BAYESIAN LEARNING - LECTURE 6

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LECTURE OVERVIEW

- ► Flexible nonlinear regression and splines
- ► Smoothness/shrinkage priors
- ► Gaussian process regression

POLYNOMIAL REGRESSION

Recall the linear regression model with a single covariate

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i, \qquad \varepsilon_i \stackrel{iid}{\sim} N(0, \sigma^2).$$

► Polynomial regression:

$$y_i = \beta_0 + \beta_1 x_i + \beta_2 x_i^2 + ... + \beta_k x_i^k + \varepsilon_i.$$

This can be written as a linear regression

$$y = X_P \beta + \varepsilon$$
,

where

$$X_P = (1, x, x^2, ..., x^k).$$

▶ The posterior of β is obtained like any linear regression.

SPLINES

► A more local basis: truncated polynomials

$$b_{ij} = \begin{cases} (x_i - k_j)^p & \text{if } x_i > k_j \\ 0 & \text{otherwise} \end{cases}$$

- \triangleright $k_1, k_2, ..., k_m$ are the knots.
- Splines are nonlinear in x, but linear in basis space

$$y = X_b \beta + \varepsilon$$
,

where X_b is the basis regression matrix

$$X_b = (b_1, ..., b_m).$$

Common extension

$$X_b = (1, x, b_1, ..., b_m).$$

 \triangleright Still just linear in X_b . Linear regression fitting.

SMOOTHNESS PRIOR FOR SPLINES

- Problem: too many knots leads to over-fitting.
- Solution: smoothness/shrinkage/regularization prior

$$\beta_i | \sigma^2 \stackrel{iid}{\sim} N\left(0, \frac{\sigma^2}{\lambda}\right)$$

- ▶ Larger λ gives smoother fit. Note: here we have $\Omega_0 = \lambda I$.
- Equivalent to a penalized likelihood:

$$-2 \cdot LogPost \propto RSS(\beta) + \lambda \beta' \beta$$

▶ Posterior mean gives ridge regression estimator

$$\tilde{\beta} = \left(X'X + \lambda I \right)^{-1} X' y$$

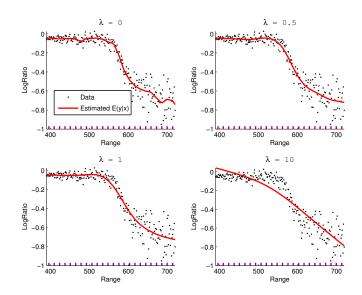
► Shrinkage toward zero

As
$$\lambda o \infty$$
, $ilde{eta} o 0$

▶ When X'X = I

$$\tilde{\beta} = \frac{1}{1+\lambda}\hat{\beta}_{OLS}$$

BAYESIAN SPLINE WITH SMOOTHNESS PRIOR



SMOOTHNESS PRIOR FOR SPLINES, CONT.

► The famous **Lasso** variable selection method is equivalent to using the posterior mode estimate under the prior:

$$\beta_i | \sigma^2 \stackrel{iid}{\sim} \text{Laplace} \left(0, \frac{\sigma^2}{\lambda} \right)$$

where the general Laplace density is

$$p(\beta_i) = \frac{1}{2b} \exp\left(-\frac{|\beta_i - \mu|}{b}\right)$$

- ► The Bayesian shrinkage prior is interpretable, not ad hoc.
- Laplace distribution have heavy tails.
- ▶ Laplace prior: many β_i close to zero, but some β_i may be very large.
- Normal distribution have light tails.
- Normal prior: most β_i are fairly equal in size, and no single β_i can be very much larger than the other ones.

ESTIMATING THE SHRINKAGE

- ▶ How do we determine the degree of smoothness, λ ? Cross-validation.
- ▶ Bayesian: I cannot specify $\lambda \Rightarrow \lambda$ is unknown \Rightarrow use a prior for λ .
- ▶ One possibility: $\lambda \sim Inv \chi^2(\eta_0, \lambda_0)$. The user specifies η_0 and λ_0 .
- ▶ Alternative approach: specify the prior on the *degrees of freedom*.
- ► Hierarchical setup:

$$\begin{aligned} y|\beta, x &\sim \textit{N}(x'\beta, \sigma^2) \\ \beta|\sigma^2 &\sim \textit{N}\left(\left(\begin{array}{c} 0\\ 0 \end{array}\right), \sigma^2\left(\begin{array}{cc} \delta_0^{-1}\textit{I}_q & 0\\ 0 & \lambda^{-1}\textit{I}_m \end{array}\right)\right) \\ \sigma^2 &\sim \textit{Inv} - \chi^2(\nu_0, \sigma_0^2) \\ \lambda &\sim \textit{Inv} - \chi^2(\eta_0, \lambda_0) \end{aligned}$$

Note: different shrinkage on the original q covariates (δ_0) and the covariates that comes from the knots (λ).

ESTIMATING THE SHRINKAGE, CONT.

► Joint posterior

$$p(\beta, \sigma^2, \lambda | y, x) = p(\beta, \sigma^2 | \lambda, y, x) p(\lambda | y, x)$$

where

$$p(\lambda|y,x) = \int \int p(\beta,\sigma,\lambda|y,x) d\beta d\sigma^2$$

is the marginal posterior of λ .

► The conditional posterior $p(\beta, \sigma^2 | \lambda, y, x)$ is a special case linear regression with conjugate prior $\mu_0 = (0, 0)'$ and

$$\Omega_0 = \left(\begin{array}{cc} \delta_0 I_q & 0\\ 0 & \lambda I_m \end{array}\right)$$

SUMMARY OF THE POSTERIOR WITH NORMAL SHRINKAGE PRIOR

▶ The joint posterior of β , σ^2 and λ is

$$\begin{split} \beta|\sigma^2, \lambda, y &\sim \textit{N}\left(\mu_n, \Omega_n^{-1}\right) \\ \sigma^2|\lambda, y &\sim \textit{Inv} - \chi^2\left(\nu_n, \sigma_n^2\right) \\ \rho(\lambda|y) &\propto \sqrt{\frac{|\Omega_0|}{|X'X + \Omega_0|}} \left(\frac{\nu_n \sigma_n^2}{2}\right)^{-\nu_n/2} \cdot \rho(\lambda) \end{split}$$

where $p(\lambda)$ is the prior for λ , and

$$\mu_n = (X'X + \Omega_0)^{-1} X'y$$

$$\Omega_n = X'X + \Omega_0$$

$$\nu_n = \nu_0 + n$$

$$\nu_n \sigma_n^2 = \nu_0 \sigma_0^2 + y'y - \mu_n' \Omega_n \mu_n$$

BAYESIAN VARIABLE SELECTION AND ESTIMATING KNOT LOCATIONS

- ▶ Selecting among a set of fixed knots $k_1, ..., k_m$ is a variable selection problem. More on Bayesian variable selection in the last module.
- ► The location of the knots can be treated as unknown, and estimated from the data.
- ▶ The joint posterior of parameters and knot locations

$$p(\beta, \sigma^2, \lambda, \xi_1, ..., \xi_q | y, x)$$

where ξ_i is the location of the *i*th knot.

► Posterior is complex but can be sampled from by Markov Chain Monte Carlo (MCMC). Li and Villani (2013, SJS).

NON-PARAMETRIC REGRESSION

► Linear regression

$$y = \beta \cdot x + \varepsilon$$

where $\varepsilon \sim N(0, \sigma^2)$ and iid over observations.

► Nonlinear regression

$$y = f(x) + \varepsilon$$

where $f(\cdot)$ is some nonlinear function (ex $f(x) = \beta_0 + \beta_1 x + \beta_2 x^2$).

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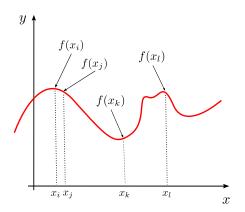
Nonlinear regression

$$y = f(x) + \varepsilon$$

where $f(\cdot)$ is some nonlinear function (ex $f(x) = \beta_0 + \beta_1 x + \beta_2 x^2$).

- ▶ Non-parametric regression: avoiding a parametric form for $f(\cdot)$.
- ► How do we put a prior over a set of functions?
- ▶ Restrict attention to a grid of (ordered) x-values: $x_1, x_2, ..., x_k$.
- We can now put a joint prior on the k function values: $f(x_1), f(x_2), ..., f(x_k)$.

Nonparametric = one parameter for every x!



NONPARAMETRIC REGRESSION - SMOOTH INTERPOLATION

► Treat all *n* ordinates as unknown parameters:

$$f(x_i) = \gamma_i$$
.

- ▶ Problem: too many parameters. Estimated curve wiggles way too much.
- ▶ Solution: use a (multivariate) prior on $\gamma = (\gamma_1, ..., \gamma_n)'$ that carries the info that the regression curve is smooth:

if x_i and x_k are close then γ_i is close to γ_k

Order the data with respect to covariates and assign the prior

$$p(\gamma_i|\gamma_{i-1}) \sim N(\gamma_{i-1}, \tau_0^2 \cdot |x_i - x_{i-1}|)$$
, for $i = 2, ..., n$.

▶ The hyperparameter τ_0^2 controls the degree of prior smoothness.

GAUSSIAN PROCESS REGRESSION

Generalization of smooth interpolation. Multivariate normal (Gaussian) prior:

$$\begin{pmatrix} f(x_1) \\ \vdots \\ f(x_k) \end{pmatrix} \sim N(\mathbf{m}, \mathbf{K})$$

▶ But how do we specify the $k \times k$ covariance matrix K?

$$Cov\left(f(x_p),f(x_q)\right)$$

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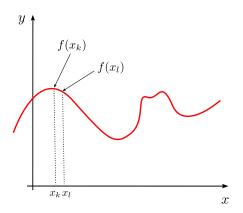
$$Cov\left(f(x_p),f(x_q)\right)$$

Squared exponential covariance function

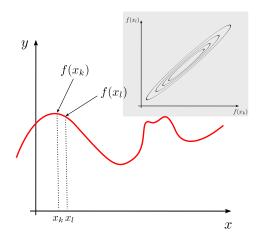
$$Cov(f(x_p), f(x_q)) = K(x_p, x_q) = \sigma_f^2 \exp\left(-\frac{1}{2} \left(\frac{x_p - x_q}{\ell}\right)^2\right)$$

- ▶ The covariance between $f(x_p)$ and $f(x_q)$ is a function of x_p and x_q .
- ▶ Nearby x's have highly correlated function ordinates f(x).
- ▶ We can compute $Cov(f(x_p), f(x_q))$ for any x_p and x_q .

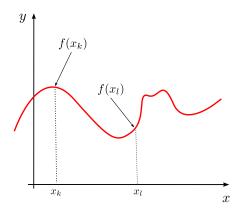
SMOOTH FUNCTION - POINTS NEARBY



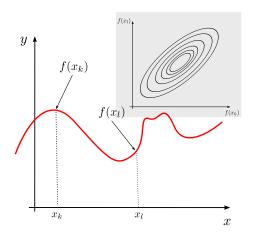
SMOOTH FUNCTION - POINTS NEARBY



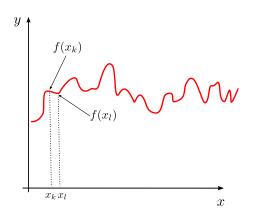
SMOOTH FUNCTION - POINTS FAR APART



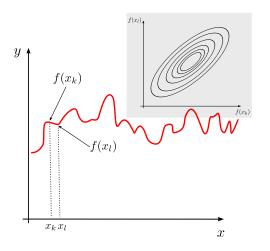
SMOOTH FUNCTION - POINTS FAR APART



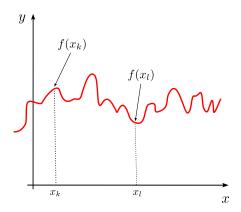
JAGGED FUNCTION - POINTS NEARBY



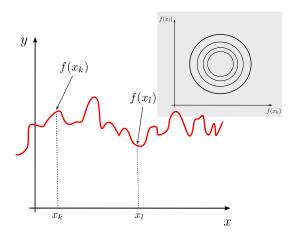
JAGGED FUNCTION - POINTS NEARBY



JAGGED FUNCTION - POINTS FAR APART



JAGGED FUNCTION - POINTS FAR APART



DEFINITION

A Gaussian process (GP) is a collection of random variables, any finite number of which have a multivariate Gaussian distribution.

- ► A Gaussian process is really a **probability distribution over functions** (curves). This is exactly what we want! No need for a grid!
- ▶ A GP is completely specified by a mean and a covariance function

$$m(x) = \mathrm{E}\left[f(x)\right]$$

$$K(x,x') = E\left[\left(f(x) - m(x) \right) \left(f(x') - m(x') \right) \right]$$

for any two inputs x and x' (note: this is *not* the transpose here).

► A Gaussian process (prior) is denoted by

$$f(x) \sim GP(m(x), K(x, x'))$$

Example:

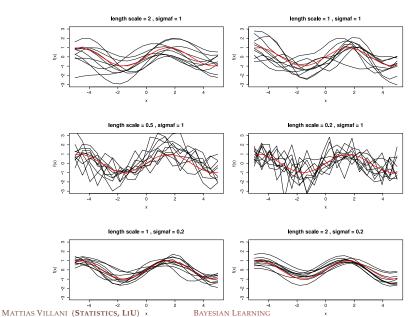
$$m(x) = \sin(x)$$
 $K(x, x') = \sigma_f^2 \exp\left(-\frac{1}{2}\left(\frac{x_p - x_q}{\ell}\right)^2\right)$

where l > 0 is the length scale.

- ▶ Larger I gives more smoothness in f(x).
- ▶ Simulate draw from $f(x) \sim GP(m(x), K(x, x'))$ over any grid $x_* = (x_1, ..., x_n)$ by using that

$$f(x_*) \sim N(m(x_*), K(x_*, x_*))$$

SIMULATING A GP - SINE MEAN AND SE KERNEL



Model

$$y_i = f(x_i) + \varepsilon_i, \quad \varepsilon \stackrel{iid}{\sim} N(0, \sigma^2)$$

Prior

$$f(x) \sim GP(0, K(x, x'))$$

- ▶ You have observed the data: $\mathbf{x} = (x_1, ..., x_n)'$ and $\mathbf{y} = (y_1, ..., y_n)'$.
- ▶ Goal: the posterior of $f(\cdot)$ over a grid of x-values: $\mathbf{f}_* = \mathbf{f}(\mathbf{x}_*)$.

Model

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- ▶ Intermediate step: joint distribution of y and f_{*}

$$\left(\begin{array}{c} y \\ f_* \end{array}\right) \sim \textit{N} \left\{ \left(\begin{array}{c} 0 \\ 0 \end{array}\right), \left[\begin{array}{ccc} \textit{K}(x,x) + \sigma^2 \textit{I} & \textit{K}(x,x_*) \\ \textit{K}(x_*,x) & \textit{K}(x_*,x_*) \end{array}\right] \right\}$$

Model

$$y_i = f(x_i) + \varepsilon_i, \quad \varepsilon \stackrel{iid}{\sim} N(0, \sigma^2)$$

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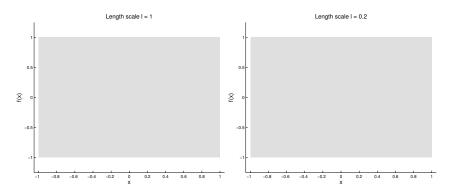
$$\left(\begin{array}{c} \mathbf{y} \\ \mathbf{f}_* \end{array}\right) \sim \textit{N}\left\{ \left(\begin{array}{c} \mathbf{0} \\ \mathbf{0} \end{array}\right), \left[\begin{array}{ccc} \textit{K}(\mathbf{x},\mathbf{x}) + \sigma^2 \textit{I} & \textit{K}(\mathbf{x},\mathbf{x}_*) \\ \textit{K}(\mathbf{x}_*,\mathbf{x}) & \textit{K}(\mathbf{x}_*,\mathbf{x}_*) \end{array}\right] \right\}$$

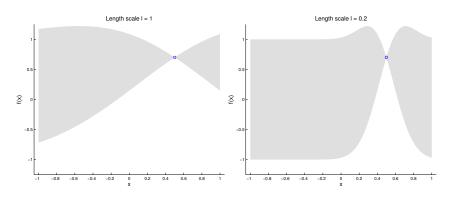
► The posterior

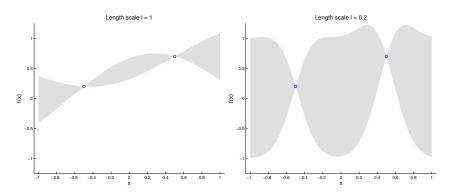
$$f_*|x,y,x_* \sim \textit{N}\left(\overline{f}_*, \text{cov}(f_*)\right)$$

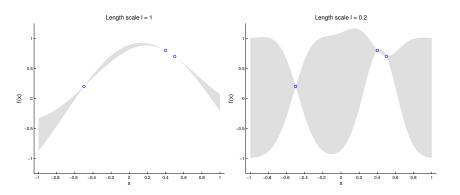
$$\mathbf{\bar{f}}_* = K(\mathbf{x}_*, \mathbf{x}) \left[K(\mathbf{x}, \mathbf{x}) + \sigma^2 I \right]^{-1} \mathbf{y}$$

$$cov(\mathbf{f}_*) = K(\mathbf{x}_*, \mathbf{x}_*) - K(\mathbf{x}_*, \mathbf{x}) \left[K(\mathbf{x}, \mathbf{x}) + \sigma^2 I \right]^{-1} K(\mathbf{x}, \mathbf{x}_*)$$

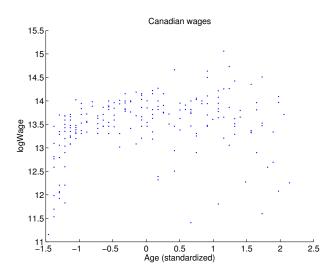








EXAMPLE - CANADIAN WAGES



Posterior of F - $\ell = 0.2, 0.5, 1, 2$

