#### 1 HP97

The Hodrick and Prescott (1997, HP) Filter can be expressed in State Space Form (SSF) using the following Unobserved Component (UC) model structure:

$$y_t = y_t^* + y_t^c \tag{1a}$$

$$\Delta^2 y_t^* = \varepsilon_{1t} \tag{1b}$$

$$y_t^c = \phi \varepsilon_{2t},$$
 (1c)

where  $y_t$  is (100 times) the log of GDP, and  $\varepsilon_{1t}$  and  $\varepsilon_{2t}$  are i.i.d. N(0,1) and mutually uncorrelated. The standard deviation  $\phi$  is the (square root of the) smoothing parameter, commonly set to 40 for quarterly macroeconomic data, implying a value of ' $\lambda$ ' of 1600.

The 'numbered shock' to 'named shock' mapping is:

$$\begin{bmatrix} \varepsilon_{1t} \\ \varepsilon_{2t} \end{bmatrix} = \begin{bmatrix} \varepsilon_t^* \\ \varepsilon_t^c \end{bmatrix}, \tag{2}$$

where  $\varepsilon_t^*$  is the trend (or permanent) shock and  $\varepsilon_t^c$  is the cycle (or transitory) shock.

# 2 Shock recovery

### 2.1 State Space Models with lagged states

Kurz (2018) adopts the following general SSF with lagged states in the measurement:

Measurement : 
$$Z_t = D_1 X_t + D_2 X_{t-1} + R \varepsilon_t$$
 (3a)

State: 
$$X_t = AX_{t-1} + C\varepsilon_t$$
, (3b)

where  $\varepsilon_t \sim MN(0_m, I_m)$ ,  $D_1, D_2, A, R$  are C are conformable system matrices,  $Z_t$  the observed variable and  $X_t$  the latent state variable, and m is the number of shocks  $\{\varepsilon_{it}\}_{i=1}^2$ .

## 2.2 HP97 in 'shock recovery' SSF

To assess shock recovery, write the model in (1) in 'shock recovery' SSF by collecting all observable variables in  $Z_t$  and all shocks (and other latent state variables) in  $X_t$ . Differencing  $y_t$  and  $y_t^c$  twice, and re-arranging the relations in (1) then yields:

$$\Delta^{2} y_{t} = \Delta^{2} y_{t}^{*} + \Delta^{2} y_{t}^{c}$$

$$= \varepsilon_{1t} + \phi \Delta^{2} \varepsilon_{2t}$$

$$= \varepsilon_{1t} + \phi \varepsilon_{2t} - 2\phi \varepsilon_{2t-1} + \phi \varepsilon_{2t-2},$$
(4)

where  $\Delta^2 y_t$  is the only observed variable.

The Measurement and State equations of the 'shock recovery' SSF corresponding to the relations in (4) are then given by:

Measurement : 
$$Z_t = D_1 X_t + D_2 X_{t-1} + R \varepsilon_t$$
  
 $\Leftrightarrow \Delta^2 y_t = \varepsilon_{1t} + \phi \varepsilon_{2t} - 2\phi \varepsilon_{2t-1} + \phi \varepsilon_{2t-2}$  (5a)

$$Z_{t} = \underbrace{\begin{bmatrix} 1 & \phi & 0 \end{bmatrix}}_{D_{1}} \underbrace{\begin{bmatrix} \varepsilon_{1t} \\ \varepsilon_{2t} \\ \varepsilon_{2t-1} \end{bmatrix}}_{X_{t}} + \underbrace{\begin{bmatrix} 0 & -2\phi & \phi \end{bmatrix}}_{D_{2}} \underbrace{\begin{bmatrix} \varepsilon_{1t-1} \\ \varepsilon_{2t-1} \\ \varepsilon_{2t-2} \end{bmatrix}}_{X_{t-1}} + \underbrace{\begin{bmatrix} 0 & 0 \end{bmatrix}}_{R} \underbrace{\begin{bmatrix} \varepsilon_{1t} \\ \varepsilon_{2t} \end{bmatrix}}_{\varepsilon_{t}}$$
(5b)

State:  $X_t = AX_{t-1} + C\varepsilon_t$ ,

$$\underbrace{\begin{bmatrix} \varepsilon_{1t} \\ \varepsilon_{2t} \\ \varepsilon_{2t-1} \end{bmatrix}}_{X_t} = \underbrace{\begin{bmatrix} 0 & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 1 & 0 \end{bmatrix}}_{A} \underbrace{\begin{bmatrix} \varepsilon_{1t-1} \\ \varepsilon_{2t-1} \\ \varepsilon_{2t-2} \end{bmatrix}}_{X_{t-1}} + \underbrace{\begin{bmatrix} 1 & 0 \\ 0 & 1 \\ 0 & 0 \end{bmatrix}}_{C} \underbrace{\begin{bmatrix} \varepsilon_{1t} \\ \varepsilon_{2t} \end{bmatrix}}_{\varepsilon_t}.$$
(5c)

### 2.3 Shock recovery

The diagonal of the steady-state variance/covariance matrix of the smoothed and filtered states  $X_t$  denoted by  $P_{t|T}^*$  and  $P_{t|t}^*$ , respectively, are:

$$\begin{array}{c|cccc}
 & P_{t|T}^* & P_{t|t}^* \\
\hline
\varepsilon_{1t} & 0.9439 & 0.9995 \\
\varepsilon_{2t} & 0.0561 & 0.2006
\end{array}$$
(6)

indicating that the trend (or permanent) shock  $\varepsilon_{1t} = \varepsilon_t^*$  cannot be recovered ( $P^* \approx 1$ ), while the cycle shock  $\varepsilon_{2t} = \varepsilon_t^c$  appears to be recoverable ( $P^* \approx 0$ ). In Figure 1, simulated states and estimated Kalman smoothed states are plotted for the two shocks of interest.

The correlation between the true (simulated) and estimated Kalman smoothed shocks can be analyzed by simply computing  $Corr(X_t, \hat{X}_{t|T})$ , where  $X_t = \begin{bmatrix} \varepsilon_{1t} & \varepsilon_{2t} & \varepsilon_{2t-1} \end{bmatrix}'$  and  $\hat{X}_{t|T} = E_T X_t = E_T \begin{bmatrix} \varepsilon_{1t} & \varepsilon_{2t} & \varepsilon_{2t-1} \end{bmatrix}'$ , yielding (for the first two elements of  $X_t$ ):

$$\frac{\operatorname{Corr}(X_{t}, \hat{X}_{t|T})}{\varepsilon_{1t} \qquad 0.2368} \\
\varepsilon_{2t} \qquad 0.9713 \tag{7}$$

As can be seen from (7), the estimated permanent shock  $\varepsilon_{1t}$  is only weakly correlated (0.2368) with the true value, while the estimated cyclical shock  $\varepsilon_{2t}$  is highly correlated (0.9713). These facts can also be seen from Figure 1 below.

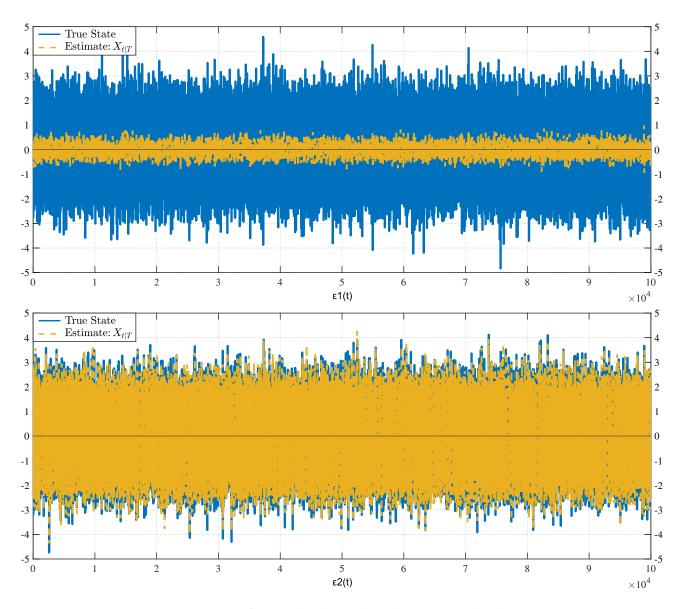


Figure 1: Comparison of true shocks and Kalman Smoothed estimates  $\varepsilon_{t|T}$ .

### 2.4 Shock Identities

Kalman Filter estimates of the permanent and transitory shocks  $\varepsilon_{1t}$  and  $\varepsilon_{2t}$  are linked by the identity:

$$E_t \varepsilon_{2t} = \phi E_t \varepsilon_{1t}, \tag{8}$$

and the corresponding Kalman Smoother estimates are linked by the *dynamic* identity:

$$\Delta^2 E_T \varepsilon_{1t} = \frac{1}{\phi} E_T \varepsilon_{2t-2}. \tag{9}$$

The filtered and smoothed estimates of the contemporaneous correlations are, respectively:

Co	$\operatorname{orr}(\hat{X}_{t t},\hat{X}_{t})$	(t t)			$\operatorname{Corr}(\hat{X}_{t T},$	$\hat{X}_{t T}$	_
	$\varepsilon_{1t}$	$\varepsilon_{2t}$	and		$\varepsilon_{1t}$	$\varepsilon_{2t}$	
$\overline{\varepsilon_{1t}}$	1	1	anu	$\varepsilon_{1t}$	1	-0.1907	•
$\varepsilon_{2t}$	1	1		$\varepsilon_{2t}$	-0.1907	1	

Note that  $\varepsilon_{1t}$  and  $\varepsilon_{2t}$  were generated *i.i.d.* N(0,1) and mutually uncorrelated.

Since  $\varepsilon_{1t} = \Delta^2 y_t^*$  and  $\varepsilon_{2t} = \frac{1}{\phi} y_t^c$ , this implies that the output from the standard HP–Filter will give the identity:

$$\Delta^4 y_t^* = \frac{1}{\phi} y_{t-2}^c$$
, or alternatively:  $\Delta^4 \text{HP-trend}_t = \frac{1}{\phi} \text{HP-cycle}_{t-2}$ .

Running regressions corresponding to (8) and (9) obtained from the simulated data (without an intercept) yields regression coefficient of 40 and 0.025 = 1/40 when the smoothing parameter was set to  $\lambda = 1600 = 40^2$  in the simulation. The regression fit is perfect, yielding an  $R^2$  of 1 and a residual sum of squares of exactly 0 (see Panels (a) and (b) in Table 1, respectively). The file HP97.m replicates the output summarized here.

<b>Variable</b>	Estimate	stderr(HAC)	t-stat(HAC)	p-value			
 Etε1(t)	40.000000	0.000000	619573092194908800.000000	0.000000			
R-squared	:			1.000000	No. of Regressors	:	1.00000
Rbar-squar	ed :			1.000000	Plus Const.(if exist	):	1.00000
SE of regr	ession :			0.000000	Mean(y)	:	0.00470
Sum Square	d Errors :			0.000000	Stdev(y)	:	0.89849
Log-likeli				355982.354905	AIC	:	-74.03414
F-statisti	.c :		38387081657196095016516104	46978854912.000000	AICc	:	-74.03414
Pr(F-stati	stic) :			0.000000	ј віс	:	-74.03342
No. of obs	ervations :			10000.000000	і ноіс	:	-74.03390
Std.err.ML	E (div by T) :			0.000000	DW-stat.	:	2.01616
	e-whitening :			0.000000	HAC Trunct.Lag.	:	12.00000
			(a) Kalma	an Filter			
	entity. Dependen		======================================			====	
======= Smoother Id ======== /ariable	entity. Dependen	t variable: Δ²E stderr(HAC)		an Filter		=====	
/ariable		stderr(HAC)	======================================			=====	
/ariable Tε2(t-2)	Estimate	stderr(HAC)	Tε1(t) 	p-value	No. of Regressors :		1.000000
/ariable	Estimate 	stderr(HAC)	Tε1(t) 	p-value 0.000000	No. of Regressors : Plus Const.(if exist) :	=====	1.000000
/ariable  Τε2(t-2) R-squared	Estimate 0.025000	stderr(HAC)	Tε1(t) 	p-value 0.000000		=====	
/ariable  TE2(t-2)  R-squared Rbar-squar SE of regr	Estimate  0.025000  : ed : ession :	stderr(HAC)	Tε1(t) 	p-value 0.000000 1.000000   1.000000	Plus Const.(if exist) :	=====	1.000000
/ariable  TE2(t-2) R-squared Rbar-squar SE of regr	Estimate  0.025000  :ed : ession : d Errors :	stderr(HAC)	Tε1(t) 	p-value 0.000000 1.000000   1.000000   0.000000	Plus Const.(if exist) : Mean(y) :	=====	1.000000 -0.000006
/ariable  TE2(t-2) R-squared Rbar-squar SE of regr Sum Square	Estimate  0.025000  : ed : ession : d Errors : hood :	stderr(HAC) 0.000000	Tε1(t) 	p-value  0.000000  1.000000   1.000000   0.000000   0.000000   353018.214591	Plus Const.(if exist) : Mean(y) : Stdev(y) :	=====	1.000000 -0.000006 0.024396
/ariable	Estimate  0.025000  : ed : ession : d Errors : hood : C :	stderr(HAC) 0.000000	TE1(t) 	p-value  0.000000  1.000000   1.000000   0.000000   0.000000   353018.214591	Plus Const.(if exist) : Mean(y) : Stdev(y) : AIC :	=====	1.000000 -0.000006 0.024396 -73.455443
/ariable	Estimate  0.025000  : ed : ession : d Errors : hood : c : stic) :	stderr(HAC) 0.000000	TE1(t) 	p-value  0.000000  1.000000   1.000000   0.000000   0.000000   353018.214591   56627712.000000	Plus Const.(if exist) : Mean(y) : Stdev(y) : AIC : AICc :	=====	1.000000 -0.000006 0.024396 -73.455443 -73.455443
/ariable	Estimate  0.025000  : ed : ession : d Errors : hood : c : stic) :	stderr(HAC) 0.000000	TE1(t) 	p-value  0.000000  1.000000    1.000000    0.000000    0.000000    353018.214591    56627712.000000    0.000000	Plus Const.(if exist) : Mean(y) : Stdev(y) : AIC : AICc : BIC :	=====	1.000000 -0.000006 0.024396 -73.455443 -73.455443 -73.454722

(b) Kalman Smoother

Table 1: Shock identity regressions using Kalman filter and smoother estimates.

In summary, the HP-Filter can reasonably accurately recover the transitory cyclical shock  $\varepsilon_{2t} = \varepsilon_t^c$ , but not the permanent shock. Moreover, the HP-Filter will produce negatively correlated trend-cycle shocks (Corr( $E_T\varepsilon_{1t}$ ,  $E_T\varepsilon_{2t}$ ) = -0.1907), despite the data being generated from a model that has zero shock correlation. The HP-Filter shock correlation is thus spurious.