

Lecture 10: Consistency of MLE, Covariance Matrices, and

课程 □ Unit 3 Methods of Estimation □ Multivariate Statistics

□ 8. Covariance Matrices

8. Covariance Matrices

Note: Now is a good time to review the matrix exercises in <u>Homework 0</u>.

Note on Notation: In this course, we assume all vectors to be column vectors. Therefore, while

$$\mathbf{X} = egin{bmatrix} X^{(1)} \ X^{(2)} \ dots \ Y^{(d)} \end{bmatrix},$$

we sometimes write it as $\mathbf{X} = \left(X^{(1)}, \dots, X^{(d)}\right)^T$ to be more compact in representation.

Example of Covariance II

4/4 points (graded)

Let X, Y be random variables such that

- X takes the values ± 1 each with probability 0.5
- (Conditioned on X) Y is chosen uniformly from the set $\{-3X-1, -3X, -3X+1\}$.

(Round all answers to 2 decimal places.)

What is Cov(X, X) (equivalent to Var(X))?

$$Cov(X,X) = 1$$

What is Cov(Y, Y) (equivalent to Var(Y))?

$$Cov(Y,Y) = 29/3 \qquad \qquad \Box \text{ Answer: 9.67}$$

What is Cov(X, Y)?

$$Cov(X,Y) =$$
 \Box Answer: -3.00

What is Cov(Y, X)?

$$\mathsf{Cov}\left(Y,X\right)=$$
 -3 \square Answer: -3.00

Solution:

Observe that $\mathbb{E}[X]$ and $\mathbb{E}[Y]$ are both zero, since X is uniformly distributed over $\{\pm 1\}$ and Y is uniformly distributed over the set $\{-4, -3, -2, 2, 3, 4\}.$

- $\mathsf{Cov}\left(X,X\right)$ is the variance of X, which equals $\mathbb{E}\left[X^2\right] \mathbb{E}[X]^2 = p + (1-p) = 1$.
- Cov (Y,Y) is the variance of Y, which equals $\mathbb{E}[Y^2] \mathbb{E}[Y]^2 = \frac{16+9+4+4+9+16}{6} = \frac{29}{3} \approx 9.67$.
- Cov(X,Y) and Cov(Y,X) are always equal, by the symmetry of the definition. Observe that the joint density of (X,Y) is uniform over the pairs (1, -4), (1, -3), (1, -2), (-1, 2), (-1, 3), (-1, 4). Thus, either covariance can be computed as $\mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y] = \frac{-4 - 3 - 2 - 2 - 3 - 4}{6} = -3$.

☐ Answers are displayed within the problem

Covariance Matrix

4/4 points (graded)

Given random variables $X^{(1)}, X^{(2)}, \ldots, X^{(d)}$, one can write down the **covariance matrix** Σ , where $\Sigma_{i,j} = \mathsf{Cov}(X^{(i)}, X^{(j)})$.

Let $X^{(1)}$, $X^{(2)}$ be random variables such that

- ullet $X^{(1)}$ takes the values ± 1 each with probability 0.5
- ullet (Conditioned on $X^{(1)}$) $X^{(2)}$ is chosen uniformly from the set $\left\{-3X^{(1)}-1,-3X^{(1)},-3X^{(1)}+1\right\}$.

What is the covariance matrix Σ ?

$$\Sigma_{1,1} = \boxed{1}$$

$$\square$$
 Answer: 1.0 $\Sigma_{1,2}=$ -3

$$\Sigma_{2,1} = \boxed{-3}$$

$$\square$$
 Answer: -3.00 $\Sigma_{2,2}=$ 29/3

Solution:

Using the answer to the previous question, the 2 imes 2 covariance matrix Σ evaluates to

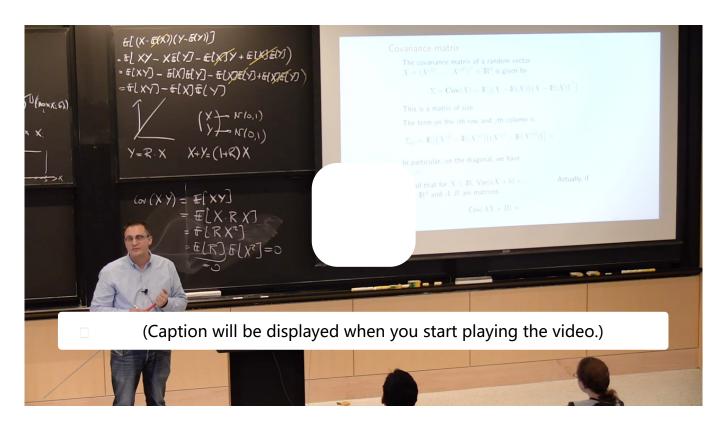
$$\Sigma = egin{pmatrix} \Sigma_{1,1} & \Sigma_{1,2} \ \Sigma_{2,1} & \Sigma_{2,2} \end{pmatrix} = egin{pmatrix} 1 & -3 \ -3 & rac{29}{3} \end{pmatrix}$$

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Covariance Matrix: Definitions



in a simple way, which is by informing the matrix that's

covariance of xx, covariance of yy, covariance of xy,

and covariance of yx.

So clearly, it's a symmetric matrix,

because these are symmetric numbers.

And so all I'm saying is that I can gather all the information about covariance and variance

in one matrix.

In this case, it's 2 by 2.

You can do that more generally, when you have, say, d by d.

If you have a vector of size d, you can talk about its covariance matrix.

And each entry on the diagonal will just be a variance term,

Here is a compact formula for the covariance matrix using vector notation.

Let
$$\mathbf{X} = \begin{pmatrix} X^{(1)} \\ \vdots \\ X^{(d)} \end{pmatrix}$$
 be a random vector of size $d imes 1$.

Let $\mu \triangleq \mathbb{E}\left[\mathbf{X}\right]$ denote the **entry-wise** mean, i.e.

$$\mathbb{E}\left[\mathbf{X}
ight] \;\;=\;\; egin{pmatrix} \mathbb{E}\left[X^{(1)}
ight] \ dots \ \mathbb{E}\left[X^{(d)}
ight] \end{pmatrix}.$$

Consider the vector outer product (refer to <u>Homework 0</u>) $(\mathbf{X} - \mu)(\mathbf{X} - \mu)^T$, which is a random $d \times d$ matrix. Then the **covariance matrix** Σ can be written as

$$\Sigma = \mathbb{E}\left[\left(\mathbf{X} - \mu
ight) \left(\mathbf{X} - \mu
ight)^T
ight].$$

Note: The following exercises will be discussed as properties of covariance in the video that follows, but we encourage you attempt these exercises before watching the video.

Covariance Matrix: Properties I

1/1 point (graded)

Let \mathbf{X} be a random vector and let $\mathbf{Y} = \mathbf{X} + \mathbf{B}$, where \mathbf{B} is a constant vector. Let $\mu_{\mathbf{X}}$ be the mean vector of \mathbf{X} and let $\Sigma_{\mathbf{X}}$ be the covariance matrix of \mathbf{X} . Select from the following all statements that are correct.

- lacksquare The covariance matrix of ${f Y}$ could potentially be equal to $\Sigma_{f X}$ only under some conditions imposed on B
- $^{f extbf{ extit{W}}}$ The covariance matrix of ${f Y}$ is the same as ${f \Sigma}_{f X}$ for all vectors ${f \it B}$ \Box
- $^{f oldsymbol{arphi}}$ The covariance matrix of ${f Y}$ has the same size as the matrix ${f \Sigma}_{f X}$ \Box
- lacksquare The covariance matrix of f Y is the same as $f \Sigma_X$ if and only if vector m B is equal to 0

Solution:

Choices 2 and 3 are correct. Let the covariance matrix of **Y** be denoted $\Sigma_{\mathbf{Y}}$. Note that $\mathbb{E}\left[\mathbf{X}+B\right]=\mu_{\mathbf{X}}+B$ for any vector B.

$$\Sigma_{\mathbf{Y}} = \mathbb{E}\left[\left(\mathbf{X} + B - \mu_{\mathbf{X}} - B
ight)\left(\mathbf{X} + B - \mu_{\mathbf{X}} - B
ight)^T
ight] = \Sigma_{\mathbf{X}}$$

Since choice 2 is correct, choices 1 and 4 that impose certain conditions on B are technically incorrect as we do not require that B satisfy some conditions for $\Sigma_{\mathbf{Y}}$ to be the same as $\Sigma_{\mathbf{X}}$.

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Covariance Matrix: Properties II

1/1 point (graded)

Let ${\bf X}$ be a random vector of size $d \times 1$ and let ${\bf Y} = A{\bf X} + B$, where ${\bf A}$ is a constant matrix of size $n \times d$ and ${\bf B}$ is a constant vector of size $n \times 1$. Let $\mu_{\bf X}$ be the mean vector of ${\bf X}$ and let $\Sigma_{\bf X}$ be the covariance matrix of ${\bf X}$. Let $\mu_{\bf Y}$ be the mean vector of ${\bf Y}$ and let $\Sigma_{\bf Y}$ be the covariance matrix of ${\bf Y}$.

Select from the following all statements that are correct.

- $extbf{ extit{ extit{\extit{\extit{\extit{ extit{ extit{ extit{ extit{ extit{ extit{ extit{\extit{ extit{ extit{ extit{\} \extit{\ext$
- $extbf{ extit{ extit{\} \extit{\ext$
- $lacksquare \Sigma_{\mathbf{Y}} = A^2 \Sigma_{\mathbf{X}}$
- $lacksquare \Sigma_{\mathbf{Y}} = A \Sigma_{\mathbf{X}} A^T \ \Box$
- $\square \ \Sigma_{\mathbf{Y}} = A^T \Sigma_{\mathbf{X}} A$

Solution:

As ${f Y}$ is an n imes 1 random vector, ${f \Sigma}_{f Y}$ is of size n imes n.

From the previous problem we know that $\Sigma_{\mathbf{Y}}$ is the same as the covariance matrix of $A\mathbf{X}$. Therefore, it suffices to find this matrix, which we denote $\Sigma_{A\mathbf{X}}$.

$$egin{aligned} \Sigma_{A\mathbf{X}} &= \mathbb{E}\left[\left(A\mathbf{X} - A\mu_X
ight)\left(A\mathbf{X} - A\mu_X
ight)^T
ight] \ &= \mathbb{E}\left[A\left(\mathbf{X} - \mu_X
ight)\left(\mathbf{X}^TA^T - \mu_X^TA^T
ight)
ight] \ &= \mathbb{E}\left[A\left(\mathbf{X} - \mu_X
ight)\left(\mathbf{X} - \mu_X
ight)^TA^T
ight] \ &= A\mathbb{E}\left[\left(\mathbf{X} - \mu_X
ight)\left(\mathbf{X} - \mu_X
ight)^T
ight]A^T \ &= A\Sigma_{\mathbf{X}}A^T. \end{aligned}$$

Therefore, choices 1, 2, and 4 are correct.

Choices 3 and 5 are not correct in general (even if A is a square matrix) because matrix multiplication is not commutative.

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Covariance Matrix: Affine Transformation



non-negative, right?

So if I give you a covariance matrix

with a negative entry on the diagonal,

you can probably just--

you can skip and say you made a mistake.

OK, now recall that for x and y, If I

look at the variance of a linear and a fine transformation of x,

so something of the form AX plus B, what is this?

a squared variance of x.

And since the covariance matrix is just

a multivariate, a matrix generalization of the variance,

it's going to have similar properties.

In particular, it should, because I

know that if I do a linear transformation, then at least

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Effect of Linear Transformations of Covariance Matrix

4/4 points (graded)

Let
$$\mathbf{X}=egin{pmatrix} X^{(1)} \ X^{(2)} \end{pmatrix}$$
 be a random vector with covariance Matrix $\Sigma_{\mathbf{X}}=egin{pmatrix} 1 & 1/2 \ 1/2 & 1 \end{pmatrix}$.

Let
$$\mathbf{Y} = M\mathbf{X}$$
, where $M = \begin{pmatrix} 1 & -1 \ 1 & 1 \end{pmatrix}$.

Observe that $Y^{(1)}=X^{(1)}-X^{(2)}$ and $Y^{(2)}=X^{(1)}+X^{(2)}$. What is the new covariance matrix Σ_Y ?

$$(\Sigma_{\mathbf{Y}})_{1,1} = \boxed{\hspace{1.5cm} 1 \hspace{1.5cm} \square \hspace{1.5cm} \mathsf{Answer:} \hspace{1.5cm} 1 \hspace{1.5cm} (\Sigma_{\mathbf{Y}})_{1,2} = \boxed{\hspace{1.5cm} 0}$$

$$\square$$
 Answer: 1 $(\Sigma_{\mathbf{Y}})_{1,2}$

$$\square$$
 Answer: 0.0

$$\left(\Sigma_{\mathbf{Y}}
ight)_{2,1}= \boxed{egin{array}{c} 0 \end{array}}$$

$$\square$$
 Answer: 0.0 $\left(\Sigma_{\mathbf{Y}}\right)_{2,2}=$ 3

Solution:

Recall from an earlier problem that for any pair of random variables $A,\,B$ with the same variance $\mathsf{Var}\,(A) = \mathsf{Var}\,(B) = \sigma^2$, $\operatorname{\mathsf{Cov}}\left(A-B,A+B
ight) = \operatorname{\mathsf{Var}}\left(A
ight) - \operatorname{\mathsf{Var}}\left(B
ight) = 0.$

Therefore, given the matrix M, $\Sigma_{\mathbf{Y}}$ must be a diagonal matrix.

$$\mathsf{Cov}\,(Y^{(1)},Y^{(1)}) = \mathsf{Cov}\,(X^{(1)}-X^{(2)},X^{(1)}-X^{(2)}) = \mathsf{Cov}\,(X^{(1)},X^{(1)}) - 2\mathsf{Cov}\,(X^{(1)},X^{(2)}) + \mathsf{Cov}\,(X^{(2)},X^{(2)}) = 1 - 1 + 1 = 1.$$

Similarly,

$$\mathsf{Cov}\,(Y^{(2)},Y^{(2)}) = \mathsf{Cov}\,(X^{(1)} + X^{(2)},X^{(1)} + X^{(2)}) = \mathsf{Cov}\,(X^{(1)},X^{(1)}) + 2\mathsf{Cov}\,(X^{(1)},X^{(2)}) + \mathsf{Cov}\,(X^{(2)},X^{(2)}) = 1 + 1 + 1 = 3.$$

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☐ Answers are displayed within the problem

讨论

显示讨论

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