

## 15. Likelihood of a Poisson Statistical Model

### Review: Statistical Model for a Poisson Distribution

2/2 points (graded)

Let  $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Poiss}(\lambda^*)$  for some unknown  $\lambda^* \in (0, \infty)$ . Let  $(E, \{\text{Poiss}(\lambda)\}_{\lambda \in \Theta})$  denote the corresponding statistical model. What is the smallest possible set that could be  $E$ ?

☐  $\mathbb{N} = \{1, 2, 3, \dots\}$

☒  $\mathbb{N} \cup \{0\}$  ☐

☐  $\mathbb{Z}$

☐  $\mathbb{R}$

The parameter space  $\Theta$  can be written as an interval  $(a, \infty)$ . What is the smallest value of  $a$  so that  $\{\text{Poiss}(\lambda)\}_{\lambda \in (a, \infty)}$  represents all possible Poisson distributions?

$a =$   ☐ Answer: 0.0

#### Solution:

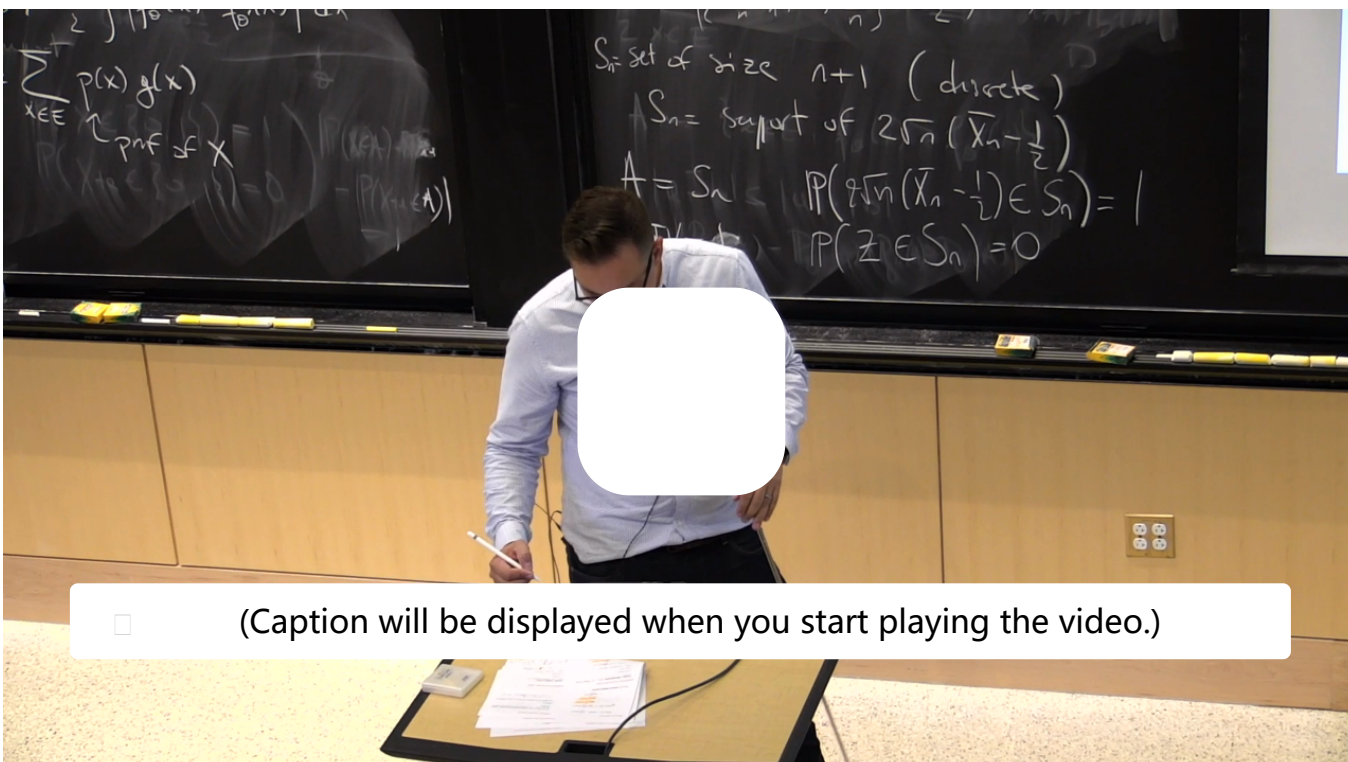
A Poisson random variable takes values on all non-negative integers  $\{0, 1, 2, \dots\}$ . Hence, the smallest possible sample space is  $\mathbb{N} \cup \{0\}$ .

A Poisson random variable is specified by its mean  $\lambda$ , which is allowed to be any positive real number. Hence,  $a = 0$  is the correct choice.

提交 你已经尝试了2次 (总共可以尝试2次)

☐ Answers are displayed within the problem

## Likelihood of a Poisson Statistical Model



I'm going to have to write-- so even let's do it for  $X_i$ .

I am going to have to take the product of those guys.

Well, it's  $\lambda$  to the  $X_i$  over  $X_i$  factorial,  $e$  to the minus  $\lambda$ .

That's just my PMF for a Poisson.

Now, if I take the product of those guys-- so my likelihood  $X_1 X_n \lambda$  is just the product.

So it's  $\lambda$  to the sum of the  $X_i$ s, the product

of my denominators,  $X_i$  factorial, and then  $e$  to the minus, and then I have a product of those guys  $n$  times.

So I get  $e$  to the minus  $n$  times  $\lambda$ .

And now if you rearrange the terms, this is exactly what's written here,

Practice: Compute Likelihood of a Poisson Statistical Model

3/3 points (graded)

Let  $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Poiss}(\lambda^*)$  for some unknown  $\lambda^* \in (0, \infty)$ . You construct the associated statistical model  $(E, \{\text{Poiss}(\lambda)\}_{\lambda \in \Theta})$  where  $E$  and  $\Theta$  are defined as in the answers to the previous question.

Suppose you observe two samples  $X_1 = 1, X_2 = 2$ . What is  $L_2(1, 2, \lambda)$ ? Express your answer in terms of  $\lambda$ .

$L_2(1, 2, \lambda) =$

exp(-2\*lambda)\*lambda^(1+2)/(1\*2)

$$\frac{\exp(-2 \cdot \lambda) \cdot \lambda^{1+2}}{1 \cdot 2}$$

Answer: e^(-2\*lambda)\*lambda^3/2

Next, you observe a third sample  $X_3 = 3$  that follows  $X_1 = 1$  and  $X_2 = 2$ . What is  $L_3(1, 2, 3, \lambda)$ ?

$L_3(1, 2, 3, \lambda) =$

exp(-3\*lambda)\*lambda^(1+2+3)/(1\*2\*3\*2)

$$\frac{\exp(-3 \cdot \lambda) \cdot \lambda^{1+2+3}}{1 \cdot 2 \cdot 3 \cdot 2}$$

Answer: e^(-3\*lambda)\*lambda^6/12

Suppose your data arrives in a different order:  $X_1 = 2, X_2 = 3, X_3 = 1$ . What is  $L_3(2, 3, 1, \lambda)$ ?

$L_3(2, 3, 1, \lambda) =$

exp(-3\*lambda)\*lambda^(1+2+3)/(1\*2\*3\*2)

$$\frac{\exp(-3 \cdot \lambda) \cdot \lambda^{1+2+3}}{1 \cdot 2 \cdot 3 \cdot 2}$$

Answer: e^(-3\*lambda)\*lambda^6/12

STANDARD NOTATION

Solution:

The probability mass function of  $\text{Poiss}(\lambda)$  is  $x \mapsto e^{-\lambda} \frac{\lambda^x}{x!}$  where  $x \in \mathbb{N} \cup \{0\}$ . Hence by definition

$$L_n(x_1, \dots, x_n, \lambda) = \prod_{i=1}^n e^{-\lambda} \frac{\lambda^{x_i}}{x_i!} = e^{-n\lambda} \frac{\lambda^{\sum_{i=1}^n x_i}}{x_1! \cdots x_n!}.$$

Hence, first we plug in  $n = 2, x_1 = 1$ , and  $x_2 = 2$ :

$$L_2(1, 2, \lambda) = e^{-2\lambda} \frac{\lambda^{1+2}}{2!1!} = e^{-2\lambda} \frac{\lambda^3}{2}.$$

When the next sample arrives, we can simply evaluate the density of a Poisson at the observation:

$$P(X_3 = 3) = e^{-\lambda} \frac{\lambda^3}{3!}, \quad X \sim \text{Poiss}(\lambda)$$

and multiply this by the previous response:

$$L_3(1, 2, 3, \lambda) = e^{-\lambda} \frac{\lambda^3}{3!} L_2(1, 2, \lambda) = e^{-3\lambda} \frac{\lambda^6}{12}.$$

**Remark 1:** Observe that we can compute the likelihood sequentially as the data arrives, updating it in the previous fashion after each new observation.

Similarly, we see that

$$L_3\left(2,3,1,\lambda\right)=e^{-3\lambda}\frac{\lambda^6}{12}.$$

**Remark 2:** Observe that the likelihood of observations  $X_1 = x_1, \dots, X_n = x_n$  is independent of the *order* in which these observations arrive.

提交

你已经尝试了1次（总共可以尝试3次）

☐ Answers are displayed within the problem

## Properties of the Likelihood

1/1 point (graded)

Let  $(E, \{P_\theta\}_{\theta \in \Theta})$  denote a discrete statistical model. Let  $p_\theta$  denote the pmf of  $P_\theta$ . Let  $X_1, \dots, X_n \stackrel{iid}{\sim} P_{\theta^*}$  where the parameter  $\theta^*$  is unknown. Then the **likelihood** is the function

$$L_n : E^n \times \Theta \rightarrow \mathbb{R} \\ (x_1, \dots, x_n, \theta) \mapsto \prod_{i=1}^n p_\theta(x_i).$$

For our purposes, we think of  $x_1, \dots, x_n$  as observations of the random variables  $X_1, \dots, X_n$ .

Which of the following are properties of the likelihood  $L_n$ ? (Choose all that apply.)

**Hint:** It may be useful to consider your responses from the previous question.

☐ The likelihood does not change with the parameter  $\theta$ .

☒ The likelihood can be updated sequentially as new samples are observed. For example,  $L_3(x_1, x_2, x_3, \theta) = L_1(x_3, \theta) L_2(x_1, x_2, \theta)$ . ☐

☒ The likelihood is symmetric: it doesn't matter the order in which we plug in the observations. For example,  $L_4(x_1, x_2, x_3, x_4, \theta) = L_4(x_2, x_3, x_1, x_4, \theta)$ , and this is true for any rearrangement of  $x_1, x_2, x_3, x_4$ . ☐

☐ If we eliminate a single observation, then the likelihood remains unchanged. For example,  $L_3(x_1, x_2, x_3, \theta) = L_2(x_1, x_2, \theta)$ .

☐

### Solution:

We examine the choices in order.

- "The likelihood does not change with the parameter  $\theta$ ." is incorrect. Rather, it is crucial that we interpret the likelihood  $L_n$  as a function of  $\theta$ . That is,  $L_n$  varies as  $\theta$  ranges over the parameter space  $\Theta$ . This is evident in the likelihoods for the Bernoulli and Poisson models in the previous problems.
- "The likelihood can be updated sequentially as new samples are observed. For example,  $L_3(x_1, x_2, x_3, \theta) = L_1(x_3, \theta) L_2(x_1, x_2, \theta)$ ." is also correct. In the previous problem, we saw that to compute the likelihood after observing  $X_3 = 3$ , we simply took the old likelihood  $L_2(1, 2, \lambda)$  and multiplied it by  $L_1(3, \lambda)$ . Note that  $L_1(x_3, \theta) = p_\theta(x_3)$ , the density of  $P_\theta$  evaluated at the new observation. Inspection of the defining formula

$$L_n(x_1, \dots, x_n, \theta) = \prod_{i=1}^n p_\theta(x_i)$$

implies that the likelihood can be updated sequentially in this fashion.

- "The likelihood is symmetric..." is correct. We observed in the previous problem that observing the samples in a different order does not affect the likelihood. This is also evident from the definition of the likelihood: we can take the product

$$\prod_{i=1}^n p_{\theta}(x_i)$$

in any order, and the result will still be the same.

- "If we eliminate a single observation, then the likelihood remains unchanged..." is incorrect. In the previous question, we saw that for a Poisson statistical model,  $L_2(\mathbf{1}, \mathbf{2}, \lambda)$  and  $L_3(\mathbf{1}, \mathbf{2}, \mathbf{3}, \lambda)$  do not have the same formula. Hence, deleting an observation from the sample will change the likelihood.

提交

你已经尝试了1次（总共可以尝试2次）

☐ Answers are displayed within the problem

讨论

显示讨论

主题： Unit 3 Methods of Estimation:Lecture 8: Distance measures between distributions / 15.  
Likelihood of a Poisson Statistical Model