

## 11-7 作业

32. (1) 由题意知

$$\begin{aligned}
 & \text{Cov}(\alpha X + \beta Y, \alpha X - \beta Y) \\
 &= \text{Cov}(\alpha X, \alpha X - \beta Y) + \text{Cov}(\beta Y, \alpha X - \beta Y) \\
 &= \text{Cov}(\alpha X, \alpha X) - \text{Cov}(\alpha X, \beta Y) + \text{Cov}(\beta Y, \alpha X) - \text{Cov}(\beta Y, \beta Y) \\
 &= \alpha^2 \text{Cov}(X, X) - 0 + 0 - \beta^2 \text{Cov}(Y, Y) \\
 &= (\alpha^2 - \beta^2) \sigma^2
 \end{aligned}$$

(2)  $X, Y \sim N(\mu, \sigma^2)$ , 且相互独立. 则  $\alpha X + \beta Y$  与  $\alpha X - \beta Y$  都服从正态分布. 由正态分布的不相关与独立等价, 有

$$\begin{aligned}
 \text{Cov}(\alpha X + \beta Y, \alpha X - \beta Y) &= (\alpha^2 - \beta^2) \sigma^2 = 0 \\
 &\Rightarrow \alpha = \pm \beta.
 \end{aligned}$$

34. 法一:

(1)  $(X, Y)$  的联合密度为

$$f_{X,Y}(x, y) = \begin{cases} \frac{1}{2}, & |x| + |y| \leq 1, \\ 0 & \text{其他.} \end{cases}$$

由对称性可知,  $EX = EY = EXY = 0$ , 所以

$$\text{Cov}(X, Y) = EXY - EXEY = 0.$$

(2) 不独立。

证: 画图通过计算面积之比易得:  $P(X \leq -\frac{1}{2}) = P(Y \leq -\frac{1}{2}) = \frac{1}{8}$ ,  
 $P(X \leq -\frac{1}{2}, Y \leq -\frac{1}{2}) = 0$ , 所以

$$P(X \leq -\frac{1}{2}, Y \leq -\frac{1}{2}) \neq P(X \leq -\frac{1}{2})P(Y \leq -\frac{1}{2}).$$

法二:

$X, Y$  的边际密度为

$$f_X(x) = \begin{cases} \int_{-x-1}^{x+1} \frac{1}{2} dy = x+1, & -1 < x < 0, \\ \int_{x-1}^{-x+1} \frac{1}{2} dy = -x+1, & 0 < x < 1. \end{cases} \quad \text{类似地, } f_Y(y) = \begin{cases} y+1, & -1 < y < 0, \\ -y+1, & 0 < y < 1. \end{cases}$$

(1)  $\text{Cov}(X, Y) = E(XY) - E(X)E(Y)$ . 其中

$$E(X) = \int_{-1}^0 x \cdot (x+1)dx + \int_0^1 x \cdot (-x+1)dx = 0, \quad \text{同理}, E(Y) = 0$$

$$E(XY) = \int_{-1}^0 \int_{-x-1}^{x+1} \frac{xy}{2} dy dx + \int_0^1 \int_{x-1}^{-x+1} \frac{xy}{2} dy dx = 0$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = 0.$$

(2)  $f(x, y) \neq f_X(x)f_Y(y)$ ,  $X, Y$  不独立.

36. (1)  $Z = \pi X + (1 - \pi)Y$ , 则方差为

$$\begin{aligned} \text{Var}(Z) &= \text{Var}(\pi X + (1 - \pi)Y) \\ &= \pi^2 \text{Var}(X) + (1 - \pi)^2 \text{Var}(Y) + 2\pi(1 - \pi) \text{Cov}(X, Y) \\ &= \pi^2 \sigma^2 + (1 - \pi)^2 \sigma^2 + 2\pi(1 - \pi) \cdot \left(-\frac{1}{2}\right) \sigma^2 \\ &= (3\pi^2 - 3\pi + 1) \sigma^2 \cdot \pi \in (0, 1).. \end{aligned}$$

在  $\pi \in (0, 1)$  上,  $3\pi^2 - 3\pi + 1 = 3\pi(\pi - 1) + 1 < 1$  恒成立,  $\text{Var}(Z) < \sigma^2$ , 即证投资组合  $Z$  的风险小于将所有资本投资于其中一个的风险.

(2)  $f(\pi) = 3\pi^2 - 3\pi + 1, \pi \in (0, 1)$  的极小值点为  $\pi = 1/2$ . 所以使得投资组合风险最小的分配比例为  $\pi = 1/2$ .

40. 由  $N(t)$  的分布律可知,  $N(T)|T = t \sim \text{Poisson}(\lambda t)$ . 记  $T$  的概率密度为  $f_T(t)$ , 由重期望公式有

$$\begin{aligned} E[N(T)] &= E\left[E(N(T)|T)\right] = \int_0^\infty E(N(T)|T = t) f_T(t) dt \\ &= \int_0^\infty \lambda t \cdot f_T(t) dt = \lambda E(T) = \lambda a \end{aligned}$$

$$\begin{aligned} \text{同理, } E[N(T)^2] &= E\left[E(N(T)^2|T)\right] = \int_0^\infty E(N(T)^2|T = t) f_T(t) dt \\ &= \int_0^\infty (\lambda^2 t^2 + \lambda t) \cdot f_T(t) dt = \lambda^2 E(T^2) + \lambda E(T) = \lambda^2 (a^2 + b) + \lambda a \end{aligned}$$

$$\begin{aligned} E[TN(T)] &= E\left[E(TN(T)|T)\right] = \int_0^\infty E(TN(T)|T = t) f_T(t) dt \\ &= \int_0^\infty t E(N(T)|T = t) \cdot f_T(t) dt = \int_0^\infty \lambda t^2 \cdot f_T(t) dt \\ &= \lambda E(T^2) = \lambda (a^2 + b) \end{aligned}$$

则可求得:

(1)

$$\text{Cov}(T, N(T)) = E(TN(T)) - E(T)E(N(T)) = \lambda(a^2 + b) - a \cdot \lambda a = \lambda b.$$

(2)

$$\text{Var}(N(T)) = E[N(T)^2] - [E(N(T))]^2 = \lambda^2(a^2 + b) + \lambda a - \lambda^2 a^2 = \lambda^2 b + \lambda a.$$

**41.** 由教材例 3.14 可知:  $(X, Y) \sim N(\mu_1, \mu_2, \sigma_1^2, \sigma_2^2, \rho)$ , 则

$$X|Y = y \sim N\left(\mu_1 + \rho \frac{\sigma_1}{\sigma_2}(y - \mu_2), \sigma_1^2(1 - \rho^2)\right).$$

本题中  $(X, Y) \sim N(1, 2, 4, 9, 0.3)$ , 则

$$\begin{aligned} E(X|Y = 2) &= \mu_1 + \rho \frac{\sigma_1}{\sigma_2}(y - \mu_2) = 1 \\ E(XY^2 + Y|Y = 1) &= E(XY^2|Y = 1) + E(Y|Y = 1) \\ &= E(X|Y = 1) + E(Y|Y = 1) \\ &= 1 + 0.3(2/3)(1 - 2) + 1 \\ &= \frac{9}{5} \end{aligned}$$