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5. Estimating the parameter of a uniform r.v.

Problem 5. Estimating the parameter of a uniform r.v.

5/5 points (graded)

The random variable X is uniformly distributed over the interval $[\theta, 2\theta]$. The parameter θ is unknown and is modeled as the value of a continuous random variable Θ , uniformly distributed between zero and one.

1. Given an observation x of X, find the posterior distribution of Θ . Express your answers below in terms of θ and x. Use 'theta" to denote θ and 'ln" to denote the natural logarithm function. For example, $\ln(\theta)$ should be entered as 'ln(theta)'.

For
$$0 \le x \le 1$$
 and $x/2 \le \theta \le x$:

2. Find the MAP estimate of Θ based on the observation X=x and assuming that $0 \le x \le 1$. Express your answer in terms of x.

For
$$0 \leq x \leq 1$$
:

3. Find the LMS estimate of Θ based on the observation X=x and assuming that $0 \le x \le 1$. Express your answer in terms of x.

For
$$0 \le x \le 1$$
:

$$\hat{\theta}_{\text{LMS}}(x) = \underbrace{\frac{1}{2} \cdot \frac{x}{\ln(x) - \ln(\frac{x}{2})}}$$
 Answer: x/(2*ln(2))

Find the linear LMS estimate $\hat{ heta}_{
m LLMS}$ of Θ based on the observation X=x. Specifically, $\hat{ heta}_{
m LLMS}$ is of the form c_1+c_2x . Find c_1 and c_2 .

Solution:

1. The prior PDF of Θ is

$$f_{\Theta}(heta) = egin{cases} 1, & ext{if } 0 \leq heta \leq 1, \ 0, & ext{otherwise}, \end{cases}$$

and the conditional PDF of the observation $oldsymbol{X}$ is

$$f_{X\mid\Theta}(x\mid heta)=\left\{egin{array}{ll} 1/ heta, & ext{if } heta\leq x\leq 2 heta,\ 0, & ext{otherwise.} \end{array}
ight.$$

Using Bayes' rule, we find that for any $x \in [0,1]$ and for $\theta \in [x/2,x]$, the posterior PDF is

$$egin{aligned} f_{\Theta|X}(heta\mid x) &= rac{f_{\Theta}(heta)f_{X|\Theta}(x\mid heta)}{\displaystyle\int_{x/2}^x f_{\Theta}(ilde{ heta})f_{X|\Theta}(x\mid ilde{ heta})d ilde{ heta}} \ &= rac{1/ heta}{\displaystyle\int_{x/2}^x rac{1}{ ilde{ heta}}d ilde{ heta}} \ &= rac{1}{ heta\cdot(\ln(x)-\ln(x/2))} \ &= rac{1}{ heta\cdot\ln(2)}. \end{aligned}$$

2. In part (1), we saw that for $x \in [0,1]$ and $x/2 \leq heta \leq x$, the posterior PDF is

$$f_{\Theta \mid X}(heta \mid x) = rac{1}{ heta \cdot \ln(2)},$$

which is decreasing in θ over the range [x/2, x] of possible values of Θ . Thus, the MAP estimate for this case is equal to x/2.

3. The LMS estimate is the conditional expectation estimate. For $x \in [0,1]$,

$$\mathbf{E}[\Theta \mid X=x] = \int_{x/2}^x heta rac{1}{ heta \cdot \ln(2)} d heta = rac{x}{2 \cdot \ln(2)}.$$

4. The LLMS estimate is of the form

$$\hat{ heta}_{LLMS}(x) = \mathbf{E}[\Theta] + rac{\mathrm{cov}(\Theta,X)}{\mathsf{Var}(X)}(x - \mathbf{E}[X]).$$

Here,

$$egin{align*} \mathbf{E}[\Theta] &= 1/2, \ \mathbf{E}[X] &= \mathbf{E}[\mathbf{E}[X \mid \Theta]] \ &= \mathbf{E}\left[rac{3}{2}\Theta
ight] \ &= rac{3}{4}, \ \mathbf{E}[X^2] &= \mathbf{E}[\mathbf{E}[X^2 \mid \Theta]] \ &= \mathbf{E}\left[rac{7}{3}\Theta^2
ight] \ &= \mathbf{E}\left[rac{7}{3}\Theta^2
ight] \ &= rac{7}{9}. \end{split}$$

Hence,

$$egin{aligned} \mathsf{Var}(X) &= \mathbf{E}[X^2] - (\mathbf{E}[X])^2 \ &= rac{31}{144}, \ \mathbf{E}[\Theta X] &= \mathbf{E}[\mathbf{E}[X\Theta \mid \Theta]] \ &= \mathbf{E}\left[rac{3}{2}\Theta^2
ight] \ &= rac{1}{2}, \ \mathsf{cov}(\Theta, X) &= \mathbf{E}[\Theta X] - \mathbf{E}[\Theta]\mathbf{E}[X] \ &= rac{1}{2} - rac{1}{2} \cdot rac{3}{4} \ &= rac{1}{8}. \end{aligned}$$

Finally, we have

$$egin{align} \hat{\Theta}_{LLMS} &= \mathbf{E}[\Theta] + rac{\mathrm{cov}(\Theta, X)}{\mathsf{Var}(X)}(x - \mathbf{E}[X]) \ &= rac{1}{2} + rac{1/8}{31/144}igg(x - rac{3}{4}igg) \ &= rac{2}{31} + rac{18}{31}x. \end{split}$$

提交

You have used 2 of 4 attempts

Answers are displayed within the problem

讨论

显示讨论

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