

## Poisson Process

(1)

This process arises in situations where we are interested in the total no. of arrivals of a particular type of events upto a specified time  $t$ ,  $t \geq 0$  such as no. of telephone calls to a call center, no. of customers arriving to a shop, no. of defective products in a lot, printing errors and so on.

In many such fields of practical studies Poisson process has found wide applications. To analyze such process in a meaningful manner we need to put certain assumptions on ~~phen~~ phenomenon under investigation.

Let  $X(t)$  denote no. of customers in the system in a time interval of length  $t$ .

We are interested to evaluate

$$P(X(t) = n) = P_n(t), \quad n = 0, 1, 2, 3, \dots \quad \text{--- ①}$$

under the following assumptions.

## Assumption for Poisson Process:

- (i) The number of arrivals of customers during disjoint time intervals are independent. That is no. of occurrence ~~of~~ during the time interval  $(t, t+h]$  is independent of the no. of occurrences in  $(0, t]$  for all  $h \geq 0$ .
- (ii) Probability of exactly one occurrence during a small time interval is proportional to the length of the interval. That is probability of exactly one occurrence in  $(t, t+h]$  is  $\lambda h$  where  $\lambda$  is ~~proportionality~~ proportionality constant. In fact  $\lambda$  is the rate at which customer arrives. [ $\because P(X(h)=1) = P_1(h) = \lambda h$ ]
- (iii) Probability of more than one occurrence during a small time interval is negligible. Thus we have

$$\begin{aligned} P(X(h) > 1) &= P(X(h) \geq 2) = o(h) \\ \Rightarrow P_2(h) + P_3(h) + \dots &= o(h) \end{aligned} \quad \left\{ \begin{array}{l} f(x) \text{ is order} \\ o(x) \text{ if} \\ \frac{f(x)}{x} \rightarrow 0 \\ \text{as } x \rightarrow 0 \end{array} \right.$$

Note that  $P_0(h) + P_1(h) + P_2(h) + P_3(h) + \dots = 1$

$$\therefore 1 - P_0(h) - P_1(h) = o(h)$$

$$\Rightarrow P_0(h) = 1 - P_1(h) - o(h) = 1 - \lambda h - o(h)$$

③

Under assumptions (i), (ii) and (iii) we find that we have  $X(t)$  distributed as Poisson dist<sup>n</sup> with parameter  $\lambda t$ , that is,

$$P(X(t)=n) = P_n(t) = \frac{e^{-\lambda t} (\lambda t)^n}{n!}, n=0,1,2,3,\dots$$

Proof: This result is proved using the mathematical induction. Let us look at the probability  $P_0(t+h)$ ,

$$\begin{aligned} P_0(t+h) &= P(\text{no customer arrived during the time interval } (0, t+h]) \\ &= P(\text{no customer arrive during } (0, t) \cap \text{no customer arrive during } (t, t+h]) \\ &= P(\text{no customer arrive in } (0, t]) \cdot P(\text{no customer arrive during the time } (t, t+h]) \\ &= P_0(t) P_0(h) = P_0(t) (1 - P_1(h) - o(h)) \\ &= P_0(t) (1 - \lambda h - o(h)) \end{aligned}$$

$$\frac{P_0(t+h) - P_0(t)}{h} = -\lambda P_0(t) - \frac{o(h)}{h} P_0(t)$$

as  $h \rightarrow 0$  we have  $\frac{dP_0(t)}{dt} = -\lambda P_0(t)$

(4)

Solving this we get

$$P_0(t) = c e^{-\lambda t}$$

find the constant c: note that  $P_0(0) = 1$   
 $\Rightarrow c = 1$

$$\therefore \boxed{P_0(t) = e^{-\lambda t}}$$

Equation (1) holds for  $n = 0$ .

Next consider

$$P_1(t+h) = P(\text{single customer arrives in time interval } (0, t+h])$$

$$= P(\{ \text{one customer arrives in } (0, t] \} \cap \{ \text{no customer arrives in } (t, t+h] \}) + P(\{ \text{no customer arrives in } (0, t] \} \cap \{ \text{one customer arrives in } (t, t+h] \})$$

$$= P(\text{one customer in } (0, t]) P(\text{no customer in } (t, t+h]) + P(\text{no customer in } (0, t]) P(\text{one customer in } (t, t+h])$$

$$= P_1(t) P_0(h) + P_0(t) P_1(h)$$

$$= P_1(t) (1 - \lambda h - o(h)) + \lambda h e^{-\lambda t}$$

$$\frac{P_1(t+h) - P_1(t)}{h} = -\lambda P_1(t) - P_1(t) \frac{o(h)}{h} + \lambda e^{-\lambda t}$$

As  $h \rightarrow 0$  we get that

$$\frac{dP_1(t)}{dt} = -\lambda P_1(t) + \lambda e^{-\lambda t}$$

$$\Rightarrow \frac{dP_1(t)}{dt} + \lambda P_1(t) = \lambda e^{-\lambda t}$$

Solving we get  $P_1(t) e^{\lambda t} = \lambda t + c$

To find the constant  $c$  we use the condition

$$P_1(0) = 0. \text{ This implies that } c = 0.$$

$$\therefore \boxed{P_1(t) = \lambda t e^{-\lambda t}}$$

Thus result stated in Eqn ① holds for  $n=1$  as well.

Let us assume that this result is true up to  $n=k$ .

Then we look ~~the~~ at following probability:

$$P_{k+1}(t+h) = P((k+1) \text{ customers arrive in } (0, t+h])$$

$$= P[\{(k+1) \text{ customers arrive in } (0, t]\} \cap \{\text{no customers arrive in } (t, t+h]\}] + P[\{k \text{ customer in } (0, t]\} \cap \{\text{one customer in } (t, t+h]\}]$$

$$+ \sum_{i=1}^k P[\{(k-i) \text{ customer in } (0, t]\} \cap \{(i+1) \text{ customer in } (t, t+h]\}]$$



(6)

$$P_{k+1}(t+h) = P_{k+1}(t) P_0(h) + P_k(t) P_1(h) + \sum_{i=1}^k P_{k-i}(t) P_{i+1}(h)$$

$$= P_{k+1}(t) (1 - \lambda h - o(h)) + P_k(t) \lambda h + \sum_{i=1}^k P_{k-i}(t) o(h)$$

$$= P_{k+1}(t) (1 - \lambda h - o(h)) + \frac{e^{-\lambda t} (\lambda t)^k}{k!} \lambda h + \left( \sum_{i=1}^k P_{k-i}(t) \right) (o(h))$$

$$\frac{P_{k+1}(t+h) - P_{k+1}(t)}{h} = -\lambda P_{k+1}(t) - \frac{o(h)}{h} P_{k+1}(t) + \frac{e^{-\lambda t} (\lambda t)^k}{k!} \lambda + \left( \sum_{i=1}^k P_{k-i}(t) \right) \frac{o(h)}{h}$$

as  $h \rightarrow 0$  we have

$$\frac{dP_{k+1}(t)}{dt} = -\lambda P_{k+1}(t) + \frac{e^{-\lambda t} (\lambda t)^k}{k!} \lambda$$

$$\Rightarrow \frac{dP_{k+1}(t)}{dt} + \lambda P_{k+1}(t) = \frac{e^{-\lambda t} (\lambda t)^k}{k!} \lambda$$

we have to solve this ODE. note that IF is  $e^{\lambda t}$

(7)

Solving it we get

$$P_{k+1}(t) e^{\lambda t} = \int e^{-\lambda t} \frac{(\lambda t)^k}{k!} \lambda \cdot e^{\lambda t} dt$$

$$= \frac{(\lambda t)^{k+1}}{(k+1)!} + c$$

To find the constant  $c$ , we note that

$$P_{k+1}(0) = 0 \text{ which implies that } c = 0.$$

$$\therefore P_{k+1}(t) = \frac{e^{-\lambda t} (\lambda t)^{k+1}}{(k+1)!}$$

So Equation (1) holds for  $n = k+1$  also.

Thus it holds for all  $n = 0, 1, 2, 3, \dots$

Ex: Let average no. of telephone calls arriving at a call center is 30 calls per hour.

- (i) What is the probability that no call arrives in a 3 minute period
- (ii) What is the probability that more than 5 calls arrive in a 5 minute period.

Solution:

$$\lambda = 30 \text{ call per hour}$$

$$= \frac{1}{2} \text{ call per minute}$$

(i)  ~~$P(X(3)=0)$~~

$X(t)$ : no of calls during time interval of length  $t$ .

$$P(X(t)=n) = \frac{e^{-\lambda t} (\lambda t)^n}{n!}, \quad n=0,1,2,3, \dots$$

$$\lambda = 1/2$$

$$(i) \quad P(X(3)=0) = \frac{e^{-\frac{1}{2} \cdot 3} \left(\frac{1}{2} \cdot 3\right)^0}{0!} = e^{-3/2}.$$

$$(ii) \quad P(X(5) > 5) = 1 - P(X(5) \leq 5)$$

$$= 1 - \sum_{j=0}^5 \frac{e^{-5/2} (5/2)^j}{j!}$$

$$\approx 0.42.$$