

Lecture 11: Fisher Information, Asymptotic Normality of MLE;

课程 □ Unit 3 Methods of Estimation □ Method of Moments

9. Properties of the Generalized ☐ Method of Moments Estimator

9. Properties of the Generalized Method of Moments Estimator

Plus Minus 1 - Method of Moments

3/3 points (graded)

Let X be a random variable that takes on values -1 and +1 with probabilities p and 1-p, respectively. Let \widehat{m}_1 be the sample average of n i.i.d. observations of X.

What is the method of moments estimator $\hat{p}_n^{ ext{MM}}$?

Use **hatm_1** for \widehat{m}_1 .

(1 - hatm_1)/2

☐ **Answer:** (1-hatm_1)/2

Assume that we observe k instances of -1 out of n outcomes. What is the ML estimator \hat{p}_n^{MLE} ?

k/n

☐ **Answer:** k/n

 \underline{k}

Are the two estimators for the ± 1 random variable equal?

● Yes □

No

STANDARD NOTATION

Solution:

The expected value of X is 1-2p.

Therefore, $\hat{p}_n^{ ext{MM}} = rac{1-\widehat{m}_1}{2}$

The ML estimator of p is $\hat{p}_n^{\mathrm{MLE}} = k/n$.

The two estimators are equal because of the following:

$$egin{array}{ll} \widehat{p}_n^{ ext{MM}} &= rac{1-\widehat{m}_1}{2} \ &= rac{1-rac{(k)\cdot -1+(n-k)\cdot 1}{n}}{2} \ &= rac{k}{n} \end{array}$$

提交

你已经尝试了1次(总共可以尝试3次)

Method of Moments - Multiple Estimators

2/2 points (graded)

Let X be a non-zero uniform random variable that we model using the distribution $\mathsf{Unif}\,[0,\theta]$, where $\{\theta\mid\theta>0\}=\Theta$. Our objective is to estimate θ using a moments estimator constructed out of n i.i.d. samples X_1,X_2,\ldots,X_n .

For a random variable $X \sim \mathsf{Unif}[0, \theta]$,

$$\mathbb{E}\left[X
ight] \ = rac{ heta}{2},$$

$$\mathbb{E}\left[X^2
ight] = rac{ heta^2}{3}.$$

We have only one parameter to estimate here, and there are two invertible moment functions that we can use to estimate the parameter. Let \widehat{m}_1 be the sample average $\frac{\sum_{i=1}^n X_i}{n}$ and let \widehat{m}_2 denote $\frac{\sum_{i=1}^n X_i^2}{n}$. By the law of large numbers, $\widehat{m}_1 \to \mathbb{E}\left[X\right]$ and $\widehat{m}_2 \to \mathbb{E}\left[X^2\right]$ as $n \to \infty$.

To enter your answers to the following, use <code>hatm_1</code> for \widehat{m}_1 , <code>hatm_2</code> for \widehat{m}_2 .

What is the method of moments estimator $\hat{ heta}_{n,1}^{ ext{MM}}$ based on \widehat{m}_1 ?

2*hatm_1

☐ **Answer:** 2*hatm_1 + 0*hatm_2

What is the method of moments estimator $\hat{ heta}_{n,2}^{ ext{MM}}$ based on \widehat{m}_2 ?

sqrt(3*hatm_2)

☐ **Answer:** sqrt(3*hatm_2) + 0*hatm_1

STANDARD NOTATION

Solution:

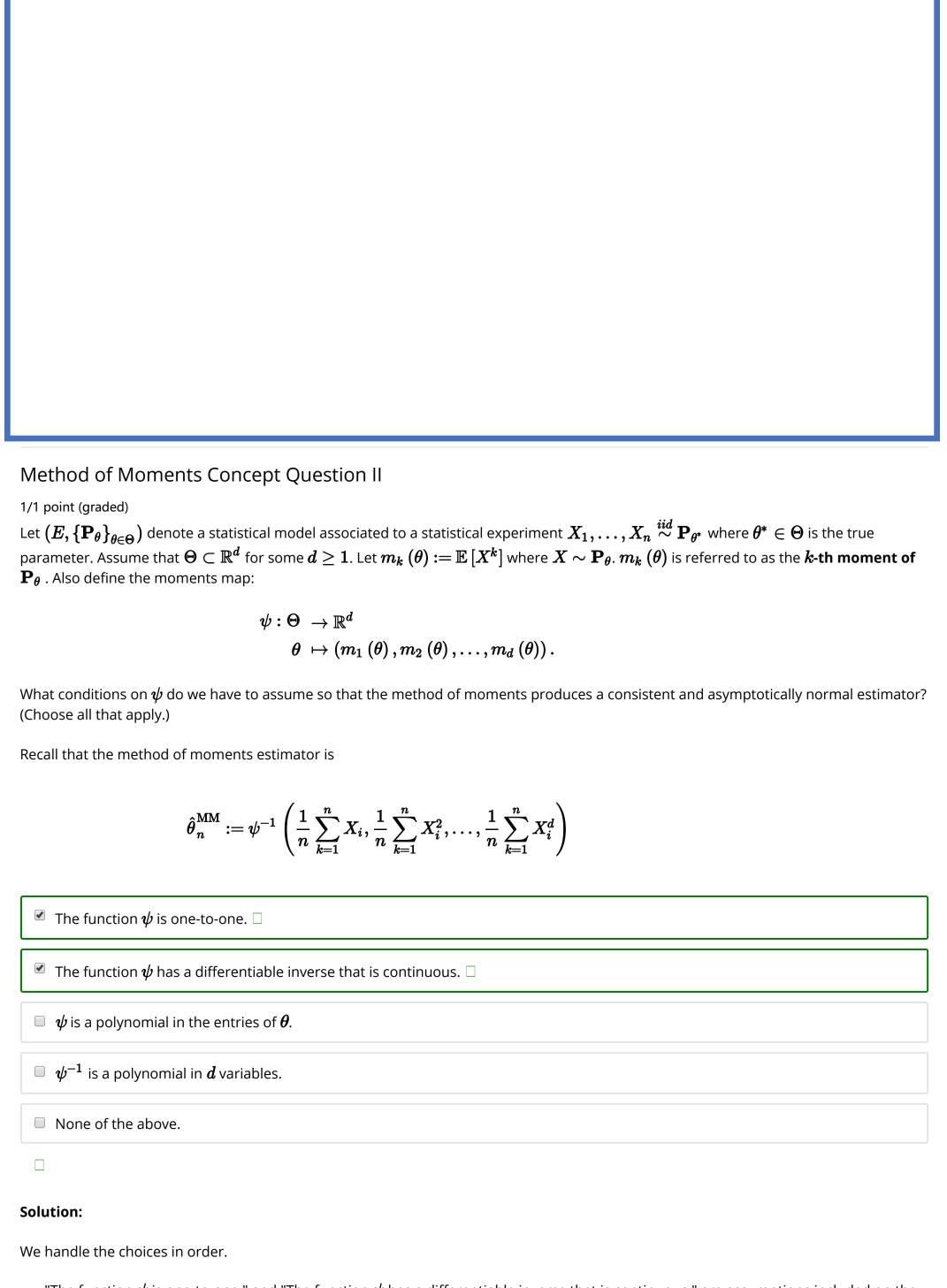
Note that both $\mathbb{E}\left[X
ight]=m_1\left(heta
ight)$ and $\mathbb{E}\left[X^2
ight]=m_2\left(heta
ight)$ are one-to-one and invertible in Θ . Therefore,

$$egin{aligned} \widehat{theta}_{n,1}^{ ext{MM}} &= 2\widehat{m}_1, \ \hat{ heta}_{n,2}^{ ext{MM}} &= \sqrt{3\widehat{m}_2}. \end{aligned}$$

提交

你已经尝试了2次(总共可以尝试3次)

☐ Answers are displayed within the problem



• "The function ψ is one-to-one." and "The function ψ has a differentiable inverse that is continuous." are assumptions included on the theorem regarding the convergence of the method of moments estimator. If ψ is not one-to-one, then we cannot even define ψ^{-1} . Also, the asymptotic covariance matrix is in terms of the inverse of the gradient of ψ , so the second assumption is certainly necessary.

• " ψ is a polynomial in the entries of θ ." and " ψ^{-1} is a polynomial in d variables." are incorrect. There are no specific assumptions needed on the form that ψ must take. However, for practical purposes, to be able to perform the method of moments, we need ψ^{-1} to be (efficiently) computable.	
提交 你已经尝试了1次(总共可以尝试3次)	
□ Answers are displayed within the problem	
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