

For

α
>
0

{\displaystyle \alpha >0,}

 and

β
>
0,

{\displaystyle \beta >0,}

E

[ln
⁡
(
X
)
]
=
ln
⁡
(
β
)
−
ψ
(
α
)

{\displaystyle \mathbb {E} [\ln(X)]=\ln(\beta)-\psi (\alpha)}

and

E

[

X

−
1

]
=

α
β

,

{\displaystyle \mathbb {E} [X^{-1}]={\frac {\alpha }{\beta }},}

The information entropy is

H
(
X
)
=

E

[
−
ln
⁡
(
p
(
X
)
)
]

=

E

[
−
α
ln
⁡
(
β
)
+
ln
⁡
(
Γ
(
α
)
)
+
(
α
+
1
)
ln
⁡
(
X
)
+

β
X

]

=
−
α
ln
⁡
(
β
)
+
ln
⁡
(
Γ
(
α
)
)
+
(
α
+
1
)
ln
⁡
(
β
)
−
(
α
+
1
)
ψ
(
α
)
+
α

=
α
+
ln
⁡
(
β
Γ
(
α
)
)
−
(
α
+
1
)
ψ
(
α
)
.

{\displaystyle {\begin{aligned}H(X)&=\mathbb {E} [-\ln(p(X))]\\&=\mathbb {E} \left[-\alpha \ln(\beta)+\ln(\Gamma (\alpha))+(\alpha +1)\ln(X)+{\frac {\beta }{X}}\right]\\&=-\alpha \ln(\beta)+\ln(\Gamma (\alpha))+(\alpha +1)\ln(\beta)-(\alpha +1)\psi (\alpha)+\alpha \\&=\alpha +\ln(\beta \Gamma (\alpha))-(\alpha +1)\psi (\alpha).\end{aligned}}

where

ψ
(
α
)

{\displaystyle \psi (\alpha)}

 is the digamma function.

The Kullback-Leibler divergence of Inverse-Gamma(

α

p

,

β

p

) from Inverse-Gamma(

α

q

,

β

q

) is the same as the KL-divergence of Gamma(

α

p

,

β

p

) from Gamma(

α

q

,

β

q

):

D

KL

(

α

p

,

β

p

;

α

q

,

β

q

)
=

E

[
log
⁡

ρ
(
X
)

π
(
X
)

]
=

E

[
log
⁡

ρ
(
1

/

Y
)

π
(
1

/

Y
)

]
=

E

[
log
⁡

ρ

G

(
Y
)

π

G

(
Y
)

]
,

{\displaystyle D_{\mathrm {KL} }(\alpha _{p},\beta _{p};\alpha _{q},\beta _{q})=\mathbb {E} \left[\log {\frac {\rho (X)}{\pi (X)}}\right]=\mathbb {E} \left[\log {\frac {\rho (1/Y)}{\pi (1/Y)}}\right]=\mathbb {E} \left[\log {\frac {\rho _{G}(Y)}{\pi _{G}(Y)}}\right],}

where

ρ
,
π

{\displaystyle \rho ,\pi }

 are the pdfs of the Inverse-Gamma distributions and

ρ

G

,
π

G

{\displaystyle \rho _{G},\pi _{G}}

 are the pdfs of the Gamma distributions,

Y

{\displaystyle Y}

 is Gamma(

α

p

,

β

p

) distributed.

Related distributions

- If

X
∼

Inv-Gamma

(
α
,
β
)

{\displaystyle X\sim {\mathrm {Inv-Gamma} }(\alpha ,\beta)}

 then

k
X
∼

Inv-Gamma

(
α
,
k
β
)

{\displaystyle kX\sim {\mathrm {Inv-Gamma} }(\alpha ,k\beta)}
- If

X
∼

Inv-Gamma

(
α
,

1
2

)

{\displaystyle X\sim {\mathrm {Inv-Gamma} }(\alpha ,{\tfrac {1}{2}})}

 then

X
∼

Inv-χ

2

(
2
α
)

{\displaystyle X\sim {\mathrm {Inv-\chi ^{2}} }(2\alpha)}

 (inverse-chi-squared distribution)
- If

X
∼

Inv-Gamma

(

α
2

,

1
2

)

{\displaystyle X\sim {\mathrm {Inv-Gamma} }({\tfrac {\alpha }{2}},{\tfrac {1}{2}})}

 then

X
∼

Scaled Inv-χ

2

(
α
,

1
α

)

{\displaystyle X\sim {\mathrm {Scaled Inv-\chi ^{2}} }(\alpha ,{\tfrac {1}{\alpha }})}

 (scaled-inverse-chi-squared distribution)
- If

X
∼

Inv-Gamma

(

1
2

,

c
2

)

{\displaystyle X\sim {\mathrm {Inv-Gamma} }({\tfrac {1}{2}},{\tfrac {c}{2}})}

 then

X
∼

Levy

(
0
,
c
)

{\displaystyle X\sim {\mathrm {Levy} }(0,c)}

 (Lévy distribution)
- If

X
∼

Gamma

(
α
,
β
)

{\displaystyle X\sim {\mathrm {Gamma} }(\alpha ,\beta)}

 (Gamma distribution) then

1
X

∼

Inv-Gamma

(
α
,
β
)

{\displaystyle {\tfrac {1}{X}}\sim {\mathrm {Inv-Gamma} }(\alpha ,\beta)}

 (see derivation in the next paragraph for details)
- Inverse gamma distribution is a special case of type 5 Pearson distribution
- A multivariate generalization of the inverse-gamma distribution is the inverse-Wishart distribution.
- For the distribution of a sum of independent inverted Gamma variables see Witkovsky (2001)

Derivation from Gamma distribution

Let

X
∼

Gamma

(
α
,
β
)

{\displaystyle X\sim {\mathrm {Gamma} }(\alpha ,\beta)}

, and recall that the pdf of the gamma distribution is

f
(
x
)
=

β

α

Γ
(
α
)

x

α
−
1

e

−
β
x

.

{\displaystyle f(x)={\frac {\beta ^{\alpha }}{\Gamma (\alpha)}}x^{\alpha -1}e^{-\beta x}.}

Define the transformation

Y
=
g
(
X
)
=

1
X

.

{\displaystyle Y=g(X)={\tfrac {1}{X}}.}

 Then, the pdf of

Y

{\displaystyle Y}

 is

f

Y

(
y
)
=

f

X

(

g

−
1

(
y
)
)

|

d

g

−
1

(
y
)

d
y

|

=

β

α

Γ
(
α
)

(

1
y

)

α
−
1

exp
⁡
(

−
β
y

)

1

y

2

=

β

α

Γ
(
α
)

(

1
y

)

α
+
1

exp
⁡
(

−
β
y

)

=

β

α

Γ
(
α
)

(
y

)

−
α
−
1

exp
⁡
(

−
β
y

)

{\displaystyle {\begin{aligned}f_{Y}(y)&=f_{X}\left(g^{-1}(y)\right)\left|{\frac {d}{dy}}g^{-1}(y)\right|\\&={\frac {\beta ^{\alpha }}{\Gamma (\alpha)}}\left({\frac {1}{y}}\right)^{\alpha -1}\exp \left({\frac {-\beta }{y}}\right){\frac {1}{y^{2}}}\\&={\frac {\beta ^{\alpha }}{\Gamma (\alpha)}}\left({\frac {1}{y}}\right)^{\alpha +1}\exp \left({\frac {-\beta }{y}}\right)\\&={\frac {\beta ^{\alpha }}{\Gamma (\alpha)}}(y)^{-\alpha -1}\exp \left({\frac {-\beta }{y}}\right)\end{aligned}}

Occurrence

See also

- gamma distribution
- inverse-chi-squared distribution
- normal distribution

References

- "InverseGammaDistribution—Wolfram Language Documentation" (http://reference.wolfram.com/language/ref/InverseGammaDistribution.html). *reference.wolfram.com*. Retrieved 9 April 2018.
- John D. Cook (Oct 3, 2008). "InverseGammaDistribution" (https://www.johndcook.com/inverse_gamma.pdf) (PDF). Retrieved 3 Dec 2018.

- Witkovsky, V. (2001). "Computing the Distribution of a Linear Combination of Inverted Gamma Variables". *Kybernetika*. **37** (1): 79–90. MR 1825758 (<https://www.ams.org/mat-hscinet-getitem?mr=1825758>). Zbl 1263.62022 (<https://zbmath.org/?format=complete&q=an:1263.62022>).

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