

Bayesian Notebook

Outline

Part 1

1. Probability notation: Marginal probability, joint probability, conditional probability, Bayes rule
2. Maximum likelihood, Bayesian inference
3. Normalization trick
4. Quiz and programming break(Naive bayes)

Part 2

5. Graphical model and bayesian network
6. Basic structure of graphical model
7. Programming break (Linear regression)

Part 3

8. Bayesian inference problem
9. Variable elimination
10. Latent variable
11. EM algorithm
12. Variational inference
13. Sampling method

Probability notation

Random variable(RV)

Random variable is a variable that can take different values in random space. I may denote RV in a capital letter, for example, X is random variable and lower capital x is the constant values. Note that I may use X and x interchangeably but you can understand it in the context, anyway I try to write as correct as possible.

Marginal probability

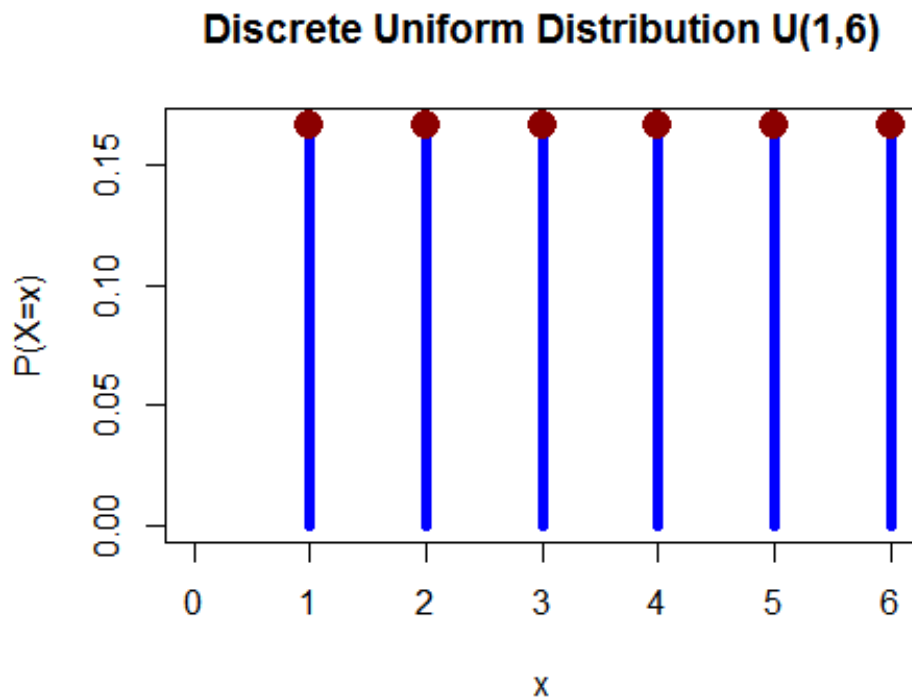
$p(X)$ denotes a marginal probability distribution of X .

Marginal probability is the probability of an event irrespective of the outcome of another variable.

Discrete case (Roll dice) $p(X) \sim \mathcal{U}(1, 6)$

You can call $p(X)$ as probability mass function (pmf).

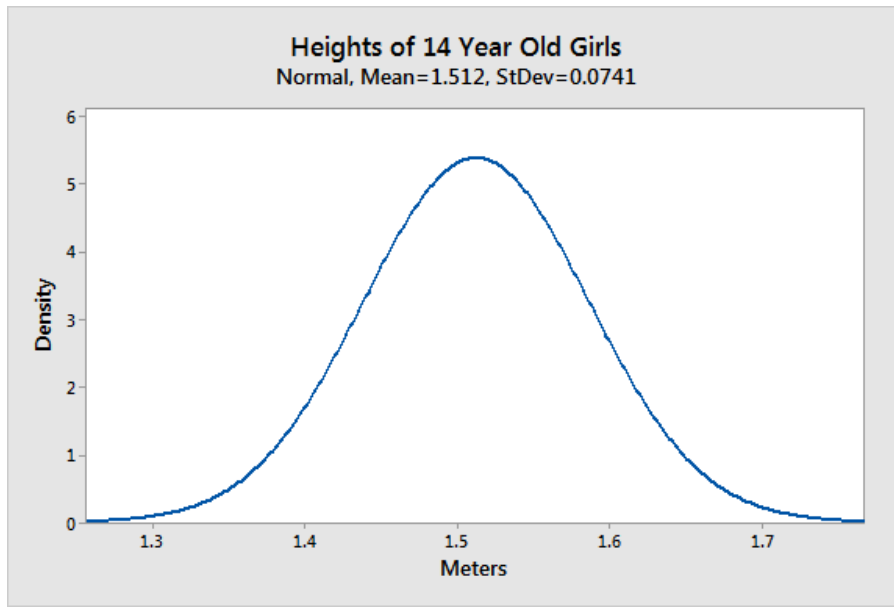
(The actual notation of pmf is $p(X = x)$)



Continuous case $p(X) \sim \mathcal{N}(\mu, \sigma)$

You can call $p(X)$ as probability density function (pdf).

(The actual notation of pdf is $p(X = x)$)



Joint probability

$p(X, Y)$ denoted a joint probability of X and Y

Independent For example, flip the coin two times, the outcome of second coin not depend on first coin.

x denote for outcome of flipping the first coin

y denote for outcome of flipping the second coin

$$p(X, Y) = p(X)p(Y)$$

Dependent For example, this is the record of cloudy and rain in 30 day

model	Rain	No rain	Total
Cloudy	10	5	15
Not cloudy	2	13	15
Total	12	18	30

How do you calculate $p(rain)$?

In this case

$$p(rain) = p(rain, cloudy) + p(rain, notcloudy)$$

model	Rain	No rain	Total
Total	12	18	30

$$p(rain) = \sum_{all\ c} p(rain, C)$$

probability of raining $p(rain)$ is $12/30 = 0.4$

probability of cloudy $p(cloudy)$ is $15/30 = 0.5$

probability of raining and cloudy $p(rain, cloudy)$ is $10/30 = 0.333$

$$p(R, C) \neq p(R)p(C)$$

Conditional probability

$p(X | Y)$ denotes a conditional probability distribution of X conditioned on Y

Let look the rain & cloudy table

model	Rain	No rain	Total
Cloudy	10	5	15
Not cloudy	2	13	15
Total	12	18	30

$$p(R | cloudy)$$

model	Rain	No rain	Total
Cloudy	10	5	15
Total	10	5	15

probability of raining in cloudy day $p(rain | cloudy)$ is $10/15 = 0.666$

probability of cloudy $p(cloudy)$ is $15/30 = 0.5$

probability of raining and cloudy $p(rain, cloudy)$ is $10/30 = 0.333$

$$\text{Bayes Theorem} \quad p(R | C)p(C) = p(C | R)p(R) = p(R, C)$$

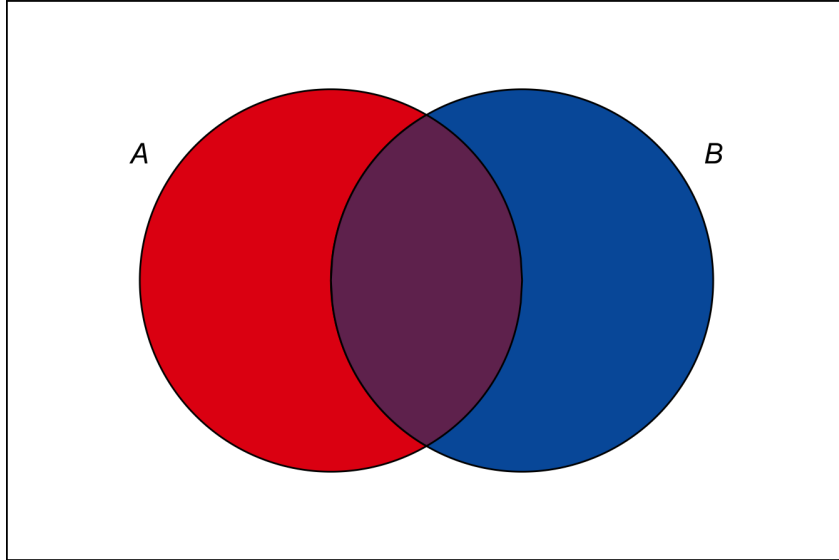
Look at the picture below

Purple color is joint probability $p(A, B)$

Red and Purple color are marginal probability $p(A)$

Blue and Purple color are marginal probability $p(B)$

$$p(A | B) = \frac{p(A, B)}{p(B)}$$

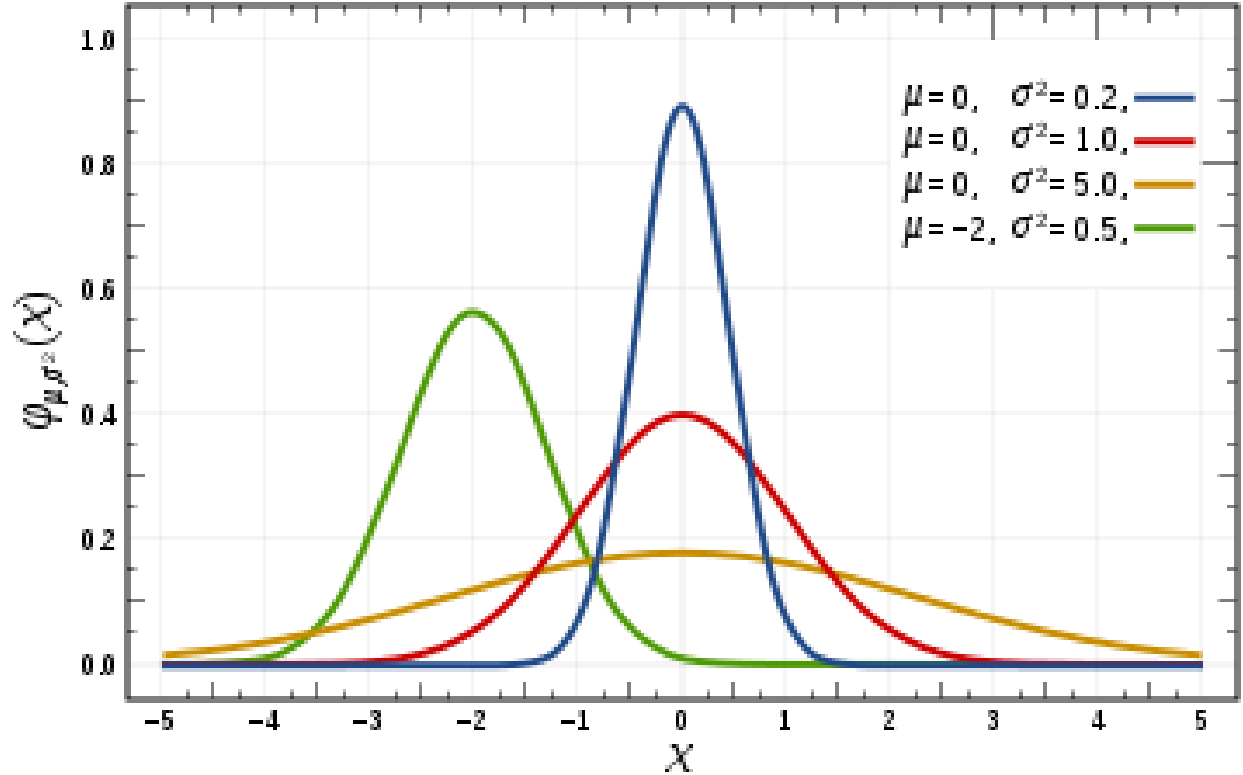


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The Bayes theorem tell how often X happen given that Y happens, written as $p(R | C)$

Maximum likelihood estimation for parameter

For the given probabilistic model $p(X) \sim \mathcal{N}(\mu, \sigma)$ and the given data distribution as below, how do we find the best parameter μ and σ that fit the data.



Because probability of $X = x$ depend on the model parameter μ and σ we can write probability function as $p(X = x; \mu, \sigma)$.

$p(X = x; \mu, \sigma)$ mean the likelihood of observed X for given model parameter μ, σ

or we can use notation $L(\mu, \sigma; X = x)$

And we can get the model parameter μ, σ by find the μ, σ that maximize $L(\mu, \sigma; X = x)$

$$parameter_{ML} = \operatorname{argmax}_{parameter} p(X | parameter)$$

Bayesian Inference

For the maximum likelihood estimation we assume the parameter is not random variable. But in bayesian we assume the model parameter as random variable. $p(X, Y) = p(Y, X)$

$$p(Y | X)p(X) = p(X | Y)p(Y)$$

$$p(Y | X) = \frac{p(X | Y)p(Y)}{p(X)}$$

$$p(parameter | X) = \frac{p(X | parameter)p(parameter)}{p(X)}$$

$$Posterior = \frac{Likelihood \times Prior}{Evident}$$

Because the evident is just constant ($p(X = x)$)

$$Posterior \propto Likelihood \times Prior$$

Prior show the probability before we know the data (distribution over possible parameter values)

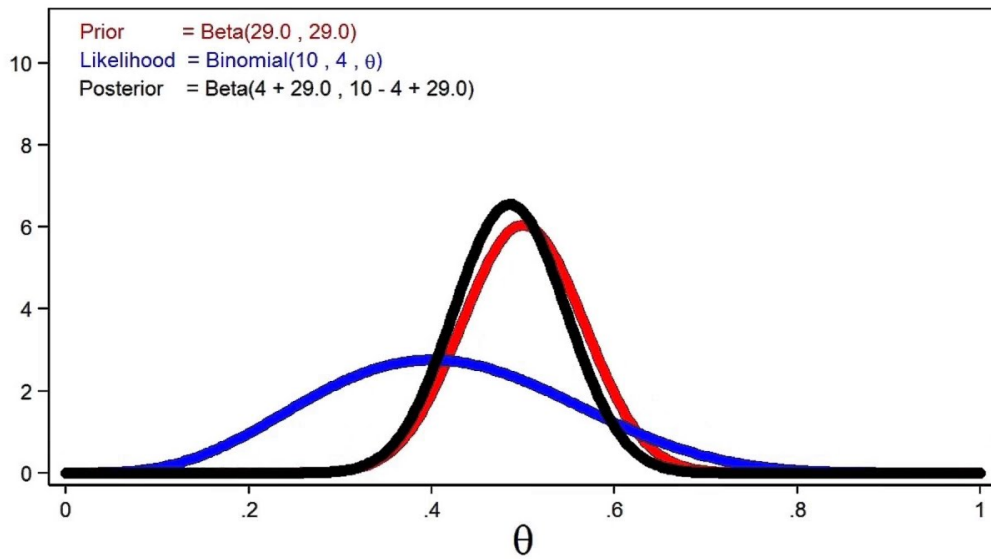
Likelihood give us a best parameter for given data.

Posterior is the distribution that combine the information from prior knowledge and the data.

And we can find the best parameter by find the maximum posterior

$$parameter_{MAP} = \operatorname{argmax}_{parameter} p(parameter | X)$$

Example : flipping coin



Normalization trick

model	Rain	No rain	Total
Cloudy	10	5	15
Not cloudy	2	13	15
Total	12	18	30

model	Rain	No rain	Total
Cloudy	0.333	0.166	0.5
Not cloudy	0.066	0.433	0.5
Total	0.4	0.6	1

We can calculate $p(R \mid \text{cloudy})$ by calculate only $p(R, \text{cloudy})$

model	Rain	No rain	Total
Cloudy	0.333	0.166	0.5

And normalize the table to get $p(R \mid \text{cloudy})$

model	Rain	No rain	Total
Cloudy	0.333/0.5	0.166/0.5	1

Because

$$p(R \mid C = c) = \frac{p(R, C = c)}{p(C = c)}$$

$p(C = c)$ is just constant

$$p(R \mid C = c) \propto p(R, C = c)$$

Quiz true or false ?

1. $\sum_{all\ x} p(X) = 1$
2. $\sum_{all\ x, all\ y} p(X, Y) \neq 1$
3. $\sum_{all\ x, all\ y} p(X \mid Y) = 1$
4. $argmax_{parameter} p(parameter \mid X)$ is only depend on $argmax_{parameter} p(X; parameter)$ sometimes.

Programming break

In this example we have the data of New China Virus and the symptom of patients. The given data will be look like table below.

Patients number	New China Virus	Fever	Vomitting
1	1	1	1
2	0	1	0
3	1	1	1
4	1	1	1
5	0	1	1
6	0	1	0
7	1	0	1
8	0	1	0
9	1	0	1
10	0	1	0

Naive bayes

```
import numpy as np

def generate():

    dataset = []
    for i in range(10):
        patient = np.random.randint(0,2, 3)
        dataset.append(patient)
    dataset = np.array(dataset)
    return dataset

dataset = generate()
```

$$\text{Bayes Theorem} \quad p(X \mid Y)p(Y) = p(X, Y)$$

$$p(virus | fever, vomit) = \frac{p(virus, fever, vomit)}{p(fever, vomit)}$$

Because naive bayes assume independent variable.

$$p(virus | fever, vomit) = \frac{p(fever, vomit | virus)p(virus)}{p(fever)p(vomit)}$$

$$p(virus | fever, vomit) = \frac{p(fever | virus)p(vomit | virus)p(virus)}{p(fever)p(vomit)}$$

```
print("dataset = \n",dataset)
```

```
## dataset =
## [[1 0 0]
##  [1 1 0]
##  [1 1 0]
##  [1 0 0]
##  [1 0 1]
##  [1 1 1]
##  [1 0 0]
##  [1 0 1]
##  [1 0 0]
##  [0 1 1]]
```

```
fever = dataset[:, 1]
virus = dataset[:, 0]
vomit = dataset[:, 2]
```

```
p_vomit = vomit.sum()/len(vomit)
p_virus = virus.sum()/len(virus)
p_fever = fever.sum()/len(fever)
```

```
dataset_virus_positive = dataset[virus == 1]
```

```
print("dataset virus positive = \n" ,dataset_virus_positive)
```

```
## dataset virus positive =
## [[1 0 0]
##  [1 1 0]
##  [1 1 0]
##  [1 0 0]
##  [1 0 1]
##  [1 1 1]
##  [1 0 0]
##  [1 0 1]
##  [1 0 0]]
```

```

vomit_virus_pos = dataset_virus_positive[:,2]
fever_virus_pos = dataset_virus_positive[:,1]

p_vomit_condition_virus = vomit_virus_pos.sum()/len(vomit_virus_pos)
p_fever_condition_virus = fever_virus_pos.sum()/len(fever_virus_pos)

posterior = p_vomit_condition_virus*p_vomit_condition_virus*p_virus/p_vomit/p_fever

print( "posterior = ",posterior)

## posterior = 0.6249999999999999

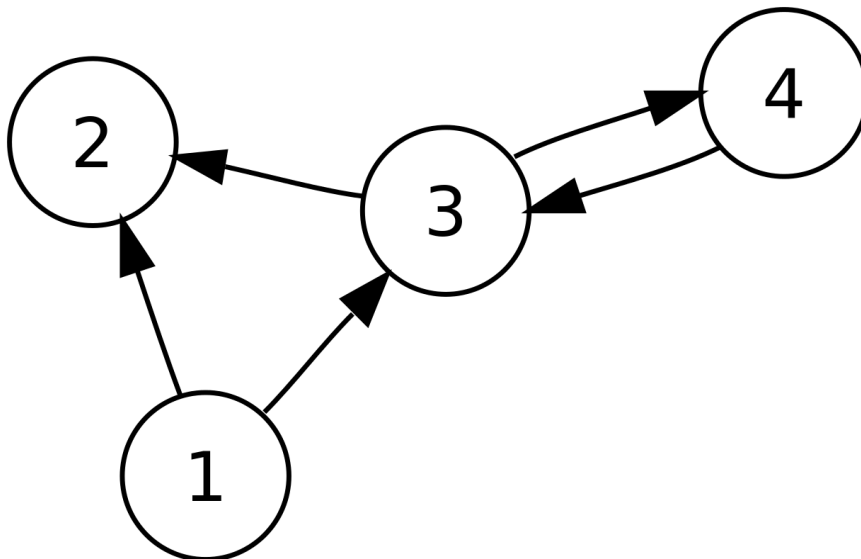
```

Example above you will know that if the patient have fever and come from china the probability that they have virus is the “posterior”

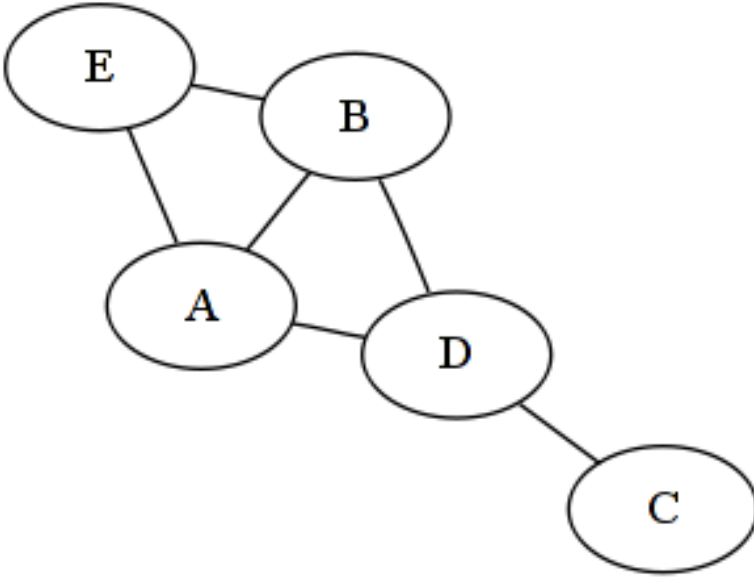
Graphical models

1. Directed graph

- Bayesian network (today)
- Naive Bayes classifier (today)



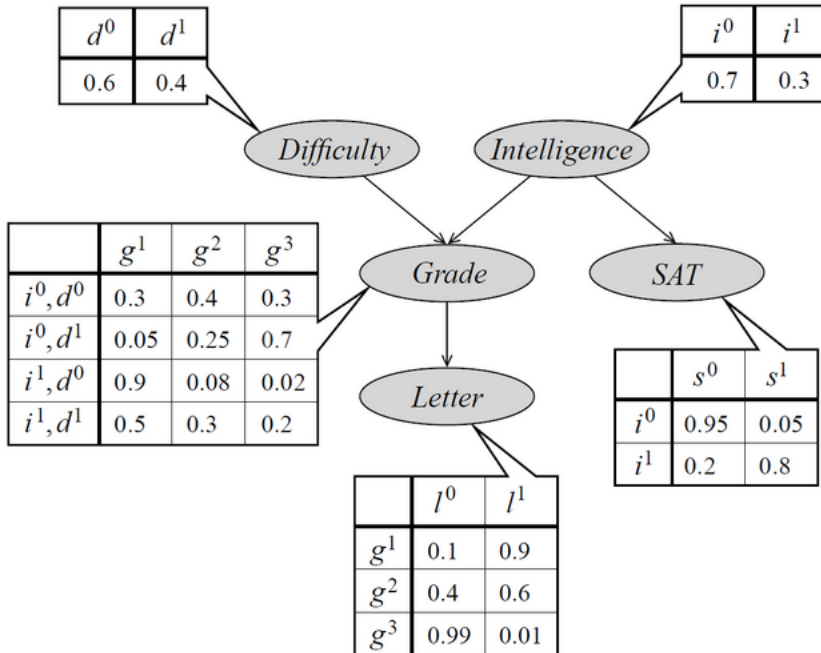
2. Undirected graph



Bayesian network (BN)

BN model the random variable(RV) using directed acyclic graph. Let start by modeling the graph of student network.

We have RV : G(grade), D(school difficulty), I(intelligence), S(SAT score), L(recommend Letter)



Bayesian network encode the joint distribution as conditional distribution.

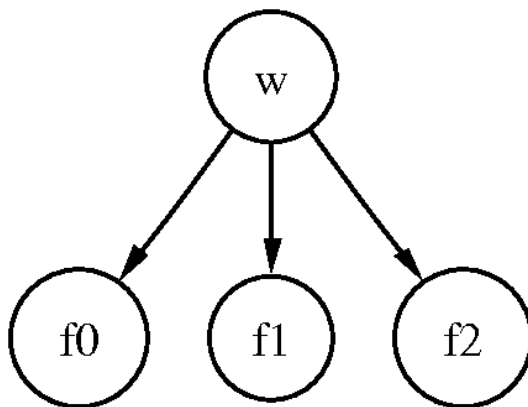
$$p(X_1, X_2, X_3, \dots, X_n) = \prod_{i=1}^n p(X_i \mid \text{parent}(X_i))$$

We can factor the joint probability of student network by looking at the graph.

$$p(G, D, I, S, L) = p(G | D, I)p(D)p(I)p(S | I)p(L | G)$$

Naive bayes (revisit)

Patients number	New China Virus	Fever	Vomitting
1	1	1	1
2	0	1	0
3	1	1	1
4	1	1	1
5	0	1	1
6	0	1	0
7	1	0	1
8	0	1	0
9	1	0	1
10	0	1	0

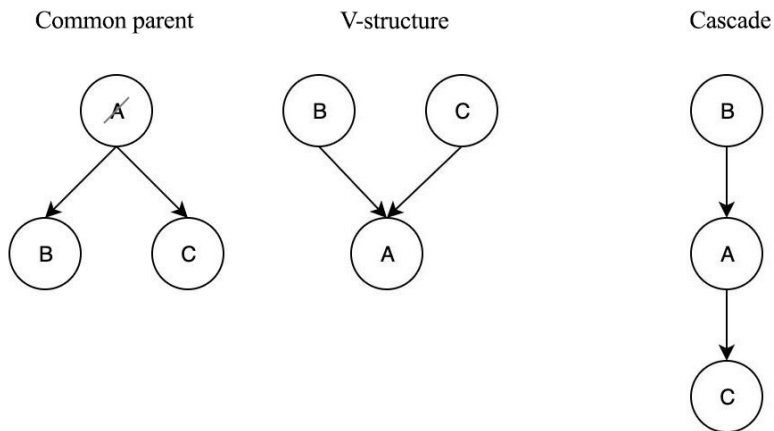


Naive Bayes model:

$$p(W | f_1, f_2, f_3) \propto p(W) \prod_i p(f_i | W)$$

Structure in graphical model

In this picture we want to know if B can influence C or C can influence B by given the basic structure of graphical model



Knowing A, we know B & C,
nothing more is needed

$$B = f_1(A) \\ C = f_2(A)$$

$$B \not\perp C \in I(p)$$

$$B \perp C \mid A \in I(p)$$

Know A, if we know C, we
know B

$$A = f(B, C)$$

$$B \perp C \in I(p)$$

$$B \not\perp C \mid A \in I(p)$$

Knowing A, we know C,
nothing more is needed

$$C = f(A)$$

$$B \not\perp C \in I(p)$$

$$B \perp C \mid A \in I(p)$$

model	B can give information about C not given A	B can give information about C given A
B->A->C	yes	no
C->A->B	yes	no
C->A<-B	no	yes
B<-A->C	yes	no

Programming break (linear regression)

I will show you bayesian approach for linear regression.

$$y = mx + c$$

$$y = \beta_1 x + \beta_2 x^2 + \beta_0$$

$$\begin{bmatrix} y_1 \\ y_2 \\ y_3 \\ \vdots \\ y_n \end{bmatrix} = \begin{bmatrix} 1 & x_1 & x_1^2 & \dots & x_1^m \\ 1 & x_2 & x_2^2 & \dots & x_2^m \\ 1 & x_3 & x_3^2 & \dots & x_3^m \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & x_n & x_n^2 & \dots & x_n^m \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \\ \vdots \\ \beta_m \end{bmatrix}$$

$$Y = X \times W$$

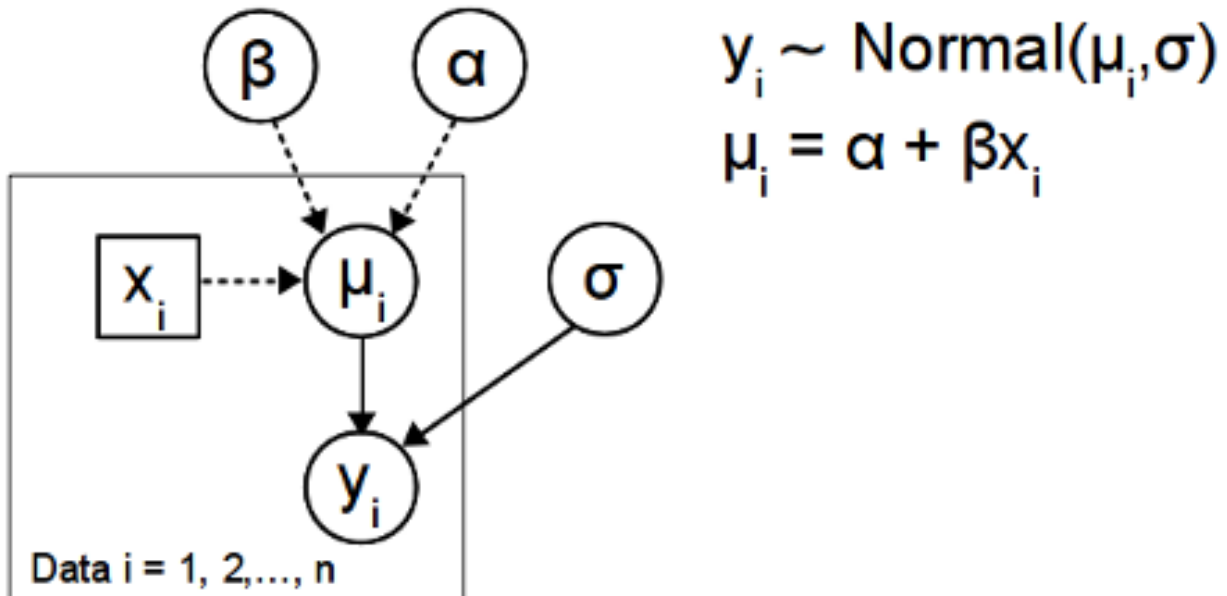
$$X \rightarrow Y \leftarrow W$$

$$p(W | X, Y) = \frac{p(W, X, Y)}{p(X, Y)}$$

$$p(W | X, Y) \propto p(W, X, Y)$$

$$p(W | X, Y) \propto p(W)p(Y | W, X)$$

We don't know anything about the actual model of $p(W)$ and $p(Y | X, W)$, but we can assume it.



$$p(Y | W, X) = \mathcal{N}(Y | \mu, \sigma I)$$

$$p(Y | W, X) = \mathcal{N}(Y | XW, \sigma I)$$

and

$$p(W) = \mathcal{N}(W | 0, \gamma I)$$

you get

$$p(W | X, Y) \propto p(W)p(Y | W, X)$$

$$p(W | X, Y) \propto \mathcal{N}(W | 0, \gamma I) \mathcal{N}(Y | \mu, \sigma I)$$

The multivariate gaussian function of Y is

$$\mathcal{N}(Y | \mu, \sigma) = \frac{1}{(2\pi)^{d/2}} |\sigma I|^{-1/2} \exp\left[-\frac{1}{2}(y - \mu)^T (\sigma I)^{-1} (y - \mu)\right]$$

The multivariate gaussian function of W is

$$\mathcal{N}(W | \mu, \sigma I) = \frac{1}{(2\pi)^{d/2}} |\gamma I|^{-1/2} \exp\left[-\frac{1}{2}(w - \mu)^T (\gamma I)^{-1} (w - \mu)\right]$$

You get

$$\mathcal{N}(W | 0, \gamma) \mathcal{N}(Y | \mu, \sigma) = \frac{1}{(2\pi)^{d/2}} |\gamma I|^{-1/2} \exp\left[-\frac{1}{2}(w - \mu)^T (\gamma I)^{-1} (w - \mu)\right] \frac{1}{(2\pi)^{d/2}} |\sigma I|^{-1/2} \exp\left[-\frac{1}{2}(y - \mu)^T (\sigma I)^{-1} (y - \mu)\right]$$

You want to find w that make the function below maximum

$$\frac{1}{(2\pi)^{d/2}} |\gamma I|^{-1/2} \exp\left[-\frac{1}{2}(w - \mu)^T (\gamma I)^{-1} (w - \mu)\right] \frac{1}{(2\pi)^{d/2}} |\sigma I|^{-1/2} \exp\left[-\frac{1}{2}(y - \mu)^T (\sigma I)^{-1} (y - \mu)\right]$$

take log and find w that give the maximum of below function

$$\log C_1 - \frac{1}{2}(w)^T (\gamma I)^{-1} (w) + \log C_2 - \frac{1}{2}(y - \mu)^T (\sigma I)^{-1} (y - \mu)$$

$$- > -\frac{1}{2}(w)^T (\gamma I)^{-1} (w) - \frac{1}{2} \frac{\|y - xw\|^2}{\sigma^2}$$

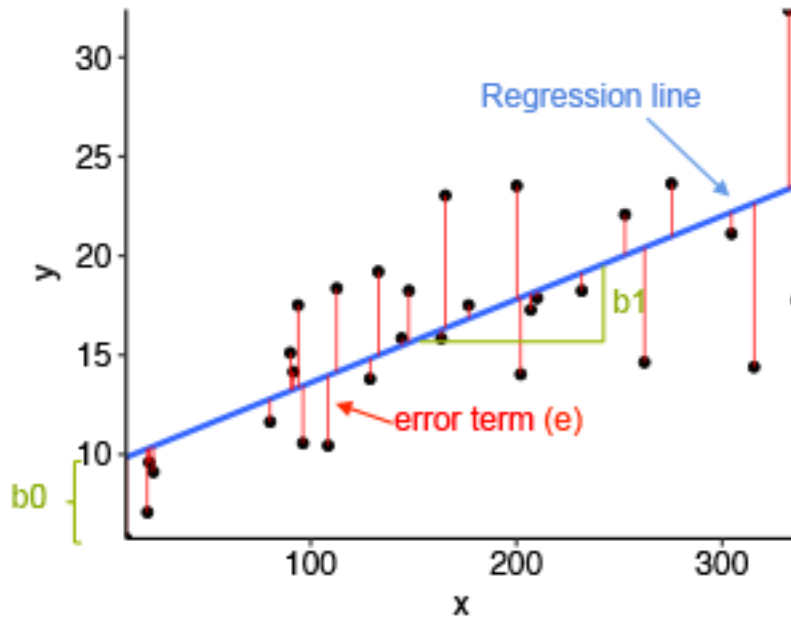
$$= -\frac{1}{2\gamma} \|w\|^2 - \frac{1}{2} \frac{\|y - xw\|^2}{\sigma^2}$$

$$- > -\|y - xw\|^2 - \lambda \|w\|^2$$

Find w that give maximum of above function or find the minimum of below function

$$\|y - xw\|^2 + \lambda \|w\|^2$$

Let look at $\|y - xw\|^2$ term



And $\lambda ||w||^2$ term.

$$||y - xw||^2 + \lambda ||w||^2 = (y - xw)^T (y - xw) + \lambda w^T w$$

$$= y^T y - 2y^T xw + w^T x^T xw + \lambda w^T w$$

When find w that give the function minimum you just differentiate it.

$$\frac{\partial y^T y - 2y^T xw + w^T x^T xw + \lambda w^T w}{\partial w} = 0$$

$$-2y^T x + 2x^T xw + 2\lambda w = 0$$

$$(2x^T x + 2\lambda I)w = 2y^T x$$

$$w = (x^T x + \lambda I)^{-1} y^T x$$

Problem about bayesian inference

Let define the random variables first

- Evidence variable: $E_1 = e_1 \dots E_k = e_k$
- Query variable: Q
- Hidden variable: $H_1 \dots H_r$

We want $p(Q \mid e_1 \dots e_k)$

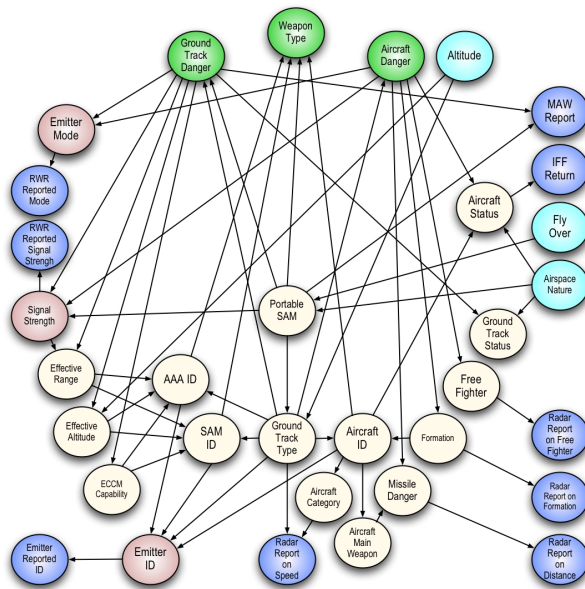
$$p(Q, e_1 \dots e_k) = \sum_{h_1 \dots h_r} p(Q, h_1 \dots h_r e_1 \dots e_k)$$

And

$$p(Q \mid e_1 \dots e_k) = \frac{p(Q, e_1 \dots e_k)}{p(e_1 \dots e_k)}$$

$$p(Q \mid e_1 \dots e_k) = \frac{\sum_{h_1 \dots h_r} p(Q, h_1 \dots h_r e_1 \dots e_k)}{p(e_1 \dots e_k)}$$

If you have a very large network, it very slow and hard to compute.



Solution Variable elimination

Latent variable

EM algorithm

Variational inference

Sampling method

Python Library

1. PyMC3, PyMC4
2. pyStan
3. pyro
4. Tensorflow probability
5. Edward 2

Reference

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3. <https://www.youtube.com/playlist?list=PLe5rNUydzV9QHe8VDStpU0o8Yp63OecdW>
4. http://ai.berkeley.edu/lecture_slides.html