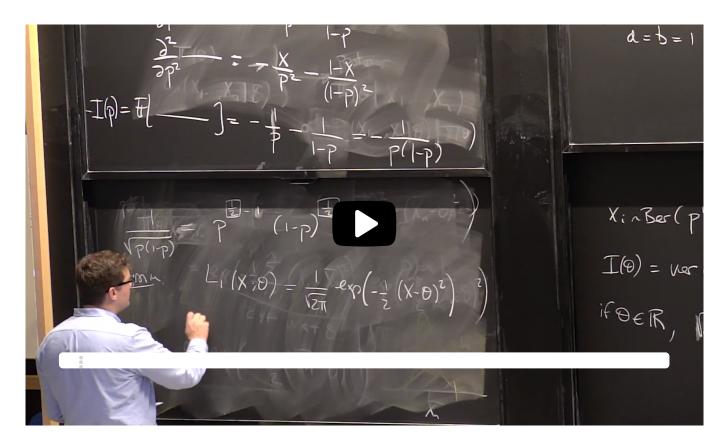
7. Jeffreys Prior II: Examples Jeffreys Prior II: Examples



is just proportional to 1.

This is the same as the one we had before. In the Gaussian case, there's no point that's harder to estimate than any other, right?

This is completely a translation invariant problem.

Nothing is harder than any other point.

And the reason why things can change is because the variance might depend on your parameter,

but if nothing depends on your parameter, then it doesn't matter where the mean is.

And so in this case we have this prior.

How did we call such priors?

Improper priors.

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Computing Jeffreys Prior

2/2 points (graded)

Let N be a Poisson random variable. That is,

$$p_{N}\left(n|\lambda
ight) =e^{-\lambda}rac{\lambda^{n}}{n!},$$

where $p_N\left(n|\lambda\right)$ denotes the conditional pmf of N given the parameter λ .

• Evaluate the Jeffreys prior, $p(\lambda)$ up to a proportionality constant, which only is a function of λ . Remove outside constants in your answer such that I(1) = 1.

1/sqrt(lambda) \checkmark Answer: 1/sqrt(lambda) $\frac{1}{\sqrt{\lambda}}$

- Is the Jeffrey's prior proper?
 - Yes

STANDARD NOTATION

Solution:

• We begin by computing Fisher information, $I(\lambda)$, as follows.

$$rac{d}{d\lambda}{\log p_N\left(n|\lambda
ight)} = -1 + rac{n}{\lambda}.$$

Therefore,

$$I\left(\lambda
ight)=\mathbb{E}\left[\left(rac{d}{d\lambda}{
m log}\,p_N\left(n|\lambda
ight)
ight)^2
ight]=\mathbb{E}\left[rac{\left(N-\lambda
ight)^2}{\lambda^2}
ight]=rac{1}{\lambda^2}{
m Var}\left(N
ight).$$

Since ${
m Var}\left(N
ight)=\lambda$, for a Poisson random variable, we arrive at,

$$I\left(\lambda
ight)=rac{1}{\lambda}\implies p\left(\lambda
ight)\propto\sqrt{I\left(\lambda
ight)}=rac{1}{\sqrt{\lambda}}.$$

• Since $\lambda \in (0, \infty)$, we can check that,

$$\int_0^\infty \frac{1}{\sqrt{\lambda}} \ d\lambda = \infty,$$

and therefore, Jeffreys prior for this problem is improper.

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You have used 2 of 3 attempts

Answers are displayed within the problem

Jeffreys Prior for Matrix Case

3/3 points (graded)

In this problem we will consider a model which has a two-dimensional parameter. Then you will calculate Jeffrey's prior using the Fisher information matrix.

Suppose that $X_1, \ldots, X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$, where μ and σ^2 are unknown. In this case, the Fisher information matrix $I(\theta)$ for $\theta = (\mu, \sigma^2)$ will be a 2×2 matrix, where the off-diagonal entries are 0.

• Find $(I(\theta))_{11}$.

• Find $(I(\theta))_{22}$.

1/(2*sigma^4)

Answer: 1/(2*sigma^4) $\frac{1}{2 \cdot \sigma^4}$

• Using your answers to the previous part, determine Jeffreys prior, $\pi(\theta)$, in terms of μ and σ . Express your answer in such a form that $\pi((1,1))=1$.



STANDARD NOTATION

Solution:

• Clearly, the likelihood model is of form,

$$p_{Y|\mu,\sigma} = rac{1}{\sqrt{2\pi\sigma^2}} \mathrm{exp}\left(-rac{(x-\mu)^2}{2\sigma^2}
ight) \implies rac{1}{2} \mathrm{log}\, p_{Y|\mu,\sigma} = \mathrm{log}\left(rac{1}{2\pi\sigma^2}
ight) - rac{(x-\mu)^2}{2\sigma^2}.$$

In particular,

$$\left(I\left(heta
ight)
ight)_{11} = -\mathbb{E}\left[rac{\partial^{2}}{\partial\mu^{2}}{
m log}\,p_{Y|\mu,\sigma}
ight] = rac{1}{\sigma^{2}},$$

using the fact that,

$$rac{\partial}{\partial \mu} {
m log} \, p_{Y|\mu,\sigma} = rac{x-\mu}{\sigma^2} \implies rac{\partial^2}{\partial \mu^2} {
m log} \, p_{Y|\mu,\sigma} = -rac{1}{\sigma^2}.$$

ullet Using the exact same strategy as above, and the fact that, $\mathbb{E}\left[\left(X-\mu
ight)^2
ight]=\mathrm{Var}\left(X
ight)=\sigma^2$, we obtain that,

$$\left(I\left(heta
ight)
ight)_{22}=rac{1}{2\sigma^{4}}.$$

• Since $\pi\left(heta
ight) \propto \sqrt{\det\!I\left(heta
ight)}$, we obtain that,

$$\det I\left(heta
ight) = rac{1}{2\sigma^6} \implies \pi\left(heta
ight) \propto rac{1}{\sigma^3}.$$

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You have used 2 of 3 attempts

Answers are displayed within the problem

Discussion

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