# **Basics of Monte Carlo simulations**

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#### **Outline**

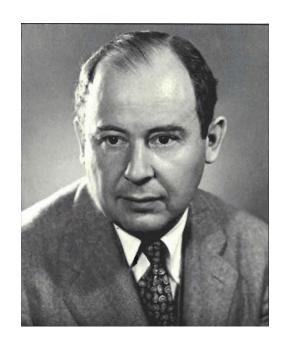
- History of Monte Carlo method
- Simple sampling Monte Carlo
- Importance sampling Monte Carlo
  - Markov process
  - Detailed balance
  - Metropolis algorithm
- Case study 2: 3D Lennard-Jones fluid

### **History of Monte Carlo methods**

- Stanislaw Ulam (re)invented the method in ~1947 during his work on Manhattan project at the Los Alamos National Laboratory.
- John von Neumann programmed the ENIAC computer to perform MC calculations.
- Nicholas Metropolis coined the name Monte Carlo, and later together with Marshall Rosenbluth reported adopting Metropolis algorithm to calculate the equation of state of 2D hard spheres.



Stanislaw Ulam (1909-1984)



John von Neumann (1903-1957)



Nicholas Metropolis (1915-1999)



Marshall Rosenbluth (1927-2003)

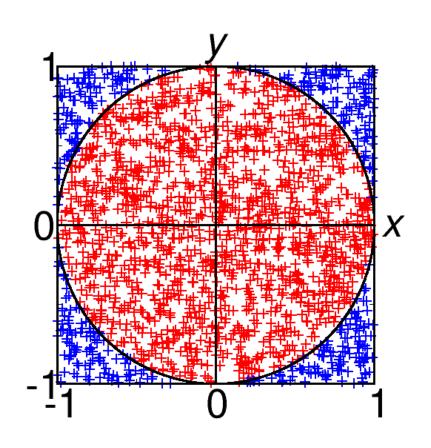
N. Metropolis & S. Ulam, J. Am. Stat. Assoc. 44(247): 335-341 (1949).

N. Metropolis, A. W. Rosenbluth, M. N. Rosenbluth, A. H. Teller & E. Teller, J. Chem. Phys. 21(6): 1087-1092 (1953).

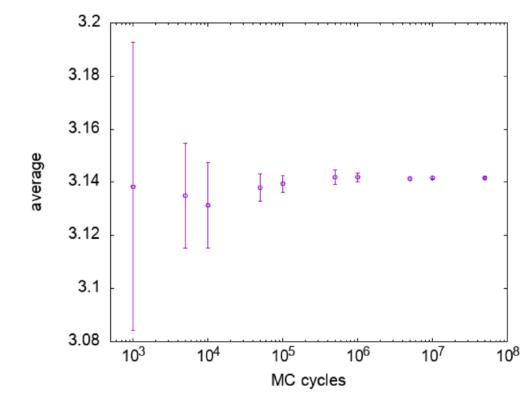
N. Metropolis, Los Alamos Science Special Issue, 15: 125-130 (1987).

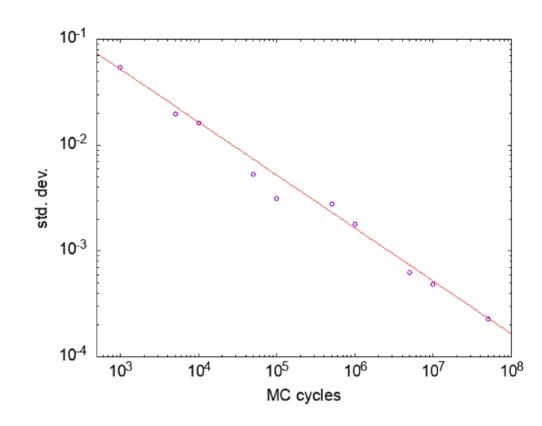
M. N. Rosenbluth, AIP Confer. Proc. 690(1): 22-30 (2003).

# Simple sampling: MC estimation of $\pi$

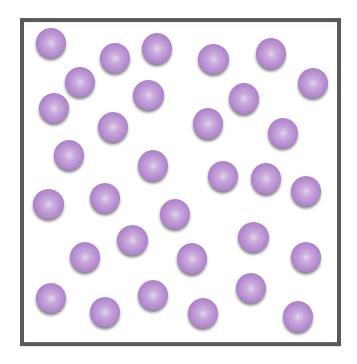


- Exact value:  $\pi = 4A_{\rm circle}/A_{\rm square}$
- MC estimate:  $\pi pprox 4N_{
  m red}/(N_{
  m red}+N_{
  m blue})$
- How many MC cycles do we need to get a good estimate?





### Simple sampling fails ...



224 hard discs

- Dimension of the **configuration space**: 448
- Each dim. sampled by 100 random points
- Total number of random configurations: 100<sup>448</sup>!
- Random config. may be statistically insignificant

### Importance sampling

- Idea: to only count the important configurations, i.e. microstates with considerable statistical weight
- Recall the basics of statistical mechanics

observable: 
$$\langle \mathcal{O} \rangle = \frac{\int d\mathbf{r}^N \mathcal{O}(\mathbf{r}^N) e^{-U(\mathbf{r}^N)/k_B T}}{\int d\mathbf{r}^N e^{-U(\mathbf{r}^N)/k_B T}}$$

config. part of partition function:  $Z = \int \mathrm{d}\mathbf{r}^N e^{-U(\mathbf{r}^N)/k_{\mathrm{B}}T}$ 

probability density function:  $p(\mathbf{r}^N) = e^{-U(\mathbf{r}^N)/k_{\rm B}T}/Z$ 

$$\langle \mathcal{O} \rangle = \int \mathcal{O}(\mathbf{r}^N) \underline{p(\mathbf{r}^N)} d\mathbf{r}^N$$

• Strategy: to generate M microstates according to the Boltzmann distribution  $p(\mathbf{r}^N)$ 

$$\langle \mathcal{O} 
angle pprox rac{1}{M} \sum_{i=1}^{M} \mathcal{O}_i$$

### Markov process for generating chain of states

- Stochastic process that has no memory, i.e. the probability for getting to a new state depends only on the present state, but NOT on the history how the present state is reached.
- Process is fully defined by the transition-probability matrix

$$P = \begin{bmatrix} P_{11} & P_{12} & \dots & P_{1M} \\ P_{21} & P_{22} & \dots & P_{2M} \\ \vdots & \vdots & \ddots & \vdots \\ P_{M1} & P_{M2} & \dots & P_{MM} \end{bmatrix}$$

with the requirements:

- each transition probability from i to j:  $0 \le P_{ii} \le 1$
- sum of each row is 1:  $\sum_{i} P_{ii} = 1$
- prob. of staying in the present state could be non-zero: P<sub>ii</sub> ≠ 0

How to construct the Markov transition-probability matrix?

### Stationary distribution of Markov process

 For a system in equilibrium, the transitions into and out of any state must be balanced:

$$\sum_{j} p_j P_{ji} = \sum_{j} p_i P_{ij} = p_i$$

**p** is the **stationary distribution** that does not change under further application of the transition-probability matrix **P**. Once the equilibrium is reached, the distribution will not change any more.

- An arbitrary initial distribution will converge to the stationary (equilibrium) distribution p if one repeats the Markov process, given that P is irreducible and aperiodic.
  - irreducible: any state is reachable from any other state in finite steps, i.e.  $[P^m]_{ij} > 0$  for any ij pair condition of ergodicity 各态历经
  - aperiodic: if there exists a state i such that  $P_{ii} > 0$ , P is aperiodic. E.g.,  $P = \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix}$  has a stationary distribution  $p = (0.5 \ 0.5)$ , but  $p^* \cdot P^m$  alternates between  $p^*$  and  $I - p^*$ .

#### Condition of detailed balance

• Detailed balance  $p_j P_{jj} = p_i P_{ij}$  is a sufficient but unnecessary condition for

 $\sum_{j} p_j P_{ji} = \sum_{j} p_i P_{ij} = p_i$ 

and widely used to construct the transition-probability matrix P.

Under the conditions of ergodicity and detailed balance, one uses
the Markov process to generate a chain of states. If the probability of
remaining in a state is non-zero, i.e. P<sub>ii</sub> ≠ 0, one will end up with
states that obey the equilibrium distribution p.

### Example of a three-state system

- Equilibrium distribution:  $p = (p_1 p_2 p_3) = (0.25 \ 0.5 \ 0.25)$
- For a transition-probability matrix

$$P = \begin{pmatrix} P_{11} & P_{12} & P_{13} \\ P_{21} & P_{22} & P_{23} \\ P_{31} & P_{32} & P_{33} \end{pmatrix} = \begin{pmatrix} 0 & 0.5 & 0.5 \\ 0.25 & 0.5 & 0.25 \\ 0.5 & 0.5 & 0 \end{pmatrix}$$

that satisfies the detailed-balance condition  $p_j P_{ji} = p_i P_{ij}$ ,

$$p \cdot P = (0.25 \ 0.5 \ 0.25) = p$$

 p is the stationary or limiting distribution of the transitionprobability matrix.

### Example of a three-state system

- Start with an initial distribution  $p' = (0 \ 0.05 \ 0.95)$
- Application of the transition-probability matrix

$$p' \cdot P = (0 \ 0.05 \ 0.95) \begin{pmatrix} 0 & 0.5 & 0.5 \\ 0.25 & 0.5 & 0.25 \\ 0.5 & 0.5 & 0 \end{pmatrix} = (0.4875 \ 0.5 \ 0.0125)$$
$$p' \cdot P^{10} = (0.249536 \ 0.5 \ 0.250464)$$

$$p' \cdot P^{20} = (0.25 \ 0.5 \ 0.25)$$

leads to the stationary distribution.

### Metropolis algorithm

 Transition probability = probability for selecting a transition X probability for accepting the transition:

$$p_{i}\underline{P_{ij}} = p_{j}\underline{P_{ji}}$$
$$p_{i}\underline{g_{ij}}\underline{A_{ij}} = p_{j}\underline{g_{ji}}\underline{A_{ji}}$$

Metropolis solution for the acceptance ratio:

$$A_{ij} = \min\left\{1, \frac{p_j g_{ji}}{p_i g_{ij}}\right\}, A_{ji} = \min\left\{1, \frac{p_i g_{ij}}{p_j g_{ji}}\right\}$$

which reduces for symmetric selection probability  $g_{ij} = g_{ji}$  to

$$A_{ij} = \min\{1, p_j/p_i\}, A_{ji} = \min\{1, p_i/p_j\}$$

Other but less common solutions exist, e.g., Glauber algorithm

$$A_{ij} = \frac{p_j g_{ji}}{p_j g_{ji} + p_i g_{ij}} = \frac{1}{1 + \frac{p_i g_{ij}}{p_j g_{ji}}}$$

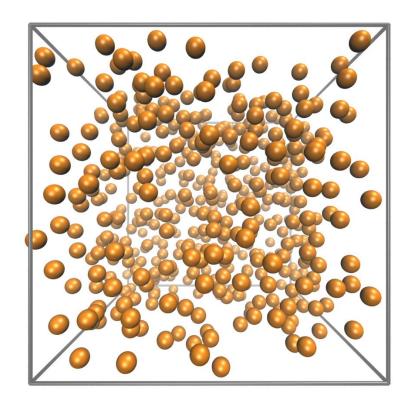
## Case study: MC simulations of 3D LJ fluid

- Forward move to generate a new state:
  - choose a particle x, prob. = 1/N
  - translate x around old position, prob.  $\propto 1/(2d)^3$
  - accept the move with prob. =  $min\{1, e^{-\beta(Un-Uo)}\}$
- Reverse move following the same protocol:
  - prob. of choosing particle x = 1/N
  - prob. of translating the particle back  $\propto 1/(2d)^3$
  - prob. of accepting the move =  $min\{1, e^{-\beta(Uo-Un)}\}$

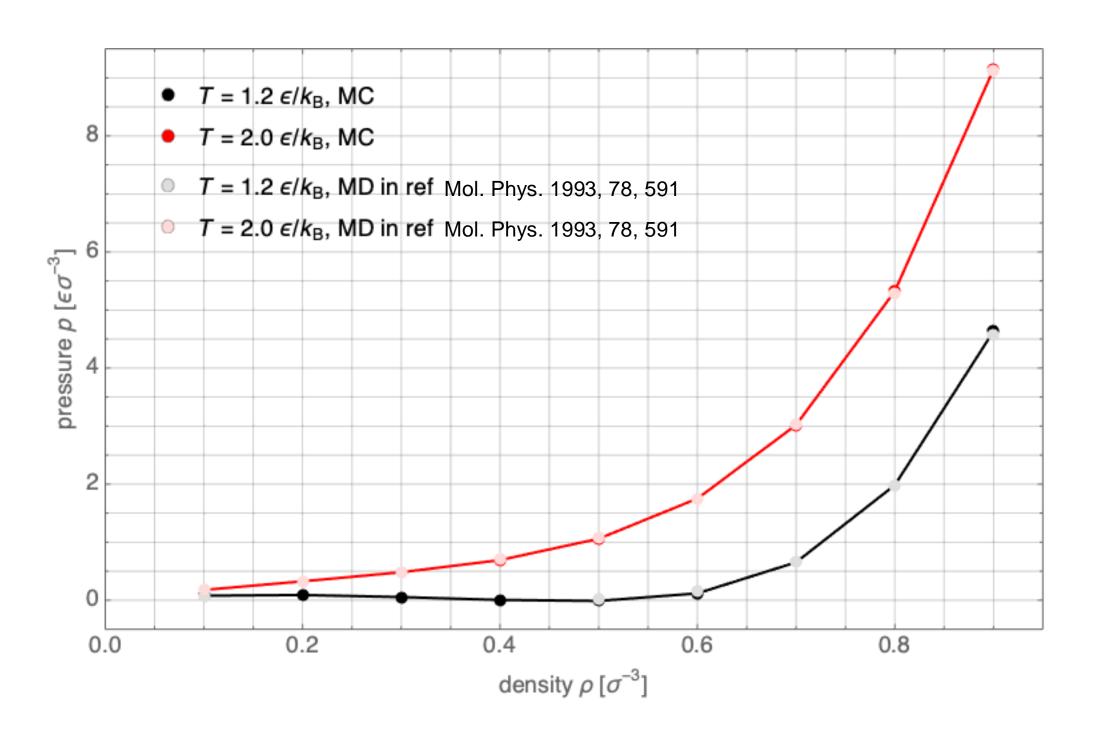


Check detailed-balance condition

$$\frac{e^{-\beta U_{\rm o}} d\mathbf{r}_{\rm o}}{Z} \cdot \frac{1}{N} \cdot \frac{1}{(2d)^3} \cdot \min\{1, e^{-\beta (U_{\rm n} - U_{\rm o})}\} = \frac{e^{-\beta U_{\rm n}} d\mathbf{r}_{\rm n}}{Z} \cdot \frac{1}{N} \cdot \frac{1}{(2d)^3} \cdot \min\{1, e^{-\beta (U_{\rm o} - U_{\rm n})}\}$$



# Case study: MC simulations of 3D LJ fluid



#### Remarks on MC simulations

- A MC move often involves a *local update* of the current configuration, in contrast to the global update in molecular dynamics simulations.
- Be careful with rotational moves in MC simulations.
- There are MC schemes that violate detailed-balance condition, but can still correctly sample the equilibrium distribution, e.g., sequential updating in Ising model.
- MC simulations can be used for the measurement of dynamical properties in a system if its dynamics is stochastic in nature and the MC moves are not unphysical.