

<u>Lecture 8: Distance measures</u>

13. Parameter Estimation via KL

课程 □ Unit 3 Methods of Estimation □ between distributions

Divergence

# 13. Parameter Estimation via KL Divergence **Deriving the Maximum Likelihood Estimator**

#### Maximum likelihood

$$\widehat{\mathsf{KL}}(\mathbb{P}_{\theta^*}, \mathbb{P}_{\theta}) = \text{"constant"} - \frac{1}{n} \sum_{i=1}^n \log p_{\theta}(X_i)$$

$$\min_{\theta \in \Theta} \widehat{\mathsf{KL}}(\mathbb{P}_{\theta^*}, \mathbb{P}_{\theta}) \qquad \frac{1}{n} \sum_{i=1}^n \log p_{\theta}(X_i)$$

(Caption will be displayed when you start playing the video.)

This is the maximum likelihood principle.

Start of transcript. Skip to the end.

OK, so what is the maximum likelihood estimator?

Well, I have this KL, I just reproduce the same thing.

So now I am just trying to minimize the KL. Now, what those equivalence signs means, what you need to pay attention is that this equivalence is not an equality.

## 视频

下载视频文件

下载 SubRip (.srt) file 下载 Text (.txt) file

The next four problems concern the following statistical set-up.

You observe discrete random variables

$$X_1,\ldots,X_n \stackrel{iid}{\sim} P_{ heta^*}$$

where  $heta^*$  is the true parameter. You construct an associated statistical model  $(E,\{P_{ heta}\}_{ heta\in\mathbb{R}})$ . The sample space E is discrete.

Intuitively, your goal is to find an estimator  $\hat{ heta}_n\in\mathbb{R}$  so that the distributions  $P_{\hat{ heta}_n}$  and  $P_{ heta^*}$  are close. Precisely, you want to find an estimator  $\hat{m{ heta}}_n \in \mathbb{R}$  so that the quantity

$$\mathrm{KL}\left(P_{ heta^*},P_{\hat{ heta}_n}
ight)$$

is as small as possible.

This approach will naturally lead to the construction of the **maximum likelihood estimator** .

#### Finding a Minimizer of KL Divergence

0/1 point (graded)

Consider the optimization problem in which we minimize the KL divergence between  $P_{ heta^*}$ , the true distribution, and  $P_{ heta}$ . Formally, we want to solve

$$\min_{ heta \in \mathbb{R}} \operatorname{KL}\left(P_{ heta^*}, P_{ heta}
ight).$$

We are not so much interested in the minimum value attained by the objective function  $\mathrm{KL}\,(P_{\theta^*},P_{\theta})$ , but rather the value of  $\theta$  where the minimum is attained. We refer to such a  $\theta$  as a **minimizer** .

Let's suppose that there is a unique minimizer for the above optimization problem– *i.e.*, if m is the minimum value of  $\mathrm{KL}\,(P_{\theta^*},P_{\theta})$ , there is only one point  $\theta_{\min}$  such that

$$m = \mathrm{KL}\left(P_{ heta^*}, P_{ heta_{\min}}
ight).$$

For which heta is the minimum value of  $\mathrm{KL}\left(P_{ heta^*},P_{ heta}
ight)$  attained? (Equivalently, what is  $heta_{\min}$ ?)

⊕* □	
$\circ$ $\theta$	
O 0	
None of the above.	

#### **Solution:**

The KL divergence is nonnegative, so  $\mathrm{KL}\left(P_{\theta^*},P_{\theta}\right)\geq 0$ . The right-hand side is achieved if we set  $\theta=\theta^*$ :  $\mathrm{KL}\left(P_{\theta^*},P_{\theta^*}\right)=0$ . Since the minimizer is unique by assumption, we conclude that the minimum value is attained at  $\theta=\theta^*$ .

**Remark:** The assumption that there is a unique minimizer holds if we are given that the parameter  $\theta$  is identified. Here is why: since KL divergence is definite,  $\mathrm{KL}\left(P_{\theta^*},P_{\theta}\right)=0$  if and only if  $P_{\theta^*}$  and  $P_{\theta}$  are the same distribution. And if  $\theta$  is identified, this implies that  $\theta=\theta^*$ 

提交 你已经尝试了2次(总共可以尝试2次)

□ Answers are displayed within the problem

## Can we Minimize KL Divergence Directly?

2/2 points (graded)

Let's use the same statistical set-up as above. Recall that you have access to the iid samples  $X_1,\ldots,X_n$ . You use these samples to build an estimator  $\hat{\theta}_n$ . Can you compute

$$\mathrm{KL}\left(P_{\hat{ heta}_n},P_{1/2}
ight)$$

without knowing  $\theta^*$ , the true parameter?

Yes □No

Can you compute

without knowing $oldsymbol{ heta}^*$ ?
O Yes
● No □
Solution:
In general, we can compute $\mathrm{KL}\left(\mathbf{P},\mathbf{Q}\right)$ if and only if we know both distributions $\mathbf{P}$ and $Q$ . Moreover, by our statistical model, we can compute $P_{\theta}$ if and only if we know the real number $\theta$ . Putting these last two facts together, we can compute

# $\mathrm{KL}\left(P_{\hat{ heta}_n},P_{1/2} ight)$

because  $\hat{\theta}_n$  is known-it is an estimator so its expression does not depend on  $\theta^*$ , the true parameter. However, regardless of how many samples we take, we cannot compute  $\mathrm{KL}\left(P_{\theta^*},P_{1/2}\right)$  exactly because the distribution  $P_{\theta^*}$  is unknown.

**Remark:** Since we cannot even compute the function  $\mathrm{KL}\left(P_{ heta^*},P_{ heta}
ight)$  for general heta, this implies that the optimization problem

$$\min_{ heta \in \mathbb{R}} \operatorname{KL}\left(P_{ heta^*}, P_{ heta}
ight)$$

cannot be solved exactly, regardless of the number of samples we have. So to estimate the minimizer of this optimization problem (which is the true parameter  $\theta^*$ ) we will have to consider an approximation for  $\mathrm{KL}\left(P_{\theta^*},P_{\theta}\right)$ .

提交

你已经尝试了1次(总共可以尝试2次)

Answers are displayed within the problem

## Finding the Minimizer for an Approximation of KL Divergence

1/1 point (graded)

We use the same statistical set-up as above. Recall that  $X_1,\ldots,X_n\stackrel{iid}{\sim}P_{ heta^*}$  . Let  $p_{ heta}$  be the pmf of  $P_{ heta}$ .

Which of the following is a (weakly) consistent estimator for

$$\mathbb{E}_{ heta^{st}}\left[\ln p_{ heta}\left(X
ight)
ight] = \sum_{x \in E} p_{ heta^{st}} \ln p_{ heta}\left(x
ight) \, ?$$

$$\stackrel{\bigcirc}{=} \frac{1}{n} \sum_{i=1}^n X_i$$

$$\stackrel{ullet}{=} rac{1}{n} \sum_{i=1}^n \ln \left( p_ heta \left( X_i 
ight) 
ight) \square$$

$$rac{1}{n}\sum_{i=1}^{n}\ln\left(p_{ heta^{st}}\left(X_{i}
ight)
ight)-rac{1}{n}\sum_{i=1}^{n}\ln\left(p_{ heta}\left(X_{i}
ight)
ight)$$

$$ullet \; heta^* - \mathbb{E}_{ heta^*} \left[ \ln p_{ heta^*} 
ight]$$

**Solution:** 

By the law of large numbers,  $\frac{1}{n}\sum_{i=1}^n\ln\left(p_{\theta}\left(X_i\right)\right) o \mathbb{E}_{\theta^*}\left[\ln p_{\theta}\right]$  in probability. Hence, the second choice is correct.

**Remark 1:** The KL divergence between  $P_{ heta^*}$  and  $P_{ heta}$  can be written

$$\mathrm{KL}\left(P_{ heta^*},P_{ heta}
ight) = \sum_{x \in E} p_{ heta^*} \ln p_{ heta^*}\left(x
ight) - \sum_{x \in E} p_{ heta^*} \ln p_{ heta}\left(x
ight) = \mathbb{E}_{ heta^*}\left[\ln p_{ heta^*}\left(X
ight)
ight] - \mathbb{E}_{ heta^*}\left[\ln p_{ heta}\left(X
ight)
ight]$$

where  $X \sim P_{ heta^*}$  .

**Remark 2:** While we can't find heta that minimizes  $\mathrm{KL}\,(P_{ heta^*},P_{ heta})$ , we can find heta that minimizes

$$\hat{\mathrm{KL}}\left(P_{ heta^*},P_{ heta}
ight):=\mathbb{E}_{ heta^*}\left[\ln p_{ heta^*}
ight]-rac{1}{n}\sum_{i=1}^n \ln \left(p_{ heta}\left(X_i
ight)
ight).$$

Here's why: the first term on the RHS,  $\mathbb{E}_{\theta^*}\left[\ln p_{\theta^*}\right]$  , does not depend on  $\theta$ . Hence, the  $\theta$  that minimizes  $\hat{\mathrm{KL}}\left(P_{\theta^*},P_{\theta}\right)$  is the same as the  $\theta$  that minimizes  $-\frac{1}{n}\sum_{i=1}^n \ln\left(p_{\theta}\left(X_i\right)\right)$ .

提交

你已经尝试了1次(总共可以尝试2次)

☐ Answers are displayed within the problem

### Deriving the Maximum Likelihood Estimator

1/1 point (graded)

We use the same statistical set-up as above. Recall that  $p_ heta$  is the pmf of  $P_ heta$  and  $X_1,\dots,X_n\stackrel{iid}{\sim}P_{ heta^*}$  .

Suppose that  $heta_{\min}$  is a minimizer for the function

$$f\left( heta
ight) := -rac{1}{n}\sum_{i=1}^{n}\ln\left(p_{ heta}\left(X_{i}
ight)
ight)$$

Which of the following functions is also minimized at  $heta_{\min}$ ?

$$\bigcirc \ g_1\left( heta
ight) = -\prod_{i=1}^n p_{ heta}\left(X_i
ight)$$

$$igcup_{0} g_{2}\left( heta
ight)=25-\prod_{i=1}^{n}p_{ heta}\left(X_{i}
ight)$$

$$g_3\left( heta
ight) = h\left( heta^*
ight) - \prod_{i=1}^n p_ heta\left(X_i
ight)$$
 where  $h$  is a function of  $heta^*$  that does **not** depend on  $heta$ .

$$\bigcirc \ g_4\left( heta
ight) = heta^* - \prod_{i=1}^n p_{ heta}\left(X_i
ight)$$

 $^{ullet}$  All of the above  $\Box$ 

#### **Solution:**

Observe that rescaling by n does not change where the minimum of a function is attained. Hence,  $f(\theta)$  and  $nf(\theta)$  have the same minimizer. Next, by the addition property of logarithms,

$$nf\left( heta
ight) = \sum_{i=1}^{n} \ln\left(p_{ heta}\left(X_{i}
ight)
ight) = \ln\left(\prod_{i=1}^{n} p_{ heta}\left(X_{i}
ight)
ight).$$

Since ln is an increasing function, the function

$$heta \mapsto \prod_{i=1}^n p_ heta\left(X_i
ight)$$

has the same minimizer as  $\ln\left(\prod_{i=1}^n p_{ heta}\left(X_i
ight)\right)$ . Thus the first choice is correct.

Moreover, the second and third choices are also correct. Whenever we have an optimization problem

$$\min_{ heta \in \mathbb{R}} C + g\left( heta
ight)$$

where  $m{C}$  does not depend on  $m{ heta}$ , then the above will have the same minimizer as the optimization problem

$$\min_{ heta \in \mathbb{R}} g\left( heta
ight).$$

In the second choice, C=25 (which is independent of heta), and in the third choice,  $C=h\left( heta^*
ight)$  (which by assumption is independent of heta).

**Remark 1:** The quantity

$$\hat{ heta}_n := ext{maximizer of } \prod_{i=1}^n p_{ heta}\left(X_i
ight)$$

is referred to as the maximum likelihood estimator . Note that this is the same as the estimator

$$\hat{ heta}_n := ext{minimizer of} \quad - \frac{1}{n} \sum_{i=1}^n \ln \left( p_{ heta} \left( X_i 
ight) 
ight)$$

considered in Remark 2 in the solution of the previous problem.

**Remark 2:** Under certain technical conditions, the maximum likelihood estimator is guaranteed to (weakly) converge to the true parameter  $\theta^*$ .

提交

你已经尝试了1次(总共可以尝试3次)

Answers are displayed within the problem

讨论

显示讨论

主题: Unit 3 Methods of Estimation:Lecture 8: Distance measures between distributions / 13. Parameter Estimation via KL Divergence