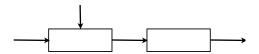
### **LECTURE 23**

- Readings: Section 9.1
   (not responsible for t-based confidence intervals, in pp. 471-473)
- Outline
- Classical statistics
- Maximum likelihood (ML) estimation
- Estimating a sample mean
- Confidence intervals (CIs)
- CIs using an estimated variance

### Classical statistics



- also for vectors x and  $\theta$ :  $p_{X_1,...,X_n}(x_1,...,x_n;\theta_1,...,\theta_m)$
- These are NOT conditional probabilities;  $\theta$  is NOT random
- mathematically: many models, one for each possible value of  $\theta$

### • Problem types:

– Hypothesis testing:

 $H_0: \theta = 1/2 \text{ versus } H_1: \theta = 3/4$ 

– Composite hypotheses:

 $H_0$ :  $\theta = 1/2$  versus  $H_1$ :  $\theta \neq 1/2$ 

- Estimation: design an **estimator**  $\hat{\Theta}$ , to keep estimation **error**  $\hat{\Theta} - \theta$  small

### Maximum Likelihood Estimation

- Model, with unknown parameter(s):  $X \sim p_X(x; \theta)$
- ullet Pick heta that "makes data most likely"

$$\hat{\theta}_{\mathsf{ML}} = \arg\max_{\theta} p_X(x;\theta)$$

• Compare to Bayesian MAP estimation:

$$\hat{\theta}_{\mathsf{MAP}} = \arg\max_{\theta} p_{\Theta\mid X}(\theta\mid x)$$

$$\hat{\theta}_{\mathsf{MAP}} = \arg\max_{\theta} \frac{p_{X|\Theta}(x|\theta)p_{\Theta}(\theta)}{p_{X}(x)}$$

• **Example:**  $X_1, \ldots, X_n$ : i.i.d., exponential( $\theta$ )

$$\max_{\theta} \prod_{i=1}^{n} \theta e^{-\theta x_i}$$

$$\max_{\theta} \left( n \log \theta - \theta \sum_{i=1}^{n} x_i \right)$$

$$\hat{\theta}_{\mathsf{ML}} = \frac{n}{x_1 + \dots + x_n}$$
  $\hat{\Theta}_n = \frac{n}{X_1 + \dots + X_n}$ 

# Desirable properties of estimators (should hold FOR ALL $\theta$ !!!)

- Unbiased:  $E[\hat{\Theta}_n] = \theta$
- exponential example, with n=1:  ${\rm E}[1/X_1] = \infty \neq \theta$  (biased)
- Consistent:  $\hat{\Theta}_n \to \theta$  (in probability)
  - exponential example:

$$(X_1 + \cdots + X_n)/n \to \mathbf{E}[X] = 1/\theta$$

– can use this to show that:

1

$$\hat{\Theta}_n = n/(X_1 + \dots + X_n) \to 1/\mathbb{E}[X] = \theta$$

• "Small" mean squared error (MSE)

$$E[(\hat{\Theta} - \theta)^{2}] = var(\hat{\Theta} - \theta) + (E[\hat{\Theta} - \theta])^{2}$$
$$= var(\hat{\Theta}) + (bias)^{2}$$

### Estimate a mean

•  $X_1, \dots, X_n$ : i.i.d., mean  $\theta$ , variance  $\sigma^2$   $X_i = \theta + W_i$ 

 $W_i$ : i.i.d., mean, 0, variance  $\sigma^2$ 

 $\hat{\Theta}_n = \text{sample mean} = M_n = \frac{X_1 + \dots + X_n}{n}$ 

## Properties:

- $E[\hat{\Theta}_n] = \theta$  (unbiased)
- WLLN:  $\hat{\Theta}_n \to \theta$  (consistency)
- MSE:  $\sigma^2/n$
- Sample mean often turns out to also be the ML estimate. E.g., if  $X_i \sim N(\theta, \sigma^2)$ , i.i.d.

The case of unknown  $\sigma$ 

- Option 1: use upper bound on  $\sigma$
- if  $X_i$  Bernoulli:  $\sigma \leq 1/2$
- ullet Option 2: use ad hoc estimate of  $\sigma$
- if  $X_i$  Bernoulli( $\theta$ ):  $\hat{\sigma} = \sqrt{\hat{\Theta}(1 \hat{\Theta})}$
- Option 3: Use generic estimate of the variance
- Start from  $\sigma^2 = \mathbf{E}[(X_i \theta)^2]$

$$\hat{\sigma}_n^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \theta)^2 \rightarrow \sigma^2$$

(but do not know  $\theta$ )

$$\hat{S}_n^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \hat{\Theta}_n)^2 \to \sigma^2$$

(unbiased:  $\mathbf{E}[\hat{S}_n^2] = \sigma^2$ )

# Confidence intervals (CIs)

- An estimate  $\hat{\Theta}_n$  may not be informative enough
- An  $1 \alpha$  confidence interval is a (random) interval  $\left[\widehat{\Theta}_n^-, \widehat{\Theta}_n^+\right]$ ,

s.t. 
$$P(\hat{\Theta}_n^- \le \theta \le \hat{\Theta}_n^+) \ge 1 - \alpha, \forall \theta$$

- often  $\alpha = 0.05$ , or 0.25, or 0.01
- interpretation is subtle
- CI in estimation of the mean  $\hat{\Theta}_n = (X_1 + \dots + X_n)/n$
- normal tables:  $\Phi(1.96) = 1 0.05/2$

$$\mathbf{P}\Big(\frac{|\hat{\Theta}_n - \theta|}{\sigma/\sqrt{n}} \le 1.96\Big) \approx 0.95$$
 (CLT)

$$\mathbf{P}\Big(\hat{\Theta}_n - \frac{1.96\,\sigma}{\sqrt{n}} \le \theta \le \hat{\Theta}_n + \frac{1.96\,\sigma}{\sqrt{n}}\Big) \approx 0.95$$

More generally: let z be s.t.  $\Phi(z) = 1 - \alpha/2$ 

$$P(\widehat{\Theta}_n - \frac{z\sigma}{\sqrt{n}} \le \theta \le \widehat{\Theta}_n + \frac{z\sigma}{\sqrt{n}}) \approx 1 - \alpha$$

6.041SC Probabilistic Systems Analysis and Applied Probability Fall 2013

For information about citing these materials or our Terms of Use, visit: http://ocw.mit.edu/terms.