

1. DEFINITIONS

All definitions given are for 2-dimensional inputs.

1.1. Assume.

$$(1) \quad [y^*, y_1^\delta, y_2^\delta, y \mid \sigma^2, l] \sim N(0, K)$$

$$(2) \quad K = \begin{bmatrix} K^{00}(x^*, x^*) & K^{10}(x_1^\delta, x^*)^T & K^{20}(x_2^\delta, x^*)^T & K^{00}(x, x^*)^T \\ K^{10}(x_1^\delta, x^*) & K^{11}(x_1^\delta, x_1^\delta) & K^{21}(x_2^\delta, x_1^\delta)^T & K^{01}(x, x_1^\delta)^T \\ K^{20}(x_2^\delta, x^*) & K^{21}(x_2^\delta, x_1^\delta) & K^{22}(x_2^\delta, x_2^\delta) & K^{02}(x, x_2^\delta)^T \\ K^{00}(x, x^*) & K^{01}(x, x_1^\delta) & K^{02}(x, x_2^\delta) & K^{00}(x, x) \end{bmatrix}$$

where

- $y = y(x)$: observed target function values (training set)
- $y^* = y(x^*)$: unobserved function values
- $y_k^\delta = \frac{\partial}{\partial x_k} y(x^\delta)$: vectors of partial derivatives in the k th input dimension
- σ^2 : constant variance parameter
- l : length-scale parameter
- $K^{d_1 d_2}(x_1, x_2)$: covariance functions. The superscripts indicate the argument where the derivative is taken (explicit formulas are in the following section).
 - Let $\dim(x) = ncol(x)$.
 - Then $d_1, d_2 \in \{0, 1, 2, \dots, \dim(x)\}$, where d_i indicates the dimension to which the derivative is taken for the i -th argument of $K(\cdot, \cdot)$

1.2. Covariance functions. Let

- $d = \dim(x)$
- $d_1, d_2 \in \{0, 1, 2, \dots, \dim(x)\}$, where d_i indicates the dimension to which the derivative is taken for the i -th argument of $K(\cdot, \cdot)$
- x_i is a row vector. That is, $x_i = \begin{bmatrix} x_{(i,1)} & x_{(i,2)} & \cdots & x_{(i,d)} \end{bmatrix}$
- $g(\cdot)$ is a correlation function.

1.2.1. $d_1 = d_2 = 0 \Rightarrow \text{"matern"/"sqexp"}$.

$$\begin{aligned} \text{Cov}(Y(x_1), Y(x_2)) &= K^{00}(x_1, x_2) \\ &= \sigma^2 \prod_{k=1}^d g(x_{(1,k)}, x_{(2,k)}; l_k) \end{aligned}$$

1.2.2. $d_1 > 0, d_2 = 0$ or $d_1 = 0, d_2 > 0 \Rightarrow \text{"matern1"/"sqexp1"}$.

$$\begin{aligned} \frac{\partial}{\partial x_{(1,i)}} \text{Cov}(Y(x_1), Y(x_2)) &= K^{i0}(x_1, x_2) \\ &= \sigma^2 \frac{\partial}{\partial x_{(1,i)}} \prod_{k=1}^d g(x_{(1,k)}, x_{(2,k)}; l_k) \\ &= \sigma^2 \left[\prod_{\substack{k=1 \\ k \neq i}}^d g(x_{(1,k)}, x_{(2,k)}; l_k) \right] \left[\frac{\partial}{\partial x_{(1,i)}} g(x_{(1,i)}, x_{(2,i)}; l_i) \right] \\ \frac{\partial}{\partial x_{(2,i)}} \text{Cov}(Y(x_1), Y(x_2)) &= K^{0i}(x_1, x_2) \\ &= -K^{i0}(x_1, x_2) \end{aligned}$$

1.2.3. $d_1, d_2 > 0 \Rightarrow \text{"matern2"/"sqexp2"}$.

$$\begin{aligned} \frac{\partial^2}{\partial x_{(2,j)} \partial x_{(1,i)}} \text{Cov}(Y(x_1), Y(x_2)) &= K^{ij}(x_1, x_2) \\ &= \begin{cases} \sigma^2 \left[\prod_{\substack{k=1 \\ k \neq i,j}}^d g(x_k, x'_k; l_k) \right] \left[\frac{\partial^2}{\partial x_{(2,i)} \partial x_{(1,i)}} g(x_{(1,i)}, x_{(2,i)}; l_i) \right], & i = j \\ \sigma^2 \left[\prod_{\substack{k=1 \\ k \neq i,j}}^d g(x_k, x'_k; l_k) \right] \left[\frac{\partial}{\partial x_{(2,j)}} g(x_{(1,j)}, x_{(2,j)}; l_j) \right] \left[\frac{\partial}{\partial x_{(1,i)}} g(x_{(1,i)}, x_{(2,i)}; l_i) \right], & i \neq j \end{cases} \end{aligned}$$

1.2.4. *Matern correlation function and its derivatives.* If $g(\cdot, \cdot)$ is from the Matern class of covariance functions with parameter $5/2$ there is a more simple form. The subscripts for the inputs in $g(\cdot, \cdot)$ are dropped for legibility.

Let $\theta = \sqrt{5}/l$:

$$(3) \quad g(x, x') = \left(1 + \theta |x - x'| + \frac{1}{3}\theta^2 |x - x'|^2\right) \exp\{-\theta |x - x'|\}$$

$$(4) \quad \frac{\partial}{\partial x} g(x, x') = -\frac{1}{3}\theta^2 (x - x') [1 + \theta |x - x'|] \exp\{-\theta |x - x'|\}$$

$$(5) \quad \frac{\partial^2}{\partial x' \partial x} g(x, x') = \frac{1}{3}\theta^2 [1 + \theta |x - x'| - \theta^2 (x - x')^2] \exp\{-\theta |x - x'|\}$$

See calculations of derivatives of $g(\cdot, \cdot)$ in section 3.

2. SETUP

2.1. The posterior density.

$$(6) \quad [l, \sigma^2, y^*, y_1^\delta, y_2^\delta | y] \propto [y, y^*, y_1^\delta, y_2^\delta | \sigma^2, l] [\sigma^2, l]$$

$$(7) \quad \propto [y^*, y_1^\delta, y_2^\delta | y, \sigma^2, l] [[y | \sigma^2, l]] [\sigma^2, l]$$

The goal is to evaluate densities on the RHS. Specifically,

$$[y^*, y_1^\delta, y_2^\delta | y, \sigma^2, l] \sim N(m, S)$$

due to the properties of the multivariate normal distribution (https://en.wikipedia.org/wiki/Multivariate_normal_distribution#Conditional_distributions). That is, the formulas from Kriging follow immediately.

To calculate m, S , start by dividing the covariance matrix K into the appropriate blocks:

$$(8) \quad K = \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix}, \text{ where}$$

$$(9) \quad \Sigma_{11} = \begin{bmatrix} K^{00}(x^*, x^*) & K^{10}(x_1^\delta, x^*)^T & K^{10}(x_2^\delta, x^*)^T \\ K^{10}(x_1^\delta, x^*) & K^{11}(x_1^\delta, x_1^\delta) & K^{11}(x_2^\delta, x_1^\delta)^T \\ K^{10}(x_2^\delta, x^*) & K^{11}(x_2^\delta, x_1^\delta) & K^{11}(x_2^\delta, x_2^\delta) \end{bmatrix}$$

$$(10) \quad \Sigma_{12} = \Sigma_{21}^T = \begin{bmatrix} K^{00}(x, x^*)^T \\ K^{01}(x, x_1^\delta)^T \\ K^{01}(x, x_2^\delta)^T \end{bmatrix} = \begin{bmatrix} K^{00}(x^*, x) \\ K^{10}(x_1^\delta, x) \\ K^{10}(x_2^\delta, x) \end{bmatrix}$$

$$(11) \quad \Sigma_{22} = K^{00}(x, x).$$

Then,

$$(12) \quad m = 0 + \Sigma_{12}\Sigma_{22}^{-1}(y - 0) = \begin{bmatrix} K^{00}(x^*, x) \\ K^{10}(x_1^\delta, x) \\ K^{10}(x_2^\delta, x) \end{bmatrix} K^{00}(x, x)^{-1}y$$

$$(13) \quad S = \Sigma_{11} - \Sigma_{12}\Sigma_{22}^{-1}\Sigma_{21}$$

2.2. Example. Given

$$(14) \quad y(x) = 11x_1^{10} + 9x_1^8 + 7x_1^6 + 10x_2^9 + 8x_2^7$$

$$(15) \quad \frac{\partial}{\partial x_1}y(x) = 110x_1^9 + 72x_1^7 + 42x_1^5$$

$$(16) \quad \frac{\partial}{\partial x_2}y(x) = 90x_2^8 + 56x_2^6$$

where

- y : 10×1
- x^* : 2×2
- x_1^δ : 8×2
- x_2^δ : 8×2

Then including the GP parameters σ^2 and l , there are 20 unknowns. Fixing GP parameters, we can calculate the covariance matrix S required for $[y^*, y_1^\delta, y_2^\delta | y, \sigma^2, l]$.

2.3. Problem. Assuming reasonable values for σ^2 and l , following the formula for S given above, the resulting matrix is negative definite. Adding small nuggets does not fix this.

It seems that if I use rather small values of l (all entries less than approximately 0.3), then this problem disappears.

3. DERIVATIVES FOR THE MATERN

Recall: $g(x, x') = [1 + \theta |x - x'| + \frac{1}{3}\theta^2 |x - x'|^2] \exp\{-\theta |x - x'|\}$

3.1. First derivative.

$$\begin{aligned}
(17) \quad \frac{\partial}{\partial x} g(x, x') &= \left[\theta \operatorname{sign}(x - x') + \frac{2}{3} \theta^2 (x - x') \right] \exp \{-\theta |x - x'|\} \\
(18) \quad &+ \left[1 + \theta |x - x'| + \frac{1}{3} \theta^2 (x - x')^2 \right] \exp \{-\theta |x - x'|\} (-\theta \operatorname{sign}(x - x')) \\
(19) \quad &= \exp \{-\theta |x - x'|\} \left[-\frac{1}{3} \theta^2 (x - x') - \frac{1}{3} \theta^3 (x - x')^2 \operatorname{sign}(x - x') \right] \\
(20) \quad &= \exp \{-\theta |x - x'|\} \left(-\frac{1}{3} \theta^2 (x - x') \right) [1 + \theta (x - x') \operatorname{sign}(x - x')] \\
(21) \quad &= -\frac{1}{3} \theta^2 (x - x') [1 + \theta |x - x'|] \exp \{-\theta |x - x'|\}
\end{aligned}$$

3.2. Second derivative.

$$\begin{aligned}
(22) \quad \frac{\partial^2}{\partial x' \partial x} g(x, x') &= \frac{\partial}{\partial x'} \left[-\frac{1}{3} \theta^2 (x - x') [1 + \theta |x - x'|] \exp \{-\theta |x - x'|\} \right] \\
(23) \quad &= \frac{\partial}{\partial x'} \left[\left[-\frac{1}{3} \theta^2 (x - x') - \frac{1}{3} \theta^3 (x - x')^2 \operatorname{sign}(x - x') \right] \exp \{-\theta |x - x'|\} \right] \\
(24) \quad &= \left[\frac{1}{3} \theta^2 - \frac{2}{3} \theta^3 (x - x') (-1) \operatorname{sign}(x - x') \right] \exp \{-\theta |x - x'|\} \\
(25) \quad &+ \left[-\frac{1}{3} \theta^2 (x - x') - \frac{1}{3} \theta^3 (x - x')^2 \operatorname{sign}(x - x') \right] \exp \{-\theta |x - x'|\} (\theta \operatorname{sign}(x - x')) \\
(26) \quad &= \exp \{-\theta |x - x'|\} \left[\frac{1}{3} \theta^2 + \frac{2}{3} \theta^3 |x - x'| - \frac{1}{3} \theta^3 |x - x'| - \frac{1}{3} \theta^4 (x - x')^2 \right] \\
(27) \quad &= \exp \{-\theta |x - x'|\} \left[\frac{1}{3} \theta^2 + \frac{1}{3} \theta^3 |x - x'| - \frac{1}{3} \theta^4 (x - x')^2 \right] \\
(28) \quad &= \frac{1}{3} \theta^2 \left[1 + \theta |x - x'| - \theta^2 (x - x')^2 \right] \exp \{-\theta |x - x'|\}
\end{aligned}$$