

機械学習 Machine Learning 確率分布 Probability Distributions

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離散変数 Discrete variables ベルヌーイ分布 Bernoulli distribution

$$x \in \{0,1\}.$$
 $0 \le \mu \le 1$, $p(x=1|\mu) = \mu$ $p(x=0|\mu) = 1 - \mu$ の $p(x=0|$

x 上の確率分布 The probability distribution over x (ベルヌーイ分布 Bernoulli distribution)

$$\mathrm{Bern}(x|\mu) = \mu^x (1-\mu)^{1-x}$$

$$\mathbb{E}[x] = \mu \qquad \qquad$$
平均 Mean
$$\mathrm{var}[x] = \mu(1-\mu) \qquad$$
分散 Variance

離散変数 Discrete variables ベルヌーイ分布 Bernoulli distribution

$$\mathcal{D} = \{x_1, \dots, x_N\}$$
 x の観測値 Observed values of x

尤度関数 Likelihood function

$$p(\mathcal{D}|\mu) = \prod_{n=1}^{N} p(x_n|\mu) = \prod_{n=1}^{N} \mu^{x_n} (1-\mu)^{1-x_n}$$

対数尤度関数 Log likelihood function

$$\ln p(\mathcal{D}|\mu) = \sum_{n=1}^{N} \ln p(x_n|\mu) = \sum_{n=1}^{N} \{x_n \ln \mu + (1-x_n) \ln(1-\mu)\}$$

最尤推定量 Maximum likelihood estimator

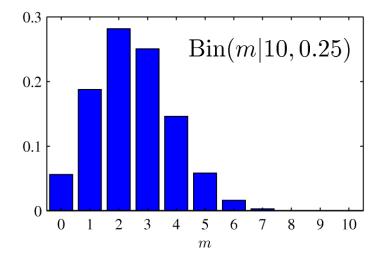
$$x=1$$
の回数

The number of observations of

$$\mu_{\rm ML} = \frac{1}{N} \sum_{n=1}^{N} x_n = \frac{m}{N}$$

m の分布 The distribution of m (二項分布 Binomial distribution)

$$Bin(m|N,\mu) = \binom{N}{m} \mu^m (1-\mu)^{N-m}$$
$$\binom{N}{m} \equiv \frac{N!}{(N-m)!m!}$$



$$\mathbb{E}[m] \equiv \sum_{m=0}^{N} m \operatorname{Bin}(m|N,\mu) = N\mu$$

$$\operatorname{var}[m] \equiv \sum_{m=0}^{N} (m - \mathbb{E}[m])^{2} \operatorname{Bin}(m|N,\mu) = N\mu(1-\mu)$$

1対K法 1-of-K scheme

K: 互いに排他的で、取ることが可能な状態の数
The number of possible mutually exclusive states

例 Example:
$$K=6$$
 $x_3=1$

$$\mathbf{x} = (0, 0, 1, 0, 0, 0)^{\mathrm{T}}$$
 $\sum_{k=1}^{K} x_k = 1$

$$\mu_k$$
: $x_k = 1$ となる確率
$$\mu_k \geqslant 0 \qquad \sum_k \mu_k = 1$$
 The probability of $x_k = 1$

x の分布 The distribution of x

$$p(\mathbf{x}|\boldsymbol{\mu}) = \prod_{k=1}^{K} \mu_k^{x_k} \qquad \boldsymbol{\mu} = (\mu_1, \dots, \mu_K)^{\mathrm{T}}$$

$$\mathbb{E}[\mathbf{x}|\boldsymbol{\mu}] = \sum_{\mathbf{x}} p(\mathbf{x}|\boldsymbol{\mu})\mathbf{x} = (\mu_1, \dots, \mu_K)^{\mathrm{T}} = \boldsymbol{\mu}$$
$$\sum_{\mathbf{x}} p(\mathbf{x}|\boldsymbol{\mu}) = \sum_{k=1}^{K} \mu_k = 1$$

確率分布

 $\mathcal{D} = \{\mathbf{x}_1, \dots, \mathbf{x}_N\}$ x の観測値 Observed values of \mathbf{x}

尤度関数 Likelihood function

$$p(\mathcal{D}|\boldsymbol{\mu}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \mu_k^{x_{nk}} = \prod_{k=1}^{K} \mu_k^{(\sum_n x_{nk})} = \prod_{k=1}^{K} \mu_k^{m_k}$$

 m_1, \ldots, m_K の分布 The distribution of m_1, \ldots, m_K (多項分布 Multinomial distribution)

$$Mult(m_1, m_2, ..., m_K | \boldsymbol{\mu}, N) = \binom{N}{m_1 m_2 ... m_K} \prod_{k=1}^K \mu_k^{m_k}$$
$$\binom{N}{m_1 m_2 ... m_K} = \frac{N!}{m_1! m_2! ... m_K!} \sum_{k=1}^K m_k = N$$

$$\mathbb{E}[m_k] = N\mu_k$$

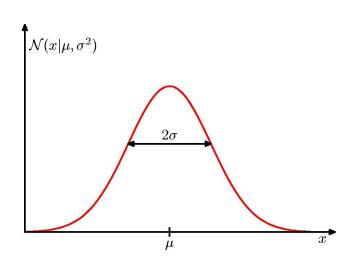
$$\operatorname{var}[m_k] = N\mu_k(1 - \mu_k)$$

$$\operatorname{cov}[m_j m_k] = -N\mu_j \mu_k$$

多変量ガウス分布 The multivariate Gaussian

単変量ガウス分布 Univariate Gaussian distribution

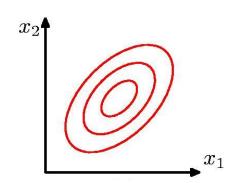
$$\mathcal{N}(x|\mu,\sigma^2) = \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left\{-\frac{1}{2\sigma^2}(x-\mu)^2\right\}$$



多変量ガウス分布

Multivariate Gaussian distribution

$$\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\boldsymbol{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right\}$$



多変量ガウス分布 The multivariate Gaussian

中心極限定理 Central limit theorem Walker (1969)

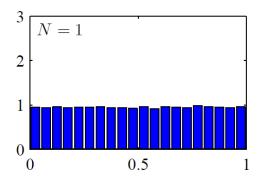
いくつかの確率変数の和は、足し合わされる変数の数が増えるに従って、徐々にガウス分布に従うようになる.

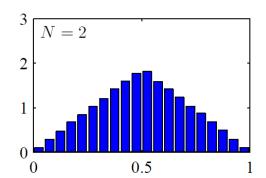
The sum of a set of random variables has a distribution that becomes increasingly Gaussian as the number of terms in the sum increases.

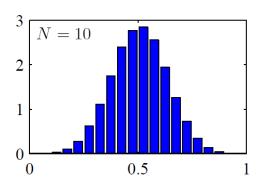
例:N個の一様に分布する量の平均

 $(x_1 + \cdots + x_N)/N$

E.g., the mean of N uniformly distributed numbers







多変量ガウス分布 The multivariate Gaussian ガウス分布の幾何 Geometry of the Gaussian

Δ:マハラノビス距離 Mahalanobis distance

$$\Delta^2 = (\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})$$

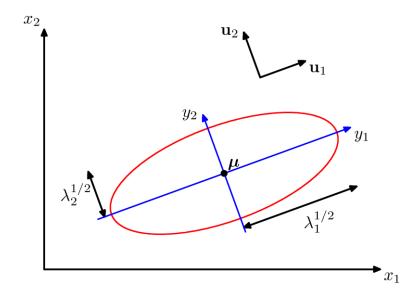
$$\mathbf{\Sigma}\mathbf{u}_i = \lambda_i \mathbf{u}_i \quad \mathbf{\Sigma} = \sum_{i=1}^D \lambda_i \mathbf{u}_i \mathbf{u}_i^{\mathrm{T}}$$

$$\mathbf{\Sigma}^{-1} = \sum_{i=1}^D \frac{1}{\lambda_i} \mathbf{u}_i \mathbf{u}_i^{\mathrm{T}}$$

$$\Delta^2 = \sum_{i=1}^{D} \frac{y_i^2}{\lambda_i}$$
 $y_i = \mathbf{u}_i^{\mathrm{T}}(\mathbf{x} - \boldsymbol{\mu})$

$$\mathbf{u}_{i}^{\mathrm{T}}\mathbf{u}_{j} = I_{ij}$$

$$I_{ij} = \begin{cases} 1, & \text{if } i = j \\ 0, & \text{otherwise} \end{cases}$$



多変量ガウス分布 The multivariate Gaussian ガウス分布の幾何 Geometry of the Gaussian

$$p(\mathbf{x}) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\mathbf{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} \mathbf{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right\}$$

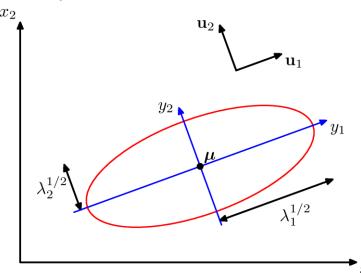
$$p(\mathbf{y}) = p(\mathbf{x})|\mathbf{J}| = \prod_{j=1}^{D} \frac{1}{(2\pi\lambda_j)^{1/2}} \exp\left\{-\frac{y_j^2}{2\lambda_j}\right\} \qquad \mathbf{y} = (y_1, \dots, y_D)^{\mathrm{T}}$$

$$J_{ij} = \frac{\partial x_i}{\partial y_j} = U_{ji}$$
 $\mathbf{y} = \mathbf{U}(\mathbf{x} - \boldsymbol{\mu})$

$$|\mathbf{J}|^2 = |\mathbf{U}^{\mathrm{T}}|^2 = |\mathbf{U}^{\mathrm{T}}| |\mathbf{U}| = |\mathbf{U}^{\mathrm{T}}\mathbf{U}| = |\mathbf{I}| = 1$$

$$|\mathbf{J}| = 1$$

$$|\mathbf{\Sigma}|^{1/2} = \prod_{j=1}^D \lambda_j^{1/2}$$



多変量ガウス分布 The multivariate Gaussian モーメント Moments

1次モーメント First-order moment

$$\mathbb{E}[\mathbf{x}] = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\mathbf{\Sigma}|^{1/2}} \int \exp\left\{-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} \mathbf{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right\} \mathbf{x} \, d\mathbf{x}$$

$$= \frac{1}{(2\pi)^{D/2}} \frac{1}{|\mathbf{\Sigma}|^{1/2}} \int \exp\left\{-\frac{1}{2} \mathbf{z}^{\mathrm{T}} \mathbf{\Sigma}^{-1} \mathbf{z}\right\} (\mathbf{z} + \boldsymbol{\mu}) \, d\mathbf{z}$$

$$= \boldsymbol{\mu}$$

2次モーメント Second-order moment

$$\mathbb{E}[\mathbf{x}\mathbf{x}^{\mathrm{T}}] = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\mathbf{\Sigma}|^{1/2}} \int \exp\left\{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} \mathbf{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right\} \mathbf{x}\mathbf{x}^{\mathrm{T}} d\mathbf{x}$$
$$= \frac{1}{(2\pi)^{D/2}} \frac{1}{|\mathbf{\Sigma}|^{1/2}} \int \exp\left\{-\frac{1}{2}\mathbf{z}^{\mathrm{T}} \mathbf{\Sigma}^{-1} \mathbf{z}\right\} (\mathbf{z} + \boldsymbol{\mu}) (\mathbf{z} + \boldsymbol{\mu})^{\mathrm{T}} d\mathbf{z}$$

多変量ガウス分布 The multivariate Gaussian モーメント Moments

2次モーメント Second-order moment

$$\mathbb{E}[\mathbf{x}\mathbf{x}^{\mathrm{T}}] = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\mathbf{\Sigma}|^{1/2}} \int \exp\left\{-\frac{1}{2}\mathbf{z}^{\mathrm{T}}\mathbf{\Sigma}^{-1}\mathbf{z}\right\} (\mathbf{z} + \boldsymbol{\mu})(\mathbf{z} + \boldsymbol{\mu})^{\mathrm{T}} d\mathbf{z} \qquad y_{j} = \mathbf{u}_{j}^{\mathrm{T}}\mathbf{z}$$

$$\frac{1}{(2\pi)^{D/2}} \frac{1}{|\mathbf{\Sigma}|^{1/2}} \int \exp\left\{-\frac{1}{2}\mathbf{z}^{\mathrm{T}}\mathbf{\Sigma}^{-1}\mathbf{z}\right\} \mathbf{z}\mathbf{z}^{\mathrm{T}} d\mathbf{z} \qquad \mathbf{z} = \sum_{j=1}^{D} y_{j}\mathbf{u}_{j}$$

$$= \frac{1}{(2\pi)^{D/2}} \frac{1}{|\mathbf{\Sigma}|^{1/2}} \sum_{i=1}^{D} \sum_{j=1}^{D} \mathbf{u}_{i}\mathbf{u}_{j}^{\mathrm{T}} \int \exp\left\{-\sum_{k=1}^{D} \frac{y_{k}^{2}}{2\lambda_{k}}\right\} y_{i}y_{j} d\mathbf{y} \qquad \mathbf{\Sigma}^{-1} = \sum_{i=1}^{D} \frac{1}{\lambda_{i}} \mathbf{u}_{i}\mathbf{u}_{i}^{\mathrm{T}}$$

$$= \sum_{i=1}^{D} \mathbf{u}_{i}\mathbf{u}_{i}^{\mathrm{T}}\lambda_{i} = \mathbf{\Sigma}$$

$$|\mathbf{\Sigma}|^{1/2} = \prod_{j=1}^{D} \lambda_{j}^{1/2}$$

$$\mathbb{E}[\mathbf{x}\mathbf{x}^{\mathrm{T}}] = \boldsymbol{\mu}\boldsymbol{\mu}^{\mathrm{T}} + \boldsymbol{\Sigma}$$

多変量ガウス分布 The multivariate Gaussian モーメント Moments

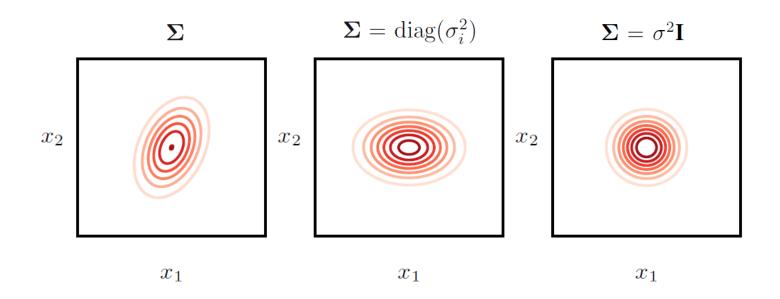
共分散 Covariance

$$\operatorname{cov}[\mathbf{x}] = \mathbb{E}\left[(\mathbf{x} - \mathbb{E}[\mathbf{x}])(\mathbf{x} - \mathbb{E}[\mathbf{x}])^{\mathrm{T}} \right]$$

$$\mathbb{E}[\mathbf{x}] = oldsymbol{\mu} \qquad \mathbb{E}[\mathbf{x}\mathbf{x}^{\mathrm{T}}] = oldsymbol{\mu}oldsymbol{\mu}^{\mathrm{T}} + oldsymbol{\Sigma}$$

$$cov[\mathbf{x}] = \mathbf{\Sigma}$$

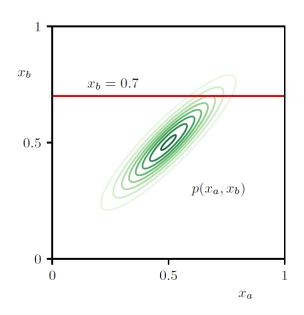
多変量ガウス分布 The multivariate Gaussian 制限 Limitations

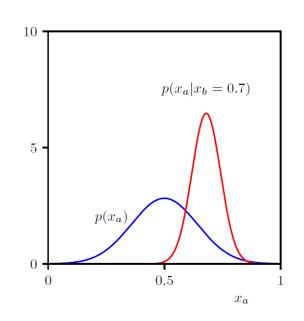


自由パラメータの数 The number of free parameters

$$D(D+1)/2 + D$$

$$D+1$$





$$egin{aligned} p(\mathbf{x}) &= \mathcal{N}(\mathbf{x}|oldsymbol{\mu}, oldsymbol{\Sigma}) \ \mathbf{x} &= egin{pmatrix} \mathbf{x}_a \ \mathbf{x}_b \end{pmatrix} \ oldsymbol{\mu} &= egin{pmatrix} oldsymbol{\mu}_a \ oldsymbol{\mu}_b \end{pmatrix} \ oldsymbol{\Sigma} &= egin{pmatrix} oldsymbol{\Sigma}_{aa} & oldsymbol{\Sigma}_{ab} \ oldsymbol{\Sigma}_{ba} & oldsymbol{\Sigma}_{bb} \end{pmatrix} \end{aligned}$$

$$p(\mathbf{x}_a|\mathbf{x}_b)$$
 ?

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) \quad \mathbf{x} = \begin{pmatrix} \mathbf{x}_a \\ \mathbf{x}_b \end{pmatrix} \quad \boldsymbol{\mu} = \begin{pmatrix} \boldsymbol{\mu}_a \\ \boldsymbol{\mu}_b \end{pmatrix} \quad \boldsymbol{\Sigma} = \begin{pmatrix} \boldsymbol{\Sigma}_{aa} & \boldsymbol{\Sigma}_{ab} \\ \boldsymbol{\Sigma}_{ba} & \boldsymbol{\Sigma}_{bb} \end{pmatrix}$$
 $\boldsymbol{\Lambda} \equiv \boldsymbol{\Sigma}^{-1} \quad \boldsymbol{\Lambda} = \begin{pmatrix} \boldsymbol{\Lambda}_{aa} & \boldsymbol{\Lambda}_{ab} \\ \boldsymbol{\Lambda}_{ba} & \boldsymbol{\Lambda}_{bb} \end{pmatrix}$

$$\begin{split} -\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu}) &= \\ -\frac{1}{2}(\mathbf{x}_{a} - \boldsymbol{\mu}_{a})^{\mathrm{T}} \boldsymbol{\Lambda}_{aa}(\mathbf{x}_{a} - \boldsymbol{\mu}_{a}) - \frac{1}{2}(\mathbf{x}_{a} - \boldsymbol{\mu}_{a})^{\mathrm{T}} \boldsymbol{\Lambda}_{ab}(\mathbf{x}_{b} - \boldsymbol{\mu}_{b}) \\ -\frac{1}{2}(\mathbf{x}_{b} - \boldsymbol{\mu}_{b})^{\mathrm{T}} \boldsymbol{\Lambda}_{ba}(\mathbf{x}_{a} - \boldsymbol{\mu}_{a}) - \frac{1}{2}(\mathbf{x}_{b} - \boldsymbol{\mu}_{b})^{\mathrm{T}} \boldsymbol{\Lambda}_{bb}(\mathbf{x}_{b} - \boldsymbol{\mu}_{b}) \end{split}$$

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) \qquad -\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu}) = -\frac{1}{2} \mathbf{x}^{\mathrm{T}} \underline{\boldsymbol{\Sigma}}^{-1} \mathbf{x} + \mathbf{x}^{\mathrm{T}} \underline{\boldsymbol{\Sigma}}^{-1} \boldsymbol{\mu} + \text{const}$$

$$p(\mathbf{x}_{a}|\mathbf{x}_{b}) = \mathcal{N}(\mathbf{x}_{a}|\boldsymbol{\mu}_{a|b}, \boldsymbol{\Sigma}_{a|b})$$

$$-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu}) =$$

$$-\frac{1}{2}(\mathbf{x}_{a} - \boldsymbol{\mu}_{a})^{\mathrm{T}} \boldsymbol{\Lambda}_{aa}(\mathbf{x}_{a} - \boldsymbol{\mu}_{a}) - \frac{1}{2}(\mathbf{x}_{a} - \boldsymbol{\mu}_{a})^{\mathrm{T}} \boldsymbol{\Lambda}_{ab}(\mathbf{x}_{b} - \boldsymbol{\mu}_{b})$$

$$-\frac{1}{2}(\mathbf{x}_{b} - \boldsymbol{\mu}_{b})^{\mathrm{T}} \boldsymbol{\Lambda}_{ba}(\mathbf{x}_{a} - \boldsymbol{\mu}_{a}) - \frac{1}{2}(\mathbf{x}_{b} - \boldsymbol{\mu}_{b})^{\mathrm{T}} \boldsymbol{\Lambda}_{bb}(\mathbf{x}_{b} - \boldsymbol{\mu}_{b})$$

$$= -\frac{1}{2} \mathbf{x}_{a}^{\mathrm{T}} \underline{\boldsymbol{\Lambda}}_{aa} \mathbf{x}_{a} + \mathbf{x}_{a}^{\mathrm{T}} \underbrace{\{\boldsymbol{\Lambda}_{aa} \boldsymbol{\mu}_{a} - \boldsymbol{\Lambda}_{ab}(\mathbf{x}_{b} - \boldsymbol{\mu}_{b})\}}_{||} + \text{const}$$

$$\boldsymbol{\Sigma}_{a|b}^{-1} \boldsymbol{\mu}_{a|b}$$

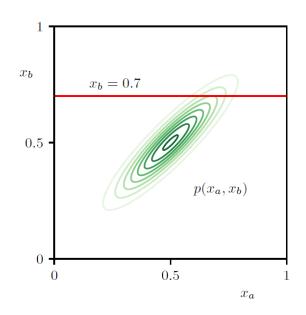
$$oldsymbol{\Lambda}_{aa}$$
 \parallel $oldsymbol{\Sigma}_{a|b}^{-1}$

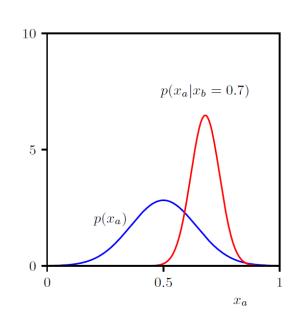
$$\Sigma_{a|b} = \Lambda_{aa}^{-1}$$

$$egin{aligned} \{oldsymbol{\Lambda}_{aa}oldsymbol{\mu}_a - oldsymbol{\Lambda}_{ab}(\mathbf{x}_b - oldsymbol{\mu}_b)\} \ & & \parallel \ & oldsymbol{\Sigma}_{a|b}^{-1}oldsymbol{\mu}_{a|b} \end{aligned}$$



$$\begin{array}{lcl} \boldsymbol{\mu}_{a|b} & = & \boldsymbol{\Sigma}_{a|b} \left\{ \boldsymbol{\Lambda}_{aa} \boldsymbol{\mu}_{a} - \boldsymbol{\Lambda}_{ab} (\mathbf{x}_{b} - \boldsymbol{\mu}_{b}) \right\} \\ & = & \boldsymbol{\mu}_{a} - \boldsymbol{\Lambda}_{aa}^{-1} \boldsymbol{\Lambda}_{ab} (\mathbf{x}_{b} - \boldsymbol{\mu}_{b}) \end{array}$$





$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma})$$
 $\mathbf{x} = \begin{pmatrix} \mathbf{x}_a \\ \mathbf{x}_b \end{pmatrix}$
 $\boldsymbol{\mu} = \begin{pmatrix} \boldsymbol{\mu}_a \\ \boldsymbol{\mu}_b \end{pmatrix}$
 $\boldsymbol{\Sigma} = \begin{pmatrix} \boldsymbol{\Sigma}_{aa} & \boldsymbol{\Sigma}_{ab} \\ \boldsymbol{\Sigma}_{ba} & \boldsymbol{\Sigma}_{bb} \end{pmatrix}$

 $p(\mathbf{x}_a)$?

$$p(\mathbf{x}_a) = \int p(\mathbf{x}_a, \mathbf{x}_b) d\mathbf{x}_b$$

$$-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) =$$

$$-\frac{1}{2} (\mathbf{x}_a - \boldsymbol{\mu}_a)^{\mathrm{T}} \boldsymbol{\Lambda}_{aa} (\mathbf{x}_a - \boldsymbol{\mu}_a) - \frac{1}{2} (\mathbf{x}_a - \boldsymbol{\mu}_a)^{\mathrm{T}} \boldsymbol{\Lambda}_{ab} (\mathbf{x}_b - \boldsymbol{\mu}_b)$$

$$-\frac{1}{2} (\mathbf{x}_b - \boldsymbol{\mu}_b)^{\mathrm{T}} \boldsymbol{\Lambda}_{ba} (\mathbf{x}_a - \boldsymbol{\mu}_a) - \frac{1}{2} (\mathbf{x}_b - \boldsymbol{\mu}_b)^{\mathrm{T}} \boldsymbol{\Lambda}_{bb} (\mathbf{x}_b - \boldsymbol{\mu}_b)$$

 \mathbf{x}_b を含む項 Terms that involve \mathbf{x}_b

$$-\frac{1}{2}\mathbf{x}_b^{\mathrm{T}}\mathbf{\Lambda}_{bb}\mathbf{x}_b + \mathbf{x}_b^{\mathrm{T}}\mathbf{m} \qquad \mathbf{m} = \mathbf{\Lambda}_{bb}\boldsymbol{\mu}_b - \mathbf{\Lambda}_{ba}(\mathbf{x}_a - \boldsymbol{\mu}_a)$$
$$= -\frac{1}{2}(\mathbf{x}_b - \mathbf{\Lambda}_{bb}^{-1}\mathbf{m})^{\mathrm{T}}\mathbf{\Lambda}_{bb}(\mathbf{x}_b - \mathbf{\Lambda}_{bb}^{-1}\mathbf{m}) + \frac{1}{2}\mathbf{m}^{\mathrm{T}}\mathbf{\Lambda}_{bb}^{-1}\mathbf{m}$$

$$\int \exp\left\{-\frac{1}{2}(\mathbf{x}_b - \boldsymbol{\Lambda}_{bb}^{-1}\mathbf{m})^{\mathrm{T}}\boldsymbol{\Lambda}_{bb}(\mathbf{x}_b - \boldsymbol{\Lambda}_{bb}^{-1}\mathbf{m})\right\} d\mathbf{x}_b \quad \mathbf{\Box} \qquad \text{const}$$

$$p(\mathbf{x}_a) = \int p(\mathbf{x}_a, \mathbf{x}_b) \, \mathrm{d}\mathbf{x}_b$$

 \mathbf{x}_a に依存する項 Terms that depend on \mathbf{x}_a

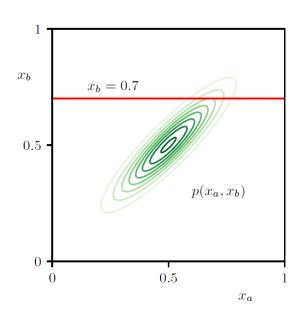
$$\frac{1}{2} \left[\mathbf{\Lambda}_{bb} \boldsymbol{\mu}_b - \mathbf{\Lambda}_{ba} (\mathbf{x}_a - \boldsymbol{\mu}_a) \right]^{\mathrm{T}} \mathbf{\Lambda}_{bb}^{-1} \left[\mathbf{\Lambda}_{bb} \boldsymbol{\mu}_b - \mathbf{\Lambda}_{ba} (\mathbf{x}_a - \boldsymbol{\mu}_a) \right] \leftarrow \frac{1}{2} \mathbf{m}^{\mathrm{T}} \mathbf{\Lambda}_{bb}^{-1} \mathbf{m}
- \frac{1}{2} \mathbf{x}_a^{\mathrm{T}} \mathbf{\Lambda}_{aa} \mathbf{x}_a + \mathbf{x}_a^{\mathrm{T}} (\mathbf{\Lambda}_{aa} \boldsymbol{\mu}_a + \mathbf{\Lambda}_{ab} \boldsymbol{\mu}_b) + \text{const}
= - \frac{1}{2} \mathbf{x}_a^{\mathrm{T}} \left(\mathbf{\Lambda}_{aa} - \mathbf{\Lambda}_{ab} \mathbf{\Lambda}_{bb}^{-1} \mathbf{\Lambda}_{ba} \right) \mathbf{x}_a + \mathbf{x}_a^{\mathrm{T}} \left(\mathbf{\Lambda}_{aa} - \mathbf{\Lambda}_{ab} \mathbf{\Lambda}_{bb}^{-1} \mathbf{\Lambda}_{ba} \right) \boldsymbol{\mu}_a + \text{const}
\mathbf{\Sigma}_a^{-1} \qquad \mathbf{\Sigma}_a^{-1} \boldsymbol{\mu}_a$$

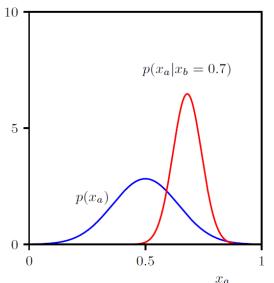
$$\Sigma_{a} = (\Lambda_{aa} - \Lambda_{ab}\Lambda_{bb}^{-1}\Lambda_{ba})^{-1} = \Sigma_{aa} \qquad \Sigma_{a}(\Lambda_{aa} - \Lambda_{ab}\Lambda_{bb}^{-1}\Lambda_{ba})\mu_{a} = \mu_{a}$$

$$\begin{pmatrix} \Lambda_{aa} & \Lambda_{ab} \\ \Lambda_{ba} & \Lambda_{bb} \end{pmatrix}^{-1} = \begin{pmatrix} \Sigma_{aa} & \Sigma_{ab} \\ \Sigma_{ba} & \Sigma_{bb} \end{pmatrix}$$

$$\begin{pmatrix} \mathbf{A} & \mathbf{B} \\ \mathbf{C} & \mathbf{D} \end{pmatrix}^{-1} = \begin{pmatrix} \mathbf{M} & -\mathbf{M}\mathbf{B}\mathbf{D}^{-1} \\ -\mathbf{D}^{-1}\mathbf{C}\mathbf{M} & \mathbf{D}^{-1} + \mathbf{D}^{-1}\mathbf{C}\mathbf{M}\mathbf{B}\mathbf{D}^{-1} \end{pmatrix}$$

$$\mathbf{M} = (\mathbf{A} - \mathbf{B}\mathbf{D}^{-1}\mathbf{C})^{-1}$$





$$p(x_a|x_b = 0.7)$$

$$p(x_a)$$

$$0.5$$

$$x_a$$

$$egin{aligned} p(\mathbf{x}_a|\mathbf{x}_b) &= \mathcal{N}(\mathbf{x}_a|oldsymbol{\mu}_{a|b}, oldsymbol{\Sigma}_{a|b}) \ oldsymbol{\mu}_{a|b} &= oldsymbol{\mu}_a - oldsymbol{\Lambda}_{aa}^{-1} oldsymbol{\Lambda}_{ab} (\mathbf{x}_b - oldsymbol{\mu}_b) \ oldsymbol{\Sigma}_{a|b} &= oldsymbol{\Lambda}_{aa}^{-1} \end{aligned}$$

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma})$$
 $\mathbf{x} = \begin{pmatrix} \mathbf{x}_a \\ \mathbf{x}_b \end{pmatrix}$
 $\boldsymbol{\mu} = \begin{pmatrix} \boldsymbol{\mu}_a \\ \boldsymbol{\mu}_b \end{pmatrix}$
 $\boldsymbol{\Sigma} = \begin{pmatrix} \boldsymbol{\Sigma}_{aa} & \boldsymbol{\Sigma}_{ab} \\ \boldsymbol{\Sigma}_{ba} & \boldsymbol{\Sigma}_{bb} \end{pmatrix}$
 $\boldsymbol{\Lambda} \equiv \boldsymbol{\Sigma}^{-1}$
 $\boldsymbol{\Lambda} = \begin{pmatrix} \boldsymbol{\Lambda}_{aa} & \boldsymbol{\Lambda}_{ab} \\ \boldsymbol{\Lambda}_{ba} & \boldsymbol{\Lambda}_{bb} \end{pmatrix}$

$$p(\mathbf{x}_a) = \mathcal{N}(\mathbf{x}_a | \boldsymbol{\mu}_a, \boldsymbol{\Sigma}_{aa})$$

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Lambda}^{-1})$$

$$p(\mathbf{y}|\mathbf{x}) = \mathcal{N}(\mathbf{y}|\mathbf{A}\mathbf{x} + \mathbf{b}, \mathbf{L}^{-1})$$

$$p(\mathbf{y}|\mathbf{y}) ?$$

$$\mathbf{z} = \begin{pmatrix} \mathbf{x} \\ \mathbf{y} \end{pmatrix} \qquad \ln p(\mathbf{z}) = \ln p(\mathbf{x}) + \ln p(\mathbf{y}|\mathbf{x})$$
$$= -\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} \boldsymbol{\Lambda}(\mathbf{x} - \boldsymbol{\mu})$$
$$-\frac{1}{2}(\mathbf{y} - \mathbf{A}\mathbf{x} - \mathbf{b})^{\mathrm{T}} \mathbf{L}(\mathbf{y} - \mathbf{A}\mathbf{x} - \mathbf{b}) + \text{const}$$

 $\ln p(\mathbf{z})$ の2次の項 The second-order terms in $\ln p(\mathbf{z})$

$$-\frac{1}{2}\mathbf{x}^{\mathrm{T}}(\mathbf{\Lambda} + \mathbf{A}^{\mathrm{T}}\mathbf{L}\mathbf{A})\mathbf{x} - \frac{1}{2}\mathbf{y}^{\mathrm{T}}\mathbf{L}\mathbf{y} + \frac{1}{2}\mathbf{y}^{\mathrm{T}}\mathbf{L}\mathbf{A}\mathbf{x} + \frac{1}{2}\mathbf{x}^{\mathrm{T}}\mathbf{A}^{\mathrm{T}}\mathbf{L}\mathbf{y}$$

$$= -\frac{1}{2}\begin{pmatrix} \mathbf{x} \\ \mathbf{y} \end{pmatrix}^{\mathrm{T}} \begin{pmatrix} \mathbf{\Lambda} + \mathbf{A}^{\mathrm{T}}\mathbf{L}\mathbf{A} & -\mathbf{A}^{\mathrm{T}}\mathbf{L} \\ -\mathbf{L}\mathbf{A} & \mathbf{L} \end{pmatrix} \begin{pmatrix} \mathbf{x} \\ \mathbf{y} \end{pmatrix} = -\frac{1}{2}\mathbf{z}^{\mathrm{T}}\mathbf{R}\mathbf{z}$$

$$= \mathbf{R}$$

$$cov[\mathbf{z}] = \mathbf{R}^{-1} = \begin{pmatrix} \mathbf{\Lambda}^{-1} & \mathbf{\Lambda}^{-1}\mathbf{A}^{\mathrm{T}} \\ \mathbf{A}\mathbf{\Lambda}^{-1} & \mathbf{L}^{-1} + \mathbf{A}\mathbf{\Lambda}^{-1}\mathbf{A}^{\mathrm{T}} \end{pmatrix}$$

 $\ln p(\mathbf{z})$ の線形の項 The linear terms in $\ln p(\mathbf{z})$

$$\mathbf{x}^{\mathrm{T}} \mathbf{\Lambda} \boldsymbol{\mu} - \mathbf{x}^{\mathrm{T}} \mathbf{A}^{\mathrm{T}} \mathbf{L} \mathbf{b} + \mathbf{y}^{\mathrm{T}} \mathbf{L} \mathbf{b} = \begin{pmatrix} \mathbf{x} \\ \mathbf{y} \end{pmatrix}^{\mathrm{T}} \frac{\left(\mathbf{\Lambda} \boldsymbol{\mu} - \mathbf{A}^{\mathrm{T}} \mathbf{L} \mathbf{b} \right)}{\mathbf{L} \mathbf{b}}$$
 $\mathbb{E}[\mathbf{z}] = \mathbf{R}^{-1} \begin{pmatrix} \mathbf{\Lambda} \boldsymbol{\mu} - \mathbf{A}^{\mathrm{T}} \mathbf{L} \mathbf{b} \\ \mathbf{L} \mathbf{b} \end{pmatrix}$
 $= \begin{pmatrix} \mathbf{\Lambda}^{-1} & \mathbf{\Lambda}^{-1} \mathbf{A}^{\mathrm{T}} \\ \mathbf{A} \mathbf{\Lambda}^{-1} & \mathbf{L}^{-1} + \mathbf{A} \mathbf{\Lambda}^{-1} \mathbf{A}^{\mathrm{T}} \end{pmatrix} \begin{pmatrix} \mathbf{\Lambda} \boldsymbol{\mu} - \mathbf{A}^{\mathrm{T}} \mathbf{L} \mathbf{b} \\ \mathbf{L} \mathbf{b} \end{pmatrix}$
 $= \begin{pmatrix} \boldsymbol{\mu} \\ \mathbf{A} \boldsymbol{\mu} + \mathbf{b} \end{pmatrix}$

$$\mathbb{E}[\mathbf{z}] = \begin{pmatrix} \boldsymbol{\mu} \\ \mathbf{A}\boldsymbol{\mu} + \mathbf{b} \end{pmatrix}$$

$$\operatorname{cov}[\mathbf{z}] = \begin{pmatrix} \boldsymbol{\Lambda}^{-1} & \boldsymbol{\Lambda}^{-1}\mathbf{A}^{\mathrm{T}} \\ \mathbf{A}\boldsymbol{\Lambda}^{-1} & \mathbf{L}^{-1} + \mathbf{A}\boldsymbol{\Lambda}^{-1}\mathbf{A}^{\mathrm{T}} \end{pmatrix}$$

$$\mathbf{E}[\mathbf{y}] = \mathbf{A}\boldsymbol{\mu} + \mathbf{b}$$

$$\operatorname{cov}[\mathbf{y}] = \mathbf{L}^{-1} + \mathbf{A}\boldsymbol{\Lambda}^{-1}\mathbf{A}^{\mathrm{T}}$$

$$\mathbb{E}[\mathbf{x}_{a}] = \boldsymbol{\mu}_{a}$$

$$\operatorname{cov}[\mathbf{x}_{a}] = \boldsymbol{\Sigma}_{aa}$$

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Lambda}^{-1})$$

$$p(\mathbf{y}|\mathbf{x}) = \mathcal{N}(\mathbf{y}|\mathbf{A}\mathbf{x} + \mathbf{b}, \mathbf{L}^{-1})$$

$$\boxed{p(\mathbf{y}) = \mathcal{N}(\mathbf{y}|\mathbf{A}\boldsymbol{\mu} + \mathbf{b}, \mathbf{L}^{-1} + \mathbf{A}\boldsymbol{\Lambda}^{-1}\mathbf{A}^{\mathrm{T}})}$$

$$\mathbb{E}[\mathbf{z}] = \begin{pmatrix} \boldsymbol{\mu} \\ \mathbf{A}\boldsymbol{\mu} + \mathbf{b} \end{pmatrix}$$

$$\operatorname{cov}[\mathbf{z}]^{-1} = \begin{pmatrix} \boldsymbol{\Lambda} + \mathbf{A}^{\mathrm{T}}\mathbf{L}\mathbf{A} & -\mathbf{A}^{\mathrm{T}}\mathbf{L} \\ -\mathbf{L}\mathbf{A} & \mathbf{L} \end{pmatrix} \qquad \mathbb{E}[\mathbf{x}|\mathbf{y}] = (\boldsymbol{\Lambda} + \mathbf{A}^{\mathrm{T}}\mathbf{L}\mathbf{A})^{-1} \left\{ \mathbf{A}^{\mathrm{T}}\mathbf{L}(\mathbf{y} - \mathbf{b}) + \boldsymbol{\Lambda}\boldsymbol{\mu} \right\}$$

$$\operatorname{cov}[\mathbf{x}|\mathbf{y}] = (\boldsymbol{\Lambda} + \mathbf{A}^{\mathrm{T}}\mathbf{L}\mathbf{A})^{-1}$$

$$\operatorname{cov}[\mathbf{x}|\mathbf{y}] = (\boldsymbol{\Lambda} + \mathbf{A}^{\mathrm{T}}\mathbf{L}\mathbf{A})^{-1}$$

$$\mathbf{\mu}_{a|b} = \boldsymbol{\Sigma}_{a|b} \left\{ \boldsymbol{\Lambda}_{aa}\boldsymbol{\mu}_{a} - \boldsymbol{\Lambda}_{ab}(\mathbf{x}_{b} - \boldsymbol{\mu}_{b}) \right\}$$

$$\boldsymbol{\Sigma}_{a|b} = \boldsymbol{\Lambda}_{aa}^{-1}$$

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Lambda}^{-1})$$
$$p(\mathbf{y}|\mathbf{x}) = \mathcal{N}(\mathbf{y}|\mathbf{A}\mathbf{x} + \mathbf{b}, \mathbf{L}^{-1})$$

条件付き分布 Conditional distribution

$$p(\mathbf{x}|\mathbf{y}) = \mathcal{N}(\mathbf{x}|\mathbf{\Sigma}\{\mathbf{A}^{\mathrm{T}}\mathbf{L}(\mathbf{y}-\mathbf{b})+\mathbf{\Lambda}\boldsymbol{\mu}\}, \mathbf{\Sigma}) \mid \mathbf{\Sigma} = (\mathbf{\Lambda}+\mathbf{A}^{\mathrm{T}}\mathbf{L}\mathbf{A})^{-1}$$

多変量ガウス分布 The multivariate Gaussian 最尤推定 Maximum likelihood

データ集合 A data set: $\mathbf{X} = (\mathbf{x}_1, \dots, \mathbf{x}_N)^{\mathrm{T}}$

対数尤度関数 Log likelihood function

$$\ln p(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = -\frac{ND}{2} \ln(2\pi) - \frac{N}{2} \ln|\boldsymbol{\Sigma}| - \frac{1}{2} \sum_{n=1}^{N} (\mathbf{x}_n - \boldsymbol{\mu})^{\mathrm{T}} \boldsymbol{\Sigma}^{-1} (\mathbf{x}_n - \boldsymbol{\mu})$$

十分統計量 Sufficient statistics

$$\sum_{n=1}^{N} \mathbf{x}_n \qquad \qquad \sum_{n=1}^{N} \mathbf{x}_n \mathbf{x}_n^{\mathrm{T}}$$

多変量ガウス分布 The multivariate Gaussian 最尤推定 Maximum likelihood

対数尤度の最大化 Maximization of the log likelihood function

$$\frac{\partial}{\partial \boldsymbol{\mu}} \ln p(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^{N} \boldsymbol{\Sigma}^{-1} (\mathbf{x}_n - \boldsymbol{\mu}) = 0$$

$$\mu_{\mathrm{ML}} = \frac{1}{N} \sum_{n=1}^{N} \mathbf{x}_{n}.$$

同様にして Similarly

$$oldsymbol{\Sigma}_{ ext{ML}} = rac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - oldsymbol{\mu}_{ ext{ML}}) (\mathbf{x}_n - oldsymbol{\mu}_{ ext{ML}})^{ ext{T}}.$$

多変量ガウス分布 The multivariate Gaussian 最尤推定 Maximum likelihood

真の分布の下での最尤推定解の期待値

Expectations of the maximum likelihood solutions under the true distribution

$$\mathbb{E}[oldsymbol{\mu}_{ ext{ML}}] = oldsymbol{\mu} \ \mathbb{E}[oldsymbol{\Sigma}_{ ext{ML}}] = rac{N-1}{N}oldsymbol{\Sigma}$$

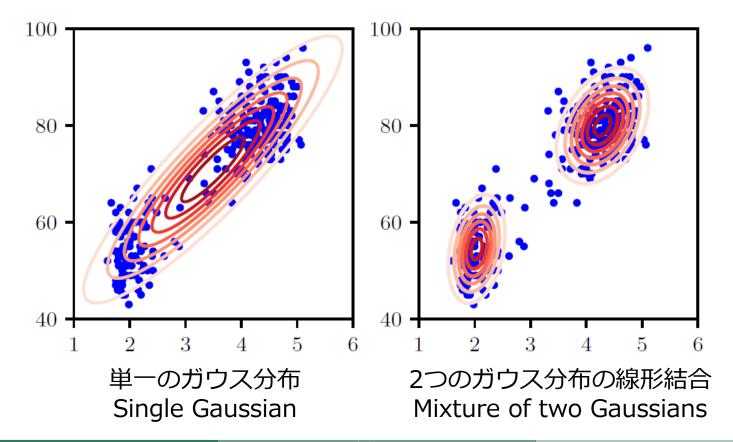
不偏推定量 Unbiased estimator

$$\widetilde{\Sigma} = rac{1}{N-1} \sum_{n=1}^{N} (\mathbf{x}_n - \boldsymbol{\mu}_{\mathrm{ML}}) (\mathbf{x}_n - \boldsymbol{\mu}_{\mathrm{ML}})^{\mathrm{T}}.$$

$$\mathbb{E}[\widetilde{\Sigma}] = \Sigma$$

多変量ガウス分布 The multivariate Gaussian 逐次推定 Sequential estimation

多変量ガウス分布 The multivariate Gaussian 混合ガウス分布 Mixture of Gaussians

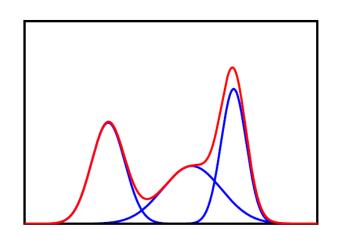


混合ガウス分布 Mixture of Gaussians

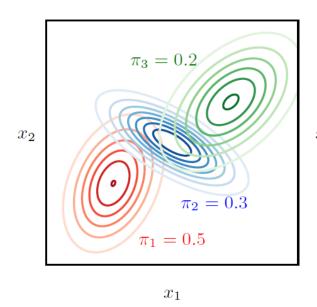
$$p(\mathbf{x}) = \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$
 混合要素 Mixture component 混合係数 Mixing coefficient

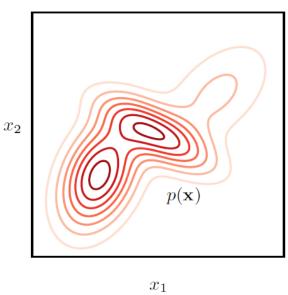
$$\forall k : \pi_k \geqslant 0 \qquad \sum_{k=1}^K \pi_k = 1$$

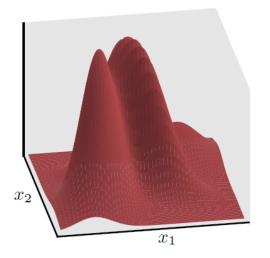
$$\Rightarrow 0 \leqslant \pi_k \leqslant 1$$



$$p(\mathbf{x}) = \sum_{k=1}^{K} p(k)p(\mathbf{x}|k)$$







混合ガウス分布 Mixture of Gaussians

$$p(\mathbf{x}) = \sum_{k=1}^{K} p(k)p(\mathbf{x}|k) = \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

負担率 Responsibility:
$$\gamma_k(\mathbf{x}) \equiv p(k|\mathbf{x})$$

$$= \frac{p(k)p(\mathbf{x}|k)}{\sum_l p(l)p(\mathbf{x}|l)}$$

$$= \frac{\pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_l \pi_l \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_l, \boldsymbol{\Sigma}_l)}$$

パラメータ Parameters

$$\boldsymbol{\pi} \equiv \{\pi_1, \dots, \pi_K\} \quad \boldsymbol{\mu} \equiv \{\boldsymbol{\mu}_1, \dots, \boldsymbol{\mu}_K\} \quad \boldsymbol{\Sigma} \equiv \{\boldsymbol{\Sigma}_1, \dots \boldsymbol{\Sigma}_K\}$$

対数尤度関数 Log likelihood function

$$\ln p(\mathbf{X}|\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^{N} \ln \left\{ \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\}$$

閉形式の解析解は得られない. No closed-form analytical solution.

指数型分布族 The exponential family

指数分布族は次式で定義される分布の集合である.

The exponential family of distributions is defined to be the set of distributions of the form

$$p(\mathbf{x}|\boldsymbol{\eta}) = h(\mathbf{x})g(\boldsymbol{\eta}) \exp\left\{\boldsymbol{\eta}^{\mathrm{T}}\mathbf{u}(\mathbf{x})\right\}$$

ただし where

$$g(\boldsymbol{\eta}) \int h(\mathbf{x}) \exp\left\{\boldsymbol{\eta}^{\mathrm{T}} \mathbf{u}(\mathbf{x})\right\} d\mathbf{x} = 1$$

 η :自然パラメータ Natural parameter

$$p(\mathbf{x}|\boldsymbol{\eta}) = h(\mathbf{x})g(\boldsymbol{\eta}) \exp\left\{\boldsymbol{\eta}^{\mathrm{T}}\mathbf{u}(\mathbf{x})\right\}$$

ベルヌーイ分布 Bernoulli distribution

$$p(x|\mu) = \operatorname{Bern}(x|\mu) = \mu^x (1-\mu)^{1-x}$$

$$= \exp\left\{x \ln \mu + (1-x) \ln(1-\mu)\right\}$$

$$= (1-\mu) \exp\left\{\ln\left(\frac{\mu}{1-\mu}\right)x\right\}$$
Logistic sigmoid function
$$\mu = \sigma(\eta) = \frac{1}{1+\exp(-\eta)}.$$

$$p(x|\eta) = \sigma(-\eta) \exp(\eta x)$$

$$u(x) = x$$

$$h(x) = 1$$

$$g(\eta) = 1-\sigma(\eta) = \sigma(-\eta).$$

ロジスティックシグモイド関数

$$\mu = \sigma(\eta) = \frac{1}{1 + \exp(-\eta)}.$$

$$p(\mathbf{x}|\boldsymbol{\eta}) = h(\mathbf{x})g(\boldsymbol{\eta}) \exp\left\{\boldsymbol{\eta}^{\mathrm{T}}\mathbf{u}(\mathbf{x})\right\}$$

多項分布 Multinomial distribution

$$p(\mathbf{x}|\boldsymbol{\mu}) = \prod_{k=1}^{M} \mu_k^{x_k} = \exp\left\{\sum_{k=1}^{M} x_k \ln \mu_k\right\} \qquad \mathbf{x} = (x_1, \dots, x_M)^{\mathrm{T}}$$
$$p(\mathbf{x}|\boldsymbol{\eta}) = \exp(\boldsymbol{\eta}^{\mathrm{T}} \mathbf{x}) \qquad \boldsymbol{\eta} = (\eta_1, \dots, \eta_M)^{\mathrm{T}}$$

$$\eta_k = \ln \mu_k$$
 $\mathbf{u}(\mathbf{x}) = \mathbf{x}$
 $h(\mathbf{x}) = 1$
 $g(\boldsymbol{\eta}) = 1.$

制約 Constraint

$$\sum_{k=1}^{M} \mu_k = 1.$$

 μ_M を消去 Eliminate μ_M

$$\exp\left\{\sum_{k=1}^{M} x_{k} \ln \mu_{k}\right\} \qquad \mu_{M} = 1 - \sum_{k=1}^{M-1} \mu_{k}
= \exp\left\{\sum_{k=1}^{M-1} x_{k} \ln \mu_{k} + \left(1 - \sum_{k=1}^{M-1} x_{k}\right) \ln \left(1 - \sum_{k=1}^{M-1} \mu_{k}\right)\right\}
= \exp\left\{\sum_{k=1}^{M-1} x_{k} \ln \left(\frac{\mu_{k}}{1 - \sum_{j=1}^{M-1} \mu_{j}}\right) + \ln \left(1 - \sum_{k=1}^{M-1} \mu_{k}\right)\right\}
\longrightarrow \eta_{k}$$

$$\mu_{k} = \frac{\exp(\eta_{k})}{1 + \sum_{j=1}^{M-1} \exp(\eta_{j})}.$$

多項分布 Multinomial distribution

$$p(\mathbf{x}|\boldsymbol{\eta}) = \left(1 + \sum_{k=1}^{M-1} \exp(\eta_k)\right)^{-1} \exp(\boldsymbol{\eta}^{\mathrm{T}}\mathbf{x})$$

$$\boldsymbol{\eta} = (\eta_1, \dots, \eta_{M-1}, 0)^{\mathrm{T}}$$

$$\mathbf{u}(\mathbf{x}) = \mathbf{x}$$

$$h(\mathbf{x}) = 1$$

$$g(\boldsymbol{\eta}) = \left(1 + \sum_{k=1}^{M-1} \exp(\eta_k)\right)^{-1}.$$

ガウス分布 Gaussian distribution

$$p(x|\mu, \sigma^2) = \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left\{-\frac{1}{2\sigma^2}(x-\mu)^2\right\}$$

$$= \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left\{-\frac{1}{2\sigma^2}x^2 + \frac{\mu}{\sigma^2}x - \frac{1}{2\sigma^2}\mu^2\right\}$$

$$p(\mathbf{x}|\boldsymbol{\eta}) = h(\mathbf{x})g(\boldsymbol{\eta}) \exp\left\{\boldsymbol{\eta}^{\mathrm{T}}\mathbf{u}(\mathbf{x})\right\}$$

$$\boldsymbol{\eta} = \begin{pmatrix} \mu/\sigma^2 \\ -1/2\sigma^2 \end{pmatrix} \qquad h(\mathbf{x}) = (2\pi)^{-1/2}$$

$$\mathbf{u}(x) = \begin{pmatrix} x \\ x^2 \end{pmatrix} \qquad g(\boldsymbol{\eta}) = (-2\eta_2)^{1/2} \exp\left(\frac{\eta_1^2}{4\eta_2}\right).$$

制限のある指数型分布族 A restricted set of the exponential family

$$p(\mathbf{x}|\boldsymbol{\lambda}_k, s) = \frac{1}{s} h\left(\frac{1}{s}\mathbf{x}\right) g(\boldsymbol{\lambda}_k) \exp\left\{\frac{1}{s}\boldsymbol{\lambda}_k^{\mathrm{T}}\mathbf{x}\right\}$$

s>0:スケールパラメータ Scale parameter

指数型分布族 The exponential family 十分推定量 Sufficient statistics

$$-\nabla \ln g(oldsymbol{\eta}) = \mathbb{E}[\mathbf{u}(\mathbf{x})]$$

指数型分布族 The exponential family 十分推定量 Sufficient statistics

データ集合 A data set:
$$\mathbf{X} = \{\mathbf{x}_1, \dots, \mathbf{x}_N\}$$

尤度関数 Likelihood function

$$p(\mathbf{X}|\boldsymbol{\eta}) = \left(\prod_{n=1}^{N} h(\mathbf{x}_n)\right) g(\boldsymbol{\eta})^N \exp\left\{\boldsymbol{\eta}^T \sum_{n=1}^{N} \mathbf{u}(\mathbf{x}_n)\right\}.$$

最尤推定量が満たすべき条件 Condition to be satisfied by the maximum likelihood estimator

$$-\nabla \ln g(\boldsymbol{\eta}_{\mathrm{ML}}) = rac{1}{N} \sum_{n=1}^{N} \mathbf{u}(\mathbf{x}_{n})$$
 十分統計量 Sufficient statistics



提出課題I

Assignment I

提出期限: **10月15日(火曜日) 23:59:00** [日本標準時]

Submission deadline: October 15 (Tuesday) 23:59:00 [Japan Standard Time]

提出課題は「一般」チャネルの「ファイル」にアップロードされます. 同チャネルに出現する通知のリンク先から解答を送信(提出)してください. Assignments will be uploaded to "File" in the "General" channel. Send (submit) your answers via the link that will appear in the same channel.

- 全6回の課題への解答をもとに成績評価を行います.
 Your grades will be based on your answers to all six assignments.
- 解答送信後の解答再送信はできません。
 Once submitted, answers cannot be resubmitted.
- 提出期限は一切延長しません.
 The submission deadline will never be extended.

霍率分布