

## Homework 5

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### Exercise 1.

I 分布表如下:

Y\X	0	1	2	3
0	$\frac{10}{220}$	$\frac{30}{220}$	$\frac{15}{220}$	$\frac{1}{220}$
1	$\frac{40}{220}$	$\frac{60}{220}$	$\frac{12}{220}$	0
2	$\frac{30}{220}$	$\frac{18}{220}$	0	0
3	$\frac{4}{220}$	0	0	0

II

$$P(X=1) = \sum_{y=0}^3 P(X=1, Y=y) = \frac{30}{220} + \frac{60}{220} + \frac{18}{220} = \frac{112}{220}$$

### Exercise 2.

$$\begin{aligned}
 P(a < X \leq b, c < Y \leq d) &= \iint_{[a,b] \times [c,d]} f(x,y) dx dy \\
 &= \int_a^b \int_c^d \frac{\partial^2 F}{\partial x \partial y}(x,y) dy dx \\
 &= \int_a^b \left( \frac{\partial F}{\partial x}(x,d) - \frac{\partial F}{\partial x}(x,c) \right) dx \\
 &= F(b,d) - F(b,c) - F(a,d) + F(a,c)
 \end{aligned}$$

### Exercise 3.

I

$$f(x,y) = \begin{cases} c & , x^2 + y^2 \leq 1 \\ 0 & , x^2 + y^2 > 1 \end{cases} \quad (1)$$

又由归一化条件,

$$\iint_{\mathbb{R}^2} f(x,y) dx dy = c\pi = 1$$

所以  $c = \frac{1}{\pi}$

II

$$\begin{aligned}
 f_X(x) &= \int_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} \frac{1}{\pi} dy = \begin{cases} \frac{2}{\pi} \sqrt{1-x^2} & , -1 \leq x \leq 1 \\ 0 & , |x| > 1 \end{cases} \\
 f_Y(y) &= \int_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} \frac{1}{\pi} dx = \begin{cases} \frac{2}{\pi} \sqrt{1-y^2} & , -1 \leq y \leq 1 \\ 0 & , |y| > 1 \end{cases}
 \end{aligned}$$

$$\text{III } P(R \leq r) = P(X^2 + Y^2 \leq r^2) = r^2$$

$$\text{IV } f(r) = \frac{dP(R \leq r)}{dr} = 2r, \text{ 那么}$$

$$E(R) = \int_0^1 r \cdot 2r dr = \frac{2}{3}$$

**Exercise 4.**

$$\begin{aligned} f_X(x) &= \int_{-\infty}^{\infty} f(x, y) dy \\ &= \int_{-\infty}^{\infty} \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)} \left[ \frac{(x-\mu_1)^2}{\sigma_1^2} - 2\rho \frac{(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2} + \frac{(y-\mu_2)^2}{\sigma_2^2} \right]\right) dy \end{aligned}$$

将指数项重新整理：

$$\begin{aligned} & \frac{(x-\mu_1)^2}{\sigma_1^2} - 2\rho \frac{(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2} + \frac{(y-\mu_2)^2}{\sigma_2^2} \\ &= \frac{(x-\mu_1)^2}{\sigma_1^2} + \frac{1}{\sigma_2^2} \left[ (y-\mu_2)^2 - 2\rho \frac{\sigma_2}{\sigma_1} (x-\mu_1)(y-\mu_2) \right] \\ &= \frac{(x-\mu_1)^2}{\sigma_1^2} + \frac{1}{\sigma_2^2} \left[ (y-\mu_2) - \rho \frac{\sigma_2}{\sigma_1} (x-\mu_1) \right]^2 - \frac{\rho^2}{\sigma_1^2} (x-\mu_1)^2 \\ &= \frac{(x-\mu_1)^2}{\sigma_1^2} (1-\rho^2) + \frac{1}{\sigma_2^2} \left[ (y-\mu_2) - \rho \frac{\sigma_2}{\sigma_1} (x-\mu_1) \right]^2 \end{aligned}$$

代入原积分：

$$f_X(x) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left(-\frac{(x-\mu_1)^2}{2\sigma_1^2}\right) \int_{-\infty}^{\infty} \exp\left(-\frac{1}{2\sigma_2^2(1-\rho^2)} \left[ (y-\mu_2) - \rho \frac{\sigma_2}{\sigma_1} (x-\mu_1) \right]^2\right) dy$$

令  $u = (y-\mu_2) - \rho \frac{\sigma_2}{\sigma_1} (x-\mu_1)$ ，则  $dy = du$ ，且当  $y$  从  $-\infty$  到  $\infty$  变化时， $u$  也从  $-\infty$  到  $\infty$  变化。所以：

$$\begin{aligned} f_X(x) &= \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left(-\frac{(x-\mu_1)^2}{2\sigma_1^2}\right) \int_{-\infty}^{\infty} \exp\left(-\frac{u^2}{2\sigma_2^2(1-\rho^2)}\right) du \\ &= \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left(-\frac{(x-\mu_1)^2}{2\sigma_1^2}\right) \cdot \sigma_2 \sqrt{2\pi(1-\rho^2)} \\ &= \frac{1}{\sqrt{2\pi}\sigma_1} \exp\left(-\frac{(x-\mu_1)^2}{2\sigma_1^2}\right) \end{aligned}$$

同理，

$$f_Y(y) = \frac{1}{\sqrt{2\pi}\sigma_2} \exp\left(-\frac{(y-\mu_2)^2}{2\sigma_2^2}\right)$$

因此，二元正态分布的边际分布仍然是正态分布，且：

$$X \sim N(\mu_1, \sigma_1^2)$$

$$Y \sim N(\mu_2, \sigma_2^2)$$

**Exercise 5.**

$$f_{Y|X}(y|x) = \frac{\frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)} \left[\frac{(x-\mu_1)^2}{\sigma_1^2} - 2\rho\frac{(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2} + \frac{(y-\mu_2)^2}{\sigma_2^2}\right]\right)}{\frac{1}{\sqrt{2\pi}\sigma_1} \exp\left(-\frac{(x-\mu_1)^2}{2\sigma_1^2}\right)}$$

将指数项重新整理：

$$\begin{aligned} & \frac{1}{2(1-\rho^2)} \left[ \frac{(x-\mu_1)^2}{\sigma_1^2} - 2\rho\frac{(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2} + \frac{(y-\mu_2)^2}{\sigma_2^2} \right] - \frac{(x-\mu_1)^2}{2\sigma_1^2} \\ &= \frac{1}{2(1-\rho^2)} \left[ \frac{(x-\mu_1)^2}{\sigma_1^2} - 2\rho\frac{(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2} + \frac{(y-\mu_2)^2}{\sigma_2^2} \right] - \frac{(x-\mu_1)^2(1-\rho^2)}{2\sigma_1^2(1-\rho^2)} \\ &= \frac{1}{2(1-\rho^2)} \left[ \frac{(x-\mu_1)^2\rho^2}{\sigma_1^2} - 2\rho\frac{(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2} + \frac{(y-\mu_2)^2}{\sigma_2^2} \right] \\ &= \frac{1}{2(1-\rho^2)} \left[ \frac{1}{\sigma_2^2} \left( y - \mu_2 - \frac{\rho\sigma_2}{\sigma_1}(x - \mu_1) \right)^2 \right] \end{aligned}$$

代入条件密度表达式：

$$\begin{aligned} f_{Y|X}(y|x) &= \frac{1}{\sqrt{2\pi}\sigma_2\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)\sigma_2^2} \left( y - \mu_2 - \frac{\rho\sigma_2}{\sigma_1}(x - \mu_1) \right)^2\right) \\ &= \frac{1}{\sqrt{2\pi}\sigma_2\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2\sigma_2^2(1-\rho^2)} \left( y - \mu_2 - \frac{\rho\sigma_2}{\sigma_1}(x - \mu_1) \right)^2\right) \end{aligned}$$

令  $\mu_{Y|X=x} = \mu_2 + \rho\frac{\sigma_2}{\sigma_1}(x - \mu_1)$  和  $\sigma_{Y|X}^2 = \sigma_2^2(1 - \rho^2)$ ，则：

$$f_{Y|X}(y|x) = \frac{1}{\sqrt{2\pi}\sigma_{Y|X}} \exp\left(-\frac{(y - \mu_{Y|X=x})^2}{2\sigma_{Y|X}^2}\right)$$

这表明在给定  $X = x$  的条件下， $Y$  的条件分布仍是正态分布：

$$Y|X = x \sim N\left(\mu_2 + \rho\frac{\sigma_2}{\sigma_1}(x - \mu_1), \sigma_2^2(1 - \rho^2)\right)$$

同理可得，在给定  $Y = y$  的条件下， $X$  的条件分布也是正态分布：

$$X|Y = y \sim N\left(\mu_1 + \rho\frac{\sigma_1}{\sigma_2}(y - \mu_2), \sigma_1^2(1 - \rho^2)\right)$$

**Exercise 6.**

I  $P(X \leq x, Y \leq y) = 2xy$ , 所以

$$f(x, y) = \begin{cases} 2 & , 0 \leq x \leq 1, 0 \leq y \leq 1 \\ 0 & , \text{其他} \end{cases}$$

II

$$f_Y(y) = \int_0^{1-y} 2dx = \begin{cases} 2(1-y) & , 0 \leq y \leq 1 \\ 0 & , \text{其他} \end{cases}$$

III

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} = \begin{cases} \frac{1}{1-y} & , 0 \leq x \leq 1-y, 0 < y < 1 \\ 0 & , \text{其他} \end{cases}$$

**Exercise 7.**

若  $X_1 \sim P(\lambda_1), X_2 \sim P(\lambda_2)$ ,  
那么

$$\begin{aligned} P(X_1 + X_2 = k) &= \sum_{i=0}^k P(X_1 = i)P(X_2 = k-i) \\ &= \sum_{i=0}^k \frac{\lambda_1^i}{i!} \exp(-\lambda_1) \frac{\lambda_2^{k-i}}{(k-i)!} \exp(-\lambda_2) \\ &= \exp(-\lambda_1 - \lambda_2) \sum_{i=0}^k \frac{\lambda_1^i \lambda_2^{k-i}}{i!(k-i)!} \\ &= \frac{(\lambda_1 + \lambda_2)^k}{k!} \exp(-\lambda_1 - \lambda_2) \end{aligned}$$

即  $X_1 + X_2 \sim P(\lambda_1 + \lambda_2)$ .  
又由条件概率公式,

$$\begin{aligned} P(X_1 = k | X_1 + X_2 = n) &= \frac{P(X_1 = k, X_1 + X_2 = n)}{P(X_1 + X_2 = n)} \\ &= \frac{P(X_1 = k)P(X_2 = n-k)}{P(X_1 + X_2 = n)} \\ &= \frac{\frac{\lambda_1^k}{k!} \exp(-\lambda_1) \frac{\lambda_2^{n-k}}{(n-k)!} \exp(-\lambda_2)}{\frac{(\lambda_1 + \lambda_2)^n}{n!} \exp(-\lambda_1 - \lambda_2)} \\ &= \frac{n!}{k!(n-k)!} \left( \frac{\lambda_1}{\lambda_1 + \lambda_2} \right)^k \left( \frac{\lambda_2}{\lambda_1 + \lambda_2} \right)^{n-k} \end{aligned}$$

这表明  $X_1 | X_1 + X_2 = n \sim B(n, \frac{\lambda_1}{\lambda_1 + \lambda_2})$ .

合并泊松过程的每个事件按比例分配给不同来源, 这符合二项分布的形式. (*Poisson Splitting Property*)

**Exercise 8.**

令  $X$  和  $Y$  分别表示甲和乙到达的时间, 以分钟为单位, 以下午 1 点为起点, 那么  $X$  和  $Y$  是相互独立的随机变量, 且均服从  $(0, 60)$  上的均匀分布. (联合分布即相乘)

所求概率为  $P\{X + 10 < Y\} + P\{Y + 10 < X\}$ . 根据对称性, 它等于  $2P\{X + 10 < Y\}$ , 而:

$$\begin{aligned}
 2P\{X + 10 < Y\} &= 2 \iint_{x+10 < y} f(x, y) dx dy \\
 &= 2 \iint_{x+10 < y} f_X(x) f_Y(y) dx dy \\
 &= 2 \int_{10}^{60} \int_0^{y-10} \left(\frac{1}{60}\right)^2 dx dy \\
 &= \frac{2}{60^2} \int_{10}^{60} (y - 10) dy \\
 &= \frac{25}{36}
 \end{aligned}$$

**Exercise 9.**

I

$$H_X(x) = \lim_{y \rightarrow +\infty} H(x, y) = F(x), \quad H_Y(y) = \lim_{x \rightarrow +\infty} H(x, y) = G(y)$$

II  $X, Y$  相互独立表明  $H(x, y) = F(x)G(y)$ , 即

$$\theta[1 - F(x)][1 - G(y)] = 0, \quad \forall x, y.$$

又  $F(x)$  和  $G(y)$  可以取值在  $(0, 1)$  内,  $[1 - F(x)][1 - G(y)]$  不总为零, 故  $\theta = 0$ .

III 在  $F$  和  $G$  分别取  $[0, 1]$  上的均匀分布, 并分别选取  $\theta = -1$  和  $\theta = 1$  即可构造.

**Exercise 10.**

$$H(X, Y) = C(F(X), G(Y))$$

那么, 由 *Copula* 函数的边际一致性,

$$H_X(x) = \lim_{y \rightarrow +\infty} C(F(x), G(y)) = C(F(x), 1) = F(x)$$

$$H_Y(y) = \lim_{x \rightarrow +\infty} C(F(x), G(y)) = C(1, G(y)) = G(y)$$

**Exercise 11.****I 情形 1:**  $X$  和  $Y$  均为离散型随机变量

全概率公式:

$$P(X = x) = \sum_y P(X = x|Y = y)P(Y = y)$$

贝叶斯公式:

$$P(Y = y|X = x) = \frac{P(X = x|Y = y)P(Y = y)}{P(X = x)} = \frac{P(X = x|Y = y)P(Y = y)}{\sum_t P(X = x|Y = t)P(Y = t)}$$

**II 情形 2:**  $X$  为连续型随机变量,  $Y$  为离散型随机变量

全概率公式:

$$f_X(x) = \sum_y f_{X|Y}(x|y)P(Y = y)$$

贝叶斯公式:

$$P(Y = y|X = x) = \frac{f_{X|Y}(x|y)P(Y = y)}{f_X(x)} = \frac{f_{X|Y}(x|y)P(Y = y)}{\sum_t f_{X|Y}(x|t)P(Y = t)}$$

**III 情形 3:**  $X$  为离散型随机变量,  $Y$  为连续型随机变量

全概率公式:

$$P(X = x) = \int_{\mathbb{R}} P(X = x|Y = y)f_Y(y)dy$$

贝叶斯公式:

$$f_{Y|X}(y|x) = \frac{P(X = x|Y = y)f_Y(y)}{P(X = x)} = \frac{P(X = x|Y = y)f_Y(y)}{\int_{\mathbb{R}} P(X = x|Y = t)f_Y(t)dt}$$

**IV 情形 4:**  $X$  和  $Y$  均为连续型随机变量

全概率公式:

$$f_X(x) = \int_{\mathbb{R}} f_{X|Y}(x|y)f_Y(y)dy$$

贝叶斯公式:

$$f_{Y|X}(y|x) = \frac{f_{X|Y}(x|y)f_Y(y)}{f_X(x)} = \frac{f_{X|Y}(x|y)f_Y(y)}{\int_{\mathbb{R}} f_{X|Y}(x|t)f_Y(t)dt}$$

**Exercise 12.**

I 由归一化条件,

$$\iint_{\mathbb{R}^2} f(x, y) dx dy = \iint_{x^2+y^2 \leq 1} \frac{c}{1+x^2+y^2} dx dy = \int_0^{2\pi} \int_0^1 \frac{c}{1+r^2} r dr d\theta = \int_0^{2\pi} \frac{c}{2} \ln 2 d\theta = \pi c \ln 2 = 1$$

$$\text{所以 } c = \frac{1}{\pi \ln 2}$$

II  $x^2 + y^2 \leq 1$

$$f_X(x) = \int_{\mathbb{R}} f(x, y) dy = \int_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} \frac{1}{\pi \ln 2 (1+x^2+y^2)} dy = \frac{2}{\pi \ln 2} \frac{1}{\sqrt{1+x^2}} \arctan \sqrt{\frac{1-x^2}{1+x^2}}$$

$$f_Y(y) = \int_{\mathbb{R}} f(x, y) dx = \int_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} \frac{1}{\pi \ln 2 (1+x^2+y^2)} dx = \frac{2}{\pi \ln 2} \frac{1}{\sqrt{1+y^2}} \arctan \sqrt{\frac{1-y^2}{1+y^2}}$$

显然

$$f_X(x)f_Y(y) = \frac{4}{\pi^2 \ln^2 2} \frac{1}{\sqrt{(1+x^2)(1+y^2)}} \arctan \sqrt{\frac{1-x^2}{1+x^2}} \arctan \sqrt{\frac{1-y^2}{1+y^2}} \neq f(x, y)$$

**Exercise 13.**

I 首先,  $f(x, y) = f_X(x)f_Y(y) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) \cdot \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{y^2}{2}\right) = \frac{1}{2\pi} \exp\left(-\frac{x^2+y^2}{2}\right)$

那么

$$\begin{aligned} \iint_{\mathbb{R}^2} g(x, y) dx dy &= \iint_{x^2+y^2 \leq 1} \frac{xy}{100} dx dy + \iint_{\mathbb{R}^2} f(x, y) dx dy \\ &= \iint_{\mathbb{R}^2} \frac{1}{2\pi} \exp\left(-\frac{x^2+y^2}{2}\right) dx dy \\ &= \int_0^{2\pi} \int_0^1 \frac{1}{2\pi} r \exp\left(-\frac{r^2}{2}\right) dr d\theta \\ &= \int_0^{+\infty} \exp\left(-\frac{r^2}{2}\right) d\left(\frac{r^2}{2}\right) \\ &= 1 \end{aligned}$$

II

$$f_U(x) = g_X(x) = \int_{\mathbb{R}} g(x, y) dy = \int_{\mathbb{R}} f(x, y) dy + \int_{x^2+y^2 \leq 1} \frac{xy}{100} dy = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right)$$

故  $U \sim N(0, 1)$ , 同理  $V \sim N(0, 1)$

显然  $g(x, y)$  的形式不符合二重正态分布.



#### Exercise 14.

```
1  import numpy as np
2  import matplotlib.pyplot as plt
3
4  y = np.random.uniform(0, 1, 10000)
5  x = -np.log(1 - y)
6
7  plt.hist(x, bins=50, density=True, alpha=0.6, color='pink')
8
9  lam = 1
10 xmin, xmax = 0, 8
11 x_pdf = np.linspace(xmin, xmax, 100)
12 p = lam * np.exp(-lam * x_pdf)
13 plt.plot(x_pdf, p, 'k', linewidth=2)
14
15 plt.xlim(xmin, xmax)
16 plt.ylim(0, 1)
17
18 plt.xlabel('x')
19 plt.ylabel('Density')
20 plt.title('Histogram of x and Exponential Distribution PDF')
21 plt.show()
```

