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1 Intro

- My OH is X.
- Email is first.last@
- X year cs + math major
- hobbies?

1.1 Some CS70 advice

- Goal: enhance problem solving techniques/approach
- Don't fall behind on content, catching up will not be fun
- problems, problems, more problems
- Ask lots of questions (imperative for strong foundation)
- Don't stress, we're in this ride together

1.2 Propositional Logic

Relevant notation:

- \wedge = and
- \(\text{\text{or}} = \text{or} \)
- ¬ = not
- \implies = implies
- \exists = there exists
- \forall = forall
- \mathbb{N} = natural numbers $\{0, 1, \ldots\}$
- a|b = a divides b

 $P \Longrightarrow Q$ is an example of an implication. We can read this as "If P, then Q." An implication is false only when P is true and Q is false. If P is false, the implication is vacuously true.

Definition 1.1 (Contrapositive)

If $P \implies Q$ is an implication, then the implication $\neg Q \implies \neg P$ is known as the **contrapositve**.

An important identity is that $P \implies Q \equiv \neg Q \implies \neg P$.

1.3 Proofs

Induction will be in its own section.

Different methods.

1.3.1 Direct proof

Want to show $P \implies Q$ by assuming P and logically concluding Q.

1.3.2 Contraposition

Want to show $P \implies Q$ by equivalently proving $\neg Q \implies \neg P$.

1.3.3 Contradiction

Want to show *P*. We do this by assuming $\neg P$ and concluding $R \wedge \neg R$.

Why? Idea is that if we can show the implication $\neg P \implies (R \land \neg R)$ is True, this is the same as showing $\neg P \implies F$ is True. The contraposition gives $T \implies P$.

1.3.4 Cases

Break up a problem into multiple cases i.e. odd vs even.

2 Induction

Goal of induction is to show $\forall nP(n)$.

2.1 (Weak) Induction

- Prove P(0) is true (or relevant base cases), then $\forall n \in \mathbb{N} (P(n) \implies P(n+1))$.
- Induction dominoes analogy!
- Sometimes you might have multiple base cases (Problem about 4x + 5y in Notes 3)

2.2 Strengthening the Hypothesis

Sometimes proving $P(n) \implies P(n+1)$ is not straightforward with induction. In such a scenario, we can try to introduce a (stronger) statement Q(n). We want to construct Q such that $Q(n) \implies P(n)$. Inducting on Q proves P.

2.3 Strong Induction

- Prove P(0) is true (or relevant base cases), then $\forall n ((P(0) \land P(1) \land \cdots \land P(n)) \implies P(n+1))$.
- Dominoes analogy, but emphasis on the difference between weak and strong induction (assuming middle domino works vs everything from start to middle).

2.4 Weak vs Strong

A common point of confusion is when one should use strong induction in lieu of weak induction. Strong induction **always** works whenever weak induction works. However, there may be scenarios in which the induction hypothesis to prove n = k + 1 requires more information than just n = k. A scenario like this requires strong induction.

3 Stable Matching

Cool application of induction.

3.1 The Propose and Reject Algorithm

Suppose jobs proposes to candidates.

- both jobs and candidates have a list of preferences
- every day a job that doesn't have a deal with a candidate will propose to the next best candidate on its preference list
- every candidate will tentatively "waitlist" the offer from the job (put it on a string)
- if a candidate has multiple offers, they will choose the one they prefer the most
- the algorithm ends when every candidate has a job on their "waitlist" (all these WLs becomes acceptances)

(walk through q1 of dis as a class to visualize this)

3.2 Stability

Definition 3.1 (Rogue Couple)

A job-candidate pair (J,C) is denoted as a **rogue couple** if they prefer each other over their final assignment in a stable matching instance.

Definition 3.2 (Unstable)

A matching that has at least one rogue couple is considered unstable.

Conversely, a **stable** matching is one that has no rogue couples.

Some tricky vocab stuff like stable matching instance.

Lemma 1 (Improvement) *If a candidate has a job offer, then they will always have an offer from a job at least as good as the one they have right now.*

Matchings produced by the algorithm are always stable.

3.3 Optimality

The propose and reject algorithm is proposer optimal and receiver pessimal.

Definition 3.3 (optimal)

A pairing is optimal for a group if each entity is paired with who it most prefers while maintaining stability.

Can be thought of a (well that's the best I could do) analogy.

Definition 3.4 (pessimal)

A pairing is pessimal for a group if each entity is paired with who it least prefers while maintaining stability.

Can be thought of a (well it can't get worse than this) analogy.

3.4 Potpourri

It is possible that there exists a stable matching instance that is neither job optimal nor candidate pessimal. Consider the following preferences

Jobs	Preferences	
A	1 > 2 > 3	
В	2 > 3 > 1	
C	3 > 1 > 2	

Candidates	Preferences
1	B > C > A
2	C > A > B
3	A > B > C

The matching above can generate (at least) 3 stable matching instances

$$S = \{(A,1), (B,2), (C,3)\}$$

$$T = \{(A,3), (B,1), (C,2)\}$$

$$U = \{(A,2), (B,3), (C,1)\}.$$

We see

- *S* is job-optimal/candidate-pessimal (result of running propose and reject with jobs proposing to candidates)
- T is candidate-optimal/job-pessimal (result of running propose and reject with candidates proposing to jobs)
- U is neither optimal nor pessimal for both candidates and jobs (S and T) corroborate that.

Also some other important facts that can be seen (from discussion worksheet questions):

- There is at least one candidate that will receive only one proposal (that too on the last day)
- We can upper bound the number of days needed by P&R algorithm to $(n-1)^2 + 1 = n^2 2n + 2$ (think about why)
- As a consequence of above, we can upper bound the number of rejections needed by P&R algorithm to $(n-1)^2 = n^2 2n + 1$ rejections.

4 Graphs

4.1 Notation

- V denotes set of vertices (points)
- E denotes set of edges (lines)
- |V| denotes size of set of vertices i.e number of vertices; |E| similarly
- Graph G with vertices V and edges E is denoted G = (V, E).

4.2 Vocabulary

Definition 4.1 (Path)

A path is a sequence of edges. In CS70, we assume a path is *simple* which means no repeated vertices.

Definition 4.2 (Cycle)

A **cycle** is a simple path that starts and ends at the same vertex.

Definition 4.3 (Walk)

A walk is any arbitrary connected sequence of edges.

Definition 4.4 (Tour)

A **tour** is a walk that starts and end at the same vertex.

Definition 4.5 (Connected)

A graph is **connected** if there exists a path between any two distinct vertices.

Definition 4.6 (Eulerian Walk)

An Eulerian walk is a walk covering all edges without repeating any.

Definition 4.7 (Eulerian Tour)

An Eulerian tour is an Eulerian walk that starts and ends at the same vertex.

To summarize,

	no repeated vertices	no repeated edges	start = end	all edges	all vertices
Walk					
Path	\checkmark	✓			
Tour			✓		
Cycle	√ *	✓	✓		
Eulerian Walk		✓		✓	
Eulerian Tour		✓	✓	√	
Hamiltonian Tour	✓	✓	✓		✓

(*except for start and end vertices)

Theorem 4.1 (Euler's Theorem)

An undirected graph G has an Eulerian tour iff G is connected and all its vertices have even degree.

The requires condition for an Eulerian walk is that we have exactly 2 vertices of odd degree. (Of course, the case of 0 odd vertices trivially works since we claim from Euler's Theorem that we can find an Eulerian tour which is a stronger statement than an Eulerian walk)

Definition 4.8 (Bipartite)

A graph is considered bipartite if V can be partitioned into two sets L and R where $V = L \cup R$ such that there are no edges between vertices in L and no edges between vertices in R.

4.3 The holy grail for graph proofs

Induct, induct, induct, and induct.

- Think about what you want to induct on (edges or vertices???)
- Base case (read the problem carefully!)
- Prove for *n* by going from $n \to n-1 \to I.H. \to n$.
 - **DO NOT** go from n 1 → n directly.
 - Why? Build-up error!
 - Good example of build-up error when trying to prove "if every vertex of a graph has degree at least 2, then there exists a cycle of length 3." Any attempt at induction will give us a false proof but we cannot make square from triangle!
 - It's also a logistical nightmare lol (in the times it might accidentally work). Try generating all 5-vertex trees from all 4-vertex trees yikes.

4.4 Relevant Potpourri

Some other relevant information.

Definition 4.9 (Degree)

The **degree** of a vertex v denoted deg(v) is defined to be the number of incident edges to v.

Lemma 2 (Handshake)

$$\sum_{v \in V} \deg(v) = 2|E|.$$

The idea of a degree (with no adjective) is only well-defined for undirected graphs. We see for directed graphs it's a little funky; we need to introduce the concept of indegree and outdegree.

In a directed graph, the number of outgoing edges equals the number of ingoing edges.

We will discuss trees, planarity, coloring, and hypercubes in the next discussion.

5 More Graphs

5.1 Trees

A graph G = (V, E) is a Tree if any of the statements below is true. TFAE (The following are equivalent):

- G is connected and has no cycles
- G is connected and |E| = |V| 1
- G is connected and removing a single edge disconnects G
- G has no cycles and adding a single edge creates a cycle

Definition 5.1

A leaf is a node of degree 1.

A consequence of above is that every tree has at least 2 leaves.

5.2 Planarity

Definition 5.2 (planar)

A graph is **planar** if it can be drawn without any edge crossings.

Theorem 5.1 (Euler)

For every connected planar graph, f + v = e + 2.

Corollary 1 *If a graph is planar, then* $e \le 3v - 6$.

Theorem 5.2 (Kuratowski)

A graph is non-planar iff it contains K_5 or $K_{3,3}$.

(draw the two above graphs on the board)

The notation K_x denotes a complete graph with x vertices.

Definition 5.3 (complete graph)

A **complete graph** is a graph where all possible edges exist. Formally, in graph G = (V, E), for any distinct $u, v \in V$, then $\{u, v\} \in E$.

5.3 Coloring

Two types: edge and vertex

- edge: color edges so that no two adjacent edges have the same color
- vertices: color vertices so that no two adjacent vertices have the same color

Theorem 5.3 (4 color theorem)

If a graph is planar, then it can be colored with 4 (or less) colors.

5.4 Hypercubes

A hypercube of dimension n is a graph whose vertices are bitstrings of length n. An edge between two vertices exists iff the two vertices differ at exactly 1 bit.

(draw n = 1, 2, 3 on the board)

We can see that $|V| = 2^n$ and $|E| = n2^{n-1}$.

Give some motivation on induction on hypercubes.

6 Modular Arithmetic

The relevant notation we'll be using for this section is expressions of the form

$$a \equiv b \pmod{x}$$

reads "a is equivalent to $b \mod x$ ". It means that the remainder of a when divided by x equals the remainder of b when divided by x.

An important identity is that

$$a \equiv b \pmod{x} \iff (\exists k \in \mathbb{Z})(a = b + kx).$$

Talk about the "clock analogy".

Example 6.1

We can see a display of some of the properties:

- Addition: $7 + 4 \equiv 1 \pmod{5}$
- Subtraction: $7 4 \equiv 1 \pmod{2}$
- Multiplication: $2 \cdot 3 \equiv 0 \pmod{6}$.
- Division??

In modular arithmetic, division is not well-defined. The opposite of multiplication is multiplying by the modular inverse.

Definition 6.1 (modular inverse)

The value a is the **modular inverse** of x with respect to mod m if

$$ax \equiv 1 \pmod{m}$$
.

Does an inverse always exist? No.

Theorem 6.1

Let x and m be positive integers. Then $x^{-1} \pmod{m}$ exists and is unique only if gcd(x,m) = 1.

Definition 6.2 (Greatest Common Divisor)

The **greatest common divisor** (gcd) of two integers a, b is the greatest $d \in \mathbb{Z}$ such that d|a and d|b.

How does one efficiently calculate the GCD?

Algorithm 6.1 (Euclidean Algorithm)

```
function GCD(a,b)

if b = 0 then

return a

return GCD(b, a \mod b)
```

7 More Modular Arithmetic

7.1 Modular inverse

Lemma 3 (Bézout) For integers x, y such that gcd(x,y) = d, there exist integers a and b that obey

$$ax + by = d$$
.

We care about the case when gcd(x,y) = d = 1.

Why? This is how we can find the modular inverse.

If ax + by = 1, taking mod x gives us

$$by \equiv 1 \pmod{x} \implies b \equiv y^{-1} \pmod{x}$$
.

Similarly, taking mod y gives us

$$ax \equiv 1 \pmod{y} \implies a \equiv x^{-1} \pmod{y}.$$

Takeaway: the values of a and b we will solve for (Q1 on discussion) give us the inverse of x with respect to y and vice versa.

7.2 Chinese Remainder Theorem (CRT)

Theorem 7.1 (CRT)

For pairwise relatively prime integers m_1, m_2, \dots, m_n , the modular system

$$x \equiv a_1 \pmod{m_1}$$

 $x \equiv a_2 \pmod{m_2}$
 \vdots
 $x \equiv a_n \pmod{m_n}$

has a unique solution $x \pmod{m_1 m_2 \cdots m_n}$.

To clarify, the term pairwise relatively prime means for any distinct i, j, it follows $gcd(m_i, m_j) = 1$.

How do we solve the system above? Discussion Q2...

...or we can solve them a faster way (not taught in the course lol)

Example 7.1

Suppose we take the first two systems from Q2 on discussion.

$$x \equiv 1 \pmod{3}$$

 $x \equiv 3 \pmod{7}$.

Since gcd(3,7) = 1, CRT tells us x has a unique solution mod 21. The first equation tells us there exists some integer k such that x = 1 + 3k. Plugging this into the second equation we have

$$1 + 3k \equiv 3 \pmod{7} \implies k \equiv 3 \pmod{7}$$
.

Plugging in k = 3 gives $x \equiv 10 \pmod{21}$.

If we wanted to solve entirety of Q2 this way, we then apply the same trick above to the systems

$$x \equiv 10 \pmod{21}$$

 $x \equiv 4 \pmod{11}$.

8 FLT, RSA

8.1 Fermat's Little Theorem

A relevant theorem in modular arithmetic that will help us with RSA is Fermat's Little Theorem (FLT).

Theorem 8.1 (Fermat's Little Theorem (FLT))

For prime p and $a \in \{1, 2, ..., p - 1\}$, it follows

$$a^p \equiv a \pmod{p}$$
.

Special case, if a is not divisible by p, then

$$a^{p-1} \equiv 1 \pmod{p}$$
.

8.2 RSA

Objective: Alice transfers info to Bob without Eve cracking it.

8.2.1 The algorithm

Here's a detailed outline of how the scheme works for RSA with 2 primes:

- 1. Entire world knows about a public key (N,e) where N=pq for primes p and q such that gcd(e,(p-1)(q-1))=1.
- 2. Alice and Bob meet in private, and Alice tells Bob what p and q are.
- 3. On his own time, Bob computes (p-1)(q-1) and then calculates

$$d = e^{-1} \pmod{(p-1)(q-1)}$$
.

(Think about why we know such a d must exist)

4. To encrypt her message x, Alice sends E(x) to Bob where

$$E(x) = x^e \pmod{N}$$
.

5. To decrypt the message received y, Bob calculate D(y) where

$$D(y) = y^d \pmod{N}$$
.

High level idea of why this works:

$$D(E(x)) = D(x^{e}) \pmod{N}$$
$$= x^{ed} \pmod{N}$$
$$= x \pmod{N}.$$

More detailed proof by cases in page 3? of Note 7.

8.3 Why does RSA work?

- N is too large to brute force solve x where $y = x^e \pmod{N}$.
- *N* is too large to factor into $p \cdot q$. Factorization is an intractable problem!

9 Polynomials

A single variable expression of the form

$$p(x) = a_n x^n + a_{n-1} x^{n-1} + \ldots + a_1 x + a_0$$

for reals a_i and x is denoted a polynomial.

Definition 9.1 (degree)

The degree of a polynomial p(x), often denoted deg(p), is the value of the largest exponent of p(x).

For example, any quadratic function has degree 2.

We mainly explore two relevant properties in this section.

Note 9.1 (Proprety 1)

If deg(p) = d, then p(x) has at most d roots.

Note 9.2 (Property 2)

Given d + 1 distinct (x,y) points, we can find/compute a unique degree d polynomial.

The concept of secret sharing follows directly from property 2.

9.1 Finite Fields

We will be using notation GF(p) which represents a finite field (aka Galois Field) with respect to modulo p. All operations in this field are done in $\mod p$. We want to convert all fractions to their modular inverse equivalents.

Example 9.1

If we're working in GF(5), we remark

$$7x^2 \equiv 2x^2 \pmod{5}$$

and

$$\frac{1}{8} \equiv 8^{-1} \equiv 3^{-1} \equiv 2 \pmod{5}.$$

9.2 Lagrange Interpolation

For d+1 points of the form $(x_1, y_1), \ldots, (x_{d+1}, y_{d+1})$, we can construct a unique degree d polynomial

$$p(x) = \sum_{i=1}^{d+1} y_i p_i(x)$$

where

$$p_i(x) = \frac{\prod_{j \neq i} x - x_j}{\prod_{j \neq i} x_i - x_j}.$$

10 Error Correcting Codes

Objective: transmit *n* packets of data (integers).

Two problems may arise.

- 1. Packets get erased/lost (erasure errors)! If we know we have up to k packet erasures, we fix this by sending n + k packets.
- 2. Packets get corrupted (general errors)! If we know we have up to k packets corrupted, we fix this by sending n + 2k packets.

If we run into erasure errors, we simply use interpolation to recover the lost packets.

If we run into general errors on the other hand, we need a more powerful tool.

10.1 Berlekamp-Welch Algorithm

We need to identify which indices the error occurs at. Messages are encoded by some polynomial P(x). Our goal is to retrieve P(x).

1. Suppose we know error at *k* bits. Define the error polynomial

$$E(x) = (x - e_1)(x - e_2) \cdots (x - e_k).$$

- 2. Denote the *i*th packet info we see as r_i . Note, r_i may not be the actual value (might be a corrupted value).
- 3. Solve the equations $P(i)E(i) = r_i E(i)$.
- 4. Define polynomial Q(x) := P(x)E(x).
- 5. Substituting, we have

$$Q(i) = P(i)E(i) = r_i E(i).$$

- 6. We solve linear equations generated by $Q(i) = r_i E(i)$ in step 5 to find the polynomials E(x), Q(x).
- 7. Once we have that, we can calculate

$$P(x) = Q(x)/E(x)$$
.

11 Countability

11.1 Terminology

For all definitions, we use a function $f: A \to B$. A is called the **domain** and B is called the **codomain**.

Definition 11.1 (Injection (one-to-one))

A function f is *injective* or *one-to-one* if no two points in the domain map to the same point in the codomain. Mathematically for all $a \in A$ and $b \in A$,

$$f(a) = f(b) \implies a = b.$$

Definition 11.2 (Surjective (onto))

A function f is *surjective* or *onto* if every point in the codomain has a point in the domain that maps to it. Mathematically, for all $b \in B$ there exists an $a \in A$ such that f(a) = b.

Definition 11.3 (Bijective (one-to-one correspondence))

A function f is *bijective* or has a *one-to-one correspondence* if it is both injective (one-to-one) and surjective (onto).

Definition 11.4 (Cardinality)

The **cardinality** of a set A, denoted |A|, is equal to the number of elements in the set.

Two sets *A* and *B* have the same cardinality (size) if there exists a bijection between *A* and *B*. Another way is to show $|A| \le |B|$ and $|B| \le |A|$ (this is how we prove $|\mathbb{N}| = |\mathbb{Q}|$).

11.2 The Countable

A set S is **countable** if there exists a bijection between S and \mathbb{N} or another countable set. The main idea is this concept of enumeration. If we can find a way to "enumerate" or number a set, we say it's countable.

Note: countable sets may be infinite!

Some examples of common countable sets: \mathbb{N} , \mathbb{Q} , \mathbb{Z} , $\mathbb{Z} \times \mathbb{Z}$, set of all finite length bit strings, set of all polynomials with coefficients in \mathbb{N} .

11.3 The Uncountable

Effectively, the sets that aren't countable are considered uncountable.

Common uncountable sets: power set, \mathbb{R} , set of infinite length bit strings

How do we prove a set S is uncountable. In this class, either we show $|S| > |\mathbb{N}|$ or...

11.4 Cantor Diagonalization

...Cantor Diagonalization. The main idea of Cantor is to show that we can always create a new number that belongs in S that was not originally in S. As a result, we cannot possibly fathom how large S is and enumerate all of its entries since we can always create new entries based on all the ones already in S.

(Lot of words, walk through visual example on board).

12 Computability

Just think of a program like a piece of text (code).

12.1 The Halting Problem

Definition 12.1 (The Halting Problem)

We claim that the ability to determine if a program *P* will terminate on input *x* is uncomputable. In other words, there does not exist computer program (code) that can determine this.

In this class, the way we will prove a problem is uncomputable is by reducing the Halting Problem to this problem. You will explore in further detail in classes like CS170, that if problem A reduces to problem B, then problem B is at least as computationally hard as problem A.

In our case, if we can display that the halting problem reduces to our problem (i.e. if we can solve our problem, we can solve the halting problem), then this shows that our problem is uncomputable.

A basic template to prove some program Other is uncomputable

```
1 def TestHalt(P, x):
2    def Q(y):
3        run P(x)
4        return <whatever makes TestOther true>
5    return TestOther(Q, y)
```

12.2 Foreshadowing

We will cover counting more in depth next discussion.

13 Counting

We introduce a topic in this class called **the first rule of counting**. This effectively says if I have k boxes with n_1, n_2, \ldots, n_k items per respective box, the number of ways to choose 1 item per box is

$$n_1 \cdot n_2 \cdot \cdot \cdot n_k$$
.

Example 13.1

How many ways can we arrange n books on a bookshelf?

Definition 13.1 (factorial)

The factorial function of n denoted n! represents the quantity

$$n! = \prod_{k=1}^{n} k.$$

If you want to check your understanding for small values of n,

n	n!
0	1
1	1
2	2
3	6
4	24
5	120
6	720
7	5040

To shorthand future notation, we will introduce the binomial coefficients.

How many ways can we choose k objects from a total of n objects?

Definition 13.2 (Binomial Coefficient)

The binomial coefficient

$$\binom{n}{k} = \frac{n!}{k!(n-k)!}$$

represents the total number of ways we can choose k objects from a total of n objects for $k \le n$.

13.1 A sampling synopsis

The number of ways I can pick k objects from a total of n objects is...

	sampling with replacement	sampling without replacement
order matters	n^k	$n(n-1)\cdots(n-k+1)$
order doesn't matter	$\binom{n+k-1}{k-1}$	$\binom{n}{k}$

13.2 Stars and Bars

The case when order doesn't matter and we are sampling with replacement is coined stars and bars.

The number of ways to throw n balls into k distinguishable bins is

$$\binom{n+k-1}{k-1}$$

Tip: When doing stars and bars problems, if the question says the bins must have a minimum number of balls, add the minimum # of balls to each bin and do stars and bars with the remaining balls.

Example 13.2

If a problem effectively says that each bin has a positive number of balls, we can add 1 ball to each bin and figure out how to distribute the remaining n - k balls.

13.3 Inclusion-Exclusion

Think of a venn diagram!

For two set case:

$$|A \cup B| = |A| + |B| - |A \cap B|$$
.

For three set case:

$$|A \cup B \cup C| = |A| + |B| + |C| - |A \cap B| - |A \cap C| - |B \cap C| + |A \cap B \cap C|.$$

13.4 Combinatorial Proofs

Focuses on proving mathematical expressions in words with a story.

Example 13.3

Prove that

$$\binom{n}{r} = \binom{n}{n-r}.$$

Pick the side that looks easier and think about what it means. Now create a story that's equivalently portrayed by the other side.

To read on your own time: https://drive.google.com/file/d/1Nzbdno6c_6n-T3A7rmqhqdIUWsYck6b0/view?usp=share_link

14 Intro to Probability

Definition 14.1 (sample space)

The *sample space* denoted Ω is the set of all possible outcomes.

Two main properties of probability:

- $0 \le \mathbb{P}[x] \le 1$ for all $x \in \Omega$.
- $\sum_{x \in \Omega} \mathbb{P}[x] = 1$. In words, the sum of the probabilities of all outcomes is 1.

What really is $\mathbb{P}[x]$?

$$\mathbb{P}[x] = \frac{\text{# of outcomes satisfying } x}{\text{total # of outcomes}} = \frac{|X|}{|\Omega|}.$$

Definition 14.2 (Complement)

The *complement* of an event X is denoted \overline{X} where $\overline{X} = \Omega \setminus X$.

Example 14.1

I flip a fair coin 10 times. What's the probability I get at least 1 head?

14.1 Inclusion-Exclusion

We saw them for counting. We now see them for probability.

$$\mathbb{P}\left[A \cup B\right] = \mathbb{P}\left[A\right] + \mathbb{P}\left[B\right] - \mathbb{P}\left[A \cap B\right]$$

and

$$\mathbb{P}\left[A \cup B \cup C\right] = \mathbb{P}\left[A\right] + \mathbb{P}\left[B\right] + \mathbb{P}\left[C\right] - \mathbb{P}\left[A \cap B\right] - \mathbb{P}\left[A \cap C\right] - \mathbb{P}\left[B \cap C\right] + \mathbb{P}\left[A \cap B \cap C\right].$$

15 Conditional Probability

So far we've looked at the likelihood of some event *A* occurring. What if I want to look at the likelihood of some even *A* occurring given that some event *B* occurred?

This is what spurs the insight into conditional probability. The notation "A|B" should be read as "A given B".

Theorem 15.1 (Bayes Rule)

$$\mathbb{P}\left[A|B\right] = \frac{\mathbb{P}\left[A \cap B\right]}{\mathbb{P}\left[B\right]}.$$

A nice corollary of Bayes Rule is

$$\mathbb{P}\left[A|B\right] = \frac{\mathbb{P}\left[B|A\right]\mathbb{P}\left[A\right]}{\mathbb{P}\left[B\right]}.$$

15.1 Total Probability

We explore an idea called the law of total probability.

Theorem 15.2 (Law of Total Probability)

$$\begin{split} \mathbb{P}\left[B\right] &= \mathbb{P}\left[A \cap B\right] + \mathbb{P}\left[\overline{A} \cap B\right] \\ &= \mathbb{P}\left[B|A\right] \mathbb{P}\left[A\right] + \mathbb{P}\left[B|\overline{A}\right] \mathbb{P}\left[\overline{A}\right]. \end{split}$$

15.2 Independence

Two events *A* and *B* are independent if the occurrence of one does not affect the likelihood of the other. Concretely, for independence TFAE:

- 1. $\mathbb{P}[A \cap B] = \mathbb{P}[A] \mathbb{P}[B]$
- 2. $\mathbb{P}[A|B] = \mathbb{P}[A]$.

16 Random Variables

Recall from last week, we looked at the concept of independence: two events A and B are independent if knowing whether B occurred tells us nothing about whether A occurred.

16.1 Unionized Events

Definition 16.1 (mutually exclusive)

Events *A* and *B* are **mutually exclusive** if $\mathbb{P}[A \cap B] = 0$.

- Generally, mutually exclusive events are (almost) never independent.
- Never want to assume a sequence of events are exclusive or independent for that matter.

Definition 16.2 (Union bound)

For events A_1, A_2, \ldots, A_n , union bound approximation claims

$$\mathbb{P}\left[A_1 \cup A_2 \cup \cdots \cup A_n\right] \leq \mathbb{P}\left[A_1\right] + \mathbb{P}\left[A_2\right] + \ldots + \mathbb{P}\left[A_n\right].$$

The intuition of above should follow from Inclusion-Exclusion.

16.2 Intro to Random Variables

For a sample space Ω , a random variable X is a function $X:\Omega\to\mathbb{R}$; it maps $X(\omega)\to\mathbb{R}$ for every $\omega\in\Omega$.

16.2.1 Distribution of R.V.

There are two important things for any R.V.

- The set of all values it can take (the ω values)
- Probabilities with which it takes on each of those values.

For example, X = a is an *event* that is a set modeled by $S = \{\omega \in \Omega \mid X(\omega) = a\}$. As a result, if I wanted to compute the probability of an event it would follow the form $\mathbb{P}[X = a] = \frac{|S|}{|\Omega|}$.

We will work with some well known distributions in discussion today, but I'll formally define them on Thursday.

17 Distributions and Expectations

We will be looking at discrete random variables for the time being.

17.1 Well known distributions

17.1.1 Bernoulli

Bernoulli RVs either output 0 or 1. If $X \sim \text{Bernoulli}(p)$ then

$$\mathbb{P}\left[X=i\right] = \begin{cases} p & i=1\\ 1-p & i=0 \end{cases}.$$

17.1.2 Binomial

Binomial RVs take in a fixed number of trials n and probability of success p and calculate the probability of i successes for all $i \le n$. If $X \sim \text{Binomial } (n, p)$, then

$$\mathbb{P}\left[X=i\right] = \binom{n}{i} p^{i} (1-p)^{n-i}.$$

17.1.3 Geometric

Geometric RVs for a fixed probability of success p determine the probability that it takes a certain number of trials i until we see our **first** success. If $X \sim \text{Geometric}(p)$, then

$$\mathbb{P}\left[X=i\right] = (1-p)^{i-1}p.$$

These variables are also cool because they are memoryless!

17.2 Joint Distributions

For two RVs X and Y, their *joint distribution* is denoted by the values $\mathbb{P}[X = a, Y = b]$ for all possible a that X can output and all possible b that Y can output.

Suppose from a joint distribution we want to retrieve a distribution of a single RV. How?

Definition 17.1 (Marginal distribution)

From a joint distribution, the distribution of a single RV is defined as its marginal distribution. To calculate it,

$$\mathbb{P}\left[X=a\right] = \sum_{b} \mathbb{P}\left[X=a, Y=b\right]$$

Definition 17.2 (Independence)

For independence in a joint distribution setting, we must have

$$\mathbb{P}\left[X=a,Y=b\right]=\mathbb{P}\left[X=a\right]\mathbb{P}\left[Y=b\right].$$

17.3 Expectation

Definition 17.3 (Expected Value)

The *expected value* of a random variable X, denoted $\mathbb{E}[X]$, is the anticipated average value of X. Mathematically,

$$\mathbb{E}\left[X\right] = \sum_{i \in \Omega} i \cdot \mathbb{P}\left[X = i\right].$$

Example 17.1

The expected value of a single fair dice roll is

$$\mathbb{E}[X] = \sum_{i \in \Omega} i \cdot \mathbb{P}[X = i] = 1 \cdot \mathbb{P}[X = 1] + 2 \cdot \mathbb{P}[X = 2] + \dots + 6 \cdot \mathbb{P}[X = 6]$$
$$= 1 \cdot \frac{1}{6} + 2 \cdot \frac{1}{6} + \dots + 6 \cdot \frac{1}{6}$$
$$= \frac{7}{2} = 3.5$$

Theorem 17.1 (Tail Sum Formula)

The Tail-Sum formula for expectation is of the form

$$\mathbb{E}\left[X\right] = \sum_{i=1}^{\infty} \mathbb{P}\left[X \ge i\right].$$

17.3.1 Linearity of Expectation

For any random variables X and Y, the property

$$\mathbb{E}\left[X+Y\right] = \mathbb{E}\left[X\right] + \mathbb{E}\left[Y\right]$$

is always true and is often denoted linearity of expectation.

A result of above is that for RV X and constant c, we have $\mathbb{E}[cX] = c\mathbb{E}[X]$. Additionally, by definition $\mathbb{E}[c] = c$.

18 Variance

18.1 Recall

Last time we looked at

Definition 18.1 (Expected Value)

The expected value of a random variable X, denoted $\mathbb{E}[X]$, is the anticipated average value of X. Mathematically,

$$\mathbb{E}\left[X\right] = \sum_{i \in \Omega} i \cdot \mathbb{P}\left[X = i\right].$$

We extend this definition slightly to incorporate for a random variable defined by a function $g(\cdot)$. If X is a RV, then so is g(X).

Theorem 18.1 (Law of the Unconscious Statistician (LOTUS))

$$\mathbb{E}\left[g(X)\right] = \sum_{i \in \Omega} g(i) \mathbb{P}\left[X = i\right].$$

18.2 Variance

Today we introduce a new topic variance.

Definition 18.2 (Variance)

The *variance* of a random variable *X* measures how much on average the variable deviates from its expectation (the mean).

The main way we will compute variance is

$$\operatorname{Var}\left(X\right) = \mathbb{E}\left[X^{2}\right] - \mathbb{E}\left[X\right]^{2}.$$

If you have taken a statistics class before, you may remark that the standard deviation σ is defined as

$$\sigma(X) = \sqrt{\operatorname{Var}(X)}.$$

We looked at how scalars behaved with expectation. For variance, with a constant c,

$$Var\left(cX\right) = c^2 Var\left(X\right)$$

and

$$Var(X + c) = Var(X)$$
.

18.3 Independence

If two RVs X and Y are independent, then

$$Var(X + Y) = Var(X) + Var(Y)$$
.

19 Covariance, More Distributions

Last time we looked at variance, today we look at

19.1 Covariance

Covariance measures the association between two (or more) RVs X and Y.

Mathematically,

$$Cov (X,Y) = \mathbb{E} [XY] - \mathbb{E} [X] \mathbb{E} [Y].$$

We see that if

$$Cov(X,Y): \begin{cases} <0 & \Longrightarrow X \text{ and } Y \text{ are inversely correlated} \\ =0 & \Longrightarrow \text{ no correlation (does not imply independence)} \\ >0 & \Longrightarrow X \text{ and } Y \text{ are directly correlated} \end{cases}$$

19.2 Some general principles for Var and Cov

- 1. Cov(X,Y) = Cov(Y,X)
- 2. Cov(X + Y, Z) = Cov(X, Z) + Cov(Y, Z)
- 3. Cov(aX, Y) = aCov(X, Y)
- 4. Var(X) = Cov(X,X)
- 5. Cov $(X,Y) = \mathbb{E}[XY] \mathbb{E}[X]\mathbb{E}[Y]$
- 6. Var(X + Y) = Var(X) + Var(Y) + 2Cov(X,Y)

A consequence of above is **when X and Y are independent**:

- Cov (X,Y) = 0
- $\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y]$
- $\operatorname{Var}(X + Y) = \operatorname{Var}(X) + \operatorname{Var}(Y)$

README: When two variables are independent their covariance is 0 but just because their covariance is 0 doesn't mean they're independent

19.3 Poisson Distribution

We use this distribution when the data tends to fluctuate around some rate or "average". We denote this distribution as $X \sim \text{Poisson}(\lambda)$ where λ is the rate.

We claim

$$\mathbb{P}\left[X=i\right] = \frac{\lambda^i}{i!}e^{-\lambda}.$$

An interesting application is that Binomial distribution as $n \to \infty$ approaches Poisson.

Another interesting fact is that for independent X and Y where $X \sim \text{Poisson}(\lambda)$ and $Y \sim \text{Poisson}(\mu)$, then $X + Y \sim \text{Poisson}(\lambda + \mu)$.

19.4 EV/Variance Recap

X	$\mathbb{E}\left[X\right]$	Var (X)
Bernoulli(p)	p	p(1 - p)
Binomial (n,p)	np	np(1-p)
Geometric (p)	$\frac{1}{p}$	$\frac{1-p}{p^2}$
Poisson(λ)	λ	λ

20 Conditional PMFs and Expectation

Recall that in a joint setting, the principal of marginal distribution revolved around the idea of

$$\mathbb{P}\left[X=a\right] = \sum_{b} \mathbb{P}\left[X=a|Y=b\right] \mathbb{P}\left[Y=b\right].$$

We extend this idea (of total probability) now to expectation. For conditional expectation, we first require the conditional PMF.

Definition 20.1 (Conditional PMF)

The conditional PMF of a variable X describes how it behaves conditioned with respect to Y. Mathematically,

$$\mathbb{P}\left[X = x | Y = y\right] = \frac{\mathbb{P}\left[X = x, Y = y\right]}{\mathbb{P}\left[Y = y\right]}.$$

Definition 20.2 (Conditional Expectation)

The *conditional expectation* of X given Y = y is defined as

$$\mathbb{E}[X|Y=y] = \sum_{x} x \cdot \mathbb{P}[X=x|Y=y]$$

If we apply the principal of marginal distribution to the conditional expectation above, we get

$$\sum_{y} \mathbb{E}[X|Y=y] \mathbb{P}[Y=y] = \mathbb{E}[X].$$

The result above is known as law of iterated expectation. Formally, it's defined as

$$\mathbb{E}\left[X\right] = \mathbb{E}\left[\mathbb{E}\left[X|Y\right]\right].$$

Remark: $\mathbb{E}[X|Y]$ is a function in terms of Y.

20.1 EV/Variance Recap

X	$\mathbb{E}\left[X\right]$	Var(X)
Bernoulli(p)	p	p(1 - p)
Binomial (n,p)	np	np(1-p)
Geometric (p)	$\frac{1}{p}$	$\frac{1-p}{p^2}$
Poisson(λ)	λ	λ
Uniform (a,b)	<u>a+b</u> 2	?

21 Concentration Inequalities

21.1 Markov's Inequality

For nonnegative random variable X, Markov's inequality tells us

$$\mathbb{P}\left[X \geq a\right] \leq \frac{\mathbb{E}\left[X\right]}{a}.$$

To use Markov's inequality we only need two pieces of information:

- X is nonnegative
- Know what $\mathbb{E}[X]$ is

21.2 Chebyshev's Inequality

For positive constant c, Chebyshev's Inequality tells us

$$\mathbb{P}\left[\left|X - \mathbb{E}\left[X\right]\right| \ge c\right] \le \frac{\operatorname{Var}\left(X\right)}{c^2}.$$

Chebyshev's inequality is stronger than Markov's inequality. To use Chebyshev's inequality we need three pieces of information:

- c is positive
- Know what $\mathbb{E}[X]$ is
- Know what Var (X) is

An interesting fact we should consider when trying to use Chebyshev is the following.

$$\mathbb{P}\left[X-a\geq b\right]\leq \mathbb{P}\left[|X-a|\geq b\right].$$

This is because the solutions to $x - a \ge b$ for x are a subset of the solutions to $|x - a| \ge b$ for x.

21.3 Law of Large Numbers (LLN)

Theorem 21.1 (Law of Large Numbers)

For a sequence of i.i.d (independent and identically distributed) random variables X_i with finite expectation $\mathbb{E}[X] < \infty$, we note the following: for any $\epsilon > 0$, defining $S_n = X_1 + X_2 + \ldots + X_n$,

$$\mathbb{P}\left[\left|\frac{S_n}{n} - \mathbb{E}\left[X\right]\right| < \epsilon\right] \to 1$$

as $n \to \infty$.

22 Continuous Probability

So far we've looked at discrete systems and distributions. Today, we turn our focus more towards continuous models.

22.1 Important Definitions

Definition 22.1 (Probability Density Function (PDF))

The *probability density function (pdf)* of a random variable X is a function $f : \mathbb{R} \to \mathbb{R}$ where

- $(\forall x \in \mathbb{R}) f(x) \ge 0$
- $\int_{-\infty}^{\infty} f(x) dx = 1$.

Visually, $\mathbb{P}[a \le X \le b] = \int_a^b f(x) dx$.

BEWARE: Tricky nuance about continuous distributions is that $\mathbb{P}[X = a] = 0$. This is why we always look at probability of X in a range and not at a singular point. **THE PDF IS NOT A PROBABILITY**.

Definition 22.2 (Cumulative Distribution Function (CDF))

The *cumulative distribution function (cdf)* of a random variable *X* is a function *F* where

$$F(x) = \mathbb{P}\left[X \le x\right] = \int_{-\infty}^{x} f(z) \, \mathrm{d}z.$$

Remark that as $x \to \infty$, $F(x) \to 1$ and by construction F(x) must be a non-decreasing function since $f(z) \ge 0$. (draw the visuals for both cdf and pdf)

IMPORTANT RELATION: $f(x) = \frac{d}{dx}F(x)$. In words, the PDF is equal to the derivative of the CDF.

22.2 Expectation

We extend the definition of expectation to

$$\mathbb{E}\left[X\right] = \int_{-\infty}^{\infty} x \cdot f_X(x) \, \mathrm{d}x$$

where $f_X(x)$ is the pdf of X.

22.3 Discrete vs Continuous

Concept	Discrete	Continuous
PDF/PMF	$\mathbb{P}\left[X=x\right]$	$f_X(x)$
CDF	$\mathbb{P}\left[X \leq x\right]$	F(x)
Expectation	$\sum_{x} x \mathbb{P} \left[x = x \right]$	$\int_{-\infty}^{\infty} x \cdot f_X(x) \mathrm{d}x$
Variance	$\sum_{x} x^{2} \mathbb{P} \left[x = x \right] - \left(\sum_{x} x \mathbb{P} \left[x = x \right] \right)^{2}$	$\int_{-\infty}^{\infty} x^2 \cdot f_X(x) \mathrm{d}x - \left(\int_{-\infty}^{\infty} x \cdot f_X(x) \mathrm{d}x \right)^2$
Joint Distribution	$\mathbb{P}\left[X=a,Y=b\right]$	$\mathbb{P}\left[a \le X \le b, c \le Y \le d\right]$
Independence	$\mathbb{P}\left[X=x,Y=y\right]=\mathbb{P}\left[X=x\right]\mathbb{P}\left[Y=y\right]$	$f_{X,Y}(x,y) = f_X(x)f_Y(y)$

22.4 Exponential Distribution

The exponential distribution is the continuous analog of the geometric distribution. Represents the idea of a "first arrival" and also adheres to the memoryless property. Note that λ here can be interpreted as the rate of arrival.

For a random variable $X \sim \text{Exp}(\lambda)$, the relevant things we have are

$$F(x) = \mathbb{P}[X \le x] = 1 - e^{-\lambda x}$$

$$f_X(x) = \frac{d}{dx}F(x) = \lambda e^{-\lambda x}$$

$$\mathbb{E}[X] = \frac{1}{\lambda}$$

$$Var(X) = \frac{1}{\lambda^2}.$$

22.5 Big Recap

22.5.1 EV/Variance

X	$\mathbb{E}\left[X\right]$	Var(X)
Bernoulli(p)	p	p(1 - p)
Binomial (n,p)	np	np(1-p)
Geometric (p)	$\frac{1}{p}$	$\frac{1-p}{p^2}$
Poisson(λ)	λ	λ
DUniform(a,b)	<u>a+b</u> 2	?
CUniform(a,b)	<u>a+b</u> 2	$\frac{(b-a)^2}{12}$
Exponential(λ)	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$

22.5.2 Discrete distributions

X	$\mathbb{P}\left[X=i\right]$	
Binomial (n,p)	$\binom{n}{i}p^i(1-p)^n-i$	
Geometric (p)	$p(1-p)^{i-1}$	
$Poisson(\lambda)$	$\frac{\lambda^i}{i!}e^{-\lambda}$	
DUniform(a,b)	$\frac{1}{b-a}$	

22.5.3 Continuous distributions

X	$f_X(X)$	F(X)
Uniform (a,b)	$\frac{1}{b-a}$	$\frac{x}{b-a}$
Exponential(λ)	$\lambda e^{-\lambda x}$	$1 - e^{-\lambda x}$

23 Gaussian Distribution and CLT

23.1 Gaussian Distribution

This is the last form of continuous distributions we will look at in this class.

A random variable of the form $X \sim \mathcal{N}(\mu, \sigma^2)$ is called a normal/gaussian random variable. It has the following properties

$$\mathbb{E}[X] = \mu$$

$$\operatorname{Var}(X) = \sigma^{2}$$

$$f_{X}(x) = \frac{1}{\sqrt{2\pi\sigma^{2}}} e^{-(x-\mu)^{2}/(2\sigma^{2})}$$

where $f_X(x)$ is the pdf of X.

Definition 23.1 (Unit Normal Distribution)

Special case of the normal distribution where $\mu = 0$ and $\sigma^2 = 1$. We generally use the letter Z to denote a distribution of this form: $Z \sim \mathcal{N}(0, 1)$.

We use a special symbol for dealing with the CDF of unit normal distributions,

$$F(t) = \mathbb{P}\left[Z \le t\right] = \Phi(t).$$

We can talk about normalizing a distribution X, this is trying to write it in terms of Z.

If
$$X \sim \mathcal{N}(\mu, \sigma^2)$$
, then

$$\frac{X-\mu}{\sigma}=Z\sim\mathcal{N}\left(0,1\right).$$

The expression above can be rewritten as $X = \sigma Z + \mu$. This just highlights another way to decompose a normal distribution X in terms of the unit normal distribution.

Scaling a distribution: for a constant c, if $X \sim \mathcal{N}(\mu, \sigma^2)$, then $cX \sim \mathcal{N}(c\mu, c^2\sigma^2)$.

Theorem 23.1

The sum of independent normal distributions is also a normal distribution. Namely, if $X \sim \mathcal{N}\left(\mu_x, \sigma_x^2\right)$ and $Y \sim \mathcal{N}\left(\mu_y, \sigma_y^2\right)$ are independent, then $X + Y \sim \mathcal{N}\left(\mu_x + \mu_y, \sigma_x^2 + \sigma_y^2\right)$.

23.2 Central Limit Theorem (CLT)

Last week, we looked at LLN which said that the probability of any small ϵ deviation of the sample average from the mean tends to 0 as the number of samples tends to infinity. Today, we look at something *stronger*.

Theorem 23.2 (Central Limit Theorem)

Let X_1, \ldots, X_n be i.i.d random variables with $\mathbb{E}[X_i] = \mu < \infty$ and $\text{Var}(X_i) = \sigma^2 < \infty$. Define $S_n = X_1 + \ldots + X_n$. Then,

$$\frac{S_n - n\mu}{\sigma\sqrt{n}} \to \mathcal{N}(0,1)$$

as $n \to \infty$.

Remark: The CLT is not only for normal distributions. It works for any i.i.d random variables with the same (finite) expectation and same (finite) variance.

24 Regression and Least Squares

Has a lot of interesting applications. Consider classes like EECS127 and CS189.

Definition 24.1 (LLSE)

The linear least squares estimate of X given Y, also denoted as LLSE(X|Y) is generally expressed by the formula

$$\mathrm{LLSE}(X|Y) = \mathbb{E}\left[X\right] + \frac{\mathrm{Cov}\left(X,Y\right)}{\mathrm{Var}\left(Y\right)}\left(Y - \mathbb{E}\left[Y\right]\right).$$

LLSE focuses on finding a linear model to solve the mean square error problem.

Definition 24.2 (MMSE)

The Minimum Mean Squares Estimate of X given Y, also denoted $\mathrm{MMSE}(X|Y)$ is generally expressed by the formula

$$\mathsf{MMSE}(X|Y) = \mathbb{E}\left[X|Y\right].$$

25 Markov Chains

25.1 Basic Terms

We start with some notation.

Definition 25.1 (State)

The state of a Markov Chain X_n refers to where in the state space we are at at the *n*th iteration.

Definition 25.2 (State Space)

The state space χ is the set of all possible states of our random variable

Definition 25.3 (Transition Matrix)

The transition matrix, commonly denoted as P, is a matrix whose entries demonstrate a mapping: $P(i,j) \mapsto$ probability of going to state j from state i.

Definition 25.4 (Distribution)

At some point in the state X_n , the distribution π_n represents $\pi_n(i) = \mathbb{P}[X_n = i]$.

Remark: In this class we generate represent vectors as row vectors and not column vectors like in other math courses.

Some other nice properties:

- For all i, $\sum_{j} P(i,j) = 1$. In english, the probability you go to another state from state i is 1. This feels rather trivial. The implication is that every row in any Transition Matrix must sum to 1.
- Markov Chains are memoryless!

25.2 Hitting Time

Objective: Focus on expected number of "steps" or "transitions" to get to an end state E starting from state S. If $\beta(i)$ is expected number of steps to get from state i to the end, then

$$\beta(i) = \begin{cases} 0 & i = \text{ end} \\ 1 + \sum_{j \in \chi} P(i,j)\beta(j) & \text{ otherwise} \end{cases}$$

The motivation of the summation is law of total probability. To explain the nonzero term of the piecewise:

- 1: take one step/turn at state i.
- P(i,j): the probability we end up in state j from state i
- $\beta(j)$: how much further to go to the end (in expectation) from state j

25.3 A before B (Absorbing States)

Objective: Focus on the probability of reaching state A before reaching state B.

Let $\alpha(i)$ denote the probability of reaching state A before state B starting at state i. Then,

$$\begin{cases} 0 & i = B \\ 1 & i = A \\ \sum_{j \in Y} P(i,j)\alpha(j) & \text{otherwise} \end{cases}.$$

We once again breakdown the nonzero term:

- P(i,j): probability end up in state j from state i
- $\alpha(j)$: probability reaching state *A* before *B* from state *j*.

25.4 More Definitions

Definition 25.5 (Invariant)

The distribution π is invariant for transition matrix P if $\pi = \pi P$.

We can conceptually think of the above as net flow in and out of states being equal. In fact, $\pi_n = \pi_0$ iff π_0 is invariant. This follows from the definition that $\pi_n = \pi_0 P^n$.

Definition 25.6 (Irreducibility)

A Markov Chain is irreducible if it is possible to go from every state i to every other state j.

This is relevant because irreducible chains have unique stationary distributions. The fraction of time spent in each state (in the long run) converges to this unique distribution irrespective of the initial distribution.

Definition 25.7 (Periodicity)

Suppose we have some function d where d(i) is defined as the GCD across all possible lengths of paths starting and ending at state i.

- If d(i) = 1, then the chain is aperiodic.
- Otherwise, it's periodic with period d(i).

Remark: The notion of periodicity is only defined for irreducible chains.

If Markov Chain is finite, irreducible, and aperiodic, then every initial distribution converges to the stationary distribution.