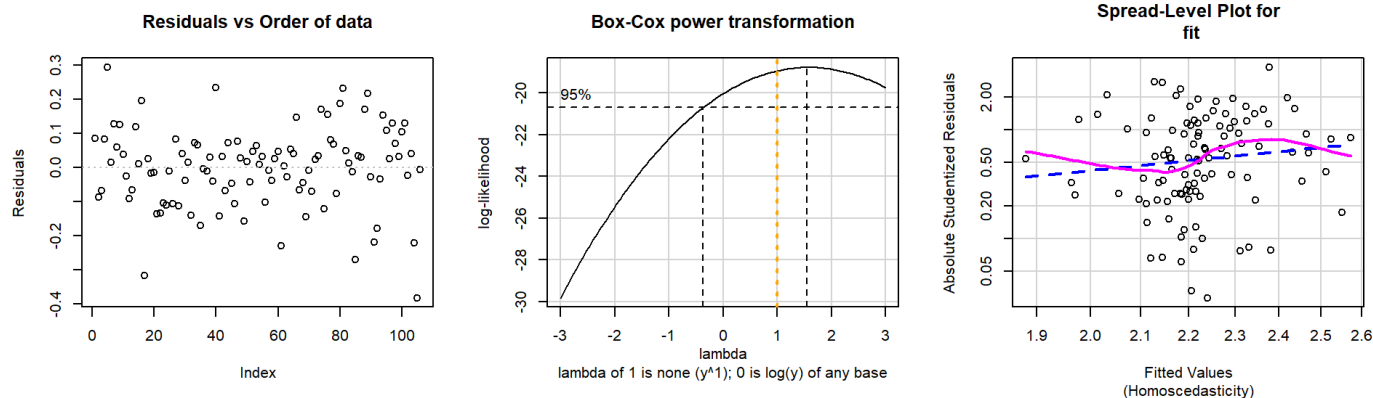
Non-constant Variance Score Test

Variance formula:  $\sim$  fitted.values

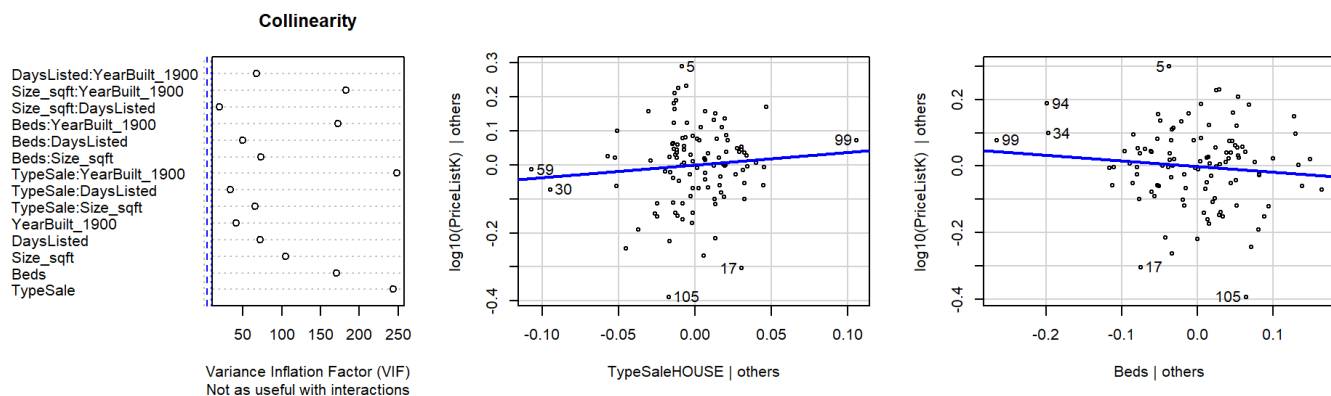
Chisquare = 0.6971503, Df = 1, p = 0.40374

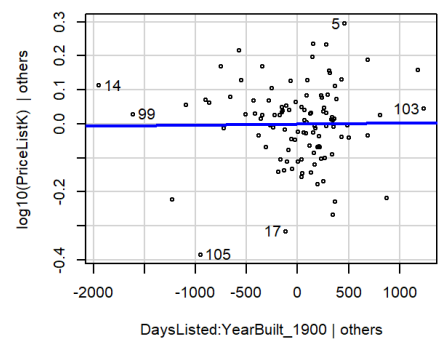
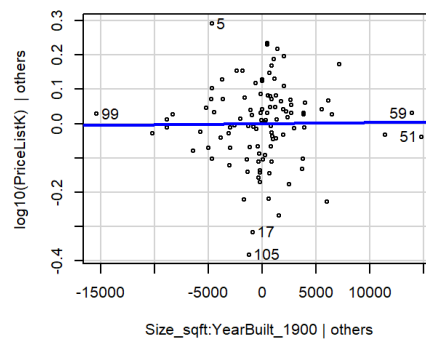
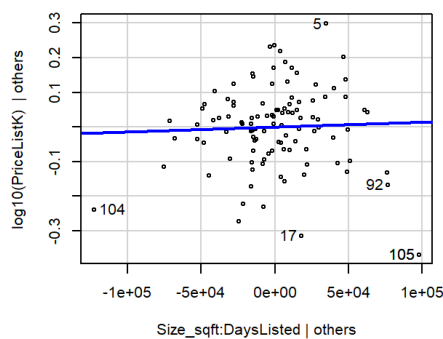
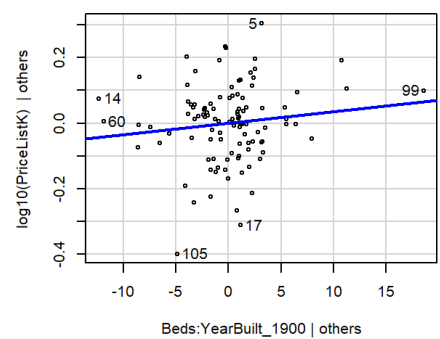
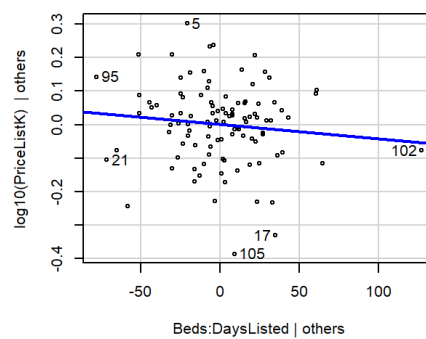
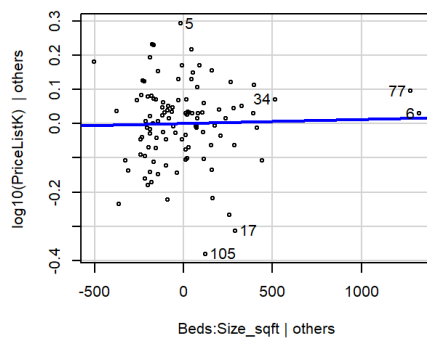
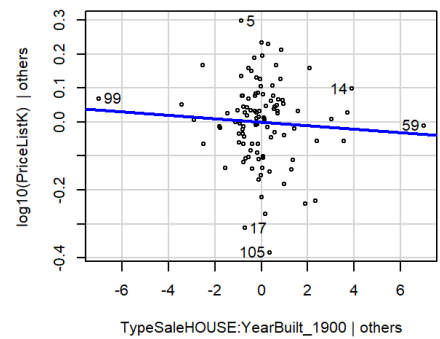
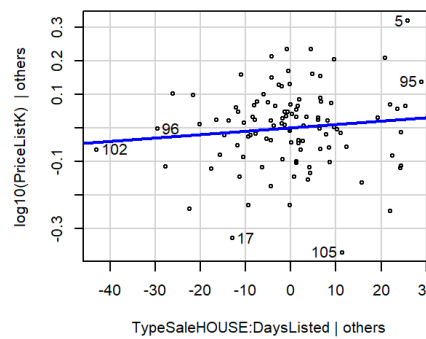
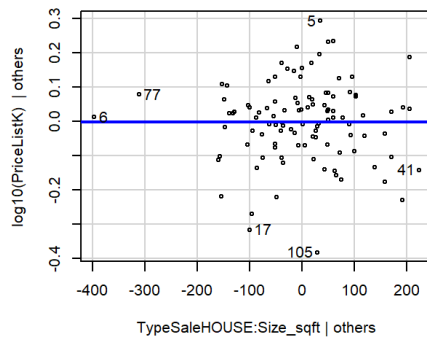
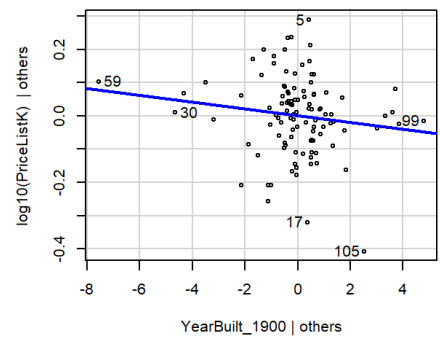
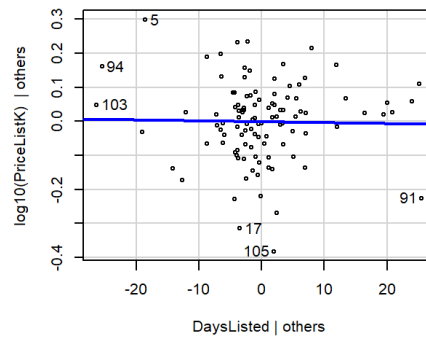
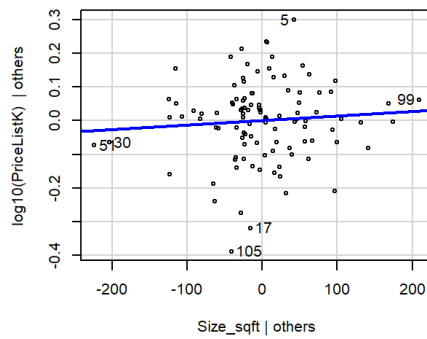
there are higher-order terms (interactions) in this model

consider setting type = 'predictor'; see ?vif



Warning in e\_plot\_lm\_diagnostics(lm\_full): Note: Collinearity plot unreliable for predictors that also have interactions in the model.





```
# List the row numbers with id numbers
# The row numbers appear in the residual plots.
```

```
# The id number can be used to exclude values in code above.
```

```
dat_sub %>% select(id) %>% print(n = Inf)
```

```
# A tibble: 106 × 1
```

	id
	<int>
1	6
2	7
3	9
4	10
5	12
6	13
7	14
8	15
9	16
10	17
11	20
12	21
13	22
14	23
15	24
16	25
17	26
18	27
19	28
20	29
21	30
22	31
23	33
24	34
25	35
26	36
27	38
28	39
29	40
30	41
31	42
32	43
33	44
34	45
35	46
36	47
37	48
38	49
39	51
40	52
41	53
42	54
43	55
44	56
45	57

46	58
47	60
48	61
49	62
50	64
51	65
52	67
53	68
54	69
55	70
56	71
57	72
58	73
59	74
60	75
61	76
62	77
63	78
64	79
65	81
66	83
67	84
68	85
69	86
70	87
71	88
72	89
73	91
74	92
75	94
76	97
77	98
78	99
79	100
80	101
81	102
82	103
83	104
84	105
85	106
86	108
87	109
88	110
89	111
90	112
91	113
92	114
93	115
94	116
95	118
96	119
97	121



```

98 123
99 124
100 126
101 127
102 128
103 131
104 133
105 134
106 135

```

```
shapiro.test(lm_full$residuals)
```

Shapiro-Wilk normality test

```

data:  lm_full$residuals
W = 0.98268, p-value = 0.183

```

After Step 2, interpret the residual plots. What are the primary issues in the original model?

## Solution

[answer] The Residual plot is roughly normal however the tail and head is skewed. by Shapiro test the residual p-value is greater than .05 which means we have not enough evidence that say residuals is not normal. so we met normality assumption. the residuals vs fitted value look acceptable. the plots for residuals vs other features are also acceptable

because there was high correlation between bath and beds and also lotsize we remove baths and lotsize from our model and also APARTMENTS don't have more than 1 Beds or Baths,so those interaction terms need to be excluded from the model. so we fixed the collinearity problem in the model.

## (2 p) (Step 5) Model selection, check model assumptions.

Using `step(..., direction="both")` with the BIC criterion, perform model selection.

## Solution

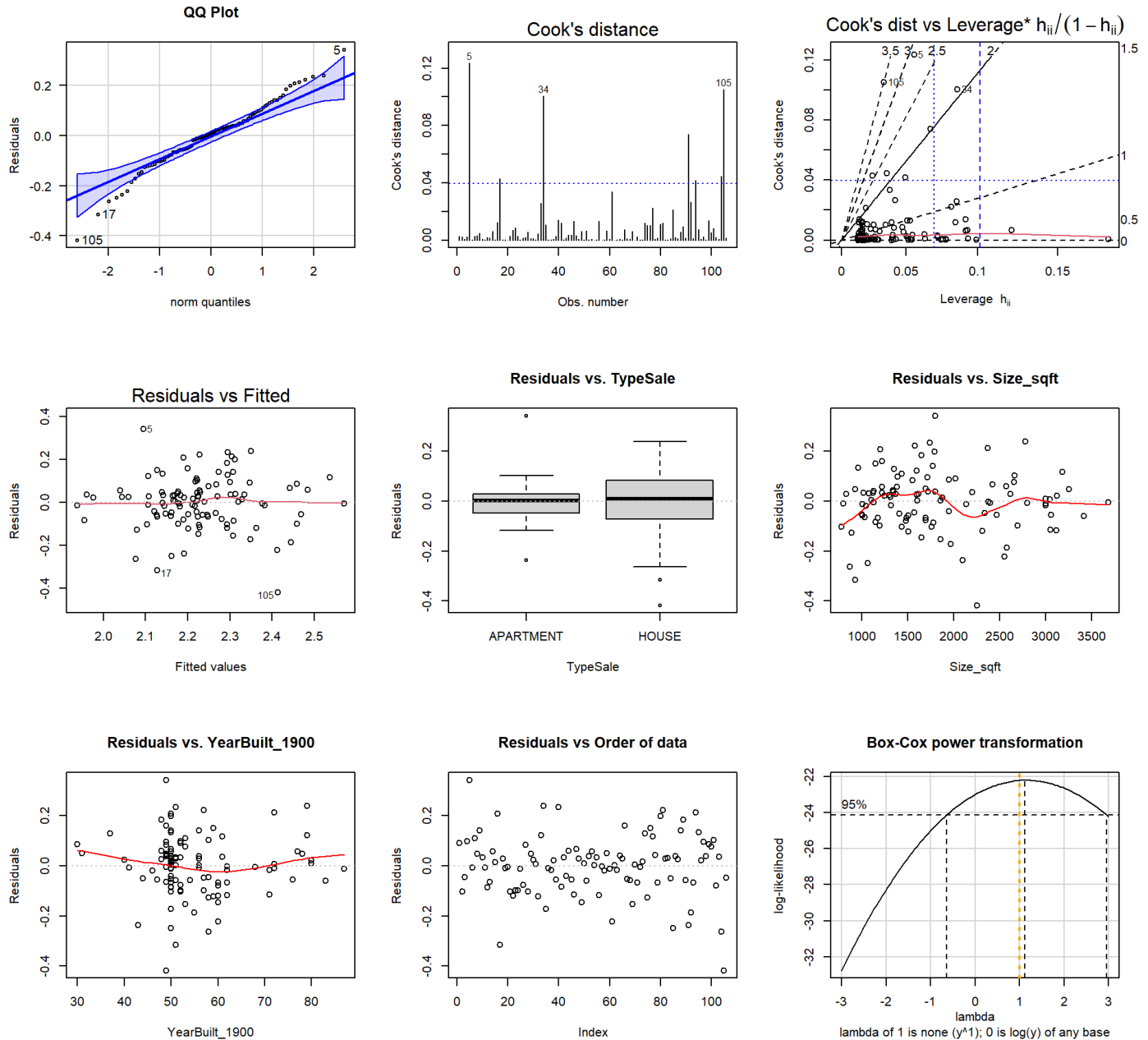
```

## BIC
# option: test="F" includes additional information
#           for parameter estimate tests that we're familiar with
# option: for BIC, include k=log(nrow( [data.frame name] ))
lm_red_BIC <-
  step(
    lm_full
    , direction = "both"
    , test = "F"
    , trace = 0
    , k = log(nrow(dat_sub))
  )

```

```
lm_final <- lm_red_BIC
lm.final = lm_red_BIC
```

```
## Uncomment this line when you're ready to assess the model assumptions
# plot diagnostics
e_plot_lm_diagnostics(lm_final)
```

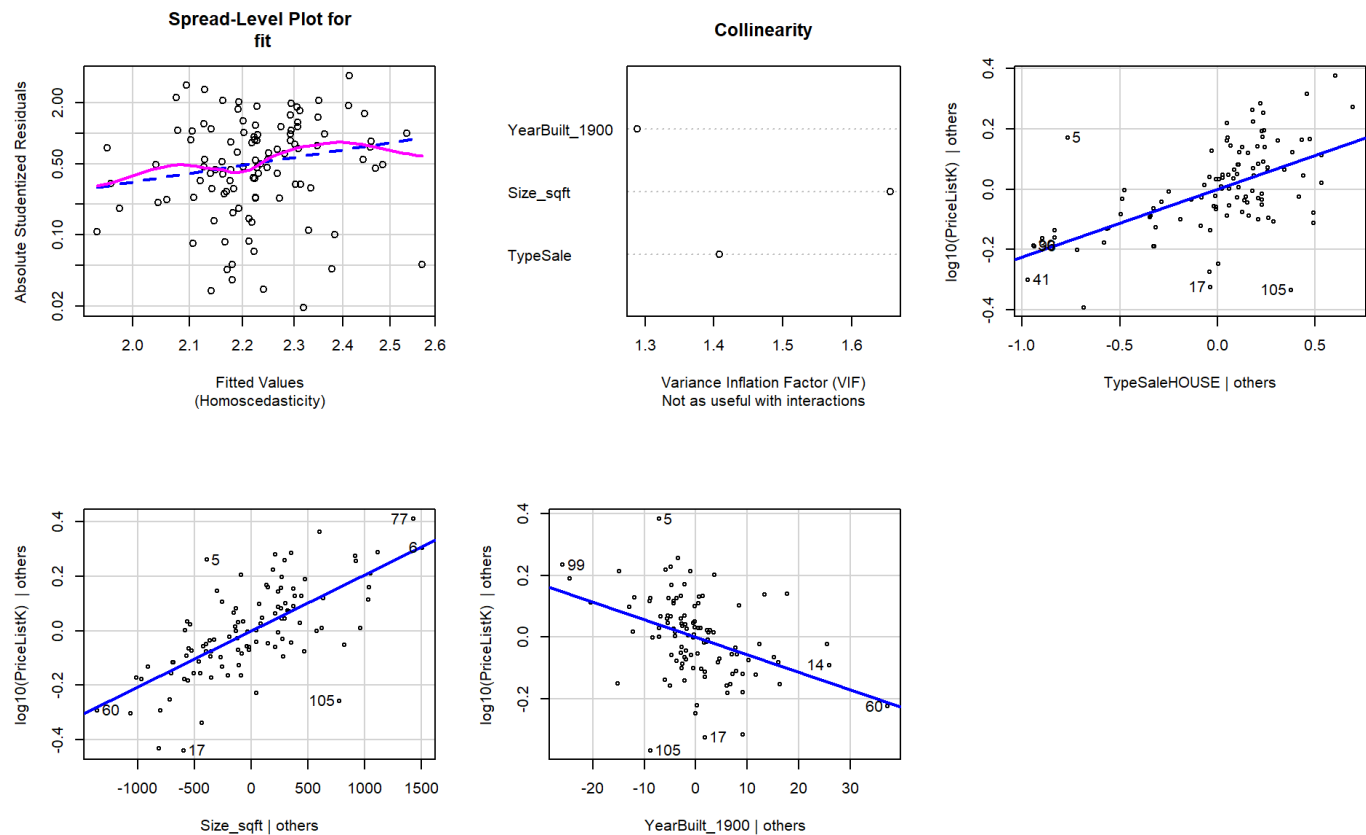


Non-constant Variance Score Test

Variance formula:  $\sim$  fitted.values

Chisquare = 1.789649, Df = 1, p = 0.18097

Warning in e\_plot\_lm\_diagnostics(lm\_final): Note: Collinearity plot unreliable for predictors that also have interactions in the model.

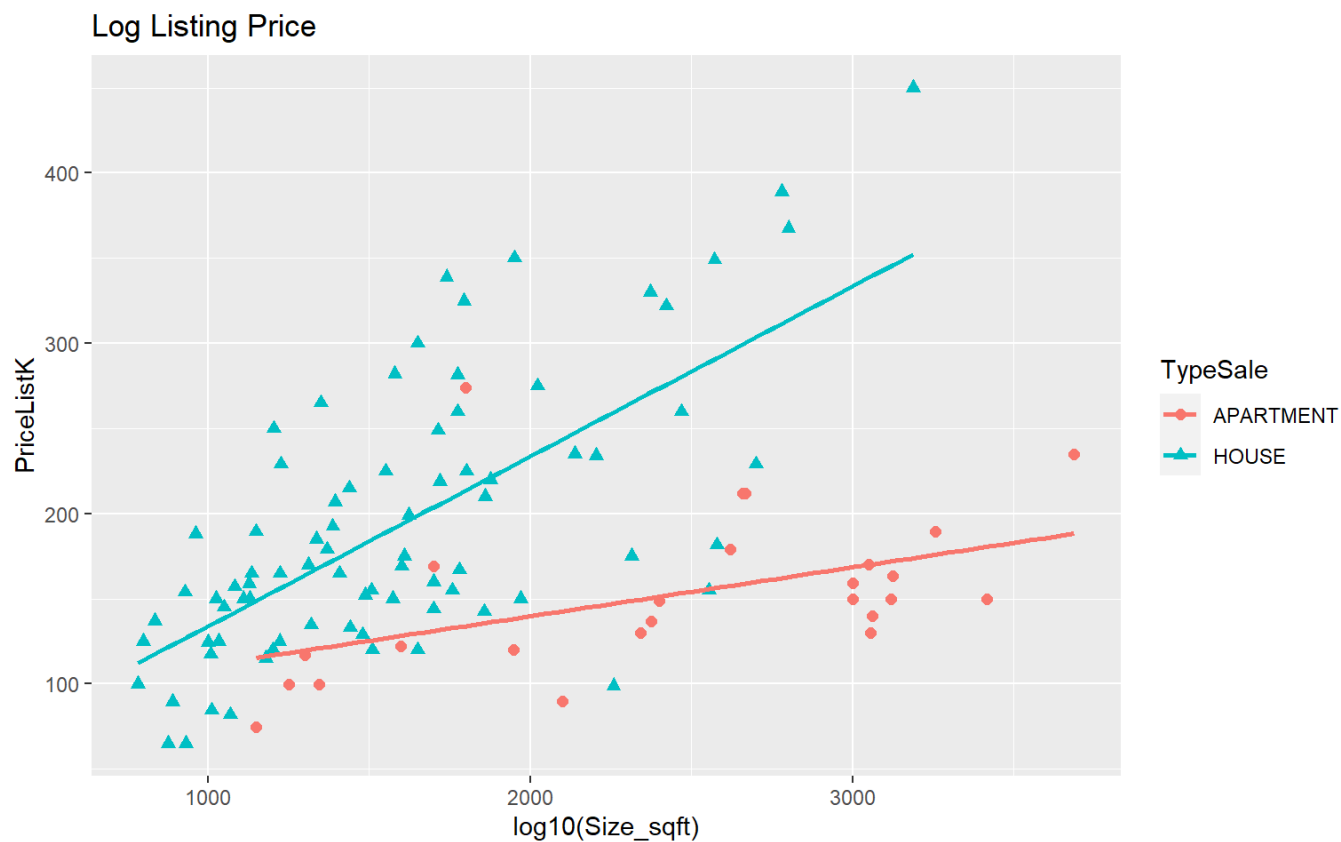


Model assumptions appear to be reasonably met. A few influential observations exist. The residuals are roughly distributed normal based on QQplot (there is a little bit skewness, but it is not that much severe). A few influential observations exist. The variances looks constant. based on box-cox plot we do not need transformation. residuals look acceptable

## (4 p) (Step 6) Plot final model, interpret coefficients.

If you arrived at the same model I did, then the code below will plot it. Eventually (after Step 7), the fitted model equations will describe the each dwelling `TypeSale` and interpret the coefficients.

```
`geom_smooth()` using formula = 'y ~ x'
```



```
library(car)
Anova(lm.final, type=3)
```

#### Anova Table (Type III tests)

Response: log10(PriceListK)

	Sum Sq	Df	F value	Pr(>F)
(Intercept)	9.4443	1	633.368	< 2.2e-16 ***
TypeSale	0.6950	1	46.611	6.363e-10 ***
Size_sqft	1.3438	1	90.118	1.079e-15 ***
YearBuilt_1900	0.2879	1	19.309	2.729e-05 ***
Residuals	1.5210	102		

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

```
summary(lm.final)
```

Call:

```
lm(formula = log10(PriceListK) ~ TypeSale + Size_sqft + YearBuilt_1900,
    data = dat_sub)
```

Residuals:

Min	1Q	Median	3Q	Max
-0.41779	-0.06529	0.00684	0.05714	0.34219

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	2.004e+00	7.964e-02	25.167	< 2e-16 ***
TypeSaleHOUSE	2.233e-01	3.271e-02	6.827	6.36e-10 ***
Size_sqft	2.064e-04	2.174e-05	9.493	1.08e-15 ***
YearBuilt_1900	-5.718e-03	1.301e-03	-4.394	2.73e-05 ***

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.1221 on 102 degrees of freedom

Multiple R-squared: 0.4941, Adjusted R-squared: 0.4793

F-statistic: 33.21 on 3 and 102 DF, p-value: 4.639e-15

Fitted model equation is

$$\log_{10}(\widehat{\text{PriceList}}) = 2 + 0.223I(\text{TypeSale} = \text{HOUSE}) + 2.06 \times 10^{-4}(\text{Size.sqft}) + -0.00572(\text{YearBuilt})$$

## Solution

After Step 7, return and interpret the model coefficients above.

[answer] the log10 of Price will increase if

on average, by one sqft increase in size we expect  $2.06 \times 10^{-4}$  increase in  $\log(\text{PriceListK})$  assuming other variables constant.

on average, by one year increase in yearBuild we expect -0.00572 increase in  $\log(\text{PriceListK})$  assuming other variables constant.

for Apartment, on average the  $\log(\text{PriceListK})$  would be 2 if all other variable would be zero(which is not usefull for interpretation).

for House, on average the  $\log(\text{PriceListK})$  would be  $(2 + 0.223)$  if all other variable would be zero(which is not usefull for interpretation).

## (2 p) (Step 7) Transform predictors.

We now have enough information to see that a transformation of a predictor can be useful. See the curvature with `Size_sqft` ? This is one of the headaches of regression modelling, *everything depends on everything else* and you learn as you go. Return to the top and transform `Size_sqft` and `LotSize`.

A nice feature of this transformation is that the model interaction goes away. Our interpretation is now on the log scale, but it's a simpler model.

```
## SOLUTION
lm_full_logSize <-
  lm(
    log10(PricelistK) ~ (TypeSale + Beds + log10(Size_sqft) + DaysListed + YearBuilt_1900)^2
    , data = dat_sub
  )
#lm_full <-
#  lm(
#    PriceListK ~ (Beds + Baths + Size_sqft + LotSize + DaysListed + YearBuilt_1900)^2
#    , data = dat_sub
```

```
# )
summary(lm_full_logSize)
```

Call:

```
lm(formula = log10(PriceListK) ~ (TypeSale + Beds + log10(Size_sqft) +
  DaysListed + YearBuilt_1900)^2, data = dat_sub)
```

Residuals:

```
      Min       1Q   Median       3Q      Max
-0.37568 -0.06478  0.01052  0.06831  0.27809
```

Coefficients: (1 not defined because of singularities)

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	1.086e+00	2.381e+00	0.456	0.649
TypeSaleHOUSE	1.177e+00	1.421e+00	0.829	0.409
Beds	-5.668e-01	5.710e-01	-0.993	0.323
log10(Size_sqft)	4.733e-01	7.519e-01	0.629	0.531
DaysListed	-1.487e-03	4.480e-03	-0.332	0.741
YearBuilt_1900	-2.247e-02	4.348e-02	-0.517	0.607
TypeSaleHOUSE:Beds	NA	NA	NA	NA
TypeSaleHOUSE:log10(Size_sqft)	-2.403e-01	4.396e-01	-0.546	0.586
TypeSaleHOUSE:DaysListed	9.654e-04	9.530e-04	1.013	0.314
TypeSaleHOUSE:YearBuilt_1900	-4.997e-03	8.291e-03	-0.603	0.548
Beds:log10(Size_sqft)	1.268e-01	1.759e-01	0.721	0.473
Beds:DaysListed	-4.139e-04	4.056e-04	-1.020	0.310
Beds:YearBuilt_1900	3.439e-03	2.997e-03	1.147	0.254
log10(Size_sqft):DaysListed	2.995e-04	1.356e-03	0.221	0.826
log10(Size_sqft):YearBuilt_1900	3.890e-03	1.325e-02	0.294	0.770
DaysListed:YearBuilt_1900	1.184e-05	2.455e-05	0.482	0.631

Residual standard error: 0.1221 on 91 degrees of freedom

Multiple R-squared: 0.5485, Adjusted R-squared: 0.479

F-statistic: 7.895 on 14 and 91 DF, p-value: 1.132e-10

```
try(Anova(lm_full_logSize, type=3))
```

Error in Anova.III.lm(mod, error, singular.ok = singular.ok, ...) :  
there are aliased coefficients in the model

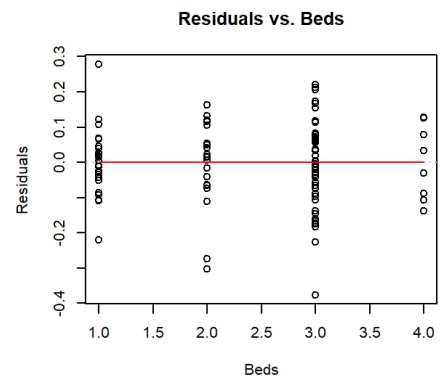
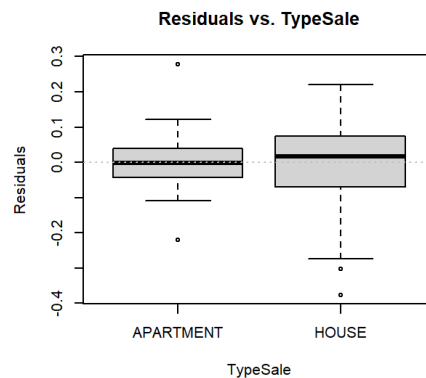
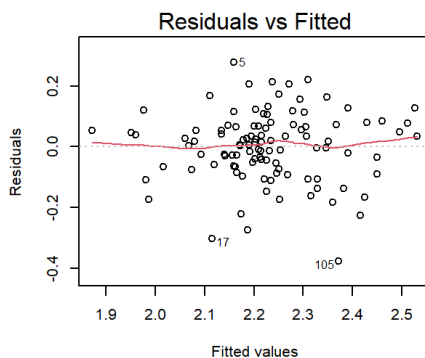
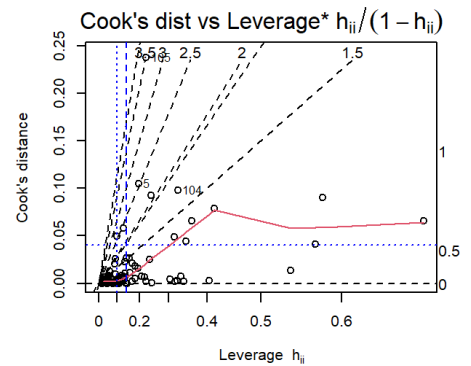
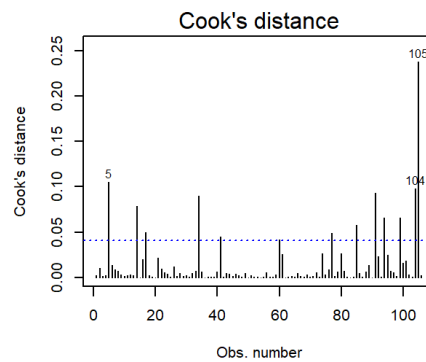
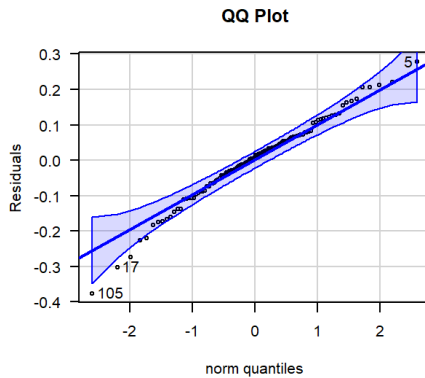
```
## Note that this doesn't work because APARTMENTS only have 1 bed and 1 bath.
## There isn't a second level of bed or bath to estimate the interaction.
## Therefore, remove those two terms
lm_full_logSize <-
  update(
    lm_full_logSize
    , . ~ . - TypeSale:Beds
  )
try(Anova(lm_full_logSize, type=3))
```

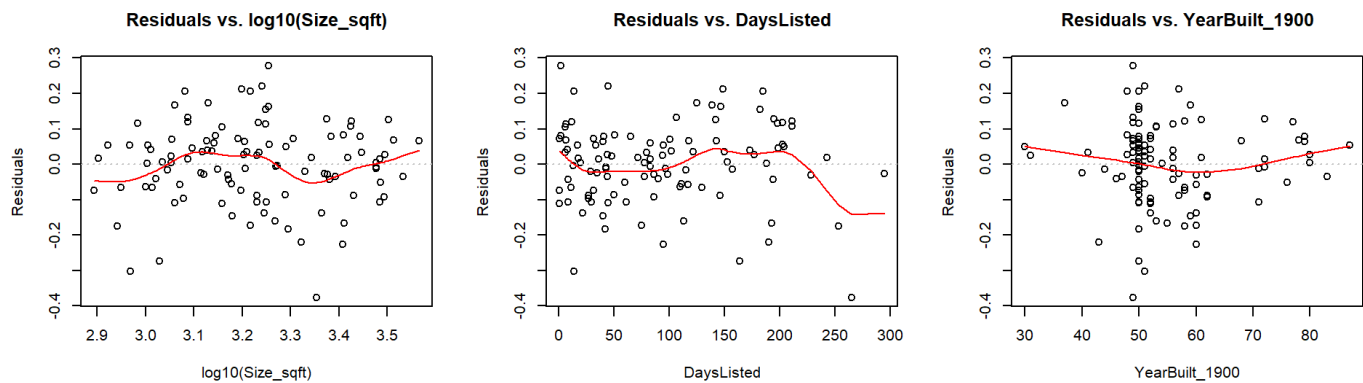
## Anova Table (Type III tests)

Response: log10(PriceListK)

	Sum Sq	Df	F value	Pr(>F)
(Intercept)	0.00310	1	0.2080	0.6494
TypeSale	0.01024	1	0.6866	0.4095
Beds	0.01470	1	0.9856	0.3235
log10(Size_sqft)	0.00591	1	0.3962	0.5306
DaysListed	0.00164	1	0.1102	0.7407
YearBuilt_1900	0.00398	1	0.2671	0.6065
TypeSale:log10(Size_sqft)	0.00446	1	0.2986	0.5861
TypeSale:DaysListed	0.01531	1	1.0263	0.3137
TypeSale:YearBuilt_1900	0.00542	1	0.3633	0.5482
Beds:log10(Size_sqft)	0.00775	1	0.5196	0.4729
Beds:DaysListed	0.01554	1	1.0414	0.3102
Beds:YearBuilt_1900	0.01964	1	1.3167	0.2542
log10(Size_sqft):DaysListed	0.00073	1	0.0488	0.8257
log10(Size_sqft):YearBuilt_1900	0.00129	1	0.0862	0.7698
DaysListed:YearBuilt_1900	0.00347	1	0.2324	0.6309
Residuals	1.35764	91		

```
## Uncomment this line when you're ready to assess the model assumptions
# plot diagnostics
e_plot_lm_diagnostics(lm_full_logSize)
```





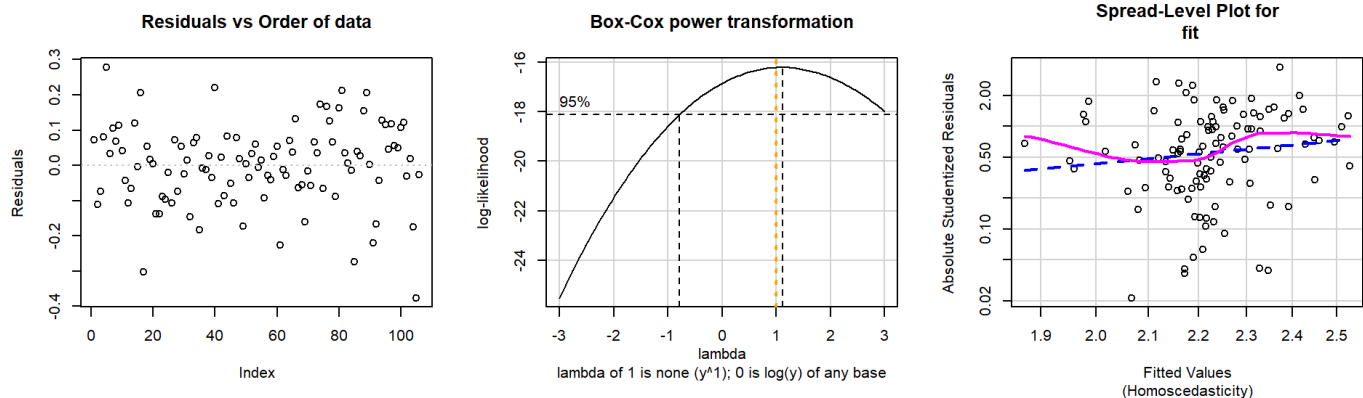
Non-constant Variance Score Test

Variance formula:  $\sim$  fitted.values

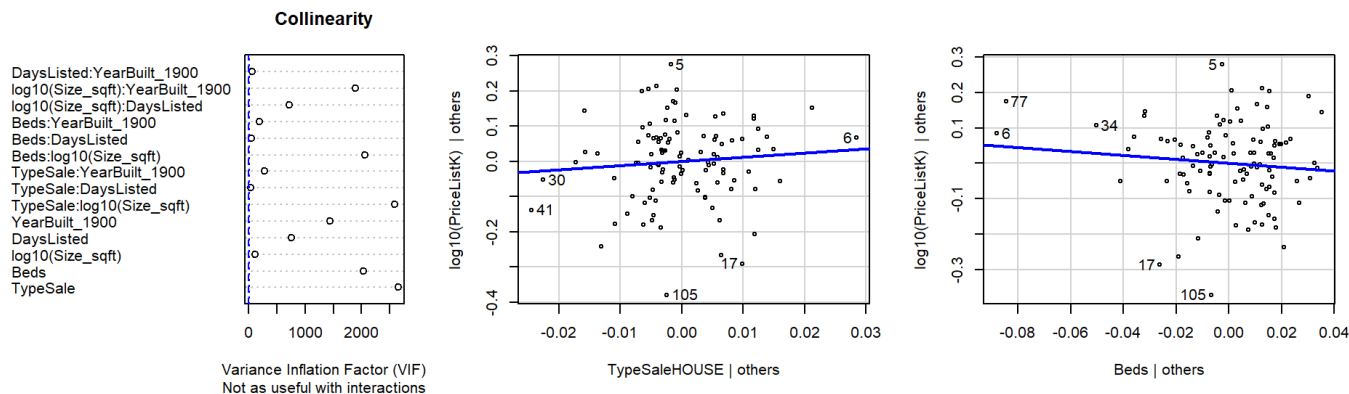
Chisquare = 1.591387, Df = 1, p = 0.20713

there are higher-order terms (interactions) in this model

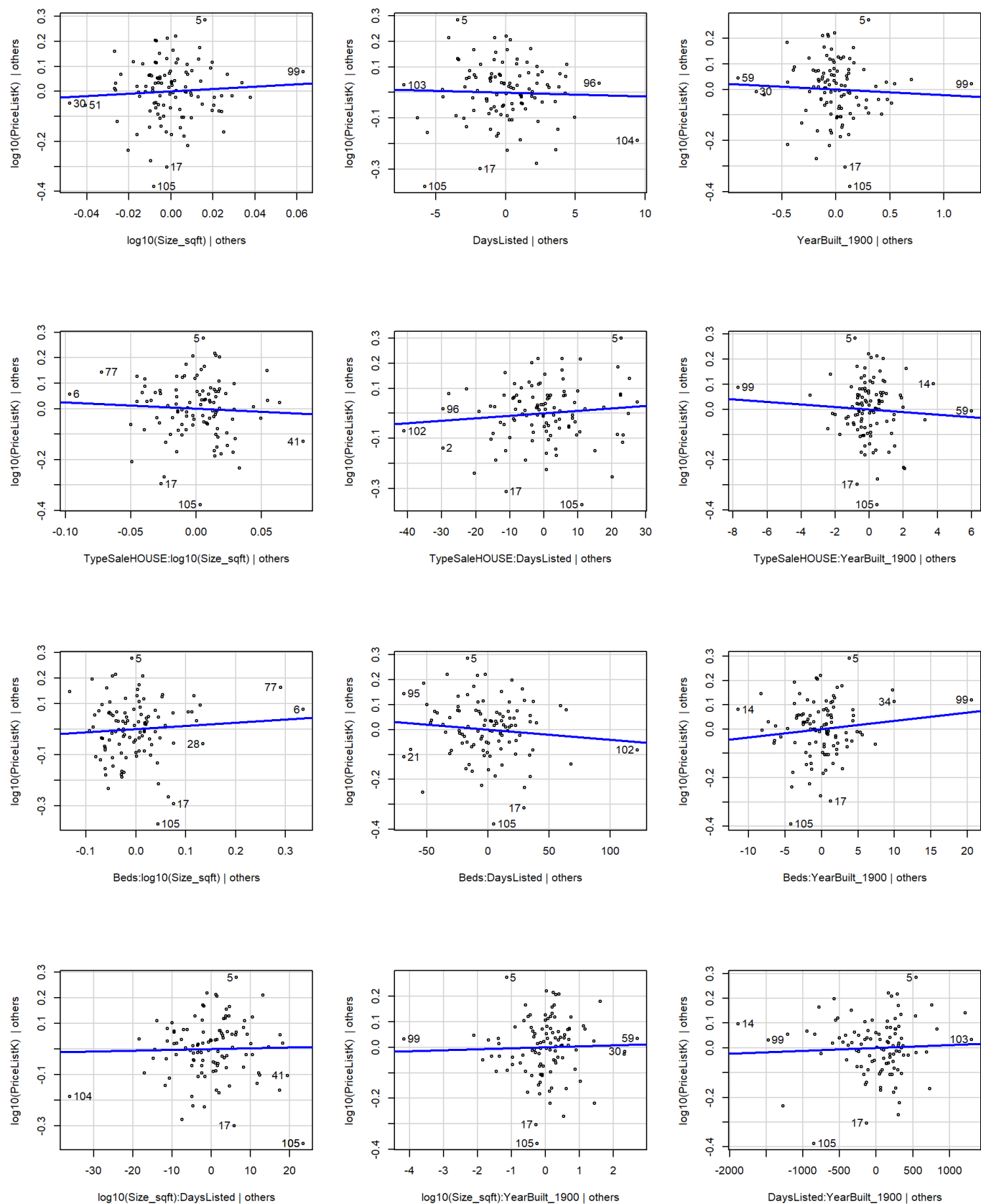
consider setting type = 'predictor'; see ?vif



Warning in e\_plot\_lm\_diagnostics(lm\_full\_logSize): Note: Collinearity plot unreliable for predictors that also have interactions in the model.







```
# List the row numbers with id numbers
# The row numbers appear in the residual plots.
```

```
# The id number can be used to exclude values in code above.
shapiro.test(lm_full_logSize$residuals)
```

Shapiro-Wilk normality test

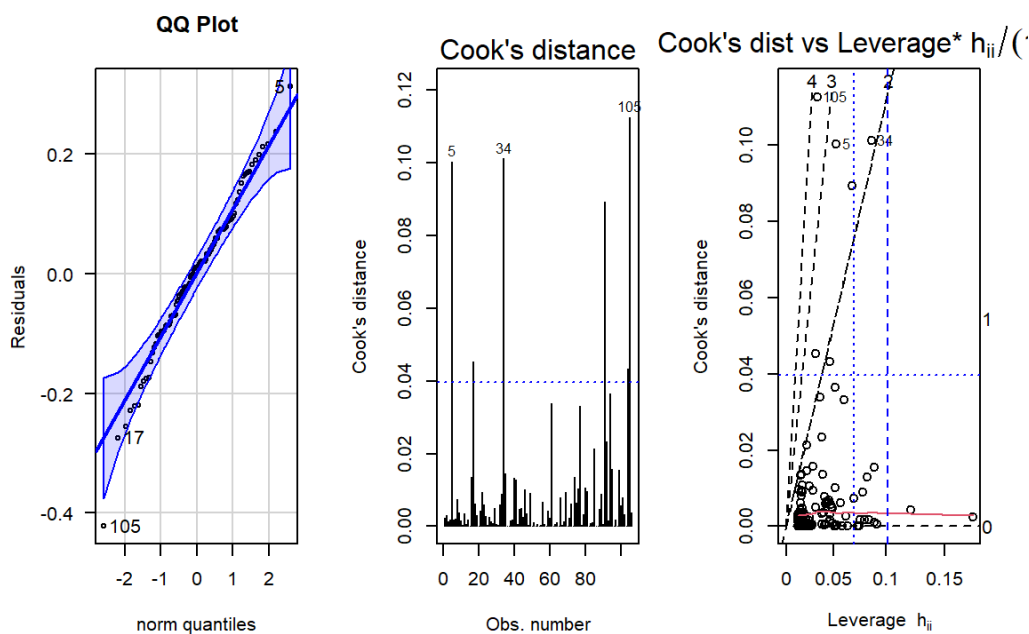
data: lm\_full\_logSize\$residuals

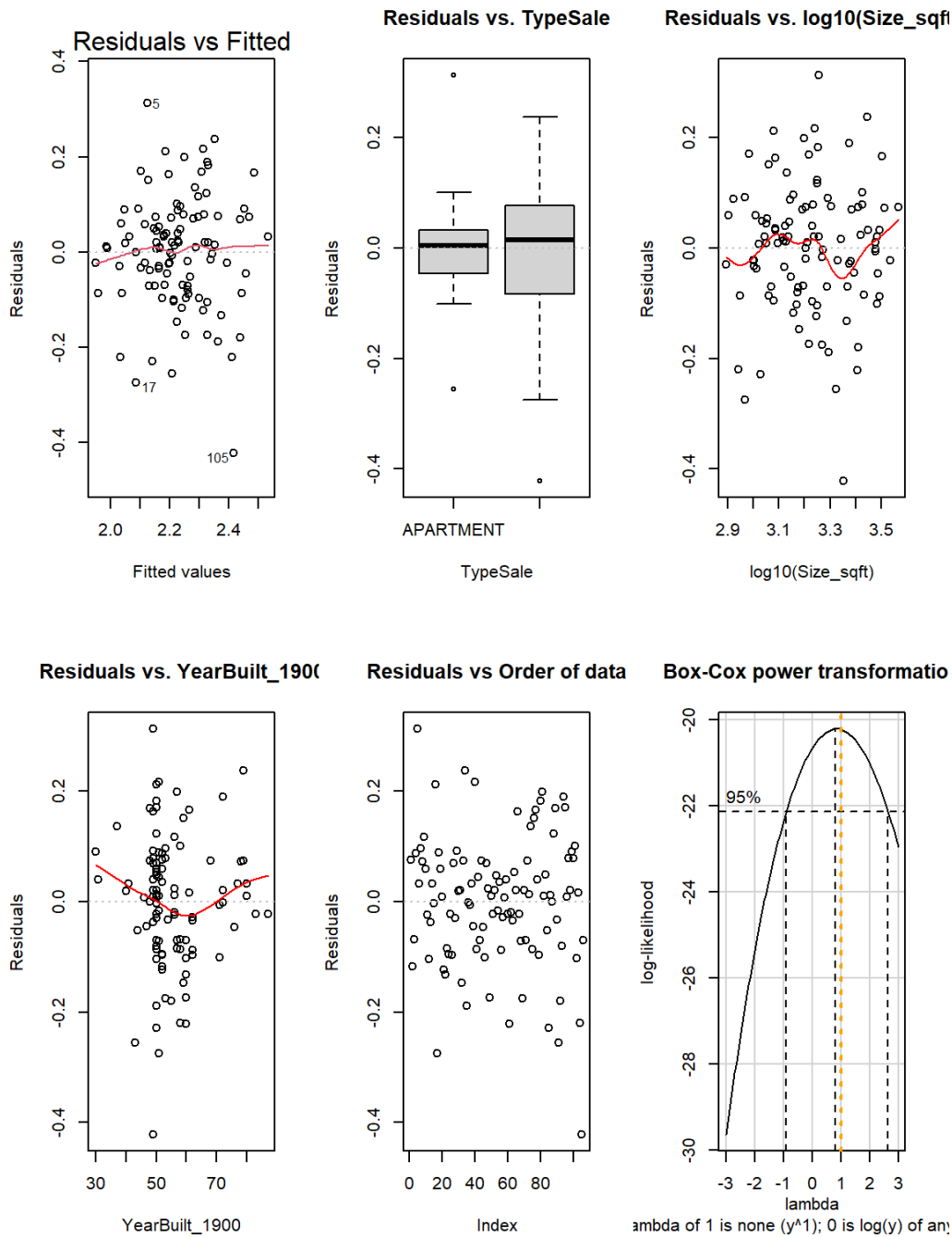
W = 0.98387, p-value = 0.2277

```
lm_red_BIC_logSize <-
  step(
    lm_full_logSize
    , direction = "both"
    , test = "F"
    , trace = 0
    , k = log(nrow(dat_sub))
  )

lm.final_logSize = lm_red_BIC_logSize

e_plot_lm_diagnostics(lm.final_logSize)
```



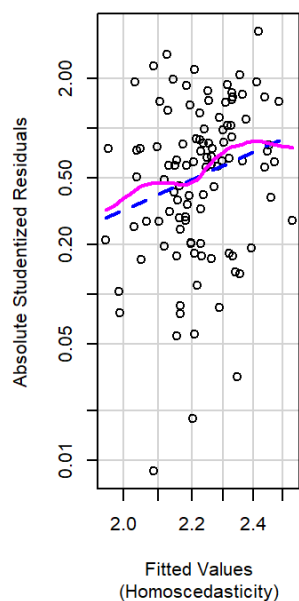
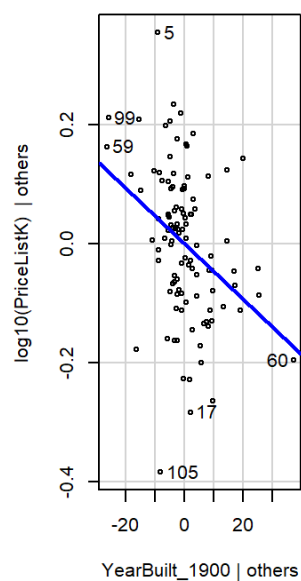
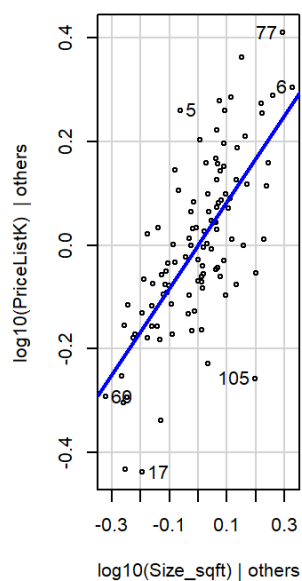
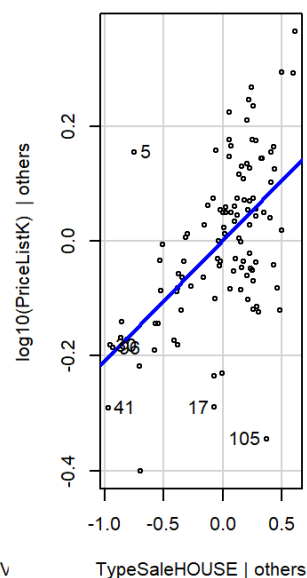


Non-constant Variance Score Test

Variance formula:  $\sim \text{fitted.values}$

Chisquare = 4.091626, Df = 1, p = 0.043096

Warning in `e_plot_lm_diagnostics(lm.final.logSize)`: Note: Collinearity plot unreliable for predictors that also have interactions in the model.

**Spread-Level Plot for fit****Collinearity**

```
summary(lm.final.logSize)
```

Call:

```
lm(formula = log10(PriceListK) ~ TypeSale + log10(Size_sqft) +  
    YearBuilt_1900, data = dat_sub)
```

Residuals:

Min	1Q	Median	3Q	Max
-0.42236	-0.07062	0.00990	0.07286	0.31216

Coefficients:

Estimate	Std. Error	t value	Pr(> t )
(Intercept)			
TypeSale			
log10(Size_sqft)			
YearBuilt_1900			

```
(Intercept)      -0.360661    0.271947   -1.326  0.187729
TypeSaleHOUSE      0.210319    0.031363    6.706  1.13e-09 ***
log10(Size_sqft)    0.833836    0.084437    9.875  < 2e-16 ***
YearBuilt_1900     -0.004655    0.001236   -3.768  0.000276 ***
```

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.1198 on 102 degrees of freedom  
 Multiple R-squared: 0.5129, Adjusted R-squared: 0.4986  
 F-statistic: 35.8 on 3 and 102 DF, p-value: 6.868e-16

$$\log 10(\widehat{\text{PriceList}}) = -0.361 + 0.21I(\text{TypeSale} = \text{HOUSE}) + 0.834(\log 10(\text{Size.sqft})) + -0.00466(\text{YearBuilt})$$

on average, by one unit  $\log(\text{size\_sqft})$  increase we expect 0.834 increase in  $\log(\text{PriceListK})$  assuming other variables constant.

on average, by one year decrease in  $\text{yearBuild}$  we expect -0.00466 increase in  $\log(\text{PriceListK})$  assuming other variables constant.

for Apartment, on average the  $\log(\text{PriceListK})$  would be -0.361 if all other variable would be zero(which is not usefull for interpretation).

for House, on average the  $\log(\text{PriceListK})$  would be (-0.361+0.21) if all other variable would be zero(which is not usefull for interpretation).

## (4 p) (Step 8) Predict new observations, interpret model's predictive ability.

Using the `predict()` function, we'll input the data we held out to predict earlier, and use our final model to predict the `PriceListK` response. Note that `10^lm_pred` is the table of values on the scale of "thousands of dollars".

Interpret the predictions below the output.

How well do you expect this model to predict? Justify your answer.

```
# predict new observations, convert to data frame
lm_pred <-
  as.data.frame(
    predict(
      lm.final
      , newdata = dat_pred
      , interval = "prediction"
    )
  ) %>%
  mutate(
    # add column of actual list prices
    PriceListK = dat_pred$PriceListK_true
  )
lm_pred
```

	fit	lwr	upr	PriceListK
1	2.188175	1.944126	2.432224	186.9
2	2.250483	2.001429	2.499536	305.0
3	2.076528	1.820203	2.332853	244.0

```
# on "thousands of dollars" scale
10^lm_pred
```

	fit	lwr	upr	PriceListK
1	154.2321	87.92769	270.5354	7.943282e+186
2	178.0256	100.32957	315.8902	1.000000e+305
3	119.2691	66.10022	215.2053	1.000000e+244

```
# attributes of the three predicted observations
dat_pred %>% print(n = Inf, width = Inf)
```

```
# A tibble: 3 × 8
```

	id	TypeSale	Beds	Size_sqft	Dayslisted	YearBuilt_1900	PriceListK
	<int>	<fct>	<dbl>	<dbl>	<dbl>	<dbl>	<lg1>
1	1	HOUSE	3	1305	0	54	NA
2	2	APARTMENT	1	2523	0	48	NA
3	3	APARTMENT	1	2816	0	89	NA

	PriceListK_true
	<dbl>
1	187.
2	305
3	244

## Solution

```
# predict new observations, convert to data frame
lm_pred <-
  as.data.frame(
    predict(
      lm.final
      , newdata = dat_pred
      , interval = "prediction"
    )
  ) %>%
  mutate(
    # add column of actual list prices
    PriceListK = dat_pred$PriceListK_true
  )
lm_pred
```

	fit	lwr	upr	PriceListK
1	2.188175	1.944126	2.432224	186.9
2	2.250483	2.001429	2.499536	305.0
3	2.076528	1.820203	2.332853	244.0

```
# on "thousands of dollars" scale
#10^lm_pred
pre.df = lm_pred %>%
  mutate(fit = 10^fit,
         lwr = 10^lwr,
         upr = 10^upr)
dat_pred$PriceListK = pre.df$fit
dat_predfinal = pre.df
# attributes of the three predicted observations
```

[answer] for a with beds and size\_sqft and yearBuild we predict the 154.2321451 PriceListK with interval (87.9276911, 270.5354172 ).

for a with beds and size\_sqft and yearBuild we predict the 178.0256314 PriceListK with interval (100.3295663, 315.8901868 ). for a with beds and size\_sqft and yearBuild we predict the 119.2691104 PriceListK with interval (87.9276911, 215.2053479 ).

the model did a good job on prediction the first observation (apartment) price with just 32.67 error. for the second observation it predict price 178 with true price of 305. however the prediction is inside the interval but its close to upper interval. for the third observation the model could not predict well. it predict the price 119 but true price is almost two time bigger and is not in the 95% interval. overall it seems the model can not predict the price precisely.

```
# predict new observations, convert to data frame
lm_pred <-
  as.data.frame(
    predict(
      lm.final.logSize
      , newdata = dat_pred
      , interval = "prediction"
    )
  ) %>%
  mutate(
    # add column of actual list prices
    PriceListK = dat_pred$PriceListK_true
  )
lm_pred
```

	fit	lwr	upr	PriceListK
1	2.196173	1.956783	2.435563	186.9
2	2.252520	2.008123	2.496917	305.0
3	2.101434	1.850131	2.352737	244.0

```
# on "thousands of dollars" scale
#10^lm_pred
pre.df = lm_pred %>%
  mutate(fit = 10^fit,
         lwr = 10^lwr,
         upr = 10^upr)
dat_pred$PriceListK = pre.df$fit
```

```
dat_predfinal = pre.df  
# attributes of the three predicted observations
```

The model with  $\log(\text{Size\_sqft})$  in prediction did slightly a better job however overall it did not predict precisely either.