

10. 解: 电子元件为一等品的概率为:

$$e^{-2} \approx 0.1353$$

设1000件中该电子元件的一等品数为 X , 则

$$X \sim b(1000, 0.1353)$$

由中心极限定理,

$$\frac{X - 1000 \times 0.1353}{\sqrt{1000 \times 0.1353 \times 0.8647}} \sim N(0, 1)$$

故

$$P\{X > 150\}$$

$$\approx P\left\{\frac{X - 1000 \times 0.1353}{\sqrt{1000 \times 0.1353 \times 0.8647}} > 1.36\right\}$$

$$\approx 1 - \Phi(1.36)$$

$$= 0.0869$$

$$2. \text{ 解: } E z_1 = E(X) \cdot E\left(\frac{1}{Y}\right) = 0$$

$$E z_2 = E(X) \cdot E(Y) = 0$$

$$E(z_1 \cdot z_2) = E(x^2) = D(X) + (EX)^2 = 3 + 0^2 = 3$$

$$\text{cov}(z_1, z_2) = E(z_1 z_2) - E z_1 \cdot E z_2 = 3$$

$$E\left(\frac{1}{Y^2}\right) = \int_1^{\infty} \frac{1}{x^2} \cdot 3x^{-4} dx = \int_1^{\infty} 3x^{-6} dx = \frac{3}{5}$$

$$E(z_1^2) = E(x^2) \cdot E\left(\frac{1}{Y^2}\right) = 3 \times \frac{3}{5} = \frac{9}{5}$$

$$E(Y^2) = \int_1^{\infty} x^2 \cdot 3x^{-4} dx = \int_1^{\infty} 3x^{-2} dx = 3$$

$$E(z_2^2) = E(x^2) \cdot E(Y^2) = 3 \times 3 = 9$$

$$D(z_1) = E(z_1^2) - (E z_1)^2 = \frac{9}{5} - 0^2 = \frac{9}{5}$$

$$D(z_2) = E(z_2^2) - (E z_2)^2 = 9$$

$$\rho = \frac{\text{cov}(z_1, z_2)}{\sqrt{D(z_1) \cdot D(z_2)}} = \frac{3}{\sqrt{\frac{9}{5} \times 9}} = \frac{\sqrt{5}}{3}$$

3. 解: (1) 设 x_1, x_2, \dots, x_n 为样本观测值, 则

$$\begin{aligned} L(\theta) &= \prod_{i=1}^n f(x_i) = \prod_{i=1}^n (2x_i \theta^{-1} \exp\{-x_i^2 \theta^{-1}\}) \\ &= 2^n \left(\prod_{i=1}^n x_i\right) \theta^{-n} \exp\left\{-\theta^{-1} \sum_{i=1}^n x_i^2\right\} \end{aligned}$$

$$\ln L(\theta) = \left(n \ln 2 + \sum_{i=1}^n \ln x_i\right) - n \ln \theta - \theta^{-1} \sum_{i=1}^n x_i^2$$

$$\text{令 } (\ln L(\theta))' = -\frac{n}{\theta} + \theta^{-2} \sum_{i=1}^n x_i^2 = 0, \text{ 解得}$$

$$\hat{\theta} = \frac{\sum_{i=1}^n x_i^2}{n}$$

$$\text{故 } \theta \text{ 的最大似然估计量为 } \hat{\theta} = \frac{\sum_{i=1}^n X_i^2}{n}.$$

$$(2) E(X^2) = \int_0^{\infty} 2x^3 \cdot \theta^{-1} \exp\{-x^2 \theta^{-1}\} dx.$$

$$\xrightarrow{y=x \cdot \theta^{-\frac{1}{2}}} 2\theta \int_0^{\infty} y^3 \exp\{-y^2\} dy.$$

$$\xrightarrow{t=y^2} \theta \int_0^{\infty} t e^{-t} dt = \theta$$

$$\text{故 } E\hat{\theta} = \frac{\sum_{i=1}^n E(X_i^2)}{n} = E(X^2) = \theta.$$

即 $\hat{\theta}$ 是 θ 的无偏估计量.

4. 解: $\bar{x} = 19$, $S^2 = \frac{40}{3} \approx 13.3333$

$S = 3.65$, $n = 10$, $\alpha = 0.05$.

血压增高均值的 95% 置信区间为

$$\begin{aligned} & \left(\bar{x} - \frac{S}{\sqrt{n}} t_{\alpha/2}^{(n-1)}, \bar{x} + \frac{S}{\sqrt{n}} t_{\alpha/2}^{(n-1)} \right) \\ &= \left(19 - \frac{3.65}{\sqrt{10}} \times 2.2622, 19 + \frac{3.65}{\sqrt{10}} \times 2.2622 \right) \\ &= (16.39, 21.61). \end{aligned}$$

血压增高^{标准差}~~均值~~的 95% 置信区间为

$$\begin{aligned} & \left(\frac{\sqrt{n-1} S}{\sqrt{\chi_{\alpha/2}^2(n-1)}}, \frac{\sqrt{n-1} S}{\sqrt{\chi_{1-\alpha/2}^2(n-1)}} \right) \\ &= \left(\frac{\sqrt{9} \times 3.65}{\sqrt{19.022}}, \frac{\sqrt{9} \times 3.65}{\sqrt{2.7}} \right) \\ &= (2.51, 6.66). \end{aligned}$$

$$X \sim N(\mu_1, \sigma^2), Y \sim N(\mu_2, \sigma^2)$$

5. 解: 检验 $H_0: \mu_1 = \mu_2$, $H_1: \mu_1 \neq \mu_2$.

检验统计量:

$$t = \frac{\bar{x} - \bar{y}}{S_w \cdot \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}, \text{ 其中 } S_w = \sqrt{\frac{(n_1-1)S_1^2 + (n_2-1)S_2^2}{n_1+n_2-2}}$$

拒绝域:

$$t < -t_{\alpha/2}(n_1+n_2-2), \text{ 或 } t > t_{\alpha/2}(n_1+n_2-2).$$

$$\alpha = 0.05, \quad n_1 = 6, \quad n_2 = 4,$$

$$t_{0.025}(8) = 2.3060$$

$$S_w^2 = 3.28, \quad S_w = \sqrt{3.28} = 1.81,$$

$$t = 1.37 \in (-2.3060, 2.3060)$$

未落入拒绝域, 故两种农药残留时间无显著差异.

6. 解: $H_0: P\{X=k\} = e^{-\lambda} \cdot \frac{\lambda^k}{k!}, k=0, 1, 2, \dots$

$\hat{\lambda} = \bar{x} = 0.6, n=200$

x_i	f_i	\hat{p}_i	$n\hat{p}_i$
0	116	0.5488	109.76
1	56	0.3293	65.76
2	22	0.0988	19.76
3	4	0.0198	3.96
4	2	0.003	0.6
≥ 5	0	0.0003	0.06

$\left. \begin{array}{l} 4 \\ 2 \end{array} \right\} 28$ $\left. \begin{array}{l} 3.96 \\ 0.6 \end{array} \right\} 24.38$

合并后组数为 $K=3$,

$$\chi^2 = \sum_{i=0}^2 \frac{f_i^2}{n\hat{p}_i} - n = 2.37 < 3.843 = \chi_{0.05}^2(1)$$

无法拒绝原假设, 故认为次品件数服从泊松分布.

7. 解: (1) $\bar{x} = 19.5$, $\bar{y} = 69.5$, $n = 10$.

$$S_{xx} = \sum_{i=1}^n x_i^2 - n\bar{x}^2 = 82.5$$

$$S_{xy} = \sum_{i=1}^n x_i y_i - n\bar{x}\bar{y} = 233$$

$$\hat{b} = \frac{S_{xy}}{S_{xx}} = 2.8242$$

$$\hat{a} = \bar{y} - \hat{b}\bar{x} \approx 13.53$$

$$\hat{y} = 13.53 + 2.8242x$$

(2) $H_0: b=0$, $H_1: b \neq 0$.

检验统计量: $t = \frac{\hat{b}}{\hat{\sigma}} \sqrt{S_{xx}}$,

其中 $\hat{\sigma} = \sqrt{\frac{Q_e}{n-2}}$, $Q_e = S_{yy} - \hat{b} S_{xy}$

拒绝域 $|t| > t_{\frac{\alpha}{2}}(n-2)$.

计算得 $S_{yy} = \sum_{i=1}^n y_i^2 - n\bar{y}^2 = 724.4$

$$Q_e = 66.3614, \quad \hat{\sigma} = 2.88,$$

$$|t| = 8.907 > 2.3060 = t_{0.025}^{(8)}$$

落入拒绝域, 故回归效果显著。

8. 解: $\bar{y}_1 = 8, \bar{y}_2 = 6.8, \bar{y}_3 = 7.2, \bar{y}_4 = 8.5$

$$\bar{y} = 7.6154.$$

方差分析表

方差来源	平方和	自由度	均方	F比
因素	4.8139	3	1.6	3.2
误差	4.5	9	0.5	
总和	9.3139	12		

$$F = 3.2 < 3.86 = F_{0.05}(3, 9)$$

未落入拒绝域, 故认为无显著影响.