

$$p(\Theta|D) = \frac{p(D|\Theta)p(\Theta)}{p(D|\Theta)p(\Theta) + p(D|\neg\Theta)p(\neg\Theta)}$$

## Bayesian Learning 732A46: Lecture 2

Matias Quiroz<sup>1,2</sup>

<sup>1</sup>Division of Statistics and Machine Learning, Linköping University

<sup>2</sup>Research Division, Sveriges Riksbank

March 2015

- ▶ The Poisson model
- ▶ Conjugate priors
- ▶ Prior elicitation
- ▶ Non-informative priors

# The Poisson model with a Gamma prior

## ► Model:

$$y_1, \dots, y_n | \theta \stackrel{iid}{\sim} \text{Poisson}(y_i | \theta) = \frac{1}{y_i!} \theta^{y_i} \exp(-\theta), \quad \theta > 0.$$

## ► Likelihood

$$p(y|\theta) = \prod_{i=1}^n p(y_i|\theta) \propto \theta^{\sum_{i=1}^n y_i} \exp(-\theta n),$$

## ► Prior

$$p(\theta) \propto \theta^{\alpha_0 - 1} \exp(-\theta \beta_0) \propto \text{Gamma}(\theta | \alpha_0, \beta_0)$$

**Interpretation:** contains the info:  $\alpha_0 - 1$  counts in  $\beta_0$  observations.

## ► Posterior

$$\begin{aligned} p(\theta|y) &\propto \left[ \prod_{i=1}^n p(y_i|\theta) \right] p(\theta) \\ &\propto \theta^{\sum_{i=1}^n y_i} \exp(-\theta n) \theta^{\alpha_0 - 1} \exp(-\theta \beta_0) \\ &= \theta^{(\alpha_0 + \sum_{i=1}^n y_i) - 1} \exp[-\theta(\beta_0 + n)] \propto \text{Gamma}(\theta | \underbrace{\alpha_0 + \sum_{i=1}^n y_i}_{\alpha_n}, \underbrace{\beta_0 + n}_{\beta_n}). \end{aligned}$$

# Poisson example - Bomb hits in London

$$n = 576, \sum_{i=1}^n y_i = 229 \cdot 0 + 211 \cdot 1 + 93 \cdot 2 + 35 \cdot 3 + 7 \cdot 4 + 1 \cdot 5 = 537.$$

**Average number of hits** per region  $= \bar{y} = 537/576 \approx 0.9323$ .

$$p(\theta|y) \propto \theta^{\alpha_0+537-1} \exp[-\theta(\beta_0 + 576)]$$

$$E(\theta|y) = \frac{\alpha_0 + \sum_{i=1}^n y_i}{\beta_0 + n} \approx \bar{y} \approx 0.9323,$$

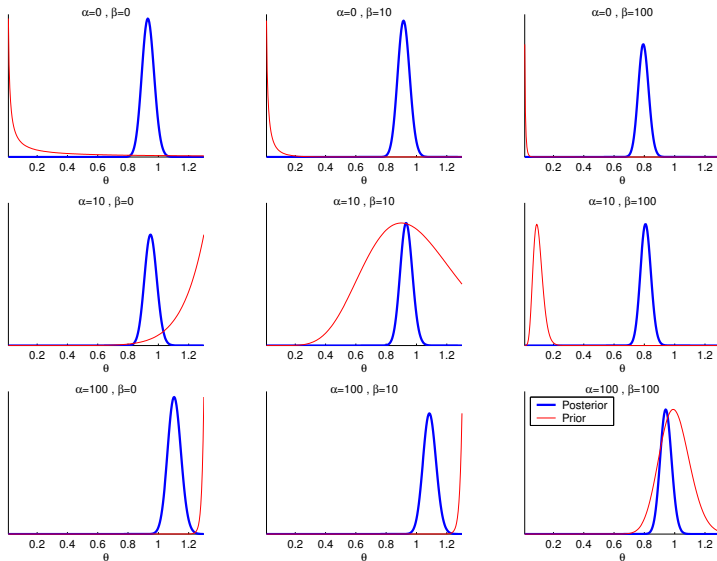
and

$$SD(\theta|y) = \left( \frac{\alpha_0 + \sum_{i=1}^n y_i}{(\beta_0 + n)^2} \right)^{1/2} = \frac{(\alpha_0 + \sum_{i=1}^n y_i)^{1/2}}{(\beta_0 + n)} \approx \frac{(537)^{1/2}}{576} \approx 0.0402.$$

if  $\alpha$  and  $\beta$  **are small compared** to  $\sum_{i=1}^n y_i$  and  $n$ .

# Poisson bomb hits in London

Analysis of bomb hits in regions of London – Poisson model with Gamma prior



# Poisson example - posterior intervals

- ▶ **Bayesian 95% interval**: the probability that the **unknown parameter**  $\theta$  lies in the interval is 0.95. **What an easy and logical interpretation!**
- ▶ *Approximate* 95% **credible interval** for  $\theta$  (for small  $\alpha_0$  and  $\beta_0$ ):

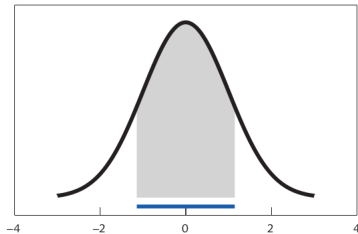
$$E(\theta|y) \pm 1.96 \cdot SD(\theta|y) = [0.8535; 1.0111]$$

**Assumes that**  $p(\theta|y)$  is (approximately) normal.

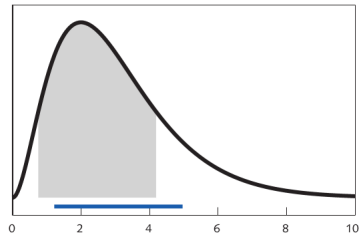
- ▶ An exact 95% **equal-tail interval** is  $[0.8550; 1.0125]$  (assuming  $\alpha_0 = \beta_0 = 0$ )
- ▶ **Highest Posterior Density (HPD)** interval contains the  $\theta$  values with highest pdf. Here  $[0.8525; 1.0144]$ , assuming  $\alpha = \beta = 0$ .

# Illustration of different interval types

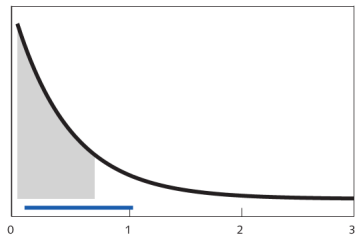
Symmetrical distribution



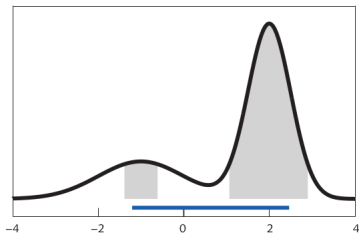
Skewed distribution



Skewed monotonous distribution



Bimodal distribution



# Conjugate priors

- **Models** we have seen

Model	Prior	→	Posterior
Bernoulli	$\theta \sim \text{Beta}(\alpha_0, \beta_0)$	→	$\theta y \sim \text{Beta}(\alpha_n, \beta_n)$
Normal ( $\sigma^2$ known)	$\theta \sim \mathcal{N}(\mu_0, \tau_0^2)$	→	$\theta y \sim \mathcal{N}(\mu_n, \tau_n^2)$
Poisson	$\theta \sim \text{Gamma}(\alpha_0, \beta_0)$	→	$\theta y \sim \text{Gamma}(\alpha_n, \beta_n)$

- **Conjugate priors:** A prior is conjugate to a model (likelihood) if the prior and posterior belong to the same distributional family.
- **Formally:** Let  $\mathcal{F} = \{p(y|\theta), \theta \in \Theta\}$  be a class of sampling distributions. A family of distributions  $\mathcal{P}$  is conjugate for  $\mathcal{F}$  if

$$p(\theta) \in \mathcal{P} \Rightarrow p(\theta|y) \in \mathcal{P}$$

holds for all  $p(y|\theta) \in \mathcal{F}$ .

- A Conjugate prior is **computationally convenient**.



# Prior elicitation

- ▶ The prior should (ideally) be elicited by an **expert** ( $\neq$  statistician, often)
- ▶ Elicit the prior on a **quantity that she knows well** (maybe log odds  $\log \frac{\theta}{1-\theta}$  when the model is  $\text{Bern}(\theta)$ ).
- ▶ The statistician can compute the **implied prior** on  $\theta$  by transformation of variables.

**Recall:** Let  $p_u(u)$  be continuous and let  $v = h(u)$  be a one-to-one transform.

$$p_v(v) = p_u(h^{-1}(v))|J|, \quad |J| = \text{determinant of } h^{-1}(v) \left[ 1 - \dim : \frac{d}{dv} h^{-1}(v) \right].$$

- ▶ **Example:** expert believes  $\phi = \log \frac{\theta}{1-\theta} \sim \mathcal{N}(0, 20)$ . The implied prior on  $\theta$  is  $[u = \phi, v = \theta, h^{-1}(v) = \log \frac{v}{1-v}]$

$$p_\theta(\theta) = \mathcal{N} \left( \log \frac{\theta}{1-\theta} \middle| 0, 20 \right) \frac{1}{\theta(1-\theta)}, \quad 0 < \theta < 1.$$

- ▶ The example works out a **full distribution**.

- ▶ Working out **hyper-parameters from expert information**.
- ▶ Elicit the prior by asking the expert simple questions: What is  $E(\theta)$ ? or  $V(\theta)$ ?
- ▶ The hyper-parameters are "backed out". **Example:** The prior is

$p(\theta) = \text{Gamma}(\theta|\alpha_0, \beta_0)$ , expert believes  $E(\theta) = 2$  and  $V(\theta) = 0.25$ .

$$E(\theta) = \frac{\alpha_0}{\beta_0}, \quad V(\theta) = \frac{\alpha_0}{\beta_0^2} \implies p(\theta) = \text{Gamma}(\theta|16, 8).$$

- ▶ **Show the expert some consequences** of her elicited prior.

# Prior elicitation - AR(p) example

- ▶ **Autoregressive process** of order  $p$

$$y_t = \mu + \phi_1 \cdot (y_{t-1} - \mu) + \dots + \phi_p \cdot (y_{t-p} - \mu) + \varepsilon_t, \quad \varepsilon_t \stackrel{iid}{\sim} N(0, \sigma^2)$$

- ▶ **Informative prior** on the unconditional mean:  $\mu \sim N(\mu_0, \tau_0^2)$ .
- ▶ **"Non-informative"** prior on  $\sigma^2$ :

$$p(\sigma^2) \propto 1/\sigma^2 \quad [\text{uniform in the parameterization } p(\log(\sigma^2)) \propto c]$$

- ▶ **Assume** for simplicity that all  $\phi_i, i = 1, \dots, p$  are independent a priori, and  $\phi_i \sim N(\mu_i, \psi_i^2)$ .
- ▶ Prior on  $\phi = (\phi_1, \dots, \phi_p)$  centered on a persistent AR(1) process:

$$\mu_1 = 0.8, \mu_2 = \dots = \mu_p = 0.$$

- ▶ **Prior variance**  $\psi_i^2$  of the  $\phi_i$  decay towards zeros:  $\text{Var}(\phi_i) = \frac{c}{i^\lambda}$ , so that "longer" lags are **more concentrated around zero** (less likely a priori).
- ▶  $\lambda$  is a parameter that can be used to determine the rate of decay.  
**Shrinkage/regularization/smoothness** prior.

# Different types of prior information

- ▶ Real **expert information**. Combo of previous studies and experience.
- ▶ Vague prior information, or even **non-informative priors**. **Beware of improper priors - make sure the posterior is proper!**
- ▶ **Smoothness priors**. Regularization. Shrinkage. Big thing in modern statistics/machine learning.
- ▶ **Hierarchical priors**. Model the uncertainty in the hyper-parameters. **Bayesian estimation of hyper-parameters.**

# Non-informative priors

- ▶ **Do not exist!** The "flatness" depends on the parametrization of the model.
- ▶ Can be improper but still lead to a **proper posterior**.
- ▶ **Reference prior**: A prior that plays a "minimal role". "Let the data speak for themselves".
- ▶ Jeffreys' **invariance principle**: The prior should contain the same information **regardless of the parametrization** of the model.
- ▶ **Jeffreys'** prior (1-dim)

$$p(\theta) \propto |I(\theta)|^{1/2}, \quad I(\theta) = -E_y \left( \frac{d^2}{d\theta^2} \log p(y|\theta) \right),$$

where  $I(\theta)$  is the **Fisher information** for  $\theta$ .

- ▶ The expectation is **w.r.t data**... an **unconditional** (frequentist) feature!
- ▶ ... consequently, Jeffreys' prior **does not respect** the likelihood principle.
- ▶ Can give **dubious results** in multivariate (parameter) models.

# Jeffreys' prior for Bernoulli trial data

Let  $y = (y_1, \dots, y_n)$

$$y_1, \dots, y_n | \theta \stackrel{iid}{\sim} \text{Bern}(\theta) \quad \text{and} \quad \log p(y|\theta) = s \log \theta + f \log(1 - \theta).$$

$$\begin{aligned} \frac{d \log p(y|\theta)}{d\theta} &= \frac{s}{\theta} - \frac{f}{(1-\theta)} \\ \frac{d^2 \log p(y|\theta)}{d\theta^2} &= -\frac{s}{\theta^2} - \frac{f}{(1-\theta)^2} \\ I(\theta) &= \frac{E_y(s)}{\theta^2} + \frac{E_{y|\theta}(f)}{(1-\theta)^2} \\ &= \frac{n\theta}{\theta^2} + \frac{n(1-\theta)}{(1-\theta)^2} = \frac{n}{\theta(1-\theta)} \end{aligned}$$

Thus, **the Jeffreys' prior** is

$$p(\theta) = |I(\theta)|^{1/2} \propto \theta^{-1/2} (1-\theta)^{-1/2} \propto \text{Beta}(\theta|1/2, 1/2).$$

# Non-informative priors - my two cents

- ▶ **OVERRATED.** Likelihood **dominates the prior** as more data becomes available.
- ▶ **State-of-the-art** models are **very complex** these days.  
**Regularization/shrinkage/smoothness priors** to avoid over-fitting.
- ▶ Non-informative priors **do not shrink**.

**Non-informative prior  $\implies$  no shrinkage  $\implies$  no fun.**