

STAT GU4261/GR5261 STATISTICAL METHODS IN FINANCE

FALL 2018

HOMEWORK 1 SUGGESTED SOLUTION

If you have any questions about the solution, please contact Brian at hl2902@columbia.edu

R lab section. Problem 4 on page 13:

Just copy the code in the book and run it. The answer is 0.50988.

R lab section. Problem 5, 6 on page 14:

(a) 0.3982

(b) 0.5833

R code:

```
# Problem 5
# parameter setting
P0 = 1e6
mu = 0.05/253
sigma = 0.23/sqrt(253)
sell_price = 1100000
cut_loss_price = 950000

# simulation parameter
n = 10000
profit = rep(0,n)
loss = rep(0,n)
set.seed(1999)
for (i in 1:n) {
  R = rnorm(100, mu, sigma)
  P = P0*exp(cumsum(R))

  # the first day of having loss
  loss_day = which(P < cut_loss_price)[1]

  # the first day of having profit
  profit_day = which(P > sell_price)[1]

  if (max(P) >= sell_price & min(P) < cut_loss_price) {
    profit[i] = as.numeric(loss_day > profit_day)
    loss[i] = as.numeric(loss_day < profit_day)
  }

  if ( max(P) >= sell_price & min(P) >= cut_loss_price) {
    profit[i] = 1
  }
}
```

```

}

if ( max(P) < sell_price & min(P) < cut_loss_price) {
loss[i] = 1
}

if (max(P) < sell_price & min(P) >= cut_loss_price) {
loss[i] = as.numeric(P[100] < 1e6)
}
}
mean(profit)
mean(loss)

```

Problem 3.11.1 in textbook:

(a)  $y_{20} = \frac{1}{20} \int_0^{20} (0.028 + 0.00042t) dt = 0.0322$ .

(b)  $P = 1000e^{-\int_0^{15} (0.028 + 0.00042t) dt} = 626.7$ .

Problem 3.11.3 in textbook:

- (a) Coupon rate > current yield if and only if price > par. Therefore, it is selling above par.
- (b) Price > par if and only if coupon rate > current yield > yield to maturity. Hence, yield to maturity is below 2.8%.

Remark: Let  $y$  be the yield. The bond price is

$$P = \sum_{t=1}^T \frac{C}{(1+y)^t} + \frac{\text{PAR}}{(1+y)^T}.$$

Divided both sides by  $P$  and rearranging the terms, we obtain

$$\frac{C}{P} = \frac{1 - \frac{\text{PAR}}{P} \frac{1}{(1+y)^T}}{\sum_{t=1}^T \frac{1}{(1+y)^t}}.$$

Since

$$\sum_{t=1}^T \frac{1}{(1+y)^t} = \frac{\frac{1}{(1+y)^{T+1}} - \frac{1}{1+y}}{\frac{1}{1+y} - 1} = \frac{1 - \frac{1}{(1+y)^T}}{y},$$

we have

$$\frac{C}{P} = y \times \frac{1 - \frac{\text{PAR}}{P} \frac{1}{(1+y)^T}}{1 - \frac{1}{(1+y)^T}}.$$

As  $\frac{C}{P}$  is the current yield, we see from the above equation that current yield > yield to maturity if and only if price > PAR.

Problem 3.11.4 in textbook:

(a)  $y_5 = \frac{1}{5} \int_0^5 (0.032 + 0.001t + 0.0002t^2) dt = 0.0362$

(b)  $P = \text{PAR}e^{-5y_5} = 0.834\text{PAR}$ .

Problem 3.11.7 in textbook:

(a)  $C = \frac{0.085}{2} \times 1000 = 42.5.$

(b)

$$P = \sum_{t=1}^{38} \frac{42.5}{(1 + \frac{0.076}{2})^t} + \frac{1000}{(1 + \frac{0.076}{2})^{38}} = 1089.717908.$$

(c)

$$P = 42.5 + \sum_{t=1}^{38} \frac{42.5}{(1 + \frac{0.076}{2})^t} + \frac{1000}{(1 + \frac{0.076}{2})^{38}} = 1132.217908.$$

Problem 3.11.13 in textbook:

$$\begin{aligned} y_{20} &= \frac{1}{20} \int_0^{20} r(t) dt \\ &= \frac{1}{20} \int_0^{10} 0.03 + 0.001t \, dt + \frac{1}{20} \int_{10}^{20} 0.03 + 0.001t - 0.00021(t - 10) \, dt \\ &= \frac{1}{20} \left[ 0.03t + \frac{0.001}{2} t^2 \right]_0^{20} - \frac{1}{20} \left[ \frac{0.00021}{2} (t - 10)^2 \right]_{10}^{20} \\ &= 0.039475. \end{aligned}$$

Problem 3.11.15 in textbook:

First,

$$\frac{d}{d\delta} \sum_{i=1}^N C_i e^{-T_i(y_{T_i} + \delta)} = - \sum_{i=1}^N T_i C_i e^{-T_i(y_{T_i} + \delta)}.$$

Setting  $\delta = 0$ ,

$$\begin{aligned} \left. \frac{d}{d\delta} \sum_{i=1}^N C_i e^{-T_i(y_{T_i} + \delta)} \right|_{\delta=0} &= - \sum_{i=1}^N T_i C_i e^{-T_i y_{T_i}} = - \left( \sum_{i=1}^N \frac{C_i e^{-T_i y_{T_i}}}{\sum_{j=1}^N C_j e^{-T_j y_{T_j}}} T_i \right) \sum_{j=1}^N C_j e^{-T_j y_{T_j}} \\ &= -\text{DUR} \sum_{j=1}^N C_j e^{-T_j y_{T_j}}. \end{aligned}$$

Note that

$$\left. \frac{d}{d\delta} \sum_{i=1}^N C_i e^{-T_i(y_{T_i} + \delta)} \right|_{\delta=0} = \lim_{\delta \rightarrow 0} \frac{\sum_{i=1}^N C_i e^{-T_i(y_{T_i} + \delta)} - \sum_{i=1}^N C_i e^{-T_i y_{T_i}}}{\delta}$$

Hence, when  $\delta$  is small,

$$\frac{\sum_{i=1}^N C_i e^{-T_i(y_{T_i} + \delta)} - \sum_{i=1}^N C_i e^{-T_i y_{T_i}}}{\delta} \approx -\text{DUR} \sum_{i=1}^N C_i e^{-T_i y_{T_i}}.$$

Finally, as  $\sum_{i=1}^N C_i e^{-T_i y_{T_i}}$  is the bond price, we have

$$\frac{\text{change bond price}}{\text{bond price}} \approx -\text{DUR} \times \delta.$$

Problem 3.11.21 in textbook:

We need to solve  $y$  in

$$1015 = \frac{25}{y} + (1000 - \frac{25}{y})(1+y)^{-8}.$$

It can be found that  $y = 0.0229$ . R code:

```
f = function(y) {
  return(25/y + (1000- 25/y)*(1+y)^(-8) - 1015)
}
uniroot(f, interval = c(0.001,0.05))
```

Other questions:

(6)  $\mathbb{P}(U \leq u) = \mathbb{P}(F(X) \leq u) = \mathbb{P}(X \leq F^{-1}(u)) = F(F^{-1}(u)) = u$ . Hence,  $U \sim \text{Unif}(0, 1)$ .

(7) (a) By chain rule and fundamental theorem of calculus,

$$f_Y(y) = \frac{d}{dy} \mathbb{P}(Y \leq y) = \frac{d}{dy} \mathbb{P}(X \leq \log Y) = \frac{d \log y}{dy} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(\log y - \mu)^2}{2\sigma^2}} = \frac{1}{y\sqrt{2\pi\sigma^2}} e^{-\frac{(\log y - \mu)^2}{2\sigma^2}}.$$

(b) We first prove the claim in the hint. Note that

$$\begin{aligned} \mathbb{E}(e^{Xt}) &= \int_{-\infty}^{\infty} e^{xt} \frac{1}{\sqrt{2\pi}} \exp \left\{ -\frac{(x - \mu)^2}{2\sigma^2} \right\} dx \\ &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \exp \left\{ -\frac{x^2 - 2x\mu + \mu^2 - 2x\sigma^2 t}{2\sigma^2} \right\} dx \\ &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \exp \left\{ -\frac{[x - (\mu + \sigma^2 t)]^2 - 2\mu\sigma^2 t - \sigma^4 t^2}{2\sigma^2} \right\} dx \\ &= \exp \left\{ \mu t + \frac{1}{2} \sigma^2 t^2 \right\} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \exp \left\{ -\frac{[x - (\mu + \sigma^2 t)]^2}{2\sigma^2} \right\} dx \\ &= \exp \left\{ \mu t + \frac{1}{2} \sigma^2 t^2 \right\}. \end{aligned}$$

Hence, the mean is  $\mathbb{E}(e^X) = e^{\mu + \frac{1}{2}\sigma^2}$  and

$$\begin{aligned} \text{Var}(Y) &= \text{Var}(e^X) \\ &= \mathbb{E}(e^X)^2 - [\mathbb{E}(e^X)]^2 \\ &= \mathbb{E}(e^{2X}) - [\mathbb{E}(e^X)]^2 \\ &= e^{2\mu + 2\sigma^2} - [e^{\mu + \sigma^2/2}]^2 \\ &= e^{2\mu + \sigma^2} (e^{\sigma^2} - 1). \end{aligned}$$

(8) (a)  $\mathbb{E}(F_n(x)) = \frac{1}{n} \sum_{i=1}^n \mathbb{P}(X_i \leq x) = F(x)$ .

(b)  $\text{Var}(F_n(x)) = \frac{1}{n^2} \sum_{i=1}^n \text{Var}(I(X_i \leq x)) = \frac{F(x)(1-F(x))}{n}$ .

(c) Note that  $F_n(x) = \frac{1}{n} \sum_{i=1}^n (X_i \leq x)$ , which is the average of iid random variables with mean  $F(x)$  and variance  $F(x)(1-F(x))$ . Hence, by central limit theorem, the term converges (in distribution) to the standard normal distribution  $\mathcal{N}(0, 1)$  as  $n \rightarrow \infty$ .