Ta)
$$X_1,..., X_n \mid \theta$$
, $\sigma^2 \sim N \left(\theta, \sigma^2\right)$

Thrown

Thrown

Posterior: (implicit conditions) on θ)

P($\sigma^2 \mid X_1,..., X_n$) $\propto P(X_1,..., X_n \mid \sigma^2) P(\sigma^2)$

= $\prod_{i=1}^{n} \frac{1}{(2\pi\sigma^2)^{N_2}} \exp\left(-\frac{1}{2\sigma^2}(X_i - \theta)^2\right) \cdot P(\sigma^2)$

Figure $S^2 = \frac{\sum_{i=1}^{n} (X_i - \theta)^2}{N} \exp\left(-\frac{NS^2}{2\sigma^2}\right)$
 $= \exp\left(-\frac{NS^2}{2\sigma^2}\right) \exp\left(-\frac{NS^2}{2\sigma^2}\right) \exp\left(-\frac{NS^2}{2\sigma^2}\right)$
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Non-informative : Vo > 0

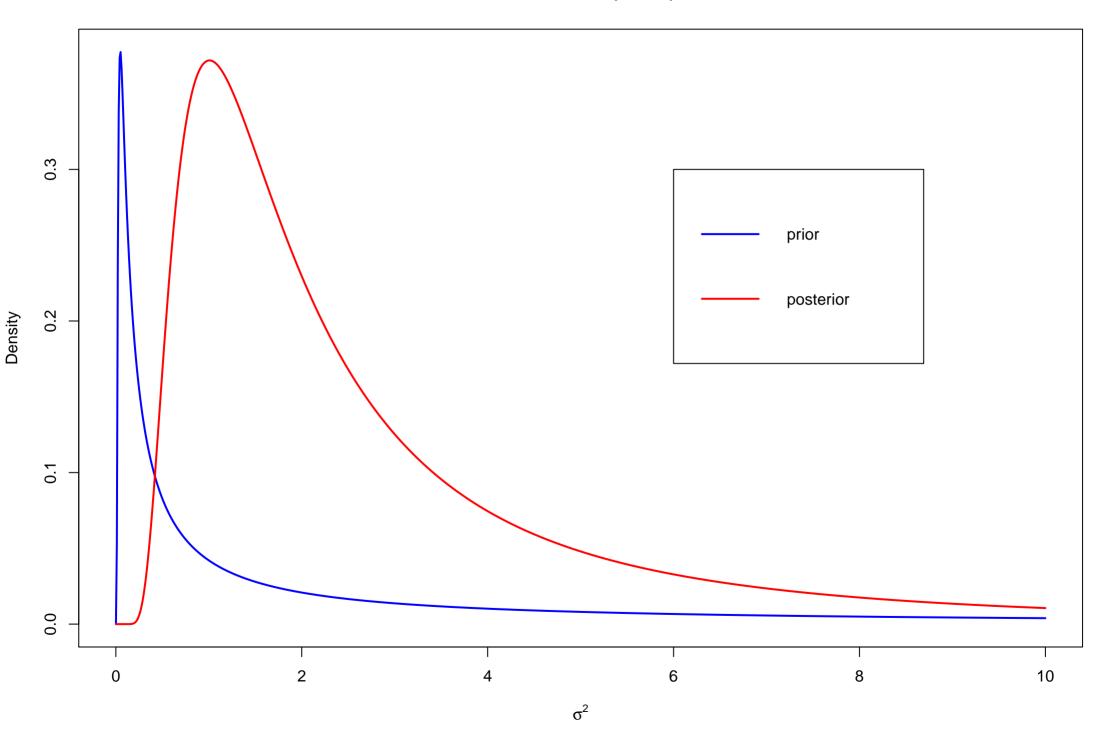
why is this non-informative?

Reason 1: Vn becomes u

Reason 2: $\ln \sqrt{\chi^2} (v_0, o_0^1)$ becomes $\frac{7}{\sigma^2}$ when $v_0 \rightarrow 0$.

Note that as $v_0 \Rightarrow 0$ the posterior approaches the $Inv X^2 (u, s^2)$ density. So,

02 (X1, X2, X3 ~ Inv X2 (3, 1.68)



40 Solutions

Problems of Chapter 6

6.1 Prediction of Bernoulli data

The predictive distribution of x_{n+1} given the first n trials $(x_{1:n})$ is

$$p(x_{n+1}|x_{1:n}) = \int p(x_{n+1}|\theta)p(\theta|x_{1:n})d\theta \qquad x_{n+1} \text{ is indep. of } x_{1:n} \text{ given } \theta$$

$$= \int \theta^{x_{n+1}}(1-\theta)^{1-x_{n+1}}p(\theta|x_{1:n})d\theta \qquad \theta|x_{1:n} \sim \text{Beta}(\alpha+s,\beta+f)$$

$$= \int \theta^{x_{n+1}}(1-\theta)^{1-x_{n+1}} \frac{\Gamma(\alpha+\beta+n)}{\Gamma(\alpha+s)\Gamma(\beta+f)} \theta^{\alpha+s-1}(1-\theta)^{\beta+f-1}d\theta$$

$$= \frac{\Gamma(\alpha+\beta+n)}{\Gamma(\alpha+s)\Gamma(\beta+f)} \int \theta^{x_{n+1}+\alpha+s-1}(1-\theta)^{1-x_{n+1}+\beta+f-1}d\theta$$

$$= \frac{\Gamma(\alpha+\beta+n)}{\Gamma(\alpha+s)\Gamma(\beta+f)} \frac{\Gamma(x_{n+1}+\alpha+s)\Gamma(1-x_{n+1}+\beta+f)}{\Gamma(1+\alpha+\beta+n)}$$

$$= \frac{\Gamma(x_{n+1}+\alpha+s)\Gamma(1-x_{n+1}+\beta+f)}{\Gamma(\alpha+s)\Gamma(\beta+f)(\alpha+\beta+n)} \text{ using } \Gamma(y+1) = y\Gamma(y)$$

So,

$$p(x_{n+1} = 1 | x_{1:n}) = \frac{\Gamma(1 + \alpha + s)}{\Gamma(\alpha + s)(\alpha + \beta + n)} = \frac{(\alpha + s)\Gamma(\alpha + s)}{\Gamma(\alpha + s)(\alpha + \beta + n)} = \frac{\alpha + s}{\alpha + \beta + n}$$

and therefore [since $p(x_{n+1} = 0|x_{1:n}) = 1 - p(x_{n+1} = 1|x_{1:n})$]

$$p(x_{n+1} = 0|x_{1:n}) = \frac{\beta + f}{\alpha + \beta + n}.$$

The predictive distribution is therefore

$$x_{n+1}|x_{1:n} \sim \operatorname{Bern}\left(\frac{\alpha+s}{\alpha+\beta+n}\right).$$

7.1 Umbrella decision

(a) Let x_{11} be the binary variable indicating rain on the 11th day. From Problem 6.1, the predictive distribution for the (n+1)th Bernoulli trial is

$$x_{n+1}|x_{1:n} \sim \operatorname{Bern}\left(\frac{\alpha+s}{\alpha+\beta+n}\right).$$

and the predictive probability for rain is therefore here

$$\Pr(x_{11} = 1 | x_{1:10}) = \frac{1+2}{1+1+10} = 0.25.$$

The expected utility from the decision to bring the umbrella is then

 $EU_{\rm bring} = \Pr({\rm sunny}) \cdot U({\rm bring, sunny}) + \Pr({\rm rain}) \cdot U({\rm bring, rain}) = 0.75 \cdot 20 + 0.25 \cdot 10 = 17.5$ and the expected utility of leaving the umbrella at home is

$$EU_{leave} = Pr(sunny) \cdot U(leave, sunny) + Pr(rain) \cdot U(leave, rain) = 0.75 \cdot 50 + 0.25 \cdot (-50) = 25.0.$$

The expected utility is therefore maximized by leaving the umbrella at home. This is the Bayesian decision.

- (b) Figure 15.1 shows how the optimal Bayesian decision varies for different combinations of the prior hyperparameters.
- (c) Figure 15.2 shows how the optimal Bayesian decision varies for different combinations of the prior hyperparameters when s=16 and f=64.

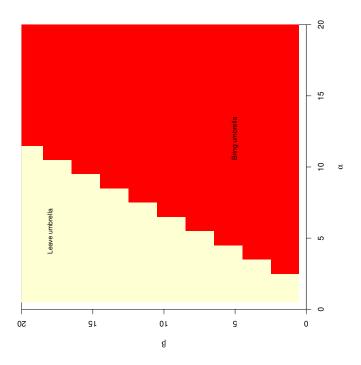


Fig. 15.1. How the Bayesian decision depends on the prior hyperparameters when s=2 and f=8

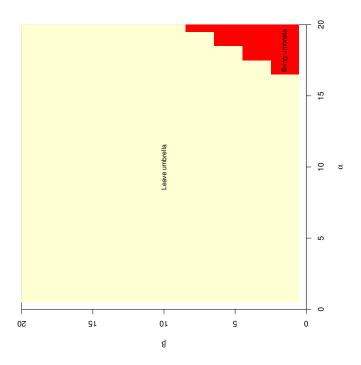


Fig. 15.2. How the Bayesian decision depends on the prior hyperparameters when s=16 and f=64