Bayesian Linear Regression

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Bayesian Linear Regression

- In this class, which is mostly based on chapter 4 of [McElreath, 2020], we are going to revisit the linear regression model from a Bayesian point of view.
- The idea is the same: to model the relationship of a numerical dependent variable y with n independent variables x₁, x₂,...,x_n from a dataset d.
- The response vaiable **y** is again modeled with a Gaussian distribution: $y_i \sim N(\mu_i, \sigma^2)$.
- We also mantain the assumption that each attribute has a linear relationship to the mean of the outcome.

$$\mu_i = \beta_0 + \beta_1 x_i + \dots \beta_n x_n$$

- However, we are not going to use least squares or maximum likelihood to obtain point estimates of the parameters.
- Instead, we are going to estimate the joint posterior distribution of all the parameters of the model:

$$f(\theta|\mathbf{d}) = f(\beta_0, \beta_1, \dots, \beta_n, \sigma|\mathbf{d})$$

Bayesian Linear Models

- The Bayesian linear regresion is more flexible than least squares as it allows incorporating prior information.
- It also allows to interpret the uncertainty of the model in a clearer way.
- Notice that the parameters of the model are $\beta_0, \beta_1, \dots, \beta_b$ and σ but not μ_i .
- ullet This is because μ_i it is determined deterministically from the linear model's coefficients.
- In order to build our posterior we need to define a likelihood function:

$$f(\mathbf{d}|\beta_0,\beta_1,\cdots,\beta_n,\sigma)=\prod_{i=1}^m f(\mathbf{d}_i|\beta_0,\beta_1,\cdots,\beta_n,\sigma)$$

- Where d_i corresponds to each data point in the dataset containing values for y and x₁,...,x_n (IID assumption).
- The likelihood of each point is modeled with a Gaussian distribution:

$$f(d_i|\beta_0,\beta_1,\cdots,\beta_n,\sigma)=N(\mu_i,\sigma^2)$$

Bayesian Linear Models

Now we need a joint prior density:

$$f(\theta) = f(\beta_0, \beta_1, \dots, \beta_n, \sigma)$$

And the posterior gets specified as follows:

$$f(\theta|d) = \frac{\prod_{i=1}^{m} f(d_i|\beta_0, \beta_1, \cdots, \beta_n, \sigma) * f(\beta_0, \beta_1, \dots, \beta_n, \sigma)}{f(d)}$$

The evidence is expressed by a multiple integral:

$$f(d) = \int \int \cdots \int \prod_{i=1}^{m} f(d_i|\beta_0, \beta_1, \cdots, \beta_n, \sigma) * f(\beta_0, \beta_1, \dots, \beta_n, \sigma) d\beta_0 d\beta_1 \cdots d\beta_n d\sigma$$

 In most cases, the priors are specified independently for each parameter, which is equivalent to assuming:

$$f(\beta_0, \beta_1, \cdots, \beta_b, \sigma) = f(\beta_0) * f(\beta_1) * \cdots * f(\beta_n) * f(\sigma).$$



A model of height revisited

- To understand this more concretely, we will rebuild the linear model relating the height and weight of the !Kung San people using a Bayesian approach.
- We will refer to each person's height and weight as y_i and x_i respectively.
- Our probabilistic model specifying all components of a Bayesian model is defined as follows:

$$\begin{array}{ll} y_i \sim \textit{N}(\mu_i, \sigma) & \text{[likelihood]} \\ \mu_i = \beta_0 + \beta_1 \textit{x}_i & \text{[linear model]} \\ \beta_0 \sim \textit{N}(150, 50) & [\beta_0 \text{ prior]} \\ \beta_1 \sim \textit{N}(0, 1) & [\beta_1 \text{ prior]} \\ \sigma \sim \text{Uniform}(0, 50) & [\sigma \text{ prior]} \end{array}$$

- Parameters β_0 and β_1 are the intercept and the slope of our linear model.
- The parameter σ is the standard deviation of all the heights.
- ullet Note that we are setting the same σ for all observations, which is equivalent to the Homoscedasticity property of the standard linear regression.

A model of height revisited

- Our priors were set independently for each parameter which implies that the joint prior density $f(\beta_0, \beta_1, \sigma)$ can be expressed as $f(\beta_0) * f(\beta_1) * f(\sigma)$.
- It should be kept in mind that the choice of priors is subjective and should be evaluated accordingly.
- Let's try to justify our choice a bit:
 - **1** The Gaussian prior for β_0 (intercept), centered on 150cm with a standard variation of 50, covers a huge range of plausible mean heights for human populations while giving very little chance for negative heights.
 - 2 The Gaussian prior for β_1 (slope), centered on 0 with a standard variation of 1, acts as a **regularizer** to prevent the model from **overfitting** the data by assigning extreme values to β_1 .¹
 - \odot The uniform prior for the standard deviation σ between 0 and 50 prohibits obtaining negative standard deviations. The upper bound (50 cm) would imply that 95% of individual heights lie within 100cm of the average height. That's a very large range.

¹Regularization and overfitting will be discussed later in the course.

Fitting the Model

- Now we need to fit the model to the data to build the posterior distribution.
- Grid approximation is not a valid option, as setting up a grid for 3 parameters would be too computationally expensive.
- We will use Laplace approximation instead.
- In this approach we obtain the MAP estimates for each parameter using a hill-climbing optimization method.
- Then we fit a multivariate Gaussian distribution centered on these values.
- This distribution is the multidimensional extension to the standard Gaussian.

 The multivariate Gaussian distribution in d-dimensions defined by the following density function (PDF):

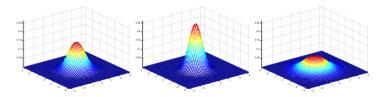
$$f_X = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp\left(-\frac{1}{2} (\vec{x} - \vec{\mu})^T \Sigma^{-1} (\vec{x} - \mu)\right)$$

- This density function allows working with a *d*-dimensional vector of random variables \vec{X} .
- The first parameter of this distributions is a mean vector $\vec{\mu} \in \mathcal{R}^d$ with the mean value of each dimension.
- The second parameter is a covariance matrix $\Sigma \in R^{d \times d}$,
- This matrix contains the variance of each variable in the diagonal and the covariance of variables X_i and X_i in the other cells Σ_{i,i}:

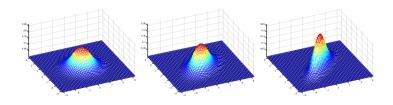
$$Cov(X) = \Sigma$$

- The matrix Σ is symmetric and positive semi-definite.
- The multivariate Gaussian $N(\vec{\mu}, \Sigma)$ is a very convenient distribution for modeling multidimensional random variables

 Here are some examples taken from [Ng, 2008] of what the density of a multivariate Gaussian distribution looks like:



- The left-most figure shows a Gaussian with mean zero (that is, the 2x1 zero-vector) and covariance matrix $\Sigma = I$ (the 2 × 2 identity matrix).
- A Gaussian with zero mean and identity covariance is also called the standard normal distribution.
- The middle figure shows the density of a Gaussian with zero mean and $\Sigma=0.6\emph{l}.$
- The rightmost figure shows one with , $\Sigma = 2I$.
- We see that as Σ becomes larger, the Gaussian becomes more "spread-out", and as it becomes smaller, the distribution becomes more "compressed".

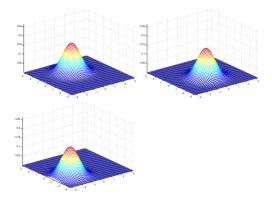


 The figures above show Gaussians with mean 0, and with covariance matrices respectively

$$\Sigma = \left[\begin{array}{cc} 1 & 0 \\ 0 & 1 \end{array} \right]; \ \ \Sigma = \left[\begin{array}{cc} 1 & 0.5 \\ 0.5 & 1 \end{array} \right]; \ \ \Sigma = \left[\begin{array}{cc} 1 & 0.8 \\ 0.8 & 1 \end{array} \right].$$

 The leftmost figure shows the familiar standard normal distribution, and we see that as we increase the off-diagonal entry in Σ, the density becomes more "compressed" towards the 45· line (given by x₁ = x₂).

• As our last set of examples, fixing $\Sigma = I$, by varying $\vec{\mu}$ we can also move the mean of the density around.



• The figures above were generated using $\Sigma = I$, and respectively

$$\mu = \left[\begin{array}{c} 1 \\ 0 \end{array} \right]; \;\; \mu = \left[\begin{array}{c} \text{-0.5} \\ 0 \end{array} \right]; \;\; \mu = \left[\begin{array}{c} \text{-1} \\ \text{-1.5} \end{array} \right].$$

Laplace approximation

- In Laplace approximation we assume that the joint posterior follows a multivariate Gaussian distribution $f(\theta_1, \dots, \theta_n) = N(\vec{\mu}, \Sigma)$.
- This approximation is convenient for unimodal and roughly symmetric posterior distributions [Gelman et al., 2013].
- The Bayesian asymptotic theory says that if the dataset is large enough, a
 posterior distribution can be approximated by a Gaussian [Gelman et al., 2013].
- The values of $\vec{\mu}$ are obtained from the posterior mode of each parameters (MAP):

$$\vec{\mu} = \vec{\theta}_{MAP}$$

These is done using numeric optimization techniques.

Laplace approximation

• The values of Σ are obtained from the curvature near these values, which are obtained from the second derivatives of the posterior:

$$\Sigma = [I(\theta_{MAP})]^{-1}$$

where

$$I(\theta) = -\frac{d^2}{d\theta^2} \log f(\theta|d)$$

- Notice that both $\vec{\mu}$ and Σ can be calculated from the unnormalized posterior $f(d|\theta)*f(\theta)$
- We don't need to calculate the evidence f(d) to perform Laplace approximation
 [?].

Fitting the Model

- Laplace approximation is implemented in the quap function from the rethinking package.
- The model for height defined above can be implemented as follows:

```
library(rethinking)
data(Howell1)
d <- Howell1
d2 <- d[ d$age >= 18 , ]

b.reg1 <- quap(
    alist(
      height ~ dnorm( b0 + b1*weight, sigma ),
      b0 ~ dnorm( 150 , 50 ) ,
      b1 ~ dnorm( 0 , 1) ,
      sigma ~ dunif( 0 , 50 )
) , data=d2 )</pre>
```

Fitting the Model

We can summarize this posterior with the command precis:

- These numbers provide Gaussian approximations for each parameter's marginal posterior distribution.
- This is the distribution of any parameter θ_i regardless of the values of the others.
- In mathematical terms this is obtained, after integrating (averaging) over $\theta_j \ \forall j \neq i$:

$$f(\theta_i|d) = \int \cdots \int f(\theta_1, \ldots, \theta_n|d)d\theta_1, \ldots, d\theta_{i-1}, d\theta_{i+1}, \ldots, d\theta_n$$

- We can compare the mean and standard deviation of β_0 and β_1 with what we obtained using least squares in previous lecture.
- The values are almost the same!

Conclusions

Blabla

References I



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